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Topic 5: Two-Way Cluster-Robust (TWCR) Standard Errors

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Key points: The validity of Two-Way Cluster-Robust (TWCR) standard errors

Disclaimer: This note is compiled by Sai Zhang.

5.1 One-Way Clustering

First, consider the case of one-way clustering. The linear model with one-way clustering

$$y_{ig} = \mathbf{x}_{ig}\boldsymbol{\beta} + u_{ig}$$

where i denotes the ith of the N individuals in the sample, j denotes the gth of the G clusters, assume that

- $\mathbb{E}\left[u_{ig} \mid \mathbf{x}_{ig}\right] = 0$
- error independence across clusters: for $i \neq j$

$$\mathbb{E}\left[u_{ig}u_{jg'}\mid\mathbf{x}_{ig},\mathbf{x}_{jg'}\right]=0\tag{5.1}$$

unless g = g', that is, errors for individuals within the same cluster may be correlated.

Grouping observations by cluster, get

$$\mathbf{y}_{g} = \mathbf{X}_{g}\boldsymbol{\beta} + \mathbf{u}$$

where \mathbf{X}_g has dimension $N_g \times K$ and \mathbf{y}_g has dimension $N_g \times 1$, with N_g observations in cluster g. Stacking over cluster, get the matrix form of the model

$$\mathbf{v} = \mathbf{X}\boldsymbol{\beta} + \mathbf{u}$$

with \mathbf{y} , \mathbf{u} being $N \times 1$ vectors, \mathbf{X} being an $N \times K$ matrix. OLS estimator gives

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} = \left(\sum_{g=1}^{G} \mathbf{X}'_{g}\mathbf{X}_{g}\right)^{-1}\sum_{g=1}^{G} \mathbf{X}'_{g}\mathbf{y}_{g}$$
(5.2)

then, by CLT, we have that $\sqrt{G}(\hat{\beta} - \beta) \xrightarrow{d} \mathcal{N}(0, \Sigma)$ where the variance matrix of the limit normal distribution Σ is

$$\left(\lim_{G\to\infty}\frac{1}{G}\sum_{g=1}^{G}\mathbf{E}\left[\mathbf{X}_{g}'\mathbf{X}_{g}\right]\right)^{-1}\left(\lim_{G\to\infty}\frac{1}{G}\sum_{g=1}^{G}\mathbf{E}\left[\mathbf{X}_{g}'\mathbf{u}_{g}'\mathbf{u}_{g}\mathbf{X}_{g}\right]\right)\times\left(\lim_{G\to\infty}\frac{1}{G}\sum_{g=1}^{G}\mathbf{E}\left[\mathbf{X}_{g}'\mathbf{X}_{g}\right]\right)^{-1}$$
(5.3)

If the primary source of clustering is due to group-level common shocks, a useful approximation is that for the jth regressor, the default OLS variance estimate based on $s^2(\mathbf{X}'\mathbf{X})^{-1}$ should be inflated by $\tau_j \simeq 1 + \rho_{x_j}\rho_u\left(\overline{N}_g - 1\right)$, where

• *s* is the estimated standard deviation of the error

- ρ_{x_i} is a measure of within-cluster correlation of x_j
- ρ_u is the within-cluster error correlation
- \overline{N}_g is the average cluster size

It's easy to see the τ_j can be large even with small ρ_u (Kloek, 1981; Scott and Holt, 1982; Moulton, 1990). If assume the model for the cluster error variance matrices $\Omega_g = \mathbb{V}\left[\mathbf{u}_g \mid \mathbf{X}_g\right] = \mathbb{E}\left[\mathbf{u}_g\mathbf{u}_g' \mid \mathbf{X}_g\right]$, and there is a consistent estimate $\hat{\Omega}_g$ of Ω_g , we can estimate $\mathbb{E}\left[\mathbf{X}_g'\mathbf{u}_g\mathbf{u}_g'\mathbf{X}_g\right] = \mathbb{E}\left[\mathbf{X}_g'\Omega_g\mathbf{X}_g\right]$ via GLS.

Cluster-robust variance matrix estimate consider

$$\hat{\mathbb{V}}\left[\hat{\boldsymbol{\beta}}\right] = (\mathbf{X}'\mathbf{X})^{-1} \left(\sum_{g=1}^{G} \mathbf{X}'_{g} \hat{\mathbf{u}}_{g} \hat{\mathbf{u}}'_{g} \mathbf{X}_{g}\right) (\mathbf{X}'\mathbf{X})^{-1}$$
(5.4)

where $\hat{\mathbf{u}}_g = \mathbf{y}_g - \mathbf{X}_g \hat{\boldsymbol{\beta}}$. This estimate is consistent if

$$G^{-1} \sum_{g=1}^{G} \mathbf{X}_{g}' \hat{\mathbf{u}}_{g} \hat{\mathbf{u}}_{g}' \mathbf{X}_{g} - G^{-1} \sum_{g=1}^{G} \mathbb{E} \left[\mathbf{X}_{g}' \mathbf{u}_{g} \mathbf{u}_{g}' \mathbf{X}_{g} \right] \xrightarrow{p} \mathbf{0}$$

as $G \to \infty$. An informal presentation of Eq.(5.4) is to rewrite the central matrix as

$$\hat{\mathbf{B}} = \sum_{g=1}^{G} \mathbf{X}_{g}' \hat{\mathbf{u}}_{g} \hat{\mathbf{u}}_{g}' \mathbf{X}_{g} = \mathbf{X}' \begin{bmatrix} \hat{\mathbf{u}}_{1} \hat{\mathbf{u}}_{1}' & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \hat{\mathbf{u}}_{2} \hat{\mathbf{u}}_{2}' & & \vdots \\ \vdots & & \ddots & \mathbf{0} \\ \mathbf{0} & \cdots & & \hat{\mathbf{u}}_{G} \hat{\mathbf{u}}_{G}' \end{bmatrix} \mathbf{X} = \mathbf{X}' \left(\hat{\mathbf{u}} \hat{\mathbf{u}}' \otimes \mathbf{S}^{G} \right) \mathbf{X}$$
(5.5)

where \otimes denotes element-wise multiplication. The (p,q)th element of this matrix is

$$\sum_{i=1}^{N} \sum_{j=1}^{N} x_{ia} x_{jb} \hat{u}_{i} \hat{u}_{j} \cdot \mathbf{1} (i, j \text{ in the same cluster})$$

with $\hat{u}_i = y_i - \mathbf{x}_i' \hat{\boldsymbol{\beta}}$.

 \mathbf{S}^G is an $N \times N$ indicator matrix with $\mathbf{S}^G_{ij} = 1$ only if the ith and jth observation belong to the same cluster: it zeros out a large amount of $\hat{\mathbf{u}}\hat{\mathbf{u}}'$ (asymptotically equivalently, $\mathbf{u}\mathbf{u}'$), specifically, only $\sum_{g=1}^G N_g^2$ out of $N^2 = \left(\sum_{g=1}^G N_g\right)^2$ terms are not zero (sub-matrices on the diagonal). Asymptotically

- for fixed N_g , $\frac{1}{N^2} \sum_{g=1}^G N_g^2 \xrightarrow{G \to \infty} 0$
- for balanced clusters $N_g = N/G$, $\frac{1}{N^2} \sum_{g=1}^G N_g^2 = \frac{1}{G} \xrightarrow{G \to \infty} 0$

A strand of literature popularizes this method:

- Liang and Zeger (1986): in a generalized estimatin equations setting
- Arellano (1987): fixed effects estimator in linear panel models
- Hansen (2007): asymptotic theory for panel data where $T \to \infty$ in addition to $N \to \infty$ (or $N_g \to \infty$ in addition to $G \to \infty$ in the notation above).

5.2 Two-Way Clustering

Now, consider the case of two-way clustering,

$$y_{i,gh} = \mathbf{x}'_{i,gh} \boldsymbol{\beta} + u$$

where each observation may belong to **two** dimension of groups: group $g \in \{1, \dots, G\}$ and $h \in \{1, \dots, H\}$, and for $i \neq j$

$$\mathbb{E}\left[u_{i,gh}u_{j,g'h'}\mid\mathbf{x}_{i,gh},\mathbf{j},\mathbf{g'h'}\right]=0$$
(5.6)

unless g = g' or h = h', that is, errors for individuals within the same group (along either g or h) may be correlated.

Cluster-robust variance matrix estimate extending the one-way clustering case, keep elements of $\hat{\mathbf{u}}\hat{\mathbf{u}}'$ where the *i*th and *j*th observations share a cluster in **any** dimension, then similar to Eq.(5.5)

$$\hat{\mathbf{B}} = \mathbf{X}' \left(\hat{\mathbf{u}} \hat{\mathbf{u}}' \otimes \mathbf{S}^{GH} \right) \mathbf{X} \tag{5.7}$$

here \mathbf{S}^{GH} is an $N \times N$ indicator matrix with $\mathbf{S}_{ij}^{GH} = 1$ only if the ith and jth observation share any cluster, the (p,q)th element of this matrix is

$$\sum_{i=1}^{N}$$

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