Econometrics June 17, 2023

Topic 17: False Discovery Rate (FDR) and Knockoffs

by Sai Zhang

Key points: Constructing knockoff variables to control FDR when estimating regression coefficients.

Disclaimer: The note is built on Prof. Jinchi Lv's lectures of the course at USC, DSO 607, High-Dimensional Statistics and Big Data Problems.

17.1 Motivation

Consider the classical linear regression setting

$$y = X\beta + \epsilon$$

where $\beta \in \mathbb{R}^p$ is the unknown vector of coefficients and $\epsilon \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$. In a high-dimensional problem, we would like to just select a subset of all variables $\hat{S} \subset \{1, \cdots, p\}$ s.t. conditional on $\{\mathbf{X}_j\}_{j \in \hat{S}}$, \mathbf{y} is **independent** of all other variables, we can define the **False Discovery Rate** (FDR) in can be defined as

Definition 17.1.1: False Discovery Rate (FDR)

$$FDR = \mathbb{E}(FDP) = \mathbb{E}\left[\frac{|\hat{S} \cap \mathcal{H}_0|}{|\hat{S}|} = \frac{\#\{j : j \in \hat{S} \setminus S\}}{\#\{j : j \in \hat{S}\}}\right]$$

where $\mathcal{H}_0 \subset \{1, \dots, p\}$ is the set of **null** variables: \mathbf{X}_j is **null** iff \mathbf{Y} is independent of \mathbf{X}_j conditional on the other variables $\mathbf{X}_{-j} = \{\mathbf{X}_1, \dots, \mathbf{X}_p\} \setminus \{\mathbf{X}_j\}$.

In this note, we consider a series of knockoff-based methods to control FDR. They all follow a common procedure:

- Step 1: Construct Knockoffs
- Step 2: Calculate test statistics for both original and knockoff variables
- Step 3: Calculate a threshold for the test statistics, controling for a desired FDR level
- Step 4: Select variables that pass the threshold

17.2 Barber and Candes (2015)

Constructing the knockoffs Barber and Candes (2015) construct the knockoffs by the following procedure

• Calculate the Gram matrix $\Sigma = \mathbf{X}'\mathbf{X}$ for the normalized original variables, where $\Sigma_{jj} = \left\|\mathbf{X}_j\right\|_2^2 = 1$

• Construct the knockoffs \tilde{X} s.t.

$$\tilde{\mathbf{X}}'\tilde{\mathbf{X}} = \mathbf{\Sigma}$$
 $\mathbf{X}'\tilde{\mathbf{X}} = \mathbf{\Sigma} - \operatorname{diag}\left\{\mathbf{s}\right\}$

where $\mathbf{s} \in \mathbb{R}^p_+$ is a p-dimensional non-negative vector (larger s_j indicates higher power) and

- \tilde{X} exhibits the same covariance structrue as the original design X
- The correlation between distinct original variables and knockoffs are the same as between the originals:

$$\mathbf{X}_{i}^{\prime}\tilde{\mathbf{X}}_{k} = \mathbf{X}_{i}^{\prime}\mathbf{X}_{k}, \ \forall j \neq k$$

- The correlation between the original variables and their own knockoffs is **less than 1**

$$\mathbf{X}_i'\tilde{\mathbf{X}}_j = \Sigma_{jj} - s_j = 1 - s_j$$

To construct such knockoffs,

- Given a proper \mathbf{s} , if $n \ge 2p$, then

$$\tilde{\mathbf{X}} = \mathbf{X}(\mathbf{I} - \mathbf{\Sigma}^{-1} \text{diag} \{\mathbf{s}\}) + \tilde{\mathbf{U}}\mathbf{C}$$

where $\tilde{\mathbf{U}} \in \mathbb{R}^{n \times p}$ is an **orthonormal** matrix s.t. $\tilde{\mathbf{U}}'\mathbf{X} = \mathbf{0}$ and $\mathbf{C}'\mathbf{C} = 2\mathrm{diag}\{\mathbf{s}\} - \mathrm{diag}\{\mathbf{s}\} \Sigma^{-1}\mathrm{diag}\{\mathbf{s}\} \geq \mathbf{0}$

- A sufficient and necessary condition for $\tilde{\mathbf{X}}$ to exist: diag $\{\mathbf{s}\} \leq 2\Sigma$
- 2 types of knockoffs can be constructed, following these procedures
- T1 <u>Equi-correlated</u> knockoffs: set $s_j = 2\lambda_{\min}(\Sigma) \wedge 1$ for all j, then $\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle = 1 2\lambda_{\min}(\Sigma) \wedge 1$ for all j. This is essentially minimizing $|\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle|$
- T2 SDP knockoffs: solve the convex problem

$$\arg\min_{\mathbf{x}} \sum_{j} (1 - s_j) \qquad \qquad s.t.0 \le s_j \le 1, \operatorname{diag} \{\mathbf{s}\} \le 2\Sigma$$

which is essentially minimizing the average of $\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle$

Calculate test statistics Define and calculate test statistics W_i for each $\beta_i \in \{1, \dots, p\}$ using $[\mathbf{X} \ \tilde{\mathbf{X}}]$:

• the test statistic W_j should be constructed s.t. large positive values are evidence against the null hypothesis $\beta_j = 0$, for example, consider a Lasso on $\begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix}$

$$\hat{\beta}(\lambda) = \arg\min_{\mathbf{b}} \left\{ \frac{1}{2} \| \mathbf{y} - \begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix} \mathbf{b} \|_{2}^{2} + \lambda \| \mathbf{b} \| \right\}_{1}$$

where λ is the point on the Lasso path at which the feature enters the model as

$$Z_i = \sup \{\lambda : \hat{\beta}_i(\lambda) \neq 0\}$$

and set
$$W_j = (Z_j \vee \tilde{Z}_j) \cdot \begin{cases} +1, & Z_j > \tilde{Z}_{j-1} \\ -1, & Z_j < \tilde{Z}_j \end{cases}$$

Other choices of
$$W_j$$
 are $W_j = \left| \mathbf{X}_j' \mathbf{y} \right| - \left| \tilde{\mathbf{X}}_j' \mathbf{y} \right|$, or $\left| \hat{\beta}_j^{\text{LS}} \right| - \left| \hat{\beta}_{j+p}^{\text{LS}} \right|$

References

Rina Foygel Barber and Emmanuel J. Candes. Controlling the false discovery rate via knockoffs. *Annals of Statistics*, 43(5):2055–2085, 2015.