## Solutions

## Final Exam - ECE 503 Fall 2016

- Date: Wednesday, December 14, 2016.
- Time: 10:30 am -12:30 pm
- Maximum Credit: 100 points
- 1. [10 points] In the Arizona State Lottery, a player picks 6 different numbers from a sequence of 1 through 15. At a lottery drawing, 6 balls are drawn at random from a box containing 15 balls numbered 1 through 15.
  - (a) What is the probability that the player has 0 matches and wins nothing?
  - (b) What is the probability that the player has exactly 6 matches and wins the jackpot?

Let 
$$n = 15$$
 be total # of balls

Suppose that the player picks  $\frac{søme}{m}$  balls  $(m = 6)$ 

Total # of ways to pick m balls out of  $n = \binom{n}{m}$ 

Total " " k matching balls  $= \binom{m}{k} \binom{n-m}{k}$ 
 $= \binom{m}{k} \binom{n-m}{k}$ 

(a) Prob of 0 matching  $= \binom{6}{6} \binom{15-6}{6-0} = \binom{9}{6}$ 

(b) Prob of 6 matching  $= \binom{6}{6} \binom{15-6}{6-6} = \frac{1}{\binom{15}{6}}$ 

- 2. [20 points] Suppose X and Y are independent random variables, and each one of them is uniformly distributed in [0, 1]. Let  $Z = \min(X, Y)$ , and  $W = \max(X, Y)$ .
  - (a) Find the joint density (PDF) of (Z, W).

independent.

- 3. [20 points] A wireless communication channel suffers from a clustered error pattern. Whenever a packet has an error, the next packet will have an error with probability 0.9. Whenever a packet is error-free, the next packet is error free with probability 0.99. We want to use a Markov chain to analyze the above process, assuming two states, E (error), and NE (no error).
  - (a) Draw a state-transition diagram showing the transition probabilities, and write the one-step transition matrix P.
  - (b) Find the stationary probability distribution (and comment on its uniqueness).
  - (c) Suppose that the initial probability distribution was  $p(0) = [p_E(0) \ p_{NE}(0)] = [0 \ 1]$ . Find the probability distribution at time n = 100.

the probability distribution at time 
$$\frac{1000}{1000}$$
  $m = 5$ .

(a)

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$$\overrightarrow{P}(0) = \begin{bmatrix} 0 & 1 \end{bmatrix} \qquad \overrightarrow{P}(5) = \begin{bmatrix} \frac{1}{11} - \frac{(0.89)^5}{11} & \frac{10}{11} + \frac{(0.89)^5}{11} \end{bmatrix}$$

$$3 = \begin{bmatrix} \frac{1}{11} - 0.05 & \frac{10}{11} + 0.05 \end{bmatrix} = \begin{bmatrix} 0.04 & 0.05 \end{bmatrix}$$

- 4. [15 points]
  - (a) Let X(t) be a wide sense stationary random process, and we define Y(t) = X(t+5). Are X(t) and Y(t) jointly wide sense stationary?
  - (b) Let  $X_n$  have the following PDF,

$$f_n(x) = g_n(x)(1 - 1/n^3) + h_n(x)/n^3$$

where  $g_n(x) \sim N(0, 1/n^2)$ , and  $h_n \sim N(n, 1)$ . Does  $X_n$  converge to 0 in the mean square sense?

(c) A W.S.S. process X(t), of mean  $\mu_X = 10$  is the input to a LTI filter, with impulse response

$$h(t) = \begin{cases} e^{t/0.2} & 0 \le t \le 0.2\\ 0 & \text{otherwise.} \end{cases}$$

What is the expected value of the output process Y(t)?

(a) 
$$Y(t) = X(t+s)$$
  $E[Y(t)] = E[X(t+s)] = const.$ 

$$E[Y(t) Y(t+e)] = E[X(t+s)X(t+t+s)] = R_X(e)$$

$$\Rightarrow Y(t) \text{ is } W.s.s.$$

$$E[X(t_1) Y(t_2)] = E[X(t_1) X(t_2+s)] = R_X((t_2-t_1)+s) \rightarrow only dep.$$

$$\Rightarrow X(t), Y(t) \text{ are jointly } W.S.s.$$

(b). 
$$E[|\times_{n}-0|^{2}] = E[|\times_{n}|^{2}] = \int_{-\infty}^{\infty} \chi_{n}^{2} f_{n}(x) dx.$$

$$= \int_{-\infty}^{\infty} \chi_{n}^{2} \left[ g_{n}(x) \left( 1 - \frac{1}{n^{3}} \right) + \frac{h_{n}(x)}{n^{3}} \right] dx = \int_{-\infty}^{\infty} \frac{1}{n^{2}} \left( \frac{1 - \frac{1}{n^{3}}}{n^{3}} \right) + \int_{-\infty}^{\infty} \frac{1 + n^{2}}{n^{3}} dx$$

$$= \sum_{n=0}^{\infty} \chi_{n}^{2} \left[ g_{n}(x) \left( 1 - \frac{1}{n^{3}} \right) + \frac{h_{n}(x)}{n^{3}} \right] dx = \int_{-\infty}^{\infty} \frac{1}{n^{2}} \left( \frac{1 - \frac{1}{n^{3}}}{n^{3}} \right) + \int_{-\infty}^{\infty} \frac{1 + n^{2}}{n^{3}} dx$$

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$$= \sum_{n=0}^{\infty} \chi_{n}^{2} \left[ g_{n}(x) \left( 1 - \frac{1}{n^{3}} \right) + \frac{h_{n}(x)}{n^{3}} \right] dx$$

(C). 
$$Y(t) = h(t) * X(t) = \int_{-\infty}^{\infty} h(z) X(t-z) dz$$
.  
 $E[Y(t)] = \int_{-\infty}^{\infty} h(z) E[X(t-z)] dz = \int_{-\infty}^{\infty} h(z) \times 10 \cdot dz$ .  
 $= 10 \times \int_{-\infty}^{\infty} e^{z/o \cdot 2} dz = 2 \int_{-\infty}^{\infty} e^{tz} dz = 2(e-1)$ .

- 5. [20 points] Weather analysis in Florida revealed that hurricanes occur according to a Poisson process of intensity  $\lambda=2$  per week.
  - (a) What is the average time between two hurricanes?
  - (b) What is the probability that no hurricanes occur during a given 2 week period?
  - (c) If the peak of hurricane season lasts 12 weeks, what is the expected number of hurricanes?
  - (d) Find the probability that during at least one of the 12 weeks of peak season, there are at least five hurricanes.

(a) 
$$E[x_{i}] = \frac{1}{2} = 0.5 \text{ weeks}$$
  
(b)  $P(N_{2} = 0) = 0.5 = 0.0183$   
(c)  $E[N_{12}] = 12 \times = 24 \text{ hurricanes}$   
(d)  $P(\frac{12}{N_{i}-N_{i-1}} > 5) = 1 - P(\frac{12}{N_{i}-N_{i-1}} < y)$   
 $= 1 - [e^{\lambda}(1+\lambda+\frac{\lambda^{2}}{2}+\frac{\lambda^{3}}{8}+\frac{\lambda^{4}}{24})]^{12}$   
 $= 1 - [e^{-2}(1+2+2+\frac{4}{3}+\frac{2}{3})]^{12}$   
 $= 1 - [7e^{-2}]^{12} = 1 - 0.523 = 0.477$ 

- 6. [15 points] For a zero-mean WSS random process X(t), with auto-correlation function  $R_X(\tau) = 1/(1+\tau^2)$ , we perform the following sequence of operations. The continuous time process is first sampled at the rate of  $f_s = 0$  samples per second. This yields a discrete time random process  $X_n$ . We then pass  $X_n$  through a discrete-time filter, whose output is denoted by  $Y_n = X_n X_{n-1}$ .
  - (a) Find the mean and auto-correlation function of  $X_n$ .
  - (b) Is the discrete time process  $X_n$  also WSS?
  - (c) Find the mean and auto-correlation function of  $Y_n$ .
  - (d) Are  $X_n$  and  $Y_n$  jointly WSS?

(a) 
$$\times_n = \times (nT_s)$$
  $T_s = \text{sampling period} = 1 \text{ sec.}$ 

$$E(\times_n) = E(\times(nT_s)) = 0$$

$$R_{\chi}(k) = E(\times_n \times_{n+k}) = E(\times(nT_s) \times ((n+k)T_s))$$

$$= R_{\chi}(kT_s) = \frac{1}{1+k^2T_s^2} = \frac{1}{1+k^2}$$

$$= \sum_{k=1}^{\infty} \sum_{n=1}^{\infty} \frac{1}{1+k^2T_s^2} = \frac{1}{1+k^2}$$

$$\begin{aligned} (E) & Y_{n} = X_{n} - X_{n-1} \\ & E[Y_{n}] = 0 \\ & E[Y_{n} Y_{n+k}] = E[(X_{n} - X_{n-1})(X_{n+k} - X_{n-1} + k)] \\ & = R_{x}(k) - R_{x}(k-1) - R_{x}(k+1) + R_{x}(k) \\ & = \frac{2}{1+k^{2}} - \frac{1}{1+(k-1)^{2}} - \frac{1}{1+(k+1)^{2}} \end{aligned}$$

$$E[Y_{n} \times_{n+k}] = E[(X_{n} - X_{n-1}) \times_{n+k}] = R_{x}(k) - R_{x}(k+1)$$

$$= \frac{1}{1+k^{2}} - \frac{1}{1+(k+1)^{2}}$$

$$\times_{n}, Y_{n} \text{ are also } W: S.S.$$