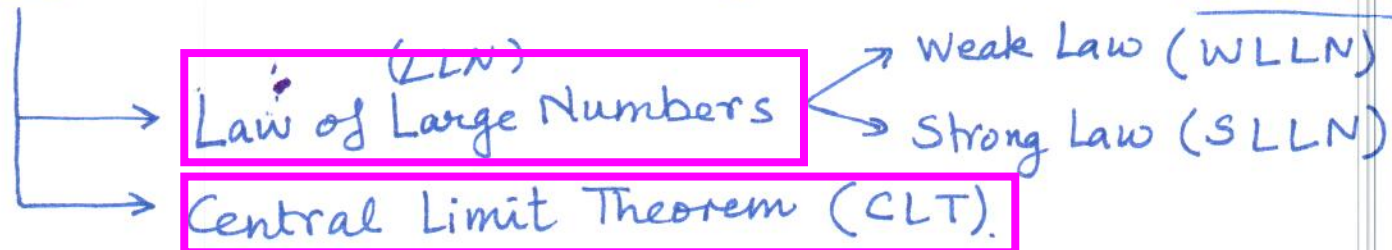


# Limit Theorems & Convergence Modes for random variables ①



LLN  $\rightarrow$  states that average of a large # of iid r.v.'s converges to the expected value.

CLT  $\rightarrow$  states that (under some conditions), sum of a large # of r.v.'s has an approximately Normal distribution. (Gaussian)

## Weak Law of Large Numbers (WLLN)

For iid r.v.'s  $X_1, X_2, \dots, X_n$ , the sample mean is denoted as

$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

\* Sample mean is also a random variable.

$$E[\bar{X}] = \frac{E[X_1] + \dots + E[X_n]}{n} = \frac{n E[X]}{n} = E[X] \quad \text{(since } X_i \text{'s are iid.)}$$

$$\begin{aligned} * \text{Var}[\bar{X}] &= \frac{\text{Var}(X_1 + X_2 + \dots + X_n)}{n^2} \quad \text{(since } \text{Var}(aX) = a^2 \text{Var}(X)) \\ &= \frac{\text{Var}(X_1) + \dots + \text{Var}(X_n)}{n^2} = \frac{n \cdot \text{Var}(X)}{n^2} \quad \text{(since } X_i \text{'s are iid)} \\ &= \frac{\text{Var}(X)}{n} \end{aligned}$$

(2)

## Weak Law of Large Numbers (WLLN)

Let  $X_1, X_2, \dots, X_n$  be iid r.v.'s with a finite expected value  $E[X_i] = \mu < \infty$ , then for any  $\epsilon > 0$

$$\lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \epsilon) = 0$$

Proof of WLLN :

$\bar{X}$  is a r.v. with mean  $\mu$   
(sample mean)

Variance  $\frac{\sigma^2}{n}$ .

Applying  
Chebyshev's  
inequality

$$P(|\bar{X} - \mu| \geq \epsilon) \leq \frac{\text{Var}(\bar{X})}{\epsilon^2}$$

$$= \frac{\sigma^2}{n \epsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty.$$

$$\Rightarrow \lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \epsilon) = 0$$

## Central Limit Theorem (CLT).

Let  $X_1, X_2, \dots, X_n$  be iid r.v.'s with mean  $\mu < \infty$ ,  
variance  $0 < \sigma^2 < \infty$ , Then, the random variable

$$Z_n = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + X_2 + \dots + X_n - n\mu}{\sqrt{n} \cdot \sigma}$$

Converges in distribution to a standard Gaussian r.v. as  $n \rightarrow \infty$ .

$$\lim_{n \rightarrow \infty} \underbrace{P(Z_n \leq z)}_{\text{CDF of } Z_n} = \underbrace{\Phi(z)}_{\text{CDF of } N(0,1)} \quad \forall z$$

CLT applies no matter what is the underlying distrib. of  $X_i$ 's (ie they need NOT be Gaussian).

(3)

Eg: Suppose  $X_i$ 's are Bernoulli( $p$ ) ;  $E[X_i] = p = \mu$   
 $\text{Var}[X_i] = p(1-p)$   
 $\sigma^2$

$$Y_n = X_1 + X_2 + \dots + X_n$$

$$Y_n \sim \text{Binomial}(n, p)$$

$$Z_n = \frac{Y_n - n\mu}{\sqrt{n}\sigma} = \frac{Y_n - np}{\sqrt{np(1-p)}}$$

Say  $p = 1/3$

$$Z_1 = \frac{X_1 - p}{\sqrt{p(1-p)}}$$

PMF of  $Z_1$



$$Z_2 = \frac{X_1 + X_2 - 2p}{\sqrt{2p(1-p)}}$$



$$Z_3 = \frac{X_1 + X_2 + X_3 - 3p}{\sqrt{3p(1-p)}}$$



increasing  
 $n$

$$Z_{30} = \frac{X_1 + X_2 + \dots + X_{30} - 30p}{\sqrt{30p(1-p)}}$$



Starts  
looking like  
 $N(0, 1)$



Ex:  $X_i$ 's  $\sim \text{unif}[0,1]$

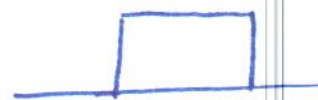
$$\mu = 1/2$$
$$\sigma^2 = 1/12$$

(4)

PDF of  $Z_1$

$$Z_n = \frac{X_1 + \dots + X_n - \frac{n}{2}}{\sqrt{n/12}}$$

$$Z_1 = \frac{X_1 - 1/2}{\sqrt{1/12}}$$



PDF of  $Z_2$

$$Z_2 = \frac{X_1 + X_2 - 1}{\sqrt{2/12}}$$

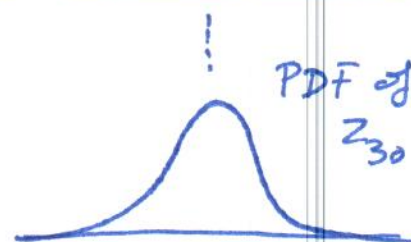


PDF of  $Z_3$

$$Z_3 = \frac{X_1 + X_2 + X_3 - 3/2}{\sqrt{3/12}}$$



$$\vdots$$
$$Z_{30} = \frac{\sum_{i=1}^{30} X_i - \frac{30}{2}}{\sqrt{30/12}}$$



PDF of  $Z_{30}$

Why is CLT useful?

- In many applications, a desired r.v. is often the sum of a large number of independent r.v.'s. & we can invoke CLT to justify the use of Normal approximation

- Examples:

- Lab. measurement errors modeled as normal
- Comms/Signal processing  $\rightarrow$  Gaussian noise is the most frequently used model.
- Finance, change/fluctuations in stocks/assets  $\sim$  Normal.
- Random sampling from a population to obtain statistical knowledge.

- CLT also significantly simplifies analysis & calculations.

How to apply CLT :

1. Write the r.v.  $Y$  (of interest) as a sum of  $n$  iid r.v.'s ( $X_i$ 's)
- $$Y = X_1 + \dots + X_n$$

2. Find  $E[Y] = n\mu$   $(\mu = E[X_i]; \sigma^2 = \text{Var}[X_i])$   
 $\text{Var}[Y] = n\sigma^2$

3. From CLT, we know  $\frac{Y - E[Y]}{\sqrt{\text{Var}[Y]}} = \frac{Y - n\mu}{\sqrt{n}\sigma} \sim \sqrt{V}(0, 1)$

To find  $P(y_1 \leq Y \leq y_2)$ , we can write

$$P(y_1 \leq Y \leq y_2) = P\left(\frac{y_1 - n\mu}{\sqrt{n}\sigma} \leq \frac{Y - n\mu}{\sqrt{n}\sigma} \leq \frac{y_2 - n\mu}{\sqrt{n}\sigma}\right)$$

$$\stackrel{\approx}{\underset{n}{\text{(for large)}}} \Phi\left(\frac{y_2 - n\mu}{\sqrt{n}\sigma}\right) - \Phi\left(\frac{y_1 - n\mu}{\sqrt{n}\sigma}\right)$$

Eg 1: In a communication system, a data packet has 1000 bits. Due to noise, each bit is received in error w. prob. 0.1. Assuming that errors occur independently, what is the prob. that there are more than 120 errors in a data packet?

Solution: Let  $X_i \rightarrow$  represent the error for  $i$ th bit

$$X_i = \begin{cases} 1 & \text{if } i\text{th bit is received with error (w.p.) } 0.1 \\ 0 & \text{error free (w.p.) } 0.9. \end{cases}$$

$$\Rightarrow X_i \sim \text{Bernoulli}(P)$$

$$(P = 0.1)$$



Let  $Y = X_1 + X_2 + \dots + X_n$

$\Rightarrow Y =$  Total number of bits received with error.

$$E[X_i] = \mu = p = 0.1$$

$$\text{Var}(X_i) = \sigma^2 = p(1-p) = 0.09$$

$$(n = 1000)$$

We are interested in

$$\begin{aligned} P(Y > 120) &= P\left(\frac{Y - n\mu}{\sqrt{n}\sigma} > \frac{120 - n\mu}{\sqrt{n}\sigma}\right) \\ &= P\left(\underbrace{\frac{Y - n\mu}{\sqrt{n}\sigma}}_{Z_n} > \frac{120 - 100}{\sqrt{90}}\right) \\ &\stackrel{\text{CLT (assuming } Z_n \sim N(0,1))}{\approx} 1 - \Phi\left(\frac{20}{\sqrt{90}}\right) = 0.0175 \end{aligned}$$

Eg 2: Packet transmission times on an internet link are iid with mean  $m$  & variance  $\sigma^2$ . Suppose that we transmit  $n$  packets & hence the total expected time is  $nm$ . Using CLT, approximate the probability that the transmission time for  $n$  packets is greater than twice the expected time.

Sol<sup>n</sup>:  $Y = X_1 + \dots + X_n$      $E[Y] = nm$

we are interested in

$$\begin{aligned} P(Y > 2E[Y]) &= P(Y > 2nm) \\ &= P\left(\frac{Y - nm}{\sqrt{n}\sigma} > \frac{2nm - nm}{\sqrt{n}\sigma}\right) \\ &= P\left(\frac{Y - nm}{\sqrt{n}\sigma} > \frac{nm}{\sqrt{n}\sigma}\right) = 1 - \Phi\left(\frac{m\sqrt{n}}{\sigma}\right) \end{aligned}$$

$$\Rightarrow \boxed{P(Y > 2E[Y]) \approx 1 - \Phi\left(\frac{m\sqrt{n}}{\sigma}\right)}$$