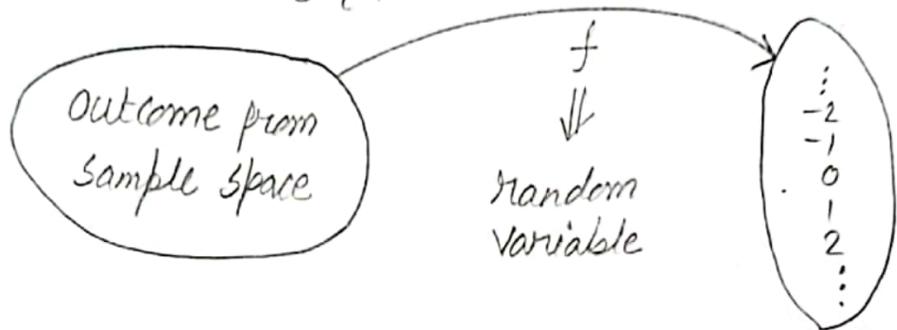


## # Random Variable

## Chapter - ②

→ यदि दो के से रोने वाली नूल जिसे ग्राम n.v x ग्राम की तरह नंबर ऑफ डोट्स ऑफ दो के द्वारा



## # Random Vector

Random vector ग्राम वेक्टर है जिसके अवस्थाएँ रॉम वारिएबल हैं।

Example:-

Biologically random vector & current random variables are weight ( $X_1$ ), Height ( $X_2$ ), blood-pressure ( $X_3$ )

## # Marginal parameters

$$\text{1) } E(X_i) = \begin{cases} \int_{-\infty}^{\infty} x_i f(x_i) dx_i & \text{If continuous} \\ \sum_{\text{all } x} x_i p_i(x_i) & \text{If discrete} \end{cases}$$

$$2) \text{Var}(X_i) = E[(X_i - \mu_i)^2] = \sigma_i^2 = \sigma_i^2$$

3) For two random variable  $X_i, X_j$

$$\text{Cov}(X_i, X_j) = E[(X_i - \mu_i)(X_j - \mu_j)] = \sigma_{ij}, i \neq j$$

50.  $\boxed{\text{Cov}(X_i, X_i) = \text{Var}(X_i)}$   $\rightarrow$  ~~3rd~~ ~~2nd~~ covariance ~~variance~~

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#  $\sigma_{ij} \rightarrow$  This denotes covariance between r.v.  $X_i$  &  $X_j$   
 $= \text{Cov}(X_i, X_j)$

$\sigma_{ii} \rightarrow$  Variance of the variable ( $X_i$ )

- $\text{Cov}(X_i, X_j)$  measures the linear association between the variable  $X_i, X_j$ .
- If there is any other relation than linear, covariance will not detect it.
- If  $X_i$  and  $X_j$  independent ~~of~~ ~~not~~,  $\text{Cov}(X_i, X_j) = 0$
- ~~If~~ pair of random variable  $X_i$  &  $X_j$  ~~of~~ ~~not~~ ~~linear~~ behaviour is described by their joint pdf.

1) population Mean Vector  $\mu = E(X) = \begin{bmatrix} E(X_1) \\ \vdots \\ E(X_p) \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \vdots \\ \mu_p \end{bmatrix}$

2) population variance covariance matrix

$$\Sigma = E[(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})^T]$$

$$= \begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} \\ \sigma_{21} & \sigma_{22} & \sigma_{23} \\ \sigma_{31} & \sigma_{32} & \sigma_{33} \end{bmatrix}$$

- P Variables ~~of~~ ~~not~~ ~~one~~ dataset ~~exists~~ ~~exists~~, there will be P variances (along diagonal) and  $\frac{P(P-1)}{2}$  distinct covariances.

### 3) Population Correlation matrix

$$\rho = \begin{bmatrix} 1 & \rho_{12} & \dots & \rho_{1p} \\ \rho_{21} & 1 & \dots & \rho_{2p} \\ \vdots & & & \\ \rho_{p1} & \rho_{p2} & \dots & 1 \end{bmatrix},$$

- correlation is defined as measure of linear relations between variables

$$\text{COR}(X_i, X_j) = \rho_{ij} = \frac{\text{Cov}(X_i, X_j)}{\sqrt{\text{Var}(X_i)} \sqrt{\text{Var}(X_j)}}$$

- Covariance of r.v.  $X_i$  and  $X_j$  after dividing by their respective standard deviation (square root of variance) we can get the correlation between random variable  $X_i$  and  $X_j$
- So If we are given variance covariance matrix  $\Sigma$ , can can easily find correlation matrix as

$$\text{If } \Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} \\ \sigma_{21} & \sigma_{22} & \sigma_{23} \\ \sigma_{31} & \sigma_{32} & \sigma_{33} \end{bmatrix}$$

$$\text{Then } \rho = \begin{bmatrix} \frac{\sigma_{11}}{\sqrt{\sigma_{11}} \sqrt{\sigma_{11}}} & \frac{\sigma_{12}}{\sqrt{\sigma_{11}} \sqrt{\sigma_{22}}} & \frac{\sigma_{13}}{\sqrt{\sigma_{11}} \sqrt{\sigma_{33}}} \\ \frac{\sigma_{21}}{\sqrt{\sigma_{22}} \sqrt{\sigma_{11}}} & \frac{\sigma_{22}}{\sqrt{\sigma_{22}} \sqrt{\sigma_{22}}} & \frac{\sigma_{23}}{\sqrt{\sigma_{22}} \sqrt{\sigma_{33}}} \\ \frac{\sigma_{31}}{\sqrt{\sigma_{33}} \sqrt{\sigma_{11}}} & \frac{\sigma_{32}}{\sqrt{\sigma_{33}} \sqrt{\sigma_{22}}} & \frac{\sigma_{33}}{\sqrt{\sigma_{33}} \sqrt{\sigma_{33}}} \end{bmatrix}$$

$$\Rightarrow \rho = \begin{bmatrix} 1 & \rho_{12} & \rho_{13} \\ \rho_{21} & 1 & \rho_{23} \\ \rho_{31} & \rho_{32} & 1 \end{bmatrix}$$

→ Note:- This  $\rho$  will be Identity matrix when all random variables are independent.

### 3) Population Correlation matrix

$$\rho = \begin{bmatrix} 1 & \rho_{12} & \dots & \rho_{1P} \\ \rho_{21} & 1 & \dots & \rho_{2P} \\ \vdots & & & \\ \rho_{P1} & \rho_{P2} & \dots & 1 \end{bmatrix},$$

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$$\Rightarrow \rho = \begin{bmatrix} 1 & \rho_{12} & \rho_{13} \\ \rho_{21} & 1 & \rho_{23} \\ \rho_{31} & \rho_{32} & 1 \end{bmatrix}$$

→ Note:- This  $\rho$  will be Identity matrix when all random variables are independent.

## R #

### ↳ Standard deviation matrix

$$V^{1/2} = \begin{bmatrix} \sqrt{\sigma_{11}} & 0 & \dots & 0 \\ 0 & \sqrt{\sigma_{22}} & \dots & 0 \\ \vdots & \ddots & & \\ 0 & 0 & \dots & \sqrt{\sigma_{pp}} \end{bmatrix}$$

- The important Relation bet<sup>n</sup> correlation matrix, standard deviation matrix and covariance matrix is :

$$\Sigma = V^{-1/2} \sum V^{-1/2} \rightarrow \sum \text{corr} \Sigma \text{ corr}, \sum \text{corr} V^{1/2} \text{ or}$$

both side ~~are~~ multiply ~~it's~~, which makes sense also

$$\Rightarrow \boxed{\Sigma = (V^{1/2})^{-1} \sum (V^{1/2})^{-1}}$$

conversely

$$\boxed{\Sigma = (V^{1/2}) \Sigma (V^{1/2})}$$

### • V. imp note

~~distribution discrete & it's expected value is replaced by summation so need to multiply with corresponding probability also~~

$$\text{Eg:- } \sigma_{11} = E[(X_1 - \mu_1)^2] = \sum_{\text{all } x_i} (x_i - \mu_1)^2 p_i(x_i)$$

$$\text{or } \sigma_{12} = E[(X_1 - \mu_1)(X_2 - \mu_2)] = \sum_{\substack{\text{all pairs} \\ (x_1, x_2)}} (x_1 - \mu_1)(x_2 - \mu_2) p_{12}(x_1, x_2)$$

## # Random Matrix.

A random matrix is a matrix whose elements are random variables.

$$\underline{X}_{n \times p} = \{x_{ij}\} = \begin{bmatrix} x_{11} & x_{12} & \dots & x_{1p} \\ \vdots & & & \\ x_{n1} & x_{n2} & \dots & x_{np} \end{bmatrix}$$

- a)  $E(\underline{X}) = \{E(x_{ij})\}$  if each expectation exists.
- b) यदि  $\underline{X}$  and  $\underline{Y}$  are random matrix of same dimension and let  $\underline{A}$  and  $\underline{B}$  be two conformable matrix of constants

$$E(\underline{X} + \underline{Y}) = E(\underline{X}) + E(\underline{Y})$$

$$E(\underline{\underline{A}} \underline{X} \underline{\underline{B}}) = \underbrace{\underline{\underline{A}} E(\underline{X}) \underline{\underline{B}}}_{K \times P \quad P \times M \quad M \times L \quad K \times L}$$

## # Partitioning random Vectors :

स्थिलांक Large dataset दिइएको ह जीत, characteristics measured on individual could naturally fall into two or more groups.

Example

Consumption and Income.

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- To handle situation like these, where distinct groups of characteristics arise naturally, a useful approach is to consider these groups as subset of larger collection of characteristics.
  - If the total collection is represented by a  $P \times 1$  dimension random vector  $\tilde{X}_{P \times 1}$ . We can regard sub-set as component of  $\tilde{X}$  and can be dealt with by partitioning

$$\tilde{X} = \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_q \\ \hline X_{q+1} \\ X_{q+2} \\ \vdots \\ X_p \end{bmatrix} = \begin{bmatrix} q X_1 \\ (P-q) X_1 \end{bmatrix} = \begin{bmatrix} X^{(1)} \\ X^{(2)} \end{bmatrix}$$

a) population mean vector ( $\tilde{\mu}$ )

$$\tilde{\mu} = E(\tilde{X}) = \begin{bmatrix} E(X_1) \\ E(X_2) \\ \vdots \\ E(X_{q+1}) \\ \vdots \\ E(X_p) \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_{q+1} \\ \vdots \\ \mu_p \end{bmatrix} = \begin{bmatrix} \tilde{\mu}^{(1)} \\ \tilde{\mu}^{(2)} \end{bmatrix}$$

b) Population Variance Covariance Matrix

$$\Sigma_{P \times P} = E[(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})^T] \rightarrow \text{by defn}$$

$$= \begin{bmatrix} \underline{\Sigma}_{11} & \underline{\Sigma}_{12} \\ \underline{\Sigma}_{21} & \underline{\Sigma}_{22} \end{bmatrix}$$

$\underline{\Sigma}_{11}$   
 $q \times q$ 
  
 $\underline{\Sigma}_{12}$   
 $q \times (P-q)$ 
  
 $\underline{\Sigma}_{21}$   
 $(P-q) \times q$ 
  
 $\underline{\Sigma}_{22}$   
 $(P-q) \times (P-q)$

Where  $\underline{\Sigma}_{11}$  = Variance covariance matrix for  $X^{(1)}$

$\underline{\Sigma}_{22}$  = Variance covariance matrix for  $X^{(2)}$

$$\underline{\Sigma}_{12} = E[(\underline{X}^{(1)} - \underline{\mu}^{(1)})(\underline{X}^{(2)} - \underline{\mu}^{(2)})^T] = \underline{\Sigma}_{21}^T$$

## # Linear Combinations of Random Vectors

• यदि  $X_1$  कोई single r.v है with mean  $\mu_1$  & variance  $\sigma_{11}$  and  $c$  be any constant number. Then

a)  $E(cX_1) = c\mu_1$

b)  $\text{Var}(cX_1) = c^2 \text{Var}(X_1) = c^2 \sigma_{11}$

<sup>4</sup> Let  $X_1$  and  $X_2$  two random variable ~~follows~~ एवं शृंखला with their respective mean and variances  $\mu_1, \mu_2$  &  $\sigma_{11}, \sigma_{22}$  and their covariance as  $\text{cov}(X_1, X_2) = \sigma_{12}$

$$\begin{aligned} a) E(ax_1 + bx_2) &= aE(X_1) + bE(X_2) \\ &= a\mu_1 + b\mu_2. \end{aligned}$$

$$b) \text{cov}(ax_1, bx_2) = ab \text{cov}(X_1, X_2) = ab \sigma_{12}$$

$$\begin{aligned} c) \text{Var}(ax_1 + bx_2) &= a^2 \text{Var}(X_1) + b^2 \text{Var}(X_2) + 2ab \text{cov}(X_1, X_2) \\ &= a^2 \sigma_{11} + b^2 \sigma_{22} + 2ab \sigma_{12} \end{aligned}$$

# Note

$$\text{Var}(X_1 + X_2) = \text{Var}(X_1) + \text{Var}(X_2) + 2 \text{cov}(X_1, X_2)$$

# For Matrix Representation

Linear combination  $ax_1 + bx_2$  can be written as

$$ax_1 + bx_2 = [a \ b] \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \tilde{C}^T \tilde{X}$$

$$E(ax_1 + bx_2) = E \left( [a \ b] \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \right)$$

$$= [a \ b] \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}$$

$$= \underline{\zeta}^T \underline{\mu}$$

If we let  $\Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix}$ . Then

$$\text{Then } \text{Var}(ax_1 + bx_2) = \text{Var}(\underline{\zeta}^T \underline{x})$$

$$= \text{Var} \left( [a \ b] \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \right)$$

$$= \underline{\zeta}^T \Sigma (\underline{\zeta}^T)^T$$

$$= \underline{\zeta}^T \Sigma \underline{\zeta}$$

$$= [a \ b] \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix} \begin{bmatrix} a \\ b \end{bmatrix}$$

$$= a^2 \sigma_{11} + 2ab \sigma_{12} + b^2 \sigma_{22}$$

• We can extend this result for a linear combination of  $P$  variables

$$\underline{C}^T \underline{X} = C_1 X_1 + C_2 X_2 + \cdots + C_p X_p . \text{ Then}$$

$$a) E(\underline{C}^T \underline{X}) = \underline{C}^T \underline{\mu} \quad \underline{\mu} = E(\underline{X})$$

Keep in mind that Transpose ~~gives~~ now matrix ~~is~~, all others are column matrix (without transpose)

$$b) \text{Var}(\underline{C}^T \underline{X}) = \underline{C}^T \Sigma (\underline{C}^T)^T$$

$$= \underline{C}^T \Sigma \underline{C}$$

where  $\Sigma = \text{variance covariance matrix of } X$ .

• Even more generally, we can consider a linear combination of  $P$  random variables  $X_1, \dots, X_p$

$$Z_1 = C_{11} X_1 + C_{12} X_2 + \cdots + C_{1p} X_p$$

$$Z_2 = C_{21} X_1 + C_{22} X_2 + \cdots + C_{2p} X_p$$

$\vdots$

$$Z_q = C_{q1} X_1 + C_{q2} X_2 + \cdots + C_{qp} X_p . \text{ So that}$$

$$\underline{Z} = \begin{bmatrix} c_{11} & c_{12} & \dots & c_{1p} \\ c_{21} & c_{22} & \dots & c_{2p} \\ \vdots & \ddots & \vdots & \vdots \\ c_{q1} & c_{q2} & \dots & c_{qp} \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{bmatrix} = \underline{C} \underline{X}$$

a) population mean of  $\underline{Z}$

$$E(\underline{Z}) = E(\underline{C} \underline{X})$$

$$= \underline{C} E(\underline{X})$$

b) variance covariance of  $\underline{Z}$

$$\Sigma_{\underline{Z}} = \text{cov}(\underline{Z}) = \text{cov}(\underline{C} \underline{X})$$

$$= \underline{C} \Sigma_X \underline{C}^T$$

- ↳ Multivariate data analysis deals with datasets containing multiple variables that interact and influence one another.
- ↳ If each data is analyzed separately, we can miss key features and interesting patterns inherent in the multivariate data.
- ↳ Multivariate Data Analysis is Important
  - 1) Data reduction or structural simplification eg GPA
  - 2) Sorting and grouping. Eg bank consumer → 'good credit' & 'bad credit'
  - 3) Investigation of the dependence among variables. Eg Consultant hiring
  - 4) Prediction
  - 5) Hypothesis construction and testing.

↳ Sample mean

$$\bar{x}_k = \frac{1}{n} \sum_{j=1}^n x_{jk}, \quad k=1, 2, \dots, p$$

↳ Sample variance

$$s_k^2 = s_{kk} = \frac{1}{n} \sum_{j=1}^n (x_{jk} - \bar{x}_k)^2, \quad k=1, 2, \dots, p$$

↳ Sample standard deviation =  $\sqrt{s_{kk}}$

↳ The sample covariance measures the degree to which two variables in a sample change together (co-vary).

$$s_{ik} = \frac{1}{n} \sum_{j=1}^n (x_{ji} - \bar{x}_i)(x_{jk} - \bar{x}_k)$$

- If large value from one variable occurs with small values for the other variable, sample covariance will be negative
- If large values for one variable are observed in conjunction with large values for the other variable, and smaller values also occur together, then sample covariance will be positive.
- If there is no particular linear association between the values for the two variables, then sample covariance will be zero. We are just talking about linear relationship, there might be other relationship like non-linear
- The correlation coefficient measures the strength and direction of the linear relationship between two variables.

$$\rho_{ik} = \frac{s_{ik}}{\sqrt{s_{ii}} \sqrt{s_{kk}}} \quad \text{i.e. } s_{ik} = \frac{\text{cov}(X_i; X_k)}{\sqrt{\text{var}(X_i)} \sqrt{\text{var}(X_k)}}$$

- Signs of the sample correlation and the sample covariance are the same, but correlation is ordinarily easier to interpret because its magnitude is bounded.

Interpretation

$\rho = 0 \rightarrow$  No linear relationship between variable

$\rho < 0 \rightarrow$  -ve linear relationship b/w variable

$\rho > 0 \rightarrow$  +ve linear relationship b/w variable

$\rho \approx 0 \Rightarrow \rho$  close to zero  $\rightarrow$  weak linear relationship b/w the two variables.

$\rho \approx 1$  or  $\rho \approx -1$  i.e.  $\rho$  is close to 1 or -1  $\rightarrow$  very strong linear relationship between the two variables.

## Interpretation of the Covariance :

→ The positive value of the covariance ( $\approx 65.22$ ) indicates that as the hours of study increases for a student, then their corresponding test score tend to increase as well. This suggest a positive relationship between hours of study and test scores. Student who study more hours tend to achieve higher test score.

## Interpretation of the Correlation coefficient :

→ Correlation value of 0.8352 indicates that there is a strong positive linear relationship between hours of study and test score in the dataset

• It suggest that the two variables are closely aligned in a linear fashion, where an increase in one variable corresponds to a substantial increase in other.

## Matrix Algebra :-

→ The sum of two vectors emanating from the origin is the diagonal of the parallelogram formed with the two original vectors as adjacent sides.

→ Inner product of vector  $x$  and  $y$  is  $x'y = x_1y_1 + x_2y_2$

→ A pair of vectors  $x$  and  $y$  of the same dimension is said to be linearly dependent if there exist constants  $c_1$  and  $c_2$ , both not zero, such that  $c_1x + c_2y = 0$

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→ Linear dependence implies that at least one vector in the set  
can be written as linear combination of the other vectors

## Matrix multiplication Rule

→ When  $A$  is  $(n \times k)$  and  $B$  is  $(k \times p)$ , so that the number of elements in a row of  $A$  is the same as the number of elements in a column of  $B$ , we can form a matrix product  $AB$ .

$\overbrace{A}$   $\overbrace{\text{row}}$   $\overbrace{\text{matrix}}$   $\overbrace{\text{row}}$   $\overbrace{\text{of element}}$   $\overbrace{B}$  (i.e. no. of columns in matrix A) should be equal to number of element in the column of matrix  $B$  (i.e. no. of rows in matrix  $B$ )

## Symmetric Matrix

→ A square matrix is said to be symmetric if  $A = A^T$  or  $a_{ij} = a_{ji}$   $\forall i, j$ .

## Inverse matrix:-

The technical condition that an inverse exists is that the  $k$  columns  $a_1, a_2, \dots, a_k$  of  $A$  are linearly independent. That is,

The existence of  $A^{-1}$  is equivalent to

$$c_1 a_1 + c_2 a_2 + \dots + c_k a_k = 0 \text{ only if } c_1 = \dots = c_k = 0$$

→  $\overbrace{A}$   $\overbrace{\text{matrix}}$   $\overbrace{\text{inverse}}$  exist  $\overbrace{\text{if and only if}}$   $\overbrace{\text{columns}}$   $\overbrace{\text{linearly independent}}$

→ And linearly independent  $\overbrace{\text{gives}}$   $\overbrace{\text{if}}$  linear combination zero

only if all constant is zero ~~is zero~~

## Orthogonal matrix

→ A matrix  $Q$  is said to be Orthogonal If

$$QQ^T = Q^TQ = I \text{ or } Q^T = Q^{-1}$$

## Eigen Value and Eigen Vector :-

→ A square matrix  $A$  is said to have an eigenvalue  $\lambda$ , with corresponding eigenvector  $x \neq 0$ , if

$$Ax = \lambda x$$

Eigen vector  
Eigen value.

↳ eigen vector of  $A$  is also vector of same matrix if multiply it by some number same result can be obtained by just multiplying by any simple number to that vector.

→ If  $A$  be a  $K \times K$  square matrix. Then  $A$  has pairs of eigenvalues and normalized eigen vectors namely,

$$\lambda_1, e_1 \quad \lambda_2, e_2 \quad \dots \quad \lambda_K, e_K$$

→ Two Eigen vectors are unique unless two or more eigenvalues are equal.

→ Eigen value & Eigen vector condition, A square matrix  $A$  if

only if all constant is zero ~~(if any)~~

orthogonal matrix

→ A matrix  $Q$  is said to be orthogonal If

$$QQ^T = Q^TQ = I \text{ or } Q^T = Q^{-1}$$

Eigen Value and Eigen Vector :-

→ A square matrix  $A$  is said to have an eigenvalue  $\lambda$ , with corresponding eigenvector  $x \neq 0$ , if

$$Ax = \lambda x$$

eigen vector  
eigen value.

↳ eigen vector given current vector of state of matrix & multiply it by same result can be obtained by just multiplying by any simple number to that vector.

→ Let  $A$  be a  $K \times K$  square matrix. Then  $A$  has pairs of eigenvalues and normalized eigen vector namely,

$$\lambda_1, e_1 \quad \lambda_2, e_2 \quad \dots \quad \lambda_K, e_K$$

→ Two Eigenvector are unique unless two or more eigenvalues are equal.

→ Eigen value & Eigen vector relationship condition, A square matrix  $A$  is

→ If vector  $\vec{v}$  is multiplied by matrix  $A$  then result is equal to vector  $\vec{v}$  number of times multiplied by  $A$  result is  $\vec{v}$ , if vector  $\vec{v}$  is eigen vector of  $A$  and eigen value  $\lambda$  is eigen value of  $A$ , corresponding to that matrix.

## Positive Definite Matrix :-

→ A square matrix is positive definite if pre-multiplying and post-multiplying it by the same vector always give a positive number as a result, independently of how we choose the vector.

→ Positive Definite symmetric matrix have the property that all their eigenvalues are positive.

→ The spectral decomposition of a  $K \times K$  symmetric matrix  $A$  is given by

$$A = \lambda_1 e_1 e_1^T + \lambda_2 e_2 e_2^T + \dots + \lambda_K e_K e_K^T$$

→ The characteristic eqn of matrix  $A$  is  $|A - \lambda I| = 0$

→ To show  $A$  is positive definite matrix

$$A = \lambda_1 e_1 e_1^T + \lambda_2 e_2 e_2^T$$

$$x^T A x = \lambda_1 x^T e_1 e_1^T x + \lambda_2 x^T e_2 e_2^T x$$

$$\text{put } y_1 = x^T e_1 + y_2 = x^T e_2$$

The behavior of any pair of random variable such as  $X_i$  and  $X_k$  is described by their joint probability function and the measure of the linear association between them is provided by the covariance.

- Joint probability can be written as the product of individual probability, then the random variable are said to be independent.
- The independent condition become  $f_{XK}(x_i, x_k) = f_i(x_i) f_k(x_k)$  for all pair  $(x_i, x_k)$ .
- If  $X_i$  and  $X_k$  are independent  $\Rightarrow \text{Cov}(X_i, X_k) = 0$   
but in general  $\text{Cov}(X_i, X_k) = 0$  does not imply  $X_i$  and  $X_k$  are independent
- Independent ~~ent~~ <sup>if</sup> covariance surely zero ~~is~~ but covariance zero ~~is~~ ~~not~~ independent ~~ent~~ ~~is~~ ~~not~~ always
- $p$  distinct variance &  $\frac{p(p-1)}{2}$  distinct covariance in  $P \times P$  variance covariance matrix.
- $\Sigma = E[(X - \mu)(X - \mu)^T]$

→ We note that the computation of mean, variances and covariances for discrete random variable involves summation, while analogous computations for continuous random variable involve integration

→ The multivariate normal distribution is completely specified once the mean vector  $\mu$  and variance-covariance matrix  $\Sigma$  are given.

$$\rightarrow S_{ik} = \frac{\sigma_{ik}}{\sqrt{\sigma_{ii}} \sqrt{\sigma_{kk}}} \quad \rightarrow \text{population correlation coefficient}$$

→ The correlation coefficient measures the amount of linear association b/w random variable  $X_i$  and  $X_k$ .

→  $S_{ik}$  fraction  $\sum \text{factors} \cancel{\sigma_{ii} \sigma_{kk}}$  (becoz  $\Sigma$  given  $\sqrt{\Sigma} V^{1/2}$  easily calculate  $\sqrt{\sigma_{ii}}$  from that) but  $\sum \cancel{\text{factors}} V^{1/2}$  and  $S$  both  $\cancel{\sigma_{ii} \sigma_{kk}}$ )

Partitioning the Covariance Matrix:

→  $\Sigma_{12}$  is a covariance matrix of partition  $X^{(1)}$  and  $X^{(2)}$  and this matrix is not necessarily symmetric or even square

$$\rightarrow \text{Note : } \Sigma_{12} = \Sigma_{21}^T$$

↳ Covariance matrix of  $X^{(1)}$  is  $\Sigma_{11}$

↳ Covariance matrix of  $X^{(2)}$  is  $\Sigma_{22}$ .

$$\hookrightarrow \text{And } \text{Cov}(X^{(1)}, X^{(2)}) = \Sigma_{12}$$

$$\hookrightarrow \text{Var}(cX_1) = c^2 \text{Var}(X_1) = c^2 \sigma_{11}$$

$$\hookrightarrow \text{Cov}(aX_1, bX_2) = ab \text{Cov}(X_1, X_2) = ab \sigma_{12}$$

$$\begin{aligned} \hookrightarrow \text{Var}(aX_1 + bX_2) &= a^2 \text{Var}(X_1) + b^2 \text{Var}(X_2) + 2ab \text{Cov}(X_1, X_2) \\ &= a^2 \sigma_{11} + b^2 \sigma_{22} + 2ab \sigma_{12} \end{aligned}$$

$$\hookrightarrow \text{Var}(C' \tilde{X}) = C' \tilde{\Sigma} C$$

$$\hookrightarrow \Sigma_Z = \begin{bmatrix} \sigma_{11} - 2\sigma_{12} + \sigma_{22} & \sigma_{11} - \sigma_{22} \\ \sigma_{11} - \sigma_{22} & \sigma_{11} + 2\sigma_{12} + \sigma_{22} \end{bmatrix} \quad \begin{array}{l} \text{for } Z = \begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix} \\ Z_1 = X_1 - X_2 \\ Z_2 = X_1 + X_2 \end{array}$$

↳ Note:-

If  $\sigma_{11} = \sigma_{22}$  that is, if  $X_1$  and  $X_2$  have equal variances

→ The off-diagonal terms in  $\Sigma_Z$  vanish. This demonstrates the well-known result that sum ( $X_1 + X_2$ ) and difference ( $X_1 - X_2$ ) of two random variable with same variance are uncorrelated.

FROM Class Note:-

→ Multivariate normal distribution at case  $\text{E}[X_1 X_2] = \text{cov}(X_1, X_2)$   
 $= 0 \iff X_1 \text{ and } X_2 \text{ independent hold from both sides.}$   
Otherwise independent ~~gives covariance zero & vice versa~~ but  
covariance zero & ~~gives~~ independent gives ~~zero~~

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\* Chapter 2, THE END \*

(Q) Find  $\mu_z$  and  $\Sigma_z$  for  $\underline{z} = \begin{bmatrix} z_1 \\ z_2 \end{bmatrix}$  If

$$z_1 = x_1 - x_2$$

$$z_2 = x_1 + x_2, \text{ Also } \underline{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}, \underline{\mu}_x = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$$

$$\Sigma_x = \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix}$$

Solution

Here we have

$$\underline{z} = \begin{bmatrix} z_1 \\ z_2 \end{bmatrix}$$

$$= \begin{bmatrix} 1 & -1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$$

$$= \bar{\Sigma} \underline{x}$$

$$\text{So, } \mu_z = E(\underline{z}) = E(\bar{\Sigma} \underline{x})$$

$$= \bar{\Sigma} E(\underline{x})$$

$$= \bar{\Sigma} \underline{\mu}_x$$

$$= \begin{bmatrix} 1 & -1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}$$

$$= \begin{bmatrix} \mu_1 - \mu_2 \\ \mu_1 + \mu_2 \end{bmatrix}$$

$$\text{and } \text{Var}(Z) = \sum_z$$

$$= \text{Var}(\bar{Z})$$

$$= \bar{C} \sum_X \bar{C}^T$$

$$= \begin{bmatrix} 1 & -1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix} \begin{bmatrix} 1 & 1 \\ -1 & 1 \end{bmatrix}$$

$$= \begin{bmatrix} \sigma_{11} - \sigma_{21} & \sigma_{12} - \sigma_{22} \\ \sigma_{11} + \sigma_{21} & \sigma_{12} + \sigma_{22} \end{bmatrix} \begin{bmatrix} 1 & 1 \\ -1 & 1 \end{bmatrix}$$

$$= \begin{bmatrix} \sigma_{11} - \sigma_{21} - \sigma_{12} + \sigma_{22} & \sigma_{11} - \sigma_{21} + \sigma_{12} - \sigma_{22} \\ \sigma_{11} + \sigma_{21} - \sigma_{12} - \sigma_{22} & \sigma_{11} + \sigma_{21} + \sigma_{12} + \sigma_{22} \end{bmatrix}$$

$$= \begin{bmatrix} \sigma_{11} - 2\sigma_{12} + \sigma_{22} & \sigma_{11} - \sigma_{22} \\ \sigma_{11} - \sigma_{22} & \sigma_{11} + 2\sigma_{12} + \sigma_{22} \end{bmatrix}$$

## Computing the Covariance Matrix

Find the covariance matrix for two random variables  $X_1$  and  $X_2$ , when their joint pmf is  $P_{12}(x_1, x_2)$ , represented by the entries in the body of the following table:

$x_2 \backslash x_1$	0	1	$P_1(x_1)$
-1	0.24	0.06	0.3
0	0.16	0.14	0.3
1	0.40	0.00	0.4
$P_2(x_2)$	0.8	0.2	1

Solu

$$\begin{aligned}
 E(X_1) &= \sum_{\text{all } x_1} x_1 P_1(x_1) \\
 &= (-1)(0.3) + (0)(0.3) + (1)(0.4) \\
 &= 0.1
 \end{aligned}$$

$$\begin{aligned}
 E(X_2) &= \sum_{\text{all } x_2} x_2 P_2(x_2) \\
 &= (0)(0.8) + (1)(0.2) \\
 &= 0.2
 \end{aligned}$$

Thus  $E(X_1) = \mu_1 = 0.1$  &  $E(X_2) = \mu_2 = 0.2$

$$\begin{aligned}
 \sigma_{11} &= E[(x_1 - \mu_1)^2] \\
 &= \sum_{\text{all } x_1} (x_1 - 0.1)^2 p_1(x_1) \\
 &= (-1 - 0.1)^2 (0.3) + (0 - 0.1)^2 (0.3) + (1 - 0.1)^2 (0.4) \\
 &= 0.69
 \end{aligned}$$

$$\begin{aligned}
 \sigma_{22} &= E[(x_2 - \mu_2)^2] \\
 &= \sum_{\text{all } x_2} (x_2 - 0.2)^2 p_2(x_2) \\
 &= (0 - 0.2)^2 (0.8) + (1 - 0.2)^2 (0.2) \\
 &= 0.16
 \end{aligned}$$

$$\begin{aligned}
 \sigma_{12} &= \sigma_{21} = E[(x_1 - \mu_1)(x_2 - \mu_2)] \\
 &= \sum_{\substack{\text{all pairs} \\ (x_1, x_2)}} (x_1 - 0.1)(x_2 - 0.2) p_{12}(x_1, x_2) \\
 &= (-1 - 0.1)(0 - 0.2)(0.24) + (-1 - 0.1)(1 - 0.2)(0.06) \\
 &\quad + (0 - 0.1)(0 - 0.2)(0.16) + (0 - 0.1)(1 - 0.2)(0.14) \\
 &\quad + (1 - 0.1)(0 - 0.2)(0.4) + (1 - 0.1)(1 - 0.2)(0.00) \\
 &= -0.08
 \end{aligned}$$

$$50, \quad \tilde{\mu} = E(X) = \begin{bmatrix} E(X_1) \\ E(X_2) \end{bmatrix} = \begin{bmatrix} 0.1 \\ 0.2 \end{bmatrix}$$

and

$$\Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix}$$
$$= \begin{bmatrix} 0.69 & -0.08 \\ -0.08 & 0.16 \end{bmatrix}$$

Q.) Computing expected values for discrete random variable

Suppose  $p=2$  and  $n=1$ , and consider the random vector  $\mathbf{X}' = [X_1, X_2]$   
 Let the discrete random variable  $X_1$  have the following probability function

$X_1$	-1	0	1
$P_1(X_1)$	0.3	0.3	0.4

Then

$$\begin{aligned} E(X_1) &= \sum_{\text{all } X_1} x_1 P_1(x_1) \\ &= (-1)(0.3) + (0)(0.3) + (1)(0.4) \\ &= 0.1 \end{aligned}$$

Similarly, Let the discrete random variable  $X_2$  has the probability function

$X_2$	0	1
$P_2(X_2)$	0.8	0.2

Then

$$\begin{aligned} E(X_2) &= \sum_{\text{all } X_2} x_2 P_2(x_2) \\ &= (0)(0.8) + (1)(0.2) \\ &= 0.2 \end{aligned}$$

Thus

$$E(\mathbf{X}) = \begin{bmatrix} E(X_1) \\ E(X_2) \end{bmatrix} = \begin{bmatrix} 0.1 \\ 0.2 \end{bmatrix}$$