EECS 545: Machine Learning Lecture 5. Classification 2

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Outline

- Probabilistic Discriminative models
 - Objective: maximize conditional likelihood over training data

$$\prod_{i} P(y^{(i)}|\mathbf{x}^{(i)},\mathbf{w})$$

- Logistic Regression (covered in previous lecture)
- Softmax Regression: Multiclass extension of logistic regression
- Probabilistic Generative models
 - Objective: maximize joint likelihood over training data

$$\prod_{i} P(\mathbf{x}^{(i)}, y^{(i)}|\mathbf{w})$$

- Gaussian Discriminant Analysis
- Naive Bayes (part 1)

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Softmax regression for multiclass classification

- For multiclass case, we can use softmax regression.
 - Softmax regression can be viewed as a generalization of logistic regression
- Recall that, logistic regression (binary classification) models class conditional probability as:

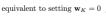
$$p(y = 1 \mid \mathbf{x}; \mathbf{w}) = \frac{\exp(\mathbf{w}^{\top} \phi(\mathbf{x}))}{1 + \exp(\mathbf{w}^{\top} \phi(\mathbf{x}))}$$
$$p(y = 0 \mid \mathbf{x}; \mathbf{w}) = \frac{1}{1 + \exp(\mathbf{w}^{\top} \phi(\mathbf{x}))}$$



- Note that these probability sum to 1.
- For multiclass classification (with K classes), we use the following model

$$p(y = k \mid \mathbf{x}; \mathbf{w}) = \frac{\exp(\mathbf{w}_k^\top \phi(\mathbf{x}))}{1 + \sum_{j=1}^{K-1} \exp(\mathbf{w}_j^\top \phi(\mathbf{x}))}$$
$$p(y = K \mid \mathbf{x}; \mathbf{w}) = \frac{1}{1 + \sum_{k=1}^{K-1} \exp(\mathbf{w}_k^\top \phi(\mathbf{x}))}$$

for
$$k = \{1, \dots, K - 1\}$$





Note that these probability sum to 1

Last updated: 2/5/2025 12pm

Softmax regression: Log-likelihood (objective function) and learning

• Defining $\mathbf{w}_K = 0$, we can write as:

$$p(y = k \mid \mathbf{x}; \mathbf{w}) = \frac{\exp\left(\mathbf{w}_k^{\top} \phi(\mathbf{x})\right)}{\sum_{j=1}^{K} \exp\left(\mathbf{w}_j^{\top} \phi(\mathbf{x})\right)}$$
or
$$p(y \mid \mathbf{x}; \mathbf{w}) = \prod_{k=1}^{K} \left[\frac{\exp\left(\mathbf{w}_k^{\top} \phi(\mathbf{x})\right)}{\sum_{j=1}^{K} \exp\left(\mathbf{w}_j^{\top} \phi(\mathbf{x})\right)} \right]^{\mathbb{I}(y=k)}$$

· Log-Likelihood

$$\begin{split} \log p(D|\mathbf{w}) &= \sum_{i} \log p(y^{(i)}|\mathbf{x}^{(i)}, \mathbf{w}) \\ &= \sum_{i} \log \prod_{k=1}^{K} \left[\frac{\exp \left(\mathbf{w}_{k}^{\top} \phi(\mathbf{x}^{(i)})\right)}{\sum_{j=1}^{K} \exp \left(\mathbf{w}_{j}^{\top} \phi(\mathbf{x}^{(i)})\right)} \right]^{\mathbb{I}(y^{(i)} = k)} \end{split}$$

 We can learn w by gradient ascent for maximizing the log-likelihood or iterative Newton's method (IRLS).

Probabilistic Generative Models

Learning the Classifier

- For classification, we want to compute $p(C_k \mid \mathbf{x})$
 - (a) **Discriminative** models: Directly model $p(C_k \mid \mathbf{x})$ and learn parameters from the training set.
 - · Logistic regression
 - · Softmax regression
 - (b) **Generative** models: Learn joint densities $p(\mathbf{x}, C_k)$ by learning $p(\mathbf{x} \mid C_k)$ and $p(C_k)$, and then use Bayes rule for predicting the class C_k given \mathbf{x} :
 - Gaussian Discriminant Analysis
 - Naive Bayes

Probabilistic Generative Models

• Bayes' theorem reduces the classification problem $p(C_k \mid \mathbf{x})$ to estimating the distribution of the data:

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})} = \frac{p(\mathbf{x}|C_k)p(C_k)}{\sum_{k'} p(\mathbf{x}|C_{k'})p(C_{k'})}$$

- Density estimation can be decomposed into learning distributions from training data.
 - $-p(C_k)$
 - $-p(\mathbf{x} \mid C_k)$
- Maximum likelihood estimation for $p(\mathbf{x}, C_k)$

Probabilistic Generative Models

• For two classes, Bayes' theorem says:

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$

• Use log odds (i.e., logit "score"):

$$a = \log \frac{p(C_1|\mathbf{x})}{p(C_2|\mathbf{x})} = \log \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)}$$

• Then we can define the posterior via the sigmoid:

$$p(C_1|\mathbf{x}) = \frac{1}{1 + \exp(-a)} = \sigma(a)$$

Gaussian Discriminant Analysis

Gaussian Discriminant Analysis

- · Probability of class label
 - $-p(C_k)$: Constant (e.g., Bernoulli)
- · Conditional probability of data given a class
 - $-p(\mathbf{x} \mid C_k)$: Gaussian distribution

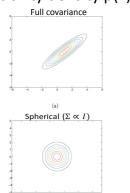
$$p(\mathbf{x} \mid C_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \mu_k)^{\top} \mathbf{\Sigma}^{-1} (\mathbf{x} - \mu_k)\right\}$$

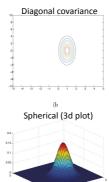
• Classification: use Bayes rule (previous slide)

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Examples of Gaussian Distributions

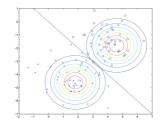
Probability density p(x) for 2 dimensional case

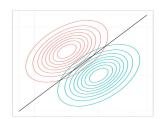




Gaussian Discriminant Analysis

- Basic GDA assumes the same covariance for all classes
 - The figure below shows class-specific density and decision boundary. Note the linear decision boundary for any types of covariance matrices!





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Prediction: Class-Conditional Densities

• Suppose we model $p(x \mid C_k)$ as Gaussians with the <u>same covariance</u> matrix.

$$p(\mathbf{x} \mid C_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \mu_k)^{\top} \mathbf{\Sigma}^{-1} (\mathbf{x} - \mu_k)\right\}$$

• This gives us $p(C_1 \mid \mathbf{x}) = \sigma(\mathbf{w}^{\top}\mathbf{x} + w_0)$ - where $\mathbf{w} = \mathbf{\Sigma}^{-1}(\mu_1 - \mu_2)$ and $w_0 = -\frac{1}{2}\mu_1^{\top}\mathbf{\Sigma}^{-1}\mu_1 + \frac{1}{2}\mu_2^{\top}\mathbf{\Sigma}^{-1}\mu_2 + \log\frac{p(C_1)}{p(C_2)}$ Derivation

$$\begin{split} P\left(x,C_{1}\right) &= P\left(x\mid C_{1}\right) P\left(C_{1}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x-\mu_{1}\right)^{\top} \Sigma^{-1}\left(x-\mu_{1}\right)\right\} P\left(C_{1}\right) \\ P\left(x,C_{2}\right) &= P\left(x\mid C_{2}\right) P\left(C_{2}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x-\mu_{2}\right)^{\top} \Sigma^{-1}\left(x-\mu_{2}\right)\right\} P\left(C_{2}\right) \end{split}$$

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Derivation

$$\begin{split} P\left(x,C_{1}\right) &= P\left(x \mid C_{1}\right) P\left(C_{1}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{1}\right)^{\top} \Sigma^{-1}\left(x - \mu_{1}\right)\right\} P\left(C_{1}\right) \\ P\left(x,C_{2}\right) &= P\left(x \mid C_{2}\right) P\left(C_{2}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{2}\right)^{\top} \Sigma^{-1}\left(x - \mu_{2}\right)\right\} P\left(C_{2}\right) \\ \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{P\left(C_{2} \mid \mathbf{x}\right)} &= \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{1 - P\left(C_{1} \mid \mathbf{x}\right)} \qquad \text{"Log-odds"} \end{split}$$

Derivation

$$\begin{split} P\left(x,C_{1}\right) &= P\left(x \mid C_{1}\right) P\left(C_{1}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{1}\right)^{\top} \Sigma^{-1}\left(x - \mu_{1}\right)\right\} P\left(C_{1}\right) \\ P\left(x,C_{2}\right) &= P\left(x \mid C_{2}\right) P\left(C_{2}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{2}\right)^{\top} \Sigma^{-1}\left(x - \mu_{2}\right)\right\} P\left(C_{2}\right) \\ \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{P\left(C_{2} \mid \mathbf{x}\right)} &= \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{1 - P\left(C_{1} \mid \mathbf{x}\right)} \quad \text{"Log-odds"} \\ &= \log \frac{\exp\left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{1}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{1}\right)\right\}}{\exp\left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{2}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{2}\right)\right\}} + \log \frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \end{split}$$

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Derivation

$$\begin{split} P\left(x,C_{1}\right) &= P\left(x \mid C_{1}\right) P\left(C_{1}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{1}\right)^{\top} \Sigma^{-1}\left(x - \mu_{1}\right)\right\} P\left(C_{1}\right) \\ P\left(x,C_{2}\right) &= P\left(x \mid C_{2}\right) P\left(C_{2}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{2}\right)^{\top} \Sigma^{-1}\left(x - \mu_{2}\right)\right\} P\left(C_{2}\right) \\ \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{P\left(C_{2} \mid \mathbf{x}\right)} &= \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{1 - P\left(C_{1} \mid \mathbf{x}\right)} &\text{"Log-odds"} \\ &= \log \frac{\exp\left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{1}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{1}\right)\right\}}{\exp\left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{2}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{2}\right)\right\}} + \log \frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \\ &= \left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{1}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{1}\right)\right\} - \left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{2}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{2}\right)\right\} + \log \frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \end{split}$$

Derivation

$$\begin{split} P\left(x,C_{1}\right) &= P\left(x \mid C_{1}\right) P\left(C_{1}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{1}\right)^{\top} \Sigma^{-1}\left(x - \mu_{1}\right)\right\} P\left(C_{1}\right) \\ P\left(x,C_{2}\right) &= P\left(x \mid C_{2}\right) P\left(C_{2}\right) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\left(x - \mu_{2}\right)^{\top} \Sigma^{-1}\left(x - \mu_{2}\right)\right\} P\left(C_{2}\right) \\ \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{P\left(C_{2} \mid \mathbf{x}\right)} &= \log \frac{P\left(C_{1} \mid \mathbf{x}\right)}{1 - P\left(C_{1} \mid \mathbf{x}\right)} \quad \text{"Log-odds"} \\ &= \log \frac{\exp\left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{1}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{1}\right)\right\}}{\exp\left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{2}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{2}\right)\right\}} + \log \frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \\ &= \left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{1}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{1}\right)\right\} - \left\{-\frac{1}{2}\left(\mathbf{x} - \mu_{2}\right)^{\top} \Sigma^{-1}\left(\mathbf{x} - \mu_{2}\right)\right\} + \log \frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \\ &= \left(\mu_{1} - \mu_{2}\right)^{\top} \Sigma^{-1}\mathbf{x} - \frac{1}{2}\mu_{1}^{\top} \Sigma^{-1}\mu_{1} + \frac{1}{2}\mu_{2}^{\top} \Sigma^{-1}\mu_{2} + \log \frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \end{split}$$

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Derivation

$$\begin{split} P\left(x,C_{1}\right) &= P\left(x\mid C_{1}\right)P\left(C_{1}\right) \\ &= \frac{1}{(2\pi)^{D/2}}\frac{1}{|\Sigma|^{1/2}}\exp\left\{-\frac{1}{2}\left(x-\mu_{1}\right)^{\top}\Sigma^{-1}\left(x-\mu_{1}\right)\right\}P\left(C_{1}\right) \\ P\left(x,C_{2}\right) &= P\left(x\mid C_{2}\right)P\left(C_{2}\right) \\ &= \frac{1}{(2\pi)^{D/2}}\frac{1}{|\Sigma|^{1/2}}\exp\left\{-\frac{1}{2}\left(x-\mu_{2}\right)^{\top}\Sigma^{-1}\left(x-\mu_{2}\right)\right\}P\left(C_{2}\right) \\ \log \frac{P\left(C_{1}\mid\mathbf{x}\right)}{P\left(C_{2}\mid\mathbf{x}\right)} &= \log\frac{P\left(C_{1}\mid\mathbf{x}\right)}{1-P\left(C_{1}\mid\mathbf{x}\right)} \quad \text{"Log-odds"} \\ &= \log\frac{\exp\left\{-\frac{1}{2}\left(\mathbf{x}-\mu_{1}\right)^{\top}\Sigma^{-1}\left(\mathbf{x}-\mu_{1}\right)\right\}}{\exp\left\{-\frac{1}{2}\left(\mathbf{x}-\mu_{1}\right)^{\top}\Sigma^{-1}\left(\mathbf{x}-\mu_{1}\right)\right\}} + \log\frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \\ &= \left\{-\frac{1}{2}\left(\mathbf{x}-\mu_{1}\right)^{\top}\Sigma^{-1}\left(\mathbf{x}-\mu_{1}\right)\right\} - \left\{-\frac{1}{2}\left(\mathbf{x}-\mu_{2}\right)^{\top}\Sigma^{-1}\left(\mathbf{x}-\mu_{2}\right)\right\} + \log\frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \\ &= \left(\mu_{1}-\mu_{2}\right)^{\top}\Sigma^{-1}\mathbf{x} - \frac{1}{2}\mu_{1}^{\top}\Sigma^{-1}\mu_{1} + \frac{1}{2}\mu_{2}^{\top}\Sigma^{-1}\mu_{2} + \log\frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \\ &= \left(\Sigma^{-1}\left(\mu_{1}-\mu_{2}\right)\right)^{\top}\mathbf{x} + w_{0} \\ \qquad \qquad \text{where } w_{0} = -\frac{1}{2}\mu_{1}^{\top}\Sigma^{-1}\mu_{1} + \frac{1}{2}\mu_{2}^{\top}\Sigma^{-1}\mu_{2} + \log\frac{P\left(C_{1}\right)}{P\left(C_{2}\right)} \end{split}$$

Prediction: Class-Conditional Densities for shared covariances

• $p(C_k | \mathbf{x})$ is a sigmoid function:

$$\sigma(a) = \frac{1}{1 + \exp(-a)}$$

0.5

– with log-odds (*logit* function):

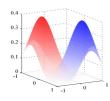
$$a = \log\left(\frac{\sigma}{1-\sigma}\right) = \left(\boldsymbol{\Sigma}^{-1}(\mu_1 - \mu_2)\right)^{\top} \mathbf{x} + w_0$$
 where $w_0 = -\frac{1}{2}\mu_1^{\top} \boldsymbol{\Sigma}^{-1} \mu_1 + \frac{1}{2}\mu_2^{\top} \boldsymbol{\Sigma}^{-1} \mu_2 + \log\frac{p(C_1)}{p(C_2)}$

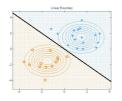
• Generalizes to normalized exponential, or softmax :

$$p_i = \frac{\exp(q_i)}{\sum_j \exp(q_j)}$$

Prediction: Linear Decision Boundaries

- At decision boundary, we have $p(C_1|\mathbf{x}) = p(C_2|\mathbf{x})$
- With the same covariance matrices, the boundary $p(C_1|\mathbf{x}) = p(C_2|\mathbf{x})$ is linear.
 - Different class $p(C_1)$, $p(C_2)$ just shift it around.





Likelihood function of generative models

• The likelihood of Data $\{(\mathbf{x}^{(n)}, y^{(n)})\}$

$$P(D|\mathbf{w}) = \prod_{i=1}^{N} P(\mathbf{x}^{(i)}, y^{(i)}|\mathbf{w}) \xrightarrow{P(\mathbf{X}, \mathbf{y}|\mathbf{w})} P(\mathbf{X}, \mathbf{y}|\mathbf{w})$$
Decomposition of the joint probability
$$= \prod_{i=1}^{N} P(\mathbf{x}^{(i)}|y^{(i)}, \mathbf{w}) P(y^{(i)}|\mathbf{w})$$

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Learning parameters via maximum likelihood

• Given training data $\{(\mathbf{x}^{(1)}, y^{(1)}), \cdots, (\mathbf{x}^{(N)}, y^{(N)})\}$ and a generative model ("shared covariance")

$$p(y) = \phi^{y} (1 - \phi)^{1-y}$$

$$p(\mathbf{x}|y=0) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (\mathbf{x} - \mu_0)^{\top} \Sigma^{-1} (\mathbf{x} - \mu_0)\right)$$

$$p(\mathbf{x}|y=1) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (\mathbf{x} - \mu_1)^{\top} \Sigma^{-1} (\mathbf{x} - \mu_1)\right)$$

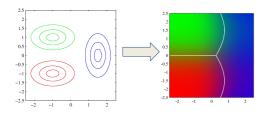
Learning via maximum likelihood

• Maximum likelihood estimation (HW2):

$$\begin{split} \phi &= \frac{1}{N} \sum_{i=1}^{N} \mathbb{I}\{y^{(i)} = 1\} \\ \mu_0 &= \frac{\sum_{i=1}^{N} \mathbb{I}\{y^{(i)} = 0\} \mathbf{x}^{(i)}}{\sum_{i=1}^{N} \mathbb{I}\{y^{(i)} = 0\}} \\ \mu_1 &= \frac{\sum_{i=1}^{N} \mathbb{I}\{y^{(i)} = 1\} \mathbf{x}^{(i)}}{\sum_{i=1}^{N} \mathbb{I}\{y^{(i)} = 1\}} \\ \Sigma &= \frac{1}{N} \sum_{i=1}^{N} (\mathbf{x}^{(i)} - \mu_{y^{(i)}}) (\mathbf{x}^{(i)} - \mu_{y_{(i)}})^{\top} \end{split}$$

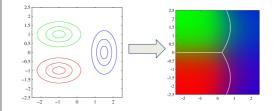
Different Covariance

 Decision boundaries between some classes can be quadratic when they have different covariances.



Different Covariance

 Decision boundaries between some classes can be quadratic when they have different covariances.





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Comparison between GDA and Logistic regression (or softmax regression)

- · Logistic regression:
 - For an M-dimensional feature space, this model has M parameters to fit.
- · Gaussian Discriminative Analysis
 - 2M parameters for the means of $p(\mathbf{x} \mid C_1)$ and $p(\mathbf{x} \mid C_2)$
 - -M(M+1)/2 parameters for the shared covariance matrix
- Logistic regression has less parameters and is more flexible about data distribution.
- GDA has a stronger modeling assumption, and works well when the distribution follows the assumption.

Naive Bayes Classifier

(Brief Intro: to be continued in the next lecture)

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Naive Bayes classifier

• Probability of class label:

- $p(C_k)$: Constant (e.g., Bernoulli)

- Conditional probability of data given the class
 - Naive Bayes assumption: $p(\mathbf{x} \mid C_k)$ is factorized (Each coordinate of \mathbf{x} is conditionally independent of other coordinates given the class label)

$$P(x_1, ..., x_M | C_k) = P(x_1 | C_k) \cdots P(x_M | C_k) = \prod_{k=1}^{M} P(x_j | C_k)$$

• Classification: use Bayes rule

(binary)
$$P(C_1|\mathbf{x}) = \frac{P(C_1,\mathbf{x})}{P(\mathbf{x})} = \frac{P(C_1,\mathbf{x})}{P(C_1,\mathbf{x}) + P(C_2,\mathbf{x})}$$

Naive Bayes classifier

• When classifying, we can simply find the class C_k that maximizes $P(C_k|\mathbf{x})$ using the Bayes rule:

$$\arg\max_{k} P(C_k|\mathbf{x}) = \arg\max_{k} P(C_k, \mathbf{x})$$

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Naive Bayes classifier

• When classifying, we can simply find the class C_k that maximizes $P(C_k|\mathbf{x})$ using the Bayes rule:

$$\arg \max_{k} P(C_k | \mathbf{x}) = \arg \max_{k} P(C_k, \mathbf{x})$$
$$= \arg \max_{k} P(C_k) P(\mathbf{x} | C_k)$$

Naive Bayes classifier

• When classifying, we can simply find the class C_k that maximizes $P(C_k|\mathbf{x})$ using the Bayes rule:

$$\arg\max_k P(C_k|\mathbf{x}) = \arg\max_k P(C_k,\mathbf{x})$$

$$= \arg\max_k P(C_k)P(\mathbf{x}|C_k)$$
 Naive Bayes assumption
$$= \arg\max_k P(C_k)\prod_{j=1}^M P(x_j|C_k)$$

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Example: Naive Bayes for real-valued inputs

- Probability of class label:
 - $-p(C_{\nu})$: Constant (e.g., Bernoulli)
- Conditional probability of data given the class
 - Naive Bayes assumption: $P(\mathbf{x} | C_{\nu})$ is factorized (e.g., 1D Gaussian)

$$\begin{split} P(x_1, \dots, x_M | C_k) &= P(x_1 | C_k) \cdots P(x_M | C_k) \\ &= \prod_{j=1}^M P(x_j | C_k) \\ &= \prod_{j=1}^M \frac{1}{\sqrt{2\pi\sigma_j^2}} \exp\left(-\frac{(x_j - \mu_j)^2}{2\sigma_j^2}\right) \end{split}$$

- Note: this is equivalent to GDA with diagonal covariance!!

Comparison: Discriminative vs. Generative

- The *generative* approach is typically model-based, and it can generate synthetic data from $p(\mathbf{x} \mid C_t)$.
 - By comparing the synthetic data and real data, we get a sense of how good the generative model is.
- The discriminative approach will typically have fewer parameters to estimate and have less assumptions about data distribution.
 - Linear (e.g. logistic regression) v/s quadratic (e.g., Gaussian discriminant analysis) in the dimension of the input.
 - Less generative assumptions about the data (however, constructing the features may require domain knowledge)

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Any feedback (about lecture, slide, homework, project, etc.)?

(via anonymous google form: https://forms.gle/fpYmiBtG9Me5qbP37



Change Log of lecture slides:

https://docs.google.com/document/d/e/2PACX-1vSSIHJiklypK7rKFSR1-5GYXyBCEW8UPtpSfCR9AR6M1l7K9ZOEmxfFwaWaW7kLDxusthsF8WlCyZJ-/pub