Econometrics May 7, 2023

Topic 14: Regularization Methods in Thresholded Parameter Space

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Key points: The connections and differences of all regularization methods and some interesting phase transition phenomena.

Disclaimer: The note is built on Prof. Jinchi Lv's lectures of the course at USC, DSO 607, High-Dimensional Statistics and Big Data Problems.

14.1 Model Setup

Now, consider a generalized linear model (GLM) linking a p-dimensional predictor \mathbf{x} to a scalar response Y. With canonical link, the conditional distribution of Y given \mathbf{x} has density

$$f(y; \theta, \phi) = \exp \left[y\theta - b(\theta) + c(y, \phi) \right]$$

where $\theta = \mathbf{x}'\boldsymbol{\beta}$ with $\boldsymbol{\beta}$ a p-dimensional regression coefficient vector, $b(\dot{\mathbf{y}})$ and $c(\cdot, \cdot)$ are know functions and ϕ is dispersion parameter. Again, $\boldsymbol{\beta} = (\beta_{0,1}, \cdots, \beta_{0,p})'$ is sparse with many zero components, and $\log p = O(n^a)$ for some 0 < a < 1.

The penalized negative log-likelihood is

$$Q_n(\boldsymbol{\beta}) = -n^{-1} \left[\mathbf{y}' \mathbf{X} \boldsymbol{\beta} - \mathbf{1}' \mathbf{b} (\mathbf{X} \boldsymbol{\beta}) \right] + \| p_{\lambda}(\boldsymbol{\beta}) \|_1$$

where

- $\mathbf{y} = (y_1, \dots, y_n)', \mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_n)', \text{ each column of } \mathbf{X} \text{ is rescaled to have } L_2\text{-norm } \sqrt{n}$
- $\mathbf{b}(\theta) = (b(\theta_1), \dots, b(\theta_n))'$ with $\theta = (\theta_1, \dots, \theta_n)'$
- $||p_{\lambda}(\boldsymbol{\beta})||_1 = \sum_{j=1}^p p_{\lambda}(|\beta_j|)$

Next, define **robust spark** κ_c

Definition 14.1.1: Robust spark κ_c

The robust spark κ_c of the $n \times p$ design matrix **X** is defined as the smallest possible positive integer s.t. there exists an $n \times \kappa_c$ submatrix of $\frac{1}{\sqrt{n}}$ **X** having a singular value less than a given positive constant c (Zheng et al., 2014), and

$$\kappa_c \leq n+1$$

Bounding sparse model size can control collinearity and ensure model identifiability and stability, and as $c \to 0+$, κ_c approaches the spark. Robust spark can be some large number diverging with n:

Proposition 14.1.2: Order of κ_c

Assume $\log p = o(n)$ and that the rows of the $n \times p$ random design matrix **X** are i.i.d. as $\mathcal{N}(\mathbf{0}, \Sigma)$, where Σ has smallest eigenvalue bounded from below by some positive constant. Then there exist

positive constants c and \tilde{c} s.t. with asymptotic probability one, $\kappa_c \geq \frac{\tilde{c}n}{\log p}$

Next, we define a thresholded parameter space

Definition 14.1.3: Thresholded parameter space

$$\mathcal{B}_{\tau,c} = \left\{ \boldsymbol{\beta} \in \mathbb{R}^p : \|\boldsymbol{\beta}\|_0 < \frac{\kappa_c}{2}, \text{ and for each } j, \beta_j = 0 \text{ or } |\beta_j| \ge \tau \right\}$$

where $\beta = (\beta_1, \dots, \beta_p)'$. τ is some positive threshold on parameter magnitude:

Here, τ is very important:

- τ is key to distinguishing between important covariates and noise covariates for the purpose of variable selection
- τ typically needs to satisfy $\tau \sqrt{n/\log p} \xrightarrow{n \to \infty} \infty$

It turns out that the solution to the regularization problem has the (very natural) hard-thresholding property:

Proposition 14.1.4: Hard-thresholding property

or the L_0 -penalty $p_{\lambda}(t) = \lambda \mathbf{1}_{t\neq 0}$, the global minimizer $\hat{\boldsymbol{\beta}} = (\hat{\beta}_1, \dots, \hat{\beta}_p)'$ of the regularization problem over \mathbb{R}^p satisfies that each component $\hat{\beta}_j$ is either 0 or has magnitude larger than some positive threshold

This hard-thresholding property is shared by many other penalties such as SICA penalties. This property guarantees sparcity of the model: weak signals are generally difficult to stand out comparing to noise variables due to impact of high dimensionality

14.2 Asymptotic Equivalence of Regularization Methods

For a universal $\lambda = c_0 \sqrt{\log p/n}$ with $c_0 > 0$ and p implicitly as $n \vee p$, consider 2 key events:

$$\mathcal{E} = \left\{ \|n^{-1} \mathbf{X}' \boldsymbol{\epsilon}\|_{\infty} \le \lambda/2 \right\} \qquad \qquad \mathcal{E}_0 = \left\{ \|n^{-1} \mathbf{X}'_{\alpha_0} \boldsymbol{\epsilon}\|_{\infty} \le c_0 \sqrt{\log n/n} \right\}$$

where $\epsilon = \mathbf{y} - \mathbb{E}\mathbf{y}$, \mathbf{X}_{α} is a submatrix of \mathbf{X} consisting of columns in α . Here, let $\alpha_0 = \operatorname{supp}(\boldsymbol{\beta}_0)$ (non-zero variables in the true model).

For this setting, consider the following technical conditions:

- C1 **Error tail distribution**: $Pr(\mathcal{E}^c) = O(p^{-c_1})$ and $Pr(\mathcal{E}^c_0) = O(n^{-c_1})$ for some positive constant c_1 that can be sufficiently large for large enough c_0
- C2 **Bounded variance**: $b(\theta)$ satisfies that $c_2 \le b''(\theta) \le c_2^{-1}$ in its domain, where c_2 is some positive constant
- C3 **Concave penalty function**: $p_{\lambda}(t)$ is increasing and concave in $t \in [0, \infty)$ with $p_{\lambda}(0) = 0$, and is differentiable with $p'_{\lambda}(0+) = c_3\lambda$ for some positive constant c_3^1
- C4 Ultra-high dimensionality: $\log p = O(n^a)$ for some constant $a \in (0, 1)$

 $^{^{1}}$ A wide class of penalties, including L_{1} -penalty in Lasso, SCAD, MCP and SICA, satisfy this condition.

C5 **True parameter vector**: $s = o(n^{1-a})$ and $\exists c > 0$ s.t. the **robust spark** $\kappa_c > 2s$. Moreover, $\min_{1 \le j \le s} |\beta_{0,j}| \gg \sqrt{\log p/n}$

Given these 5 conditions, we have that the global minimizer $\hat{\beta} = \arg\min_{\beta \in \mathcal{B}_{\tau}} Q_n(\beta)$ exists and satisfies oracle inequalities:

Theorem 14.2.1: Oracle Inequalities

Assume that Condition 1-5 hold and τ is chosen s.t. $\tau < \min_{1 \le j \le s} |\beta_{0,j}|$ and $\lambda = c_0 \sqrt{\log p/n} = o(\tau)$, then the global minimizer exists, and any such global minimizer satisfies that with probability at least $1 - O(p^{-c_1})$, it holds simultaneously that

• False sign:

$$FS(\hat{\boldsymbol{\beta}}) \le \frac{Cs\lambda^2\tau^{-2}}{1 - C\lambda^2\tau^{-2}}$$

Estimation losses:

$$\begin{aligned} \|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\|_q &\leq C\lambda s^{1/q} (1 - C\lambda^2 \tau^{-2})^{-1/q} & \forall q \in [1, 2] \\ \|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\|_{\infty} &\leq C\lambda s^{1/2} (1 - C\lambda^2 \tau^{-2})^{-1/2} \end{aligned}$$

Prediction loss:

$$\frac{1}{\sqrt{n}} \| \mathbf{X} \left(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0 \right) \|_2 \le C \lambda s^{1/2} (1 - C \lambda^2 \tau^{-2})^{-1/2}$$

where *C* is some positive constant.

How to understand Thm.14.2.1

- These results hold uniformly over the set of all possible global minimizers
- c_1 in probability bound can be chosen arbitrarily large, affecting **only** C
- $FS(\hat{\beta}) = o(s)$ since $\lambda = o(\tau)$, while $\|\hat{\beta}\|_0 = O(\phi_{\text{max}}s)$ where ϕ_{max} is the largest eigenvalue of $\frac{1}{n}X'X$
- $\forall q \in [1, 2]$, the convergence rates of estimation losses

$$\|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\|_q = O\left\{s^{1/q} \sqrt{\frac{\log p}{n}}\right\}$$
$$\frac{1}{\sqrt{n}} \|\mathbf{X} \left(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\right)\|_2 = O\left(\sqrt{\frac{s \log p}{n}}\right)$$

are consistent with Lasso.

We also have a sign consistency result:

Theorem 14.2.2: Sign Consistency and Oracle Inequalities

Assume the same conditions of Thm.14.2.1, further assume $\min_{1 \le j \le s} |\beta_{0,j}| \ge 2\tau$ and $\lambda = c_0 \sqrt{\log p/n} = o(s^{-1/2}\tau)$, and $\gamma_n = o\left(\tau \sqrt{\frac{n}{s \log n}}\right)$, then any global minimizer $\hat{\beta}$ defined satisfies that with probability at least $1 - O(n^{-c_1})$, it holds simultaneously that

- **Sign consistency**: $\operatorname{sgn}(\hat{\beta}) = \operatorname{sgn}(\beta_0)$
- **Estimation and prediction losses**: If the penalty function further satisfies $p'_{\lambda}(\tau) = O\left(\frac{\log n}{n}\right)$, then $\forall q \in [1,2]$,

$$\left\|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\right\|_q \le C s^{1/q} \sqrt{\frac{\log n}{n}} \qquad \left\|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\right\|_{\infty} \le C \gamma_n^* \sqrt{\frac{\log n}{n}} \qquad n^{-1} D\left(\hat{\boldsymbol{\beta}}\right) \le C \frac{s \log n}{n}$$

where γ_n^* is a constant showing the behavior of $\|\left[\frac{1}{n}\mathbf{X}_{\alpha_0}'\mathbf{H}\left(\boldsymbol{\beta}_1,\cdots,\boldsymbol{\beta}_n\right)\mathbf{X}_{\alpha_0}\right]^{-1}\|_{\infty}$ in a small neighborhood of $\boldsymbol{\beta}_0$, $D(\hat{\boldsymbol{\beta}})$ is the Kullback-Leibler divergence, and C is some positive constant

References

Zemin Zheng, Yingying Fan, and Jinchi Lv. High dimensional thresholded regression and shrinkage effect. *Journal of the Royal Statistical Society: Series B: Statistical Methodology*, pages 627–649, 2014.