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Topic 5: Two-Way Cluster-Robust (TWCR) Standard Errors

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Key points: The validity of Two-Way Cluster-Robust (TWCR) standard errors

Disclaimer: *This note is compiled by Sai Zhang.*

5.1 One-Way Clustering

First, consider the case of one-way clustering. The linear model with one-way clustering

$$y_{ig} = \mathbf{x}_{ig}\boldsymbol{\beta} + u_{ig}$$

where i denotes the ith of the N individuals in the sample, j denotes the gth of the G clusters, assume that

- $\mathbb{E}\left[u_{ig} \mid \mathbf{x}_{ig}\right] = 0$
- error independence across clusters: for $i \neq j$

$$\mathbb{E}\left[u_{ig}u_{jg'} \mid \mathbf{x}_{ig}, \mathbf{x}_{jg'}\right] = 0 \tag{5.1}$$

  unless $g = g'$, that is, errors for individuals within the same cluster may be correlated.

Grouping observations by cluster, get

$$\mathbf{y}_g = \mathbf{X}_g\boldsymbol{\beta} + \mathbf{u}$$

where $\mathbf{X}_g$ has dimension $N_g \times K$ and $\mathbf{y}_g$ has dimension $N_g \times 1$, with N_g observations in cluster g. Stacking over cluster, get the matrix form of the model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{u}$$

with $\mathbf{y}$, $\mathbf{u}$ being $N \times 1$ vectors, $\mathbf{X}$ being an $N \times K$ matrix. OLS estimator gives

$$\hat{\boldsymbol{\beta}} = \left(\mathbf{X}'\mathbf{X}\right)^{-1}\mathbf{X}'\mathbf{y} = \left(\sum_{g=1}^{G}\mathbf{X}_g'\mathbf{X}_g\right)^{-1}\sum_{g=1}^{G}\mathbf{X}_g'\mathbf{y}_g \tag{5.2}$$

then, by CLT, we have that $\sqrt{G}\left(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}\right) \xrightarrow{d} \mathcal{N}(0, \boldsymbol{\Sigma})$ where the variance matrix of the limit normal distribution $\boldsymbol{\Sigma}$ is

$$\left(\lim_{G\to\infty}\frac{1}{G}\sum_{g=1}^{G}\mathbf{E}\left[\mathbf{X}_g'\mathbf{X}_g\right]\right)^{-1}\left(\lim_{G\to\infty}\frac{1}{G}\sum_{g=1}^{G}\mathbf{E}\left[\mathbf{X}_g'\mathbf{u}_g'\mathbf{u}_g\mathbf{X}_g\right]\right)\times\left(\lim_{G\to\infty}\frac{1}{G}\sum_{g=1}^{G}\mathbf{E}\left[\mathbf{X}_g'\mathbf{X}_g\right]\right)^{-1} \tag{5.3}$$

If the primary source of clustering is due to group-level common shocks, a useful approximation is that for the jth regressor, the default OLS variance estimate based on $s^2\left(\mathbf{X}'\mathbf{X}\right)^{-1}$ should be inflated by $\tau_j \simeq 1 + \rho_{x_j}\rho_u\left(\overline{N}_g - 1\right)$, where

- s is the estimated standard deviation of the error

- ρ_{x_j} is a measure of within-cluster correlation of x_j
- ρ_u is the within-cluster error correlation
- $\overline{N}_g$ is the average cluster size

It's easy to see the τ_j can be large even with small ρ_u (Kloek, 1981; Scott and Holt, 1982; Moulton, 1990). If assume the model for the cluster error variance matrices $\boldsymbol{\Omega}_g = \mathbb{V}\left[\mathbf{u}_g \mid \mathbf{X}_g\right] = \mathbb{E}\left[\mathbf{u}_g\mathbf{u}_g' \mid \mathbf{X}_g\right]$, and there is a consistent estimate $\hat{\boldsymbol{\Omega}}_g$ of $\boldsymbol{\Omega}_g$, we can estimate $\mathbb{E}\left[\mathbf{X}_g'\mathbf{u}_g\mathbf{u}_g'\mathbf{X}_g\right] = \mathbb{E}\left[\mathbf{X}_g'\boldsymbol{\Omega}_g\mathbf{X}_g\right]$ via GLS.

Cluster-robust variance matrix estimate consider

$$\hat{\mathbb{V}}\left[\hat{\boldsymbol{\beta}}\right] = (\mathbf{X}'\mathbf{X})^{-1}\left(\sum_{g=1}^{G}\mathbf{X}_g'\hat{\mathbf{u}}_g\hat{\mathbf{u}}_g'\mathbf{X}_g\right)(\mathbf{X}'\mathbf{X})^{-1} \tag{5.4}$$

where $\hat{\mathbf{u}}_g = \mathbf{y}_g - \mathbf{X}_g\hat{\boldsymbol{\beta}}$. This estimate is consistent if

$$G^{-1}\sum_{g=1}^{G}\mathbf{X}_g'\hat{\mathbf{u}}_g\hat{\mathbf{u}}_g'\mathbf{X}_g - G^{-1}\sum_{g=1}^{G}\mathbb{E}\left[\mathbf{X}_g'\mathbf{u}_g\mathbf{u}_g'\mathbf{X}_g\right] \xrightarrow{\text{p}} \mathbf{0}$$

as $G \to \infty$. An informal presentation of Eq.(5.4) is to rewrite the central matrix as

$$\sum_{g=1}^{G}\mathbf{X}_g'\hat{\mathbf{u}}_g\hat{\mathbf{u}}_g'\mathbf{X}_g = \mathbf{X}'\begin{bmatrix}\hat{\mathbf{u}}_1\hat{\mathbf{u}}_1' & \mathbf{0} & \cdots & \mathbf{0}\\ \mathbf{0} & \hat{\mathbf{u}}_2\hat{\mathbf{u}}_2' & & \vdots\\ \vdots & & \ddots & \mathbf{0}\\ \mathbf{0} & \cdots & & \hat{\mathbf{u}}_G\hat{\mathbf{u}}_G'\end{bmatrix}\mathbf{X} = \mathbf{X}'\left(\hat{\mathbf{u}}\hat{\mathbf{u}}' \otimes \mathbf{S}^G\right)\mathbf{X}$$

where $\otimes$ denotes element-wise multiplication. The (p, q)th element of this matrix is

$$\sum_{i=1}^{N}\sum_{j=1}^{N} x_{ia}x_{jb}\hat{u}_i\hat{u}_j \cdot \mathbf{1}\left(i, j \text{ in the same cluster}\right)$$

with $\hat{u}_i = y_i - \mathbf{x}_i'\hat{\boldsymbol{\beta}}$.

$\mathbf{S}^G$ is an $N \times N$ indicator matrix with $\mathbf{S}_{ij}^G = 1$ only if the ith and jth observation belong to the same cluster: it zeros out a large amount of $\hat{\mathbf{u}}\hat{\mathbf{u}}'$ (asymptotically equivalently, $\mathbf{u}\mathbf{u}'$), specifically, only $\sum_{g=1}^{G} N_g^2$ out of $N^2 = \left(\sum_{g=1}^{G} N_g\right)^2$ terms are not zero (sub-matrices on the diagonal). Asymptotically

- for fixed N_g, $\frac{1}{N^2}\sum_{g=1}^{G} N_g^2 \xrightarrow{G\to\infty} 0$
- for balanced clusters $N_g = N/G$, $\frac{1}{N^2}\sum_{g=1}^{G} N_g^2 = \frac{1}{G} \xrightarrow{G\to\infty} 0$

A strand of literature popularizes this method:

- Liang and Zeger (1986): in a generalized estimatin equations setting
- Arellano (1987): fixed effects estimator in linear panel models
- Hansen (2007): asymptotic theory for panel data where $T \to \infty$ in addition to $N \to \infty$ (or $N_g \to \infty$ in addition to $G \to \infty$ in the notation above).

5.2 Two-Way Clustering

Now, consider the case of two-way clustering,

$$y_{i,gh} = \mathbf{x}'_{i,gh}\boldsymbol{\beta} + u$$

where each observation may belong to **two** dimension of groups: group $g \in \{1, \cdots, G\}$ and $h \in \{1, \cdots, H\}$, and for $i \neq j$

$$\mathbb{E}\left[u_{i,gh}u_{j,g'h'} \mid \mathbf{x}_{i,gh}, \mathbf{j}, \mathbf{g'h'}\right] = 0 \tag{5.5}$$

unless $g = g'$ or $h = h'$, that is, errors for individuals within the same group (along either g or h) may be correlated.

References

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