

## Final Exam - ECE 503 Fall 2017

- Date: Wednesday, December 13, 2017.
- Time: 10:30 am -12:30 pm
- Maximum Credit: 100 points

1. [10 points] A jet aircraft's autopilot has conditional probability  $1/3$  of failure given that it employs a faulty microprocessor chip. The autopilot has conditional probability of failure of  $1/10$  given that it employs a non-faulty chip. According to the chip manufacturer, the probability of a customer receiving a faulty chip is  $1/4$ .

(a) Find the probability of failure of autopilot.

(b) Given that an autopilot failure has occurred, find the conditional probability that a faulty chip was used.

Let  $A_F = \{\text{autopilot fails}\}$ ;  $C_F = \{\text{faulty chip was used}\}$ .

We are given:  $P(A_F | C_F) = \frac{1}{3}$ ;  $P(C_F) = \frac{1}{4}$ ;  $P(A_F | \bar{C}_F) = \frac{1}{10}$

$$\begin{aligned} \text{(a)} \quad P(A_F) &= P(C_F) \cdot P(A_F | C_F) + P(\bar{C}_F) \cdot P(A_F | \bar{C}_F) \\ &= \frac{1}{4} \times \frac{1}{3} + \frac{3}{4} \times \frac{1}{10} = \boxed{\frac{19}{120}} \end{aligned}$$

$$\begin{aligned} \text{(b)} \quad P(C_F | A_F) &= \frac{P(A_F, C_F)}{P(A_F)} = \frac{P(C_F) \cdot P(A_F | C_F)}{P(A_F)} \\ &= \frac{(1/4) \times (1/3)}{19/120} = \frac{120}{12 \times 19} = \boxed{\frac{10}{19}} \end{aligned}$$

2. [20 points]

- (a) Let  $X$  and  $Y$  be independent exponential random variables (both distributed as  $\exp(\lambda)$ ). Find  $E[\max(X, Y)]$ .
- (b) Let  $f(x)$  denote a probability density function. We define a sequence of random variables  $X_n$  which have the following density (PDF):

$$f_n(x) = n f(nx).$$

Find the CDF of  $X_n$  and its limiting behavior as  $n \rightarrow \infty$ . Does this sequence of random variables converge in distribution? If yes, then which random variable does this sequence converge to?

(a) Let  $Z = \max(X, Y)$  CDF of  $Z$ :  $F_Z(z) = P(Z \leq z)$   
 $= P(\max(X, Y) \leq z)$

$$\begin{aligned} F_Z(z) &= P(\max(X, Y) \leq z) = P(X \leq z, Y \leq z) \\ &= P(X \leq z) \cdot P(Y \leq z) \\ &= F_X(z) \cdot F_Y(z) \end{aligned}$$

$$\begin{aligned} \Rightarrow f_Z(z) &= f_X(z) F_Y(z) + f_Y(z) F_X(z) = 2 f_X(z) F_Y(z) \\ &= 2 \lambda e^{-\lambda z} (1 - e^{-\lambda z}) \end{aligned}$$

$$\begin{aligned} E[Z] &= \int_0^{\infty} z f_Z(z) dz = 2 \int_0^{\infty} \lambda z e^{-\lambda z} (1 - e^{-\lambda z}) dz = \frac{2}{\lambda} - \frac{1}{2\lambda} \quad z \geq 0. \\ &= \boxed{\frac{3}{2\lambda}} \end{aligned}$$

(b).  $F_n(x) = P(X_n \leq x) = \int_{-\infty}^x n f(nt) dt = \int_{-\infty}^{nx} f(y) dy = F(nx)$

$$\Rightarrow F_n(x) = F(nx)$$

for  $x > 0$  as  $n \rightarrow \infty$   $F_n(x) \rightarrow F(\infty) = 1$   
 for  $x < 0$  as  $n \rightarrow \infty$   $F_n(x) \rightarrow F(-\infty) = 0$

(for  $x = 0$ ,  $F_n(0) = F(0)$ ).

$$\Rightarrow X_n \rightarrow 0 \text{ in}$$

distribution as  $n \rightarrow \infty$ .

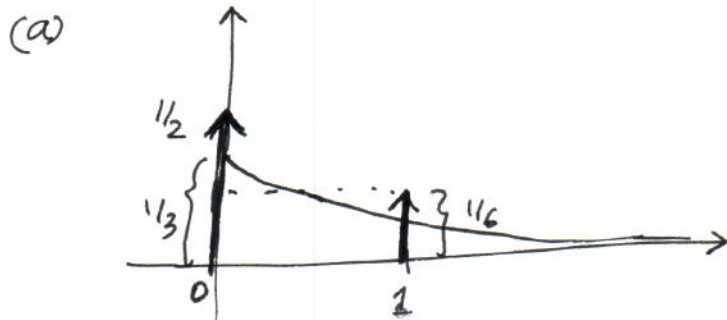


3. [15 points] A mixed random variable  $X$  has the following density:

$$f_X(x) = \frac{1}{3}e^{-x}u(x) + \frac{1}{2}\delta(x) + \frac{1}{6}\delta(x-1)$$

where  $u(x)$  is the unit step function, and  $\delta(x)$  is the Dirac delta function.

- Sketch the density function  $f_X(x)$ .
- Compute  $P(X=0)$  and  $P(X=1)$ .
- Compute  $P(0 < X < 1)$  and  $P(X > 1)$ .
- Compute  $P(0 \leq X \leq 1)$  and  $P(X \geq 1)$ .
- Compute  $E[X]$ .



(b)  $P(X=0) = \boxed{\frac{1}{2}}$  ;  $P(X=1) = \boxed{\frac{1}{6}}$

(c)  $P(0 < X < 1) = \int_{0^+}^1 f_X(x) dx = \int_{0^+}^1 \left( \frac{1}{3}e^{-x}u(x) + \frac{1}{3}\delta(x) + \frac{\delta(x-1)}{6} \right) dx$   
 $= \frac{1}{3} \int_0^1 e^{-x} dx = \boxed{\frac{1 - 1/e}{3}}$

$P(X > 1) = \int_{1^+}^{\infty} f(t) dt = \frac{1}{3} \int_1^{\infty} e^{-x} dx = \boxed{\frac{1/e}{3}} = \frac{1}{3e}$

(d)  $P(0 \leq X \leq 1) = P(X=0) + P(0 < X < 1) + P(X=1) = \frac{1}{2} + \frac{1 - 1/e}{3} + \frac{1}{6}$   
 $= \boxed{1 - \frac{1}{3e}}$

$P(X \geq 1) = P(X=1) + P(X > 1) = \boxed{\frac{1 + 2/e}{6}}$

(e)  $E[X] = \int_{-\infty}^{\infty} xf(x) dx = \int_0^{\infty} \frac{x}{3} e^{-x} dx + 0 \times P(X=0) + 1 \times P(X=1)$   
 $= \frac{1}{3} + 0 + \frac{1}{6} = \boxed{\frac{1}{2}}$



4. [20 points] Consider the random process

$$W(t) = X \cos(2\pi t) + Y \sin(2\pi t)$$

where  $X$  and  $Y$  are uncorrelated Gaussian random variables. Both  $X$  and  $Y$  have zero mean, and unit variance.

- Find the auto correlation function of the random process  $W(t)$ .
- Is  $W(t)$  a wide sense stationary random process?
- What is the average power of this random process?
- Find the probability  $P(W(1) + W(\frac{3}{4}) \leq 0)$ .

$$\begin{aligned} (a) \quad E[W(t)W(t+\tau)] &= E[(X \cos(2\pi t) + Y \sin(2\pi t))(X \cos(2\pi(t+\tau)) + Y \sin(2\pi(t+\tau)))] \\ &= E[X^2] \cdot \cos(2\pi t) \cos(2\pi(t+\tau)) + E[Y^2] \sin(2\pi t) \sin(2\pi(t+\tau)) \\ &\quad + \underbrace{E[XY]}_{=0} \times [ \dots ] \\ &= \frac{E[X^2]}{2} \cdot [\cos(2\pi\tau) + \cos(2\pi(2t+\tau))] \\ &\quad + \frac{E[Y^2]}{2} [\cos(2\pi\tau) - \cos(2\pi(2t+\tau))] \end{aligned}$$

$$E[X^2] = E[Y^2] = 1$$

$$= \frac{1}{2} \times \cos(2\pi\tau) \times 2 = \boxed{\cos(2\pi\tau)}$$

$$(b) \quad \boxed{\text{Yes, } W(t) \text{ is WSS}} \quad \left\{ E[W(t)] = 0, \quad \begin{matrix} R_X(z) \\ \downarrow \\ \text{only on time diff.} \end{matrix} \right\}$$

$$(c) \quad \text{Avg. Power} = R_W(0) = \cos(0) = \boxed{1}$$

$$(d) \quad P(W(1) + W(\frac{3}{4}) \leq 0)$$

$$\begin{aligned} W(1) + W(\frac{3}{4}) &= X \underbrace{\cos(2\pi)}_1 + Y \underbrace{\sin(2\pi)}_0 + X \underbrace{\cos(\frac{3\pi}{2})}_0 + Y \underbrace{\sin(\frac{3\pi}{2})}_{-1} \\ &= X - Y \end{aligned}$$

$$\begin{aligned} &= P(X - Y \leq 0) = P(Z \leq 0) = P\left(\frac{Z-0}{\sqrt{2}} \leq \frac{0-0}{\sqrt{2}}\right) = \boxed{\frac{1}{2}} \\ Z &= X - Y \sim N(0, 2) \end{aligned}$$

5. [20 points] Suppose  $X_n$  is a random sequence satisfying

$$X_n = cX_{n-1} + Z_{n-1}$$

where  $Z_1, Z_2, \dots$  is an i.i.d. sequence with  $E[Z_n] = 0$  and  $\text{Var}[Z_n] = \sigma^2$ , and  $c$  is a constant with  $|c| < 1$ . We also know that  $E[X_0] = 0$ , and  $\text{Var}[X_n] = \sigma^2/(1-c^2)$ . We make noisy observations as follows:

$$Y_{n-1} = dX_{n-1} + W_{n-1}$$

where  $W_1, W_2, \dots$  is a i.i.d. sequence of measurement noises, with  $E[W_n] = 0$  and  $\text{Var}[W_n] = \eta^2$ . We also know that  $W_n$ 's are independent of  $X_n$  and  $Z_n$ .

- (a) Find the optimal linear MMSE estimator  $\hat{X}_n$  of  $X_n$  using the noisy observation  $Y_{n-1}$ .  
 (b) Find the resulting mean square estimation error  $E[(X_n - \hat{X}_n)^2]$ .

$$(a) \quad \hat{X}_n = \hat{X}_n(Y_{n-1}) = \underset{\substack{\text{optimum} \\ \text{linear MMSE-Estimator}}}{\rho_{X_n, Y_{n-1}}} \left( \frac{\text{Var}(X_n)}{\text{Var}(Y_{n-1})} \right)^{1/2} (Y_{n-1} - E[Y_{n-1}]) + E[X_n]$$

$$\left. \begin{array}{l} E[Y_{n-1}] = 0 \\ E[X_n] = 0 \end{array} \right\} \Rightarrow \hat{X}_n = \left[ \rho_{X_n, Y_{n-1}} \left( \frac{\text{Var}(X_n)}{\text{Var}(Y_{n-1})} \right)^{1/2} \right] \cdot Y_{n-1}$$

$$\begin{aligned} X_n &= cX_{n-1} + Z_{n-1} \\ &= c(cX_{n-2} + Z_{n-2}) + Z_{n-1} \\ &= c^2X_{n-2} + cZ_{n-2} + Z_{n-1} \\ &\vdots \\ X_n &= c^n X_0 + \sum_{j=1}^n c^{j-1} Z_{n-j} \end{aligned}$$

$$\begin{aligned} \text{Var}(Y_{n-1}) &= d^2 \text{Var}(X_{n-1}) + \text{Var}(W_{n-1}) \\ &= \frac{d^2 \sigma^2}{1-c^2} + \eta^2 \\ \rho_{X_n, Y_{n-1}} &= \frac{E[X_n Y_{n-1}]}{\sqrt{\text{Var}(X_n) \text{Var}(Y_{n-1})}} = \frac{cd \text{Var}(X_{n-1})}{\sqrt{\text{Var}(X_n) \text{Var}(Y_{n-1})}} \end{aligned}$$

$$\begin{aligned} \text{Var}(X_n) &= c^{2n} \text{Var}(X_0) + \sum_{j=1}^n (c^2)^{j-1} \text{Var}[Z_{n-j}] \\ &= \frac{c^{2n} \sigma^2}{1-c^2} + \sigma^2 [1 + c^2 + \dots + (c^2)^{n-1}] \\ &= \frac{c^{2n} \sigma^2}{1-c^2} + \sigma^2 \frac{(1-c^{2n})}{1-c^2} = \frac{\sigma^2}{1-c^2} \end{aligned}$$

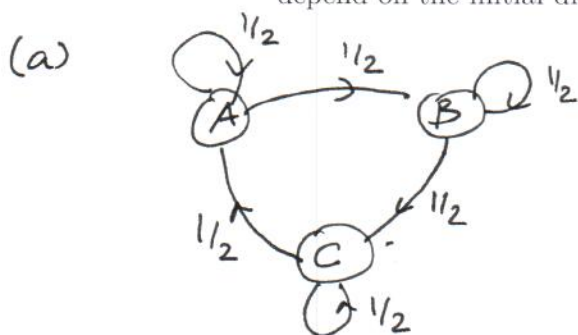
$$\Rightarrow \hat{X}_n = \left( \frac{c}{d} \right) \left( \frac{1}{1 + \beta^2(1-c^2)} \right) Y_{n-1}$$

where  $\beta^2 = \frac{\eta^2}{d^2 \sigma^2}$

6. [15 points] Consider the following Markov chain with three states  $S = \{A, B, C\}$ , which has the following one-step transition matrix:

$$P = \begin{bmatrix} 1/2 & 1/2 & 0 \\ 0 & 1/2 & 1/2 \\ 1/2 & 0 & 1/2 \end{bmatrix}$$

- (a) Draw the state transition diagram for this Markov chain.  
 (b) If you are given that  $P(X_1 = A) = 1/2$ ,  $P(X_1 = C) = 1/6$ , find the probability  $P(X_1 = B, X_2 = C, X_3 = A)$ .  
 (c) Is this MC irreducible and aperiodic?  
 (d) Find the limiting distribution of this Markov chain. Is this distribution unique? Does it depend on the initial distribution?



(b)  $P(X_1 = B, X_2 = C, X_3 = A)$

$$= \underbrace{P(X_1 = B)}_{\Downarrow} \cdot \underbrace{P(X_2 = C | X_1 = B)}_{\Downarrow} \cdot \underbrace{P(X_3 = A | X_2 = C)}_{\Downarrow}$$

$$= \left(1 - \frac{1}{2} - \frac{1}{6}\right) \cdot \frac{1}{2} \cdot \frac{1}{2}$$

$$= \frac{1}{3} \times \frac{1}{2} \times \frac{1}{2} = \frac{1}{12}$$

(c) Yes, MC is irreducible & aperiodic

(d) Unique limiting distrib  $\Rightarrow \pi = \pi P$

$\downarrow$   
 solving gives

$$\pi = \left[ \frac{1}{3} \quad \frac{1}{3} \quad \frac{1}{3} \right]$$

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