

MARKOV CHAINS

Discrete-Time Markov Chains (DTMC)

A discrete-time Markov chain $\{X_n, n=0,1,2,\dots\}$ is a discrete-time, discrete-valued random process such that given X_0, X_1, \dots, X_n , the next random variable X_{n+1} depends only on X_n through the probability

$$\begin{aligned} P(X_{n+1} = i_{n+1} \mid X_n = i_n, \dots, X_0 = i_0) \\ = P(X_{n+1} = i_{n+1} \mid X_n = i_n) \end{aligned}$$

In other words, given the sequence of values i_0, i_1, \dots, i_n , the conditional probability of what value X_{n+1} takes depends only on the value of i_n .

Example: The Random Walk is a Markov process

$$X_n = X_0 + W_1 + W_2 + \dots + W_{n-1} + W_n$$

$$\begin{aligned} P(X_6 = 11 \mid X_5 = 10, X_4 = 9, \dots, X_0 = 0) \\ = P(X_6 = 11 \mid X_5 = 10) \\ = p \end{aligned}$$

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Using the Markov Property, we can show that

$$\begin{aligned}
 & P(X_{n+m} = j_m, \dots, X_{n+1} = j_1 \mid X_n = i_n, \dots, X_0 = i_0) \\
 &= P(X_{n+m} = j_m \mid X_{n+m-1} = j_{m-1}) \\
 &\quad \times P(X_{n+m-1} = j_{m-1} \mid X_{n+m-2} = j_{m-2}) \\
 &\quad \times \dots \\
 &\quad \times P(X_{n+2} = j_2 \mid X_{n+1} = j_1) \\
 &\quad \times P(X_{n+1} = j_1 \mid X_n = i_n)
 \end{aligned}$$

To prove this claim,

$$P(\underbrace{X_{n+m} = j_m}_A, \underbrace{X_{n+m-1} = j_{m-1}, \dots, X_{n+1} = j_1}_B \mid \underbrace{X_n = i_n, \dots, X_0 = i_0}_C)$$

$$= P(A, B \mid C)$$

$$= P(A \mid B, C) \times P(B \mid C)$$

$$= \underbrace{P(X_{n+m} = j_m \mid X_{n+m-1} = j_{m-1}, \dots)}_{II} \times P(B \mid C)$$

$$P(X_{n+m} = j_m \mid X_{n+m-1} = j_{m-1}) \times \underbrace{P(B \mid C)}$$

↓
apply the
same trick
recursively
:

STATE SPACE and TRANSITION PROBABILITIES

The set of possible values that the random variables X_n can take is called the state space of the chain. We focus on the case in which the state space is either the set of integers or a subset of integers.

The conditional probabilities $P(X_{n+1} = j \mid X_n = i)$ are called transition probabilities. When the transition probabilities do not depend on time n , then the M.C. is said to be homogeneous (or have stationary transition probabilities). For a homogeneous M.C., we use the notation

$$P_{ij} = P(X_{n+1} = j \mid X_n = i).$$

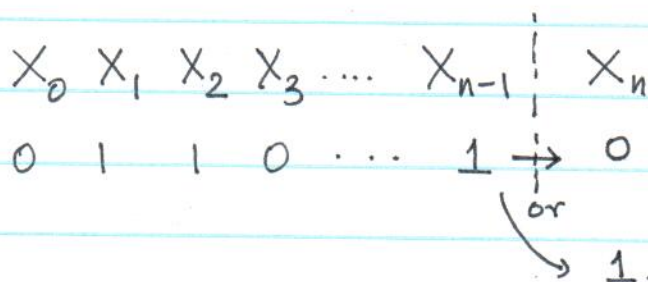
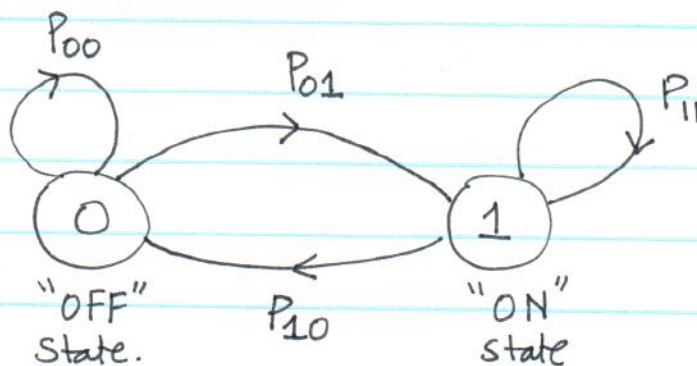
for the transition probabilities.

- * P_{ij} are also called the one-step transition probabilities since they are the prob. of going from state i to state j in one time step.
- * One of the most common ways to specify the transition probabilities is with a state transition diagram.

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Example

Two-state Markov chain.



State Space $\Rightarrow \{0, 1\} \Rightarrow$ set of values taken by $\{X_n\}$.

FACT

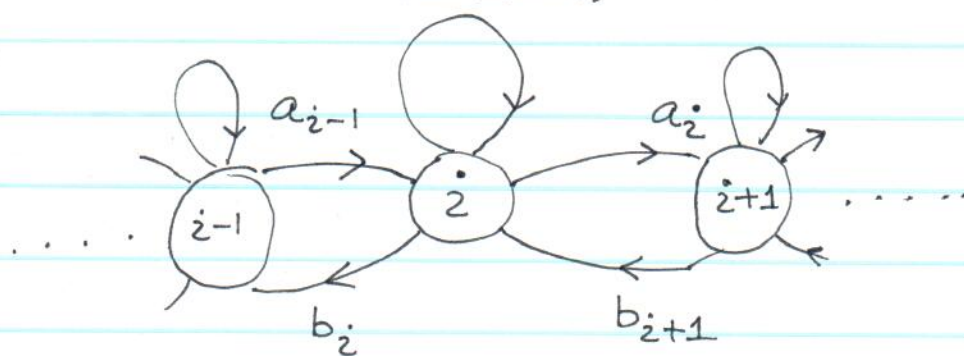
The sum of all the probabilities leaving a state must be 1

$$\sum_j P_{ij} = 1$$

WHY?

$$\sum_j P_{ij} = \sum_j P(X_{n+1}=j \mid X_n=i) = 1$$

Example: A general Random Walk.



$$P_{ij} = \begin{cases} b_i & \text{if } j = i-1 \\ 1-(a_i+b_i) & \text{if } j = i \\ a_i & \text{if } j = i+1 \\ 0 & \text{otherwise} \end{cases}$$

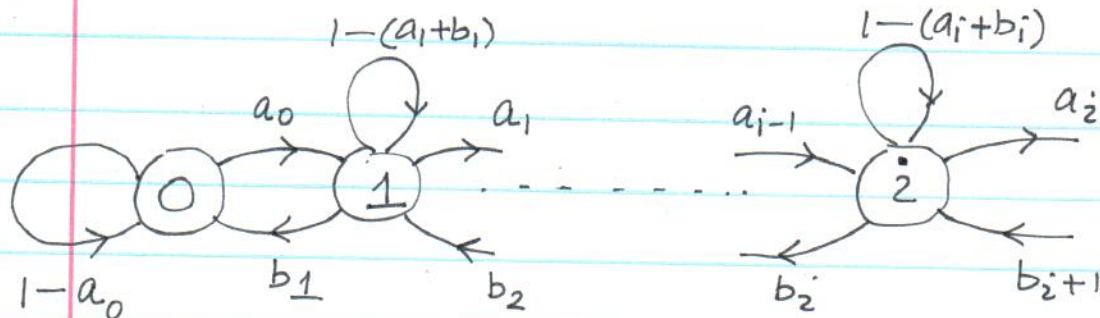
Transition Matrix

$$P = \begin{bmatrix} \dots & & & & \\ & \ddots & & & \\ & & \ddots & & \\ \dots & \dots & 0 & b_i & 1-(a_i+b_i) & a_i & 0 & \dots \\ & & & \ddots & \ddots & \ddots & \\ & & & & \ddots & \ddots & \end{bmatrix} \leftarrow i^{\text{th}} \text{ row}$$

infinite, tri-diagonal matrix.

(also known as a Birth-death process; eg population counts.)

Example: A Random Walk with a barrier at origin (6)



for $i=0$

$$P_{0j} = \begin{cases} 1-a_0 & j=0 \\ a_0 & j=1 \\ 0 & \text{otherwise.} \end{cases}$$

$$P = \begin{bmatrix} 1-a_0 & a_0 & 0 & 0 & 0 & \dots \\ b_1 & 1-(a_1+b_1) & a_1 & 0 & 0 & \dots \\ 0 & b_2 & 1-(a_2+b_2) & a_2 & 0 & \dots \\ 0 & 0 & b_3 & 1-(a_3+b_3) & a_3 & \dots \\ & & & \ddots & \ddots & \ddots \end{bmatrix}$$

semi-infinite, tri-diagonal matrix.

If $a_0 = 1 \rightarrow$ barrier is "reflecting"

If $a_0 = 0 \rightarrow$ barrier is "absorbing", i.e

once a chain hits an absorbing state,
it remains in that state from that time
onward..