

## Chapter-3

- Given my data matrix is

$$\bar{X}_{n \times p} = \begin{bmatrix} X_{11} & X_{12} & \dots & X_{1p} \\ X_{21} & X_{22} & \dots & X_{2p} \\ \vdots & & \ddots & \\ X_{n1} & X_{n2} & \dots & X_{np} \end{bmatrix}$$

$$E(S_n) = \frac{n-1}{n} \Sigma$$

$$= \Sigma - \frac{1}{n} \Sigma$$

So  $\frac{n}{n-1} S_n$  is an unbiased estimator of  $\Sigma$

Sample Mean Vector, Covariance and Correlation as matrix Operation:

Let  $\bar{X}_{n \times p}$  is a given data matrix, then

$$1 \rightarrow \text{Sample Mean Matrix: } \bar{X}_{p \times 1} = \frac{1}{n} \bar{X}^T \mathbf{1}$$

$$2 \rightarrow \text{Sample Variance matrix: } S = \frac{1}{n-1} \bar{X}' \left( \mathbf{I} - \frac{1}{n} \mathbf{1} \mathbf{1}' \right) \bar{X}$$

$$S_n = \frac{1}{n} \bar{X}' \left( \mathbf{I} - \frac{1}{n} \mathbf{1} \mathbf{1}' \right) \bar{X}$$

3.7 Sample Standard deviation Matrix:

$$D^{1/2} = \begin{bmatrix} \sqrt{S_{11}} & 0 & \dots & 0 \\ 0 & \sqrt{S_{22}} & \dots & 0 \\ \vdots & & \ddots & \vdots \\ 0 & 0 & \dots & \sqrt{S_{pp}} \end{bmatrix}$$

4.7 Sample correlation Matrix

$$R = D^{-1/2} S D^{-1/2}$$

5.7 Sample variance matrix, If  $R$  is given

$$S = D^{1/2} R D^{1/2}$$

•  $A^T A = A \implies A$  is Idempotent matrix.

Sample Values of Linear Combinations of Variable:

Consider a linear combination of vectors

$$\underline{C}^T \underline{X} = C_1 \underline{X}_1 + C_2 \underline{X}_2 + \dots + C_p \underline{X}_p$$

where  $\underline{X}_1, \underline{X}_2, \dots, \underline{X}_p$  are independent realizations from the random vector.

$$\text{Recall: } E(\underline{C}^T \underline{X}) = \underline{C}^T \underline{\mu} \text{ \& } \text{Var}(\underline{C}^T \underline{X}) = \underline{C}^T \Sigma \underline{C}$$

1) An estimate of the mean  $\underline{c}^T \underline{\mu}$  is  $\underline{c}^T \bar{\underline{x}}_{p,1}$

2) An estimate of the variance  $\underline{c}^T \Sigma \underline{c}$  is  $\underline{c}^T S_n \underline{c}$

3) An estimate for the population covariance matrix of  $\underline{b}^T \underline{x}$ ,  $\underline{c}^T \underline{x}$

which is  $\text{cov}(\underline{b}^T \underline{x}, \underline{c}^T \underline{x}) = \underline{b}^T \Sigma (\underline{c}^T)^T$

$$= \underline{b}^T \Sigma \underline{c} \text{ is } \underline{b}^T S_n \underline{c}.$$

## Chapter - 3 Idempotent matrix

From book

→ Random sampling implies

- 1) measurement taken on different items (or trials) are unrelated to one another and
- 2) The joint distribution of all  $p$  variables remains the same for all items

$$X = \begin{bmatrix} X_{11} & X_{12} & \dots & X_{1p} \\ X_{21} & X_{22} & \dots & X_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ X_{n1} & X_{n2} & \dots & X_{np} \end{bmatrix} = \begin{bmatrix} x_1' \\ x_2' \\ \vdots \\ x_n' \end{bmatrix} \begin{matrix} \leftarrow \text{1st (multivariate) observation} \\ \\ \\ \leftarrow n^{\text{th}} \text{ (multivariate) observation} \end{matrix}$$

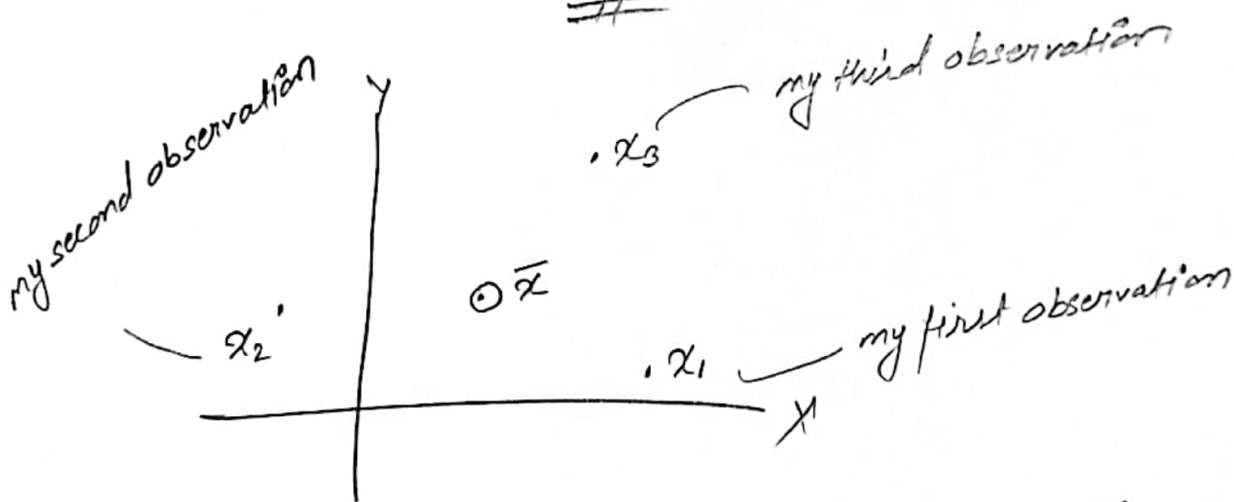
→ The rows of  $X$  represent  $n$  points in  $p$ -dimensional space (because  $n$  rows  $\& n$ )

→ If the points are regarded as solid spheres, the sample mean vector  $\bar{X}$ , is the center of balance.

→ Variability occurs in more than one direction and it is quantified by the sample variance-covariance matrix  $S_n$ .

→ A single numerical measure of variability is provided by the determinant of the sample variance-covariance matrix.

→  $\bar{x}$  is centre of balance  $\frac{1}{n} \sum x_i$   $\bar{x}$  is the point in  $n$  dimensional space  $\sum (x_i - \bar{x})^2$  (Hence no. of variable for  $\bar{x}$  should be same as data point)  $\downarrow + \downarrow + \downarrow + \downarrow$   
 $\equiv$



→ Elements of columns of the data matrix are the co-ordinates of vector (Another approach)

$$\begin{bmatrix} x & y \\ x_1 & y_1 \\ x_2 & y_2 \end{bmatrix}$$

→ If the two vectors are nearly perpendicular, the sample correlation will be approximately zero. If the two vectors are oriented in nearly opposite directions, the sample correlation will be close to -1.



→ Let  $X_1, X_2, \dots, X_n$  be a random sample from a joint distribution that has mean vector  $\mu$  and covariance matrix  $\Sigma$ . Then  $\bar{X}$  is an unbiased estimator of  $\mu$ , and its covariance matrix is  $\frac{1}{n} \Sigma$  that is

$$E(\bar{X}) = \mu$$

$$\text{cov}(\bar{X}) = \frac{1}{n} \Sigma \quad \text{--- Remember.}$$

For the covariance matrix  $S_n$

$$\begin{aligned} E(S_n) &= \frac{n-1}{n} \Sigma \\ &= \Sigma - \frac{1}{n} \Sigma \end{aligned}$$

Thus

$$E\left(\frac{n}{n-1} S_n\right) = \Sigma$$

So  $[n/(n-1)]S_n$  is an unbiased estimator of  $\Sigma$ , while

$$\begin{aligned} S_n \text{ is biased estimator with } \text{bias} &= E(S_n) - \Sigma \\ &= \Sigma - \frac{1}{n} \Sigma - \Sigma \\ &= -\left(\frac{1}{n}\right) \Sigma. \end{aligned}$$

→ Proof given  $X_1, X_2, \dots, X_n$  are random sample

$$\text{Now, } \bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\text{So } ① E(\bar{X}) = E\left(\frac{X_1 + X_2 + \dots + X_n}{n}\right)$$

$$= \frac{1}{n} (\underbrace{\mu + \mu + \dots + \mu}_{n \text{ times}})$$

$$= \mu$$

$$② \text{COV}(\bar{X}) = \frac{1}{n} \Sigma \text{ (How?)}$$

For this

$$(\bar{X} - \mu)(\bar{X} - \mu)'$$
$$= \left( \frac{1}{n} \sum_{j=1}^n X_j - \frac{n\mu}{n} \right) \left( \frac{1}{n} \sum_{l=1}^n X_l - \frac{n\mu}{n} \right)$$

$$= \left( \frac{1}{n} \sum_{j=1}^n (X_j - \mu) \right) \left( \frac{1}{n} \sum_{l=1}^n (X_l - \mu)' \right)$$

$$= \frac{1}{n^2} \sum_{j=1}^n \sum_{l=1}^n (X_j - \mu)(X_l - \mu)' \quad \text{--- *}$$

Now by def<sup>n</sup>

$$\text{COV}(\bar{X}) = E[(\bar{X} - \mu)(\bar{X} - \mu)']$$

using \*

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$$\text{COV}(\bar{X}) = E \left[ \frac{1}{n^2} \sum_{j=1}^n \sum_{l=1}^n (x_j - \mu)(x_l - \mu)' \right]$$

for  $j \neq l$ , all covariance are zero, these random variables are independent

For  $j=l$

$$\begin{aligned} \text{COV}(\bar{X}) &= E \left[ \frac{1}{n^2} \sum_{j=1}^n (x_j - \mu)(x_j - \mu)' \right] \\ &= \frac{1}{n^2} \left( \sum_{j=1}^n \underbrace{E(x_j - \mu)(x_j - \mu)'}_{\text{by def}^n \Sigma} \right) \\ &= \frac{1}{n^2} \left( \sum_{j=1}^n * \Sigma \right) \\ &= \frac{n \Sigma}{n^2} \\ &= \frac{\Sigma}{n} \end{aligned}$$

→ sample variance covariance matrix

$$S_{ik} = \frac{1}{n-1} \sum_{j=1}^n (x_{ji} - \bar{x}_i)(x_{jk} - \bar{x}_k)$$



→ Sometimes it is desirable to assign a single numerical value for the variation expressed by  $S$ . one choice for a value is the determinant of  $S$ , which reduces to the usual sample variance of a single characteristic when  $p=1$ . This determinant is called the generalized sample variance.

$$\therefore \text{Generalized sample variance} = |S|$$

Sample Mean, Covariance and Correlation as matrix Operations.

$$\rightarrow \bar{\underline{X}} = \frac{1}{n} \underline{X}' \underline{I} \quad (\text{sample mean vector})$$

$$\rightarrow (n-1) S = \underline{X}' (\underline{I} - \frac{1}{n} \underline{I} \underline{I}') \underline{X} \quad (\text{sample covariance matrix})$$

$$\rightarrow D^{1/2} = \begin{bmatrix} \sqrt{S_{11}} & 0 & \dots & 0 \\ 0 & \sqrt{S_{22}} & & \\ \vdots & & \ddots & \\ 0 & \dots & & \sqrt{S_{pp}} \end{bmatrix} \quad (\text{sample standard deviation matrix})$$

$$\rightarrow R = D^{-1/2} S D^{-1/2} \quad (\text{sample correlation matrix})$$

$$\Rightarrow \boxed{S = D^{1/2} R D^{1/2}}$$

→ sample mean of  $C'X = C'\bar{X}$

observation यात्रा (एक observation यात्रा)

From note:

$$\rightarrow \text{Cov}(\bar{X}) = \frac{1}{n} \Sigma_X$$

Watch CLT in home  
for Linear combination.

To prove  $A^T A = A$ , then  $A$  is idempotent matrix

LHS  $(A^T A)^T$