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## Topic 13: Non-convex Learning + Lasso

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**Key points**: Combining the best of the two, we can use **Lasso plus Concave** method, with Lasso screening and concave component selecting variables, achieving a coordinated intrinsic two-scale learning.

**Disclaimer**: The note is built on Prof. Jinchi Lv's lectures of the course at USC, DSO 607, High-Dimensional Statistics and Big Data Problems.

We are facing a tradeoff:

- **Convex** methods: have appealing <u>prediction power and oracle inequalities</u>, but challenging to provide tight false sign rate control
- **Concave** methods: have good <u>variable selection</u> properties, but challenging to establish <u>global</u> properties and risk properties

Here, we take advantage of the linearity of Lasso (convex *and* concave) and try to combine it with concave regularization to get the best of both.

### 13.1 Model Setup

Again, consider a linear regression model  $y = X\beta + \epsilon$ , where

- response vector  $(n \times 1)$ :  $\mathbf{y} = (y_1, \dots, y_n)'$
- design matrix  $(n \times p)$ : **X** =  $(\mathbf{x}_1, \dots, \mathbf{x}_p)$ , with each column rescaled to have  $L_2$ -norm  $n^{1/2}$

here, we consider a scenario where

- $\beta_0 = (\beta_{0,1}, \dots, \beta_{0,p})'$  is *sparse* (with many 0 components)
- ultra-**high** dimensions:  $\log p = O(n^a)$ , for some 0 < a < 1

and consider the penalized least squares

$$\min_{\beta \in \mathbb{R}^p} \left\{ (2n)^{-1} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda_0 \|\boldsymbol{\beta}\|_1 + \|p_{\lambda}(\boldsymbol{\beta})\|_1 \right\}$$
 (13.1)

where

- $\lambda_0 = c \left(\frac{\log p}{n}\right)^{1/2}$  for some c > 0
- $p_{\lambda}(\boldsymbol{\beta}) = p_{\lambda}(|\boldsymbol{\beta}|) = (p_{\lambda}(|\beta_1|), \dots, p_{\lambda}(|\beta_p|))'$ , with  $|\boldsymbol{\beta}| = (|\beta_1|, \dots, |\beta_p|)'$ ; the concave penalty  $p_{\lambda}(t)$  is defined on  $t \in [0, \infty)$ , indexed by  $\lambda \ge 0$ , increasing in **both** t and  $\lambda$ ,  $p_{\lambda}(0) = 0$

the 2 penalty components

- $L_1$ -component: minimum amount of regularization for removing noise in prediction
- concave component  $||p_{\lambda}(\beta)||_1$ : adapt model sparsity for *variable selection*

Under this set up, we can derive the hard-thresholding property as

#### Proposition 13.1.1: Hard-Thresholding Property

Assume the  $p_{\lambda}(t)$ ,  $t \ge 0$ , is **increasing and concave** with

- $p_{\lambda}(t) \ge p_{H,\lambda}(t) = \frac{1}{2} \left[ \lambda^2 (\lambda t)_+^2 \right]$  on  $[0, \lambda]$
- $p'_{\lambda}((1-c_1)\lambda) \le c_1\lambda$  for some  $c_1 \in [0,1)$
- $-p_{\lambda}''(t)$  decreasing on  $[0,(1-c_1)\lambda]$

then any <u>local minimizer</u> of 13.1 that is also a <u>global minimizer</u> in each coordinate has the **hard-thresholding** feature that each component is either 0 or of magnitude **larger** than  $(1 - c_1)\lambda$ 

Such property is shared by a wide class of concave penalties, including hard-thresholding penalty  $p_{H,\lambda}(t)$  with  $c_1 = 0$ ,  $L_0$ —penalty, and SICA (with suitable  $c_1$ ).

How to <u>understand</u> this proposition? Let  $\hat{\beta} = (\hat{\beta}_1, \dots, \hat{\beta}_p)'$ , then each  $\hat{\beta}_j$  is the glocal minimizer of the corresponding univariate penalized least-square problem along the j-th coordinate. These univariate problems share a common form with (generally) different scalars z

$$\hat{\beta}(z) = \arg\min_{\beta \in \mathbb{R}} \left\{ \frac{1}{2} (z - \beta)^2 + \lambda_0 |\beta| + p_{H,\lambda}(|\beta|) \right\}$$

after we rescale all covariates to have  $L_2$ -norm  $n^{1/2}$ . The solution to these univariate problems are

$$\hat{\beta}(z) = \operatorname{sgn}(z)(|z| - \lambda_0) \cdot \mathbf{1}_{|z| > \lambda + \lambda_0}$$

these solutions have the same feature as the hard-thresholded estimator: each component is either 0 or of magnitude larger than  $\lambda$ . This provides a better distinction between insignificant and significant covariates then soft-thresholding by  $L_1$  penalty.

With the hard-thresholding property of Prop. 13.1.1, we can prove a basic constraint for the global optimum  $\hat{\beta}$  on an event with significant probability (Fan and Lv, 2014)

$$\|\delta_2\|_1 \le 7\|\delta_1\|_1 \tag{13.2}$$

where  $\delta = \hat{\beta} - \beta_0 = (\hat{\beta}_1', \hat{\beta}_2')' - (\beta_{0,1}', \beta_{0,2}')' = (\delta_1', \delta_2')'$ , with  $\delta_1 \in \mathbb{R}^s$ . Where does this constraint come from? For the penalized least square quesion 13.1

$$\min_{\beta \in \mathbb{R}^p} \left\{ (2n)^{-1} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda_0 \|\boldsymbol{\beta}\|_1 + \|p_{\lambda}(\boldsymbol{\beta})\|_1 \right\}$$

the global minimizer  $\hat{\beta}$  leads to

$$(2n)^{-1} \|\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}\|_{2}^{2} + \lambda_{0} \|\hat{\boldsymbol{\beta}}\|_{1} + \|p_{\lambda}(\hat{\boldsymbol{\beta}})\|_{1} = (2n)^{-1} \|\mathbf{X}\boldsymbol{\beta}_{0} + \boldsymbol{\epsilon} - \mathbf{X}\hat{\boldsymbol{\beta}}\|_{2}^{2} + \lambda_{0} \|\hat{\boldsymbol{\beta}}\|_{1} + \|p_{\lambda}(\hat{\boldsymbol{\beta}})\|_{1}$$

$$= (2n)^{-1} \|\boldsymbol{\epsilon} - \mathbf{X}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_{0})\|_{2}^{2} + \lambda_{0} \|\hat{\boldsymbol{\beta}}\|_{1} + \|p_{\lambda}(\hat{\boldsymbol{\beta}})\|_{1}$$

$$\leq (2n)^{-1} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}_{0}\|_{2}^{2} + \lambda_{0} \|\boldsymbol{\beta}_{0}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1}$$

$$= (2n)^{-1} \|\boldsymbol{\epsilon}\|_{2}^{2} + \lambda_{0} \|\boldsymbol{\beta}_{0}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1}$$

then, plug in  $\delta = \hat{\beta} - \beta_0$ , we get

$$(2n)^{-1} \|\boldsymbol{\epsilon} - \mathbf{X}\boldsymbol{\delta}\|_{2}^{2} + \lambda_{0} \|\boldsymbol{\beta}_{0} + \boldsymbol{\delta}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1} \leq (2n)^{-1} \|\boldsymbol{\epsilon}\|_{2}^{2} + \lambda_{0} \|\boldsymbol{\beta}_{0}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1}$$

$$(2n)^{-1} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} - n^{-1}\boldsymbol{\epsilon}'\mathbf{X}\boldsymbol{\delta} + \lambda_{0} \|\boldsymbol{\beta}_{0} + \boldsymbol{\delta}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1} \leq \lambda_{0} \|\boldsymbol{\beta}_{0}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1}$$

since  $\beta_{0,2} = 0$ ,  $\delta_2 = \beta_{0,2} + \delta_2$ , we have

$$\|\boldsymbol{\beta}_0 + \boldsymbol{\delta}\|_1 = \|\boldsymbol{\beta}_{0,1} + \boldsymbol{\beta}_{0,2} + \boldsymbol{\delta}_1 + \boldsymbol{\delta}_2\|_1 = \|\boldsymbol{\beta}_{0,1} + \boldsymbol{\delta}_1 + \boldsymbol{\delta}_2\|_1 \le \|\boldsymbol{\beta}_{0,1} + \boldsymbol{\delta}_1\|_1 + \|\boldsymbol{\delta}_2\|_1$$

hence

$$(2n)^{-1}\|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} - n^{-1}\boldsymbol{\epsilon}'\mathbf{X}\boldsymbol{\delta} + \lambda_{0}\|\boldsymbol{\delta}_{2}\|_{1} \leq \lambda_{0}\|\boldsymbol{\beta}_{0,1}\|_{1} - \lambda_{0}\|\boldsymbol{\beta}_{0,1} + \boldsymbol{\delta}_{1}\|_{1} + \|\boldsymbol{p}_{\lambda}(\boldsymbol{\beta}_{0})\|_{1} - \|\boldsymbol{p}_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1}$$

and by the reverse triangle inequality  $\|\boldsymbol{\beta}_{0,1}\|_1 - \|\boldsymbol{\beta}_{0,1} + \boldsymbol{\delta}_1\|_1 \le \|\boldsymbol{\delta}_1\|_1$ , we get

$$(2n)^{-1} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} - n^{-1} \boldsymbol{\epsilon}' \mathbf{X} \boldsymbol{\delta} + \lambda_{0} \|\boldsymbol{\delta}_{2}\|_{1} \leq \lambda_{0} \|\boldsymbol{\delta}_{1}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1} - \|p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1}$$

If assume the distribution of the model error  $\epsilon$  as

$$\Pr\left(\left\|n^{-1}\mathbf{X}'\boldsymbol{\epsilon}\right\|_{\infty} > \frac{\lambda_0}{2}\right) = O\left(p^{-c_0}\right)$$

conditional on the event  $\mathcal{E} = \{\|n^{-1}\mathbf{X}'\boldsymbol{\epsilon}\|_{\infty} \leq \lambda_0/2\}$ , we have

$$-n^{-1}\epsilon'\mathbf{X}\delta + \lambda_0\|\delta_2\|_1 - \lambda_0\|\delta_1\|_1 \ge -\frac{\lambda_0}{2}\|\delta\|_1 + \lambda_0\|\delta_2\|_1 - \lambda_0\|\delta_1\|_1 = \frac{\lambda_0}{2}\|\delta_2\|_1 - \frac{3\lambda_0}{2}\|\delta_1\|_1$$

plug this result back, get

$$\frac{1}{2n} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} + \frac{\lambda_{0}}{2} \|\boldsymbol{\delta}_{2}\|_{1} \leq \frac{3\lambda_{0}}{2} \|\boldsymbol{\delta}_{1}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1} - \|p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1}$$
(13.3)

Now, if we further impose 2 conditions:

• **Condition 1 (eigenvalue condition)**: for some positive constant  $\kappa_0$ 

$$\min_{\|\boldsymbol{\delta}\|_2 = 1, \|\boldsymbol{\delta}\|_0 < 2s} \frac{1}{\sqrt{n}} \|\mathbf{X}\boldsymbol{\delta}\|_2 \ge \kappa_0$$
 (A)

$$\kappa = \kappa(s,7) = \min_{\delta \neq 0, \|\delta_2\|_1 \le 7\|\delta_1\|_1} \left\{ \frac{1}{\sqrt{n}} \frac{\|\mathbf{X}\boldsymbol{\delta}\|_2}{\|\delta_1\|_2 \vee \|\tilde{\delta}_2\|_2} \right\} > 0$$
 (B)

where  $\tilde{\delta}_2$  is the subvector of  $\delta_2$  consisting of the components with the *s* largest absolute values. Here

- Condition (A) is a mild sparse eigenvalue condition
- Condition **(B)** combines the restricted eigenvalue assumptions in Bickel et al. (2009)<sup>1</sup>. The intuition is, for OLS estimation, **X'X** should be **positive definite**, that is

$$\min_{0 \neq \delta \in \mathbb{R}^p} \left\{ \frac{1}{\sqrt{n}} \frac{\|\mathbf{X}\delta\|_2}{\|\delta\|_2} \right\} > 0$$

however, when p > n, this condition **never** holds, hence we replace  $\|\delta\|_2$  with the  $L_2$ -norm of  $\|\delta_1\|_2$ , a subvector of  $\delta$ 

$$\kappa = \kappa(s, 7) \min_{\delta \neq 0, \|\delta_2\|_1 \le 7\|\delta_1\|_1} \left\{ \frac{1}{\sqrt{n}} \frac{\|X\delta\|_2}{\|\delta\|_2} \right\} > 0$$

and for  $L_q$  loss with  $q \in (1,2]$ , we further bound  $\|\tilde{\delta}_2\|_2$ , which leads to condition (B).

<sup>&</sup>lt;sup>1</sup>Introduced by Candes and Tao (2007) for studying the oracle inequalities for the Lasso estimator and Dantzig selector.

• Condition 2 (hard-thresholding condition): The penalty  $p_{\lambda}(t)$  satisfies the conditions of Prop. 13.1.1 with

$$p_{\lambda}' \{ (1 - c_1)\lambda \} \le \lambda_0/4$$

$$\min_{j=1,\dots,s} |\beta_{0,j}| > \max \{ (1 - c_1)\lambda, 2\kappa_0^{-1} p_{\lambda}^{1/2}(\infty) \}$$

Now, look back at the condition 13.3, we can upper-bound  $\|p_{\lambda}(\boldsymbol{\beta}_0)\|_1 - \|p_{\lambda}(\boldsymbol{\beta}_0 + \boldsymbol{\delta})\|_1$  by  $\frac{1}{4n} \|\mathbf{X}\boldsymbol{\delta}\|_2^2 + \frac{1}{4}\lambda_0 \|\boldsymbol{\delta}\|_1$ . Consider 2 cases:

- Case 1:  $\|\hat{\boldsymbol{\beta}}\|_0 \ge s$ . By the hard-thresholding condition, we have  $|\beta_{0,j}| > (1-c_1)\lambda$  and  $p'_{\lambda}\{(1-c_1)\lambda\} \le \lambda_0/4$ . Hence, for  $j = 1, \dots, s$ ,
  - if  $\hat{\beta}_j \neq 0$ , we must have  $|\hat{\beta}_j| > (1 c_1)\lambda$ . And by the <u>mean-value theorem</u>, we have

$$|p_{\lambda}(|\beta_{0,j}|) - p_{\lambda}(|\hat{\beta}_{j}|)| = p'_{\lambda}(b)(|\hat{\beta}_{j}| - |\beta_{0,j}|) \le p'_{\lambda}(b)|\delta_{0,j}|$$

where b is between  $|\beta_{0,j}|$  and  $|\hat{\beta}_j|$ , hence,  $b > |\beta_{0,j}| > (1 - c_1)\lambda$ , by the concavity of  $p_{\lambda}$ , we have  $p'(b) < p'((1 - c_1)\lambda) \le \lambda_0/4$ , which leads to  $|p_{\lambda}(|\beta_{0,j}|) - p_{\lambda}(|\hat{\beta}_j|)| \le \frac{1}{4}\lambda_0|\delta_j|$ .

- if  $\hat{\beta}_i = 0$ , since  $\|\hat{\beta}_0\| \ge s$ , there must exist some j' > s s.t.  $\hat{\beta}_{j'} \ne 0$ , similarly

$$\begin{aligned} |p_{\lambda}(|\beta_{0,j}|) - p_{\lambda}(|\hat{\beta}_{j'}|)| &\leq |p_{\lambda}\left(|\beta_{0,j}|\right) - p_{\lambda}\left((1-c_{1})\lambda\right)| + |p_{\lambda}\left(|\hat{\beta}_{j'}|\right) - p_{\lambda}\left((1-c_{1})\lambda\right)| \\ &= p'_{\lambda}(b_{1})\left(|\beta_{0,j}| - (1-c_{1})\lambda\right) + p'_{\lambda}(b_{2})\left(|\hat{\beta}_{j'}| - (1-c_{1})\lambda\right) \\ &\leq p'_{\lambda}(b_{1})\left(|\beta_{0,j}| - \underbrace{|\hat{\beta}_{j}|}_{=0}\right) + p'_{\lambda}(b_{2})\left(|\hat{\beta}_{j'}| - \underbrace{|\beta_{0,j'}|}_{=0}\right) \\ &= p'_{\lambda}(b_{1})|\delta_{j}| + p'_{\lambda}(b_{2})|\delta_{j'}| \leq \frac{\lambda_{0}}{4}\left(|\delta_{j}| + |\delta_{j'}|\right) \end{aligned}$$

together, we have

$$||p_{\lambda}(\boldsymbol{\beta}_{0})||_{1} - ||p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})||_{1} \leq \frac{1}{4}\lambda_{0}||\boldsymbol{\delta}||_{1} \leq \frac{1}{4n}||\mathbf{X}\boldsymbol{\delta}||_{2}^{2} + \frac{1}{4}\lambda_{0}||\boldsymbol{\delta}||_{1}$$

• Case 2:  $\|\hat{\boldsymbol{\beta}}\|_0 = s - k$  for some  $k \ge 1$ . Then we must have  $\|\boldsymbol{\delta}\|_0 \le \|\hat{\boldsymbol{\beta}}\|_0 + \|\boldsymbol{\beta}_0\|_0 \le s - k + s < 2s$ , and  $\|\boldsymbol{\delta}\|_2 \ge \sqrt{k} \min_{j=1,\dots,s} |\beta_{0,j}|$ . Also, there are at least k null estimates  $(\hat{\beta}_j = 0)$ , thus

$$\underbrace{\frac{1}{4n} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} \geq \frac{\kappa_{0}^{2}}{4} \|\boldsymbol{\delta}\|_{2}^{2}}_{\text{Condition 1(A)}} \geq \underbrace{\frac{\kappa_{0}^{2}}{4} \left(\sqrt{k} \min_{j=1,\cdots,s} |\beta_{0,j}|\right)^{2}}_{\text{Condition 2}} \geq k p_{\lambda}(|\beta_{0,j}|)$$

similar to Case 1, we have the desired upper bound

$$||p_{\lambda}(\boldsymbol{\beta}_{0})||_{1} - ||p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})||_{1} \leq kp_{\lambda}(\infty) + \frac{1}{4}\lambda_{0}||\boldsymbol{\delta}||_{1} \leq \frac{1}{4n}||\mathbf{X}\boldsymbol{\delta}||_{2}^{2} + \frac{1}{4}\lambda_{0}||\boldsymbol{\delta}||_{1}$$

Combining Case 1 and 2, we have  $\|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1} - \|p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1} \le \frac{1}{4n} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} + \frac{1}{4}\lambda_{0}\|\boldsymbol{\delta}\|_{1}$ , plug this back in 13.3, get

$$\frac{1}{n} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} + \lambda_{0} \|\boldsymbol{\delta}_{2}\|_{1} \leq 3\lambda_{0} \|\boldsymbol{\delta}_{1}\|_{1} + \|p_{\lambda}(\boldsymbol{\beta}_{0})\|_{1} - \|p_{\lambda}(\boldsymbol{\beta}_{0} + \boldsymbol{\delta})\|_{1} 
\leq 3\lambda_{0} \|\boldsymbol{\delta}_{1}\|_{1} + \frac{1}{2n} \|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} + \frac{1}{2}\lambda_{0} \underbrace{\|\boldsymbol{\delta}\|_{1}}_{=\|\boldsymbol{\delta}_{1}\|_{1} + \|\boldsymbol{\delta}_{2}\|_{1}} 
\leq 7\lambda_{0} \|\boldsymbol{\delta}_{1}\|_{1}$$

which leads to the constraint in 13.2 and  $\frac{1}{n} \|\mathbf{X}\boldsymbol{\delta}\|_2^2 \le 7\lambda_0 \|\boldsymbol{\delta}_1\|_1$ .

Now, look back at Condition 1(B)

$$\kappa = \kappa(s,7) = \min_{\delta \neq 0, \|\delta_2\|_1 \le 7\|\delta_1\|_1} \left\{ \frac{1}{\sqrt{n}} \frac{\|\mathbf{X}\delta\|_2}{\|\delta_1\|_2 \vee \|\tilde{\delta}_2\|_2} \right\} > 0$$

we have

$$\frac{1}{4}\kappa^{2}(s,7)\|\boldsymbol{\delta}_{1}\|_{2}^{2} \leq \frac{1}{4}\kappa^{2}(s,7)\left(\|\boldsymbol{\delta}_{1}\|_{2}^{2} \vee \|\tilde{\boldsymbol{\delta}}_{2}\|_{2}^{2}\right) \leq \frac{1}{4n}\|\mathbf{X}\boldsymbol{\delta}\|_{2}^{2} \leq \underbrace{\frac{7}{4}\lambda_{0}\|\boldsymbol{\delta}_{1}\|_{1} \leq \frac{7}{4}\lambda_{0}\sqrt{s}\|\boldsymbol{\delta}_{1}\|_{2}}_{Cauchy-Schwartz inequality}$$

hence

$$\|\delta_1\|_2 \le \frac{7\lambda_0\sqrt{s}}{\kappa^2(s,7)} \qquad \|\delta_1\|_1 \le \sqrt{s}\|\delta_1\|_1 \le \frac{7\lambda_0s}{\kappa^2(s,7)} \qquad \|\delta_2'\|_2 \le \frac{\sqrt{7\lambda_0\sqrt{s}\|\delta_1\|_2}}{\kappa(s,7)}$$

Notice that the k-th largest absolute component of  $\delta_2$  is bounded from above by  $\|\delta_2\|_1/k$ , then for  $\delta_{2_s}$ , the subvector of  $\delta_2$  consisting of components **excluding** those with the s largest magnitudes, we have

$$\|\boldsymbol{\delta}_{2_{s}}\|_{2}^{2} \leq \sum_{k=s+1}^{p-s} \frac{1}{k^{2}} \|\boldsymbol{\delta}_{2}\|_{1}^{2} \leq s^{-1} \|\boldsymbol{\delta}_{2}\|_{1}^{2} \Rightarrow \|\boldsymbol{\delta}_{2_{s}}\|_{2} \leq \frac{1}{\sqrt{s}} \|\boldsymbol{\delta}_{2}\|_{1}^{13.2} \stackrel{\text{C-S}}{\leq} 7 \|\boldsymbol{\delta}_{1}\|_{2}$$

since  $\delta_{2_s}$  and  $\delta_2'$  are a partition of  $\delta$ , we have

$$\|\delta_2\|_2 \le \|\delta_{2_s}\|_2 + \|\delta_2'\|_2 \le 7\|\delta_1\|_2 + \frac{\sqrt{7\lambda_0\sqrt{s}}\|\delta_1\|_2}{\kappa(s,7)} \le 7\|\delta_1\|_2 \le \frac{56\lambda_0\sqrt{s}}{\kappa^2(s,7)}$$

Hence, we have the following theorem:

#### Theorem 13.1.2: 1

Assume that Condition 1 and 2 and the model error bound  $\Pr\left(\left\|n^{-1}X'\boldsymbol{\epsilon}\right\|_{\infty} > \frac{\lambda_0}{2}\right) = O\left(p^{-c_0}\right)$ , and  $p_{\lambda}(t)$  is continuously differentiable. Then the global minimizer  $\hat{\boldsymbol{\beta}}$  of 13.1 has the hard-thresholding property stated in Prop.13.1.1, and, with probability  $1 - O(p^{-c_0})$ , statisfies simultaneously that

$$\frac{1}{2} \left\| \mathbf{X} \left( \hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0 \right) \right\|_2 = O(\kappa^{-1} \lambda_0 s^{1/2}) \tag{13.4}$$

$$\|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0\|_q = O(\kappa^{-2}\lambda_0 s^{1/q}), \qquad q \in [1, 2]$$
 (13.5)

$$FS\left(\hat{\boldsymbol{\beta}}\right) = O\left(\kappa^{-4} \left(\frac{\lambda_0}{\lambda}\right)^2 s\right) \tag{13.6}$$

If in addition  $\lambda \geq \frac{56\lambda_0\sqrt{s}}{(1-c_1)\kappa^2}$ , then with probability

# References

Peter J Bickel, Ya'acov Ritov, and Alexandre B Tsybakov. Simultaneous analysis of lasso and dantzig selector. *The Annals of Statistics*, pages 1705–1732, 2009.

Emmanuel Candes and Terence Tao. The dantzig selector: Statistical estimation when p is much larger than n. *The Annals of Statistics*, 35(6):2313–2351, 2007.

Yingying Fan and Jinchi Lv. Asymptotic properties for combined 11 and concave regularization. *Biometrika*, 101(1):57–70, 2014.