

## Topic 20: Random Forest

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**Key points:** .

**Disclaimer:** The note is built on Prof. [Jinchi Lv](#)'s lectures of the course at USC, DSO 607, High-Dimensional Statistics and Big Data Problems.

### 20.1 Motivation

Denote by  $m(\mathbf{X})$  the measurable nonparametric regression function with  $p$ -dimensional random vector  $\mathbf{X}$  taking values in  $[0, 1]^p$ . The Random Forest algorithm aims to learn the regression function in a non-parametric way based on the observations  $\mathbf{x}_i \in [0, 1]^p$ ,  $y_i \in \mathbb{R}$ ,  $i = 1, \dots, n$ , from the model

$$y_i = m(\mathbf{x}_i) + \epsilon_i$$

where  $\mathbf{X}$ ,  $\mathbf{x}_i$ ,  $\epsilon_i$ ,  $i = 1, \dots, n$  are independent, and  $\{\mathbf{x}_i\}$  and  $\{\epsilon_i\}$  are two sequences of identically distributed random variables.  $\mathbf{x}_i$  is distributed identically as  $\mathbf{X}$ .

**Why Random Forest (RF)?** RF has gained significant popularity due to its

- **High accuracy:** RF consistently rank among the top performer, often surpassing more complex models
- **Robustness:** RF are less subject to overfitting due to the ensemble nature leveraging multiple decision trees
- **Interpretability:** RF provide rankings of feature importance

As illustrated in Figure 20.1, in a level-2 tree, each node (cell) defines the point where the current cell split and new cells are produced. The sets of features eligible for splitting cells at level  $k - 1$  are denoted as  $\Theta_k := \{\Theta_{k,1}, \dots, \Theta_{k,2^{k-1}}\}$ , where  $\Theta_{k,s} \subset \{1, \dots, p\}$ .



Figure 20.1: Level-2 Tree Example

Given any  $T$  (and the associated splitting criterion) and  $\Theta_{1:k}$ , the tree estimate denoted as  $\hat{m}_{T(\Theta_{1:k})}$  for a test

point  $\mathbf{c} \in [0, 1]^p$  is defined as

$$\hat{m}_{T(\Theta_{1:k})}(\mathbf{c}, \mathcal{X}_n) := \sum_{(\mathbf{t}_1, \dots, \mathbf{t}_k) \in T(\Theta_{1:k})} \mathbf{1}_{\mathbf{c} \in \mathbf{t}_k} \left( \frac{\sum_{i \in \{i: \mathbf{x}_i \in \mathbf{t}_k\}} y_i}{\# \{i: \mathbf{x}_i \in \mathbf{t}_k\}} \right)$$

where  $\mathcal{X}_n := \{\mathbf{x}_i, y_i\}_{i=1}^n$ , the fraction is defined as 0 when no sample is in the cell  $\mathbf{t}_k$ , and  $\mathbf{1}_{\mathbf{c} \in \mathbf{t}_k}$  is an indicator function = 1 if  $\mathbf{c} \in \mathbf{t}_k$  and = 0 otherwise.

## 20.2 Chi et al. (2022): High Dimensional RFs

For a RF model where

- a sequence of distinct  $\Theta_{1:k}$  results in a distinct tree
- every set of available features  $\Theta_{l,s}$ ,  $l = 1, \dots, k$ ;  $s = 1, \dots, 2^{l-1}$

**Column subsampling** Define a **column subsampling** procedure:  $\Theta_{l,s}, \forall l, s$  has  $\lceil \gamma_0 p \rceil$  distinct integers among  $1, \dots, p$ , with  $\lceil \cdot \rceil$  the ceiling function for some  $0 < \gamma_0 \leq 1$ .  $\gamma_0$  is the predetermined constant parameter of column subsampling. Introduce the boldface random mappings  $\Theta_{1:k}$ , which are independent and uniformly distributed over all possible  $\Theta_{1:k}$  for all integer  $k$ . Then random forests estimate for  $\mathbf{c}$  with observations  $\mathcal{X}_n$  is given by

$$\mathbb{E}(\hat{m}_{T(\Theta_{1:k})}(\mathbf{c}, \mathcal{X}_n) \mid \mathcal{X}_n) = \sum_{\Theta_{1:k}} \mathbb{P} \left( \bigcap_{s=1}^k \{\Theta_s = \Theta_s\} \right) \hat{m}_{T(\Theta_{1:k})}(\mathbf{c}, \mathcal{X}_n)$$

The expectation is taken over sets of available features.

**Observation resampling** Let  $A = \{a_1, \dots, a_B\}$  be a set of subsamples with each  $a_i$  consisting of  $\lceil bn \rceil$  observations (indices) drawn without replacement from  $\{1, \dots, n\}$  for some positive integer  $B$  and  $0 < b \leq 1$ ; in addition, each  $a_i$  is independent of model training. The default values of  $B$  and  $b$  are 500 and 0.632<sup>1</sup>. Then the tree estimate using subsample  $a$  is define as

$$\hat{m}_{T(\Theta_{1:k}),a}(\mathbf{c}, \mathcal{X}_n) := \sum_{(\mathbf{t}_1, \dots, \mathbf{t}_k) \in T(\Theta_{1:k})} \mathbf{1}_{\mathbf{c} \in \mathbf{t}_k} \left( \frac{\sum_{i \in a \cap \{i: \mathbf{x}_i \in \mathbf{t}_k\}} y_i}{\#(a \cap \{i: \mathbf{x}_i \in \mathbf{t}_k\})} \right)$$

the random forests estimate given  $A$  is then

$$B^{-1} \sum_{a \in A} \mathbb{E}[\hat{m}_{T,a}(\Theta_{1:k}, \mathbf{c}, \mathcal{X}_n) \mid \mathcal{X}_n] := B^{-1} \sum_{a \in A} \mathbb{E}[\hat{m}_{T(\Theta_{1:k}),a}(\mathbf{c}, \mathcal{X}_n) \mid \mathcal{X}_n]$$

**CART-split criterion** Given a cell  $\mathbf{t}$ , a subset of observation indices  $a$  and a set of available features  $\Theta \subset \{1, \dots, p\}$ , the CART-split is defined as

$$(\hat{j}, \hat{c}) = \arg \min_{j \in \Theta, c \in \{x_{ij}: \mathbf{x}_i \in \mathbf{t}, i \in a\}} \left[ \sum_{i \in a \cap P_L} (\bar{y}_L - y_i)^2 + \sum_{i \in a \cap P_R} (\bar{y}_R - y_i)^2 \right] \quad (20.1)$$

<sup>1</sup>Or,  $b = 1$  but observations are drawn with replacement.

where

$$\begin{aligned} P_L &:= \{i : \mathbf{x}_i \in \mathbf{t}, x_{ij} < c\} & P_R &:= \{i : \mathbf{x}_i \in \mathbf{t}, x_{ij} \geq c\} \\ \bar{y}_L &:= \sum_{i \in a \cap P_L} \frac{y_i}{\#(a \cap P_L)} & \bar{y}_R &:= \sum_{i \in a \cap P_R} \frac{y_i}{\#(a \cap P_R)} \end{aligned}$$

The CART-split criterion conditional on the sample is a deterministic splitting criterion; conditioning on another sample leads to another deterministic splitting criterion. Define  $\hat{T}_a$  as the sample tree growing rule that is associated with a splitting criterion following Eq. (20.1), the tree estimates using  $\hat{T}_a$  can be similarly defined as

$$\hat{m}_{\hat{T}_a(\Theta_{1:k})}(\mathbf{c}, \mathcal{X}_n) := \sum_{(\mathbf{t}_1, \dots, \mathbf{t}_k) \in \hat{T}_a(\Theta_{1:k})} \mathbf{1}_{\mathbf{c} \in \mathbf{t}_k} \left( \frac{\sum_{i \in \{i: \mathbf{x}_i \in \mathbf{t}_k\}} y_i}{\#\{i : \mathbf{x}_i \in \mathbf{t}_k\}} \right)$$

the definition is the same for  $\hat{m}_{\hat{T}_a, a}$ . Then the random forests estimate for a test point  $\mathbf{c} \in [0, 1]^p$  is given by

$$B^{-1} \sum_{a \in A} \mathbb{E} \left( \hat{m}_{\hat{T}_a, a}(\Theta_{1:k}, \mathbf{c}, \mathcal{X}_n) \mid \mathcal{X}_n \right)$$

where the average and conditional expectation correspond to the sample and column subsamplings respectively, and they are interchangeable.

**Bias-variance decomposition** For a tree growing rule  $T$  and  $\Theta_{1:k}$ , the population version is defined as

$$m_{T(\Theta_{1:k})}^*(\mathbf{c}) := \sum_{(\mathbf{t}_1, \dots, \mathbf{t}_k) \in T(\Theta_{1:k})} \mathbf{1}_{\mathbf{c} \in \mathbf{t}_k} \mathbb{E}(m(\mathbf{X}) \mid \mathbf{X} \in \mathbf{t}_k) \quad (20.2)$$

for each test point  $\mathbf{c} \in [0, 1]^p$ . And the  $\mathbb{L}^2$  prediction loss for random forests is defined as

$$\mathbb{E} \left[ m(\mathbf{X}) - B^{-1} \sum_{a \in A} \mathbb{E} \left( \hat{m}_{\hat{T}_a, a}(\Theta_{1:k}, \mathbf{X}, \mathcal{X}_n) \mid \mathbf{X}, \mathcal{X}_n \right) \right]^2 \quad (20.3)$$

if we use the full sample  $a = \{1, \dots, n\}$ , and denote  $\hat{T}_a$  and  $\hat{m}_{\hat{T}_a, a}$  as  $\hat{T}$  and  $\hat{m}_{\hat{T}}$ , the sample subsampling and average  $B^{-1} \sum_{a \in A} (\cdot)$  in the random forests estimate are no longer needed, then Eq.(20.3) can be simplified as

$$\mathbb{E} \left[ m(\mathbf{X}) - \mathbb{E}(\hat{m}_{\hat{T}}(\Theta_{1:k}, \mathbf{X}, \mathcal{X}_n) \mid \mathbf{X}, \mathcal{X}_n) \right]^2$$

By Jensen's inequality and Cauchy-Schwarz inequality,

$$\begin{aligned} & \frac{1}{2} \mathbb{E} \left[ m(\mathbf{X}) - \mathbb{E}(\hat{m}_{\hat{T}}(\Theta_{1:k}, \mathbf{X}, \mathcal{X}_n) \mid \mathbf{X}, \mathcal{X}_n) \right]^2 \\ & \leq \mathbb{E}[m(\mathbf{X})] \end{aligned}$$

Chi et al. (2022)

## References

Chien-Ming Chi, Patrick Vossler, Yingying Fan, and Jinchi Lv. Asymptotic properties of high-dimensional random forests. *The Annals of Statistics*, 50(6):3415–3438, 2022.