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Topic 17: False Discovery Rate (FDR) and Knockoffs

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Key points: Constructing knockoff variables to control FDR when estimating regression coefficients.

Disclaimer: The note is built on Prof. Jinchi Lv's lectures of the course at USC, DSO 607, High-Dimensional Statistics and Big Data Problems.

17.1 Motivation

Consider the classical linear regression setting

$$y = X\beta + \epsilon$$

where $\beta \in \mathbb{R}^p$ is the unknown vector of coefficients and $\epsilon \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$. In a high-dimensional problem, we would like to just select a subset of all variables $\hat{S} \subset \{1, \cdots, p\}$ s.t. conditional on $\{\mathbf{X}_j\}_{j \in \hat{S}}$, \mathbf{y} is **independent** of all other variables, we can define the **False Discovery Rate** (FDR) in can be defined as

Definition 17.1.1: False Discovery Rate (FDR)

$$FDR = \mathbb{E}(FDP) = \mathbb{E}\left[\frac{|\hat{S} \cap \mathcal{H}_0|}{|\hat{S}|} = \frac{\#\{j : j \in \hat{S} \setminus S\}}{\#\{j : j \in \hat{S}\}}\right]$$

where $\mathcal{H}_0 \subset \{1, \dots, p\}$ is the set of **null** variables: \mathbf{X}_j is **null** iff \mathbf{Y} is independent of \mathbf{X}_j conditional on the other variables $\mathbf{X}_{-j} = \{\mathbf{X}_1, \dots, \mathbf{X}_p\} \setminus \{\mathbf{X}_j\}$.

In this note, we consider a series of knockoff-based methods to control FDR. They all follow a common procedure:

- Step 1: Construct Knockoffs
- Step 2: Calculate test statistics for both original and knockoff variables
- Step 3: Calculate a threshold for the test statistics, controling for a desired FDR level
- Step 4: Select variables that pass the threshold

17.2 Barber and Candes (2015)

Constructing the knockoffs Barber and Candes (2015) construct the knockoffs by the following procedure

• Calculate the Gram matrix $\Sigma = \mathbf{X}'\mathbf{X}$ for the normalized original variables, where $\Sigma_{jj} = \|\mathbf{X}_j\|_2^2 = 1$

• Construct the knockoffs \tilde{X} s.t.

$$\tilde{\mathbf{X}}'\tilde{\mathbf{X}} = \mathbf{\Sigma}$$
 $\mathbf{X}'\tilde{\mathbf{X}} = \mathbf{\Sigma} - \operatorname{diag}\{\mathbf{s}\}\$

where $\mathbf{s} \in \mathbb{R}^p_+$ is a p-dimensional non-negative vector (larger s_i indicates higher power) and

- $-\tilde{X}$ exhibits the **same** covariance structrue as the original design X
- The correlation between distinct original variables and knockoffs are the same as between the originals:

$$\mathbf{X}_{i}^{\prime}\tilde{\mathbf{X}}_{k} = \mathbf{X}_{i}^{\prime}\mathbf{X}_{k}, \ \forall j \neq k$$

- The correlation between the original variables and their own knockoffs is **less than 1**

$$\mathbf{X}_{j}'\tilde{\mathbf{X}}_{j} = \Sigma_{jj} - s_{j} = 1 - s_{j}$$

To construct such knockoffs,

- Given a proper \mathbf{s} , if $n \ge 2p$, then

$$\tilde{\mathbf{X}} = \mathbf{X}(\mathbf{I} - \mathbf{\Sigma}^{-1} \operatorname{diag} \{\mathbf{s}\}) + \tilde{\mathbf{U}}\mathbf{C}$$

where $\tilde{\mathbf{U}} \in \mathbb{R}^{n \times p}$ is an **orthonormal** matrix s.t. $\tilde{\mathbf{U}}'\mathbf{X} = \mathbf{0}$ and $\mathbf{C}'\mathbf{C} = 2\mathrm{diag}\{\mathbf{s}\} - \mathrm{diag}\{\mathbf{s}\} \Sigma^{-1}\mathrm{diag}\{\mathbf{s}\} \geq \mathbf{0}$

- A sufficient and necessary condition for $\tilde{\mathbf{X}}$ to exist: diag $\{\mathbf{s}\} \leq 2\Sigma$
- 2 types of knockoffs can be constructed, following these procedures
- T1 <u>Equi-correlated</u> knockoffs: set $s_j = 2\lambda_{\min}(\Sigma) \wedge 1$ for all j, then $\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle = 1 2\lambda_{\min}(\Sigma) \wedge 1$ for all j. This is essentially minimizing $|\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle|$
- T2 SDP knockoffs: solve the convex problem

$$\arg\min_{\mathbf{x}} \sum_{j} (1 - s_j) \qquad \qquad s.t.0 \le s_j \le 1, \operatorname{diag}\{\mathbf{s}\} \le 2\Sigma$$

which is essentially minimizing the average of $\langle \mathbf{X}_i, \tilde{\mathbf{X}}_i \rangle$

Calculate test statistics Define and calculate test statistics W_j for each $\beta_j \in \{1, \dots, p\}$ using $[\mathbf{X} \ \tilde{\mathbf{X}}]$:

• the test statistic W_j should be constructed s.t. large positive values are evidence against the null hypothesis $\beta_j = 0$, for example, consider a Lasso on $\begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix}$

$$\hat{\beta}(\lambda) = \arg\min_{\mathbf{b}} \left\{ \frac{1}{2} \| \mathbf{y} - \begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix} \mathbf{b} \|_{2}^{2} + \lambda \| \mathbf{b} \| \right\}_{1}$$

where λ is the point on the Lasso path at which the feature enters the model as

$$Z_j = \sup \left\{ \lambda : \hat{\beta}_j(\lambda) \neq 0 \right\}$$

and set
$$W_j = (Z_j \vee \tilde{Z}_j) \cdot \begin{cases} +1, & Z_j > \tilde{Z}_{j_1} \\ -1, & Z_j < \tilde{Z}_j \end{cases}$$

• In general, the statistics W should satisfy the **sufficient** property and **anti-symmetry** property:

¹Other choices of
$$W_j$$
 are $W_j = \left| \mathbf{X}_j' \mathbf{y} \right| - \left| \tilde{\mathbf{X}}_j' \mathbf{y} \right|$, or $\left| \hat{\beta}_j^{\mathrm{LS}} \right| - \left| \hat{\beta}_{j+p}^{\mathrm{LS}} \right|$

Definition 17.2.1: Property of Test Statistics W_i

The test statistic W_i is said to obey

the <u>sufficient</u> property if W depends <u>only</u> on the Gram matrix and on feature-response inner products, that is

$$\mathbf{W} = f\left(\begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix}' \begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix}, \begin{bmatrix} \mathbf{X} & \tilde{\mathbf{X}} \end{bmatrix}' \mathbf{y}\right)$$

- the **antisymmetry** property if swapping the original X_j and its knockoff \tilde{X}_j has the effect of **switching the sign** of W_j , that is

$$W_j(Z_j, \tilde{Z}_j) = -W_j(\tilde{Z}_j, Z_j)$$

Calculate a threshold for the test statistics After defining the test statistic, we then

• Let *q* be the target FDR, define the data-dependent threshold *T* as

$$T = \min \left\{ t \in \mathcal{W} : \frac{\# \left\{ j : W_j \le -t \right\}}{\# \left\{ j : W_j \ge t \right\} \vee 1} \le q \right\}$$

where $W = \{|W_j| : j = 1, \dots, p\} \setminus \{0\}$ is the set of unique non-zero values attained by $|W_j|$'s.

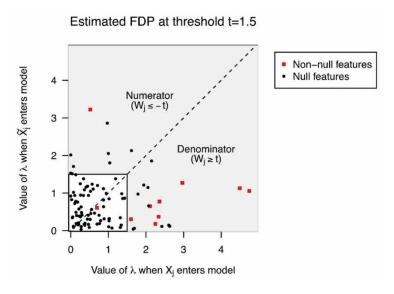


Figure 17.1: Visualizing Test Statistic Thresholding

Variable selection after building the threshold,

• for each $j = 1, \dots, p$, reject $H_{0,j} : \beta_j = 0$ if $W_j \ge T$, the knockoff filter selects the model

$$\hat{S} = \left\{ j : W_j \ge T \right\}$$

17.2.1 Intuition and Theory

Why knockoffs work?

- W is constructed (antisymmetry and sufficiency) such that the signs of the W_j 's are i.i.d. random for the null
- for any threshold *t* , we have

$$\#\{j: \beta_j = 0, W_j \ge t\} \stackrel{d}{=} \#\{j: \beta_j = 0, W_j \le -t\}$$

, and the false discovery proportion (FDP) can be estimated as

$$\frac{\#\{j: \beta_{j} = 0, W_{j} \ge t\}}{\max(\#\{j: W_{j} \ge t\}, 1)} \simeq \frac{\#\{j: \beta_{j} = 0, W_{j} \le -t\}}{\max(\#\{j: W_{j} \ge t\}, 1)}$$
$$\le \frac{\#\{j: W_{j} \le -t\}}{\max(\#\{j: W_{j} \ge t\}, 1)} := \widehat{\text{FDP}}(t)$$

then the knockoff procedure can be interpreted as finding a threshold via $T = \min \left\{ t \in \mathcal{W} : \widehat{\mathrm{FDR}}(t) \leq q \right\}$

The knockoff procedure essentially controls a quantity **nearly equal** to the FDR. To control the FDR **exactly**, we have, <u>textbfknockoff+</u>, a more conservative modification of the knockoff procedure, where the threshold is

$$T = \min \left\{ t \in \mathcal{W} : \frac{1 + \# \{ j : W_j \le -t \}}{\max (\# \{ j : W_j \ge t \}, 1)} \le q \right\}$$

the +1 part makes it harder to reject the null:

$$\begin{aligned} \text{FDP} &= \frac{\#\left\{j: \beta_{j} = 0, W_{j} \geq -T\right\}}{\#\left\{j: W_{j} \geq T\right\} \vee 1} \cdot \frac{1 + \#\left\{j: \beta_{j} = 0, W_{j} \leq -T\right\}}{1 + \#\left\{j: \beta_{j} = 0, W_{j} \leq -T\right\}} \\ &\leq \frac{1 + \#\left\{j: W_{j} \leq -T\right\}}{\#\left\{j: W_{j} \geq T\right\} \vee 1} \cdot \frac{\#\left\{j: \beta_{j} = 0, W_{j} \geq T\right\}}{1 + \#\left\{j: \beta_{j} = 0, W_{j} \leq -T\right\}} \\ &\leq q \cdot 1 \end{aligned}$$

Then, we have the following theorem

Theorem 17.2.2: Property of the Knockoff Method

For any $q \in [0, 1]$, the **knockoff** method satisfies

$$\mathbb{E}\left[\frac{\#\{j: \beta_{j} = 0, j \in \hat{S}\}}{\#\{j: j \in \hat{S}\} + q^{-1}}\right] \le q$$

and the knockoff+ method satisfies

$$\mathbb{E}\left[\frac{\#\left\{j:\beta_{j}=0,j\in\hat{S}\right\}}{\#\left\{j:j\in\hat{S}\right\}}\right]\leq q$$

in both cases, teh expectation is taken over the Gaussian noise in the model, while treating original variables X and knockoffs \tilde{X} as fixed

References

Rina Foygel Barber and Emmanuel J. Candes. Controlling the false discovery rate via knockoffs. *Annals of Statistics*, 43(5):2055–2085, 2015.