

Lec 18

Tuesday, 2 September 2025 10:10 AM

C.I.T:

 $p=1$ $X \rightarrow$ pop. random variable. No. distribution assumption.

$$E(X) = \mu < \infty$$

$$\text{Var}(X) = \sigma^2 < \infty$$

 \bar{X}_n - sample mean \rightarrow random variable.
 $n \rightarrow \infty$.

$$Z = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{n \rightarrow \infty} N(0,1).$$

$$\sqrt{n}(\bar{X}_n - \mu)\sigma^{-1} \sim N(0,1).$$

$$\sqrt{n}(\bar{X}_n - \mu) \sim N(0, \sigma^2).$$

In case if $X \sim N(\mu, \sigma^2)$ then we can drop condition of n .

$$\sqrt{n}(\bar{X}_n - \mu) \sim N(0, \sigma^2) \quad \forall n \in \mathbb{N}.$$

p>1. Under similar conditions \rightarrow

$$\sqrt{n}(\bar{X}_n - \mu) \sim N(0, \Sigma) \text{ as } n \rightarrow \infty.$$

i.e. distribution of \bar{X} is normal/gaussian if $n \rightarrow \infty$.create ways of creating samples \sim attach probs to samples & take combinations.
Simulated data: Mix up, shuffle the original & simulated & then divide 70-30.Recall from notes \sim

$$(\underline{X} - \underline{\mu})^T \Sigma^{-1} (\underline{X} - \underline{\mu}) \sim \chi^2_p.$$

here $\underline{\mu} = E(\underline{X})$ & $\underline{\Sigma} = \text{Cov}(\underline{X})$.If we take $\bar{\underline{X}}$ instead of \underline{X} .

$$(\bar{\underline{X}} - \underline{\mu})^T \left(\frac{\underline{\Sigma}}{n}\right)^{-1} (\bar{\underline{X}} - \underline{\mu}) \sim \chi^2_p. \quad \underline{\mu} = E(\bar{\underline{X}}) \text{ & } \underline{\Sigma} = \text{Cov}(\bar{\underline{X}}).$$

$$n(\bar{\underline{X}} - \underline{\mu})^T \underline{\Sigma}^{-1} (\bar{\underline{X}} - \underline{\mu}) \sim \chi^2_p$$

when $n \rightarrow \infty$ then we have deg. w.t.N $S \xrightarrow{p} \Sigma$ component wise.

$$\lim_{n \rightarrow \infty} P(|a_{ij} - s_{ij}| > \epsilon) = 0 \quad \forall (i,j).$$

If n is suff. large then Σ (if unknown) can be estimated by S .
Then $n(\bar{\underline{X}} - \underline{\mu})^T S^{-1} (\bar{\underline{X}} - \underline{\mu}) \sim \chi^2_p$ when $n \rightarrow \infty$.

$$S_n = \frac{n-1}{n} S \rightarrow \text{consistent \& unbiased estimator of } \Sigma.$$

$$MLE \hat{\Sigma} = S.$$

$$\left[S = \frac{n-1}{n-1} S_n \right] \Rightarrow S^{-1} = \frac{n-1}{n} S_n^{-1} \Rightarrow (n-1)(\bar{\underline{X}} - \underline{\mu})^T S_n^{-1} (\bar{\underline{X}} - \underline{\mu}) \sim \chi^2_p.$$

Hypothesis testing for population mean \sim Assume population $X \sim N(\mu, \Sigma)$.Null hypothesis $H_0: \mu = \mu_0$
Alternative " $H_1: \mu \neq \mu_0$ } two-tailed test.

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0,1)$$

$$\left| \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \right| > z_{\alpha/2} : \text{reject } H_0. \quad \text{2 test.}$$

If this not satisfied; we don't have enough info to reject H_0 .② σ unknown $p=1$.

$$\text{Result 1: } (n-1) \frac{S^2}{\sigma^2} \sim \chi^2_{n-1}.$$

Result 2: \bar{X} & S are indep. r.v.

$$\text{Result 3: } \sqrt{n}(\bar{X} - \mu) \sim N(0,1).$$

Recall: $p=1$.

$$X \sim N(0,1)$$

$$Y \sim \chi^2_m \Leftrightarrow Y = Z_1^2 + Z_2^2 + \dots + Z_m^2 \quad Z_i \rightarrow \text{independent.}$$

 Σ, Y are indep. r.v.

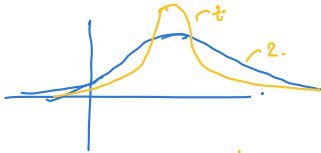
$$\Rightarrow T = \frac{\sqrt{n}(\bar{X} - \mu)}{S} = \frac{\sqrt{n}(\bar{X} - \mu)}{S} \sim t_{n-1}. \quad \text{Student } t\text{-distribution w } n\text{-degree of freedom}$$

$\sqrt{\frac{Y}{n}}$ $\sqrt{\frac{Y}{n}}$
 mean $\frac{1}{n} \sum y_i$: ratio $\frac{1}{n} \sum y_i$
 var $\frac{1}{n} \sum y_i^2 - (\frac{1}{n} \sum y_i)^2$: related to σ
 goodness of fit χ^2 test

If σ unknown can't use Z test so use t test.

It follows $\frac{\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma}}{\sqrt{\frac{(n-1)s^2/\sigma^2}{(n-1)}}} \sim t_{n-1} \Rightarrow \sqrt{n} \left(\frac{\bar{X} - \mu}{s/\sqrt{n}} \right) \sim t_{n-1}$

$H_0: \mu = \mu_0$
 $H_1: \mu \neq \mu_0$ $\left| \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \right| > t_{n-1, \alpha/2}$ reject H_0



For larger value of $n > 40$ so t dist. as good as normal bell dist.

For f , you need 2 chi square dist.

$p=1$: $y_1 \sim \chi^2_{n_1}$ $y_2 \sim \chi^2_{n_2}$
 y_1, y_2 are independent r.v.

$$F = \frac{y_1/n_1}{y_2/n_2}$$

F is a r.v. & distribution of F is called F-distribution w (n_1, n_2) degrees of freedom.

Suppose T is a r.v. which is t-dist. w n degree of freedom

$$T = \frac{\bar{X} - \mu}{s/\sqrt{n}} \quad Z \sim N(0,1), \quad Y \sim \chi^2_n$$

$$T^2 = \frac{n \frac{\bar{X} - \mu}{s/\sqrt{n}}^2}{1} = \frac{Z^2/1}{Y/n} \quad \text{if } Z \sim N(0,1) \Rightarrow Z^2 \sim \chi^2_1$$

$\sim F_{1,n}$ for distributed w.