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## Topic 3: Moving the Goalposts Approach

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Key points:

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**Disclaimer**: These notes are written by Sai Zhang (email me or check my Github page). The main reference for this topic is Armstrong, Kolesár, and Kwon (2020), I thank Prof. Armstrong for his valuable advice.

## 3.1 Finite Sample Bias-Variance Tradeoffs

## 3.1.1 **Setup**

Consider the fixed design regression model

$$y_i = w_i \beta(z_i) + h(z_i) + \epsilon_i \tag{3.1}$$

where

- $w_i$ ,  $z_i$  are treated as **fixed**
- $\epsilon_i$  is **independent**, with  $\mathbb{E}[\epsilon_i] = 0$ ,  $\mathbb{E}[\epsilon_i^2] = \sigma_i^2$
- observation:  $\left\{ \left( y_i, w_i, z_i' \right)' \right\}_{i=1}^n$

one example is the case where  $w_i$  is **binary**, then

$$\beta(z) = f(1, z) - f(0, z)$$

which is just the ATE conditional on z under the unconfoundedness assumption. This includes the RD design, where  $z_i$  is the running variable and  $w_i$  is the treatment assignment.

Now, consider for the weighted average treatment effect

$$L_{\mu}\left[\beta(\cdot)\right] = \int \beta(z) \mathrm{d}\mu(z)$$

where  $\int \mu(z) = 1$  is a **signed** measure (weight, allowing **negative** weights), construct a linear estimator

$$\hat{L}_a = \sum_{i=1}^n a_i y_i$$

where the estimation weights  $a_i$  can depend on  $\{z_i, w_i, \sigma_i^2\}_{i=1}^n$ , but **not** on  $y_i$ . Together, the bias of  $\hat{L}_a$  for  $L_{\mu}\left[\beta(\cdot)\right]$ , given the regression function  $\beta(\cdot)$ ,  $h(\cdot)$ , is

$$\mathbb{E}_{\beta(\cdot),h(\cdot)}\left[\hat{L}_a\right] - L_{\mu}\left[\beta(\cdot)\right] = \sum_{i=1}^n a_i \left[w_i\beta(z_i) + h(z_i)\right] - \int \beta(z) d\mu(z)$$

and its variance, given the regression function  $\beta(\cdot)$ ,  $h(\cdot)$ , is just

$$\operatorname{Var}_{\beta(\cdot),h(\cdot)}\left[\hat{L}_{a}\right] = \sum_{i=1}^{n} a_{i}^{2} \sigma_{i}^{2}$$

To bound the bias, assume  $h(\cdot)$  is known to belong in a class of functions  $\mathcal{H}$ , then two approaches can be adopted, for the regularity of  $\beta(\cdot)$  and the choice of  $\mu(\cdot)$ :

1 arbitrary  $\beta(\cdot)$ , optimizing weights  $\mu$  by *moving the goalposts*, s.t.  $L_{\mu}\left[\beta(\cdot)\right]$  is easy to estimate (Crump et al., 2006; Imbens and Wager, 2019) which gives the worst-case bias

$$\inf_{\mu} \sup_{\beta(\cdot),h(\cdot)} \left| \sum_{i=1}^{n} a_i \left[ w_i \beta(z_i) + h(z_i) \right] - \int \beta(z) d\mu(z) \right| \qquad \text{s.t. } h(\cdot) \in \mathcal{H}, \int d\mu(z) = 1 \qquad (3.2)$$

2 assume constant treatment effects, i.e.,  $\beta(z) = \beta$ ,  $\forall z$ , which means that  $L_{\mu} \left[ \beta(\cdot) \right] = \beta$  regardless of  $\mu$  (Armstrong et al., 2020), and the worst-case bias is

$$\sup_{\beta, h(\cdot)} \left| \sum_{i=1}^{n} a_i \left[ w_i \beta + h(z_i) \right] - \beta \right| \qquad \text{s.t. } h(\cdot) \in \mathcal{H}$$
 (3.3)

And, the two approaches can be linked as such:

• If  $\sum_{i=1}^{n} a_i w_i = 1$ , 3.2 and 3.3 are both equal to

$$\sup_{h(\cdot)} \left| \sum_{i=1}^{n} a_i h(z_i) \right| \text{ s.t. } h(\cdot) \in \mathcal{H}$$
 (3.4)

- 3.2 automatically equals 3.4
- 3.3 is optimized (w.r.t.  $\mu$ ) by setting  $\mu$  to place weight  $a_i w_i$  on observation i, i.e.,  $\mu(\mathcal{Z}) = \sum_{i:z_i \in \mathcal{Z}} a_i w_i$
- Otherwise, 3.2 and 3.3 are both infinite:
  - 3.3 can be made arbitrarily large by choosing large enough  $\beta$
  - 3.2 can be made arbitrarily large by making  $\beta(\cdot)$  constant (as in 3.3) and large enough

## References

Timothy B Armstrong, Michal Kolesár, and Soonwoo Kwon. Bias-aware inference in regularized regression models. *arXiv preprint arXiv:*2012.14823, 2020.

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Guido Imbens and Stefan Wager. Optimized regression discontinuity designs. *Review of Economics and Statistics*, 101(2):264–278, 2019.