**Econometrics** 

February 23, 2023

# Topic 11: Lasso And Beyond: Convex Learning

by Sai Zhang

Key points:

Disclaimer:

## 11.1 Lasso

Lasso (Least absolute Shrinkage and Selection Operator), proposed by Tibshirani (1996), aims to minimize the SSR (sum of residual squares) subject to the L1-norm (sum of the absolute value) of the coefficients being less than a constant.

### 11.1.1 Set up

For data  $(\mathbf{x}_i, y_i)_{i=1}^n$ , where

-  $y_i$  is the outcome for individual i

-  $\mathbf{x}_i = (x_{i1}, \dots, x_{ip})'$  is the  $p \times 1$  vector of predictors

Then the Lasso estimator  $(\hat{\alpha},\hat{\beta})$  is defined as

$$\left(\hat{\alpha}, \hat{\boldsymbol{\beta}}\right) = \arg\min_{\alpha, \boldsymbol{\beta}} \left\{ \sum_{i=1}^{n} \left( y_i - \alpha - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\}$$
 s.t. 
$$\sum_{j=1}^{p} |\beta_j| \le t$$

for the  $n \times 1$  response vector  $\mathbf{y} = (y_1, \dots, y_n)'$ , the  $n \times p$  design matrix  $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_n)'$  where  $\mathbf{x}_i = (x_{i1}, \dots, x_{ip})'$  is a  $p \times 1$  vector. Here  $\hat{\alpha} = \overline{y}$ , w.l.o.g., let  $\overline{y} = 0$  and omit  $\alpha$  for simplicity.

In matrix form, we have

• constrained form:

$$\hat{\boldsymbol{\beta}} = \arg\min_{\boldsymbol{\beta} \in \mathbb{R}^p} \left\{ \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 \right\}$$
 s.t.  $\|\boldsymbol{\beta}\|_1 \le t$ 

unconstrained form:

$$\hat{\boldsymbol{\beta}}(\lambda) = \arg\min_{\boldsymbol{\beta} \in \mathbb{R}^p} \left\{ \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1 \right\}$$

where the regularization parameter  $\lambda \geq 0$ :

- 
$$\lambda \to \infty$$
:  $\hat{\boldsymbol{\beta}}_{lasso} = \mathbf{0}$   
-  $\lambda = 0$ :  $\hat{\boldsymbol{\beta}}_{lasso} \to \hat{\boldsymbol{\beta}}_{OLS}$ 

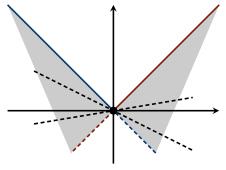
#### 11.1.2 **Solving Lasso**

Lasso is essentially a quadratic optimization problem. Hence, the solution is given by taking the derivative (of the unconstrainted question) and set it equal to 0

$$\frac{\mathrm{d}}{\mathrm{d}\boldsymbol{\beta}} \left( \frac{1}{2n} \| \mathbf{y} - \mathbf{X}\boldsymbol{\beta} \|_{2}^{2} + \lambda \| \boldsymbol{\beta} \|_{1} \right) = 0$$

$$\Rightarrow \frac{1}{n} \underbrace{\mathbf{X'}}_{p \times n} \underbrace{\left( \mathbf{y} - \mathbf{X}\boldsymbol{\beta} \right)}_{=\epsilon, n \times 1} = \lambda \begin{cases} \operatorname{sign} \left( \beta_{j} \right), & \beta_{j} \neq 0 \\ [-1, 1], & \beta_{j} = 0 \end{cases}$$

this result follows the fact the L-1 norm  $\|\beta\|$  is piecewise linear:



L1-norm (1-dimension)

For each component of the vector of the L-1 norm  $f(\beta_i) = |\beta_i|$ , we have:

$$-\beta_{i} > 0$$
:  $f'(\beta_{i}) = 1$ 

- 
$$\beta_i < 0$$
:  $f'(\beta_i) = -1$ 

-  $β_i$  = 0: d f ∈ [-1, 1] (shaded area) which gives the results stated above.

Take another look at this result

#### Proposition 11.1.1: Lasso Parameter Selection Rule

$$\frac{1}{n}\mathbf{X}'\left(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\right) = \frac{1}{n}\mathbf{X}'\boldsymbol{\epsilon} = \lambda \begin{cases} \operatorname{sign}\left(\beta_{j}\right), & \beta_{j} \neq 0 \\ [-1, 1], & \beta_{j} = 0 \end{cases}$$

which gives a parameter selection criterion: for  $\beta_i \neq 0$ ,  $sign(\beta_i)$  must agree with  $Corr(X_i, \epsilon)$ , the correlation between the *j*-th variable  $X_i$  and (full-model) residuals  $\epsilon = y - X\beta$ .

#### 11.1.3 Algorithm: from LARS to Lasso

Mathematically, Lasso is quite intuitive, but computionally, it can be quite consuming. Efron et al. (2004) propose an alogrithm that takes steps from a all-0 model to the biggest model (OLS), that is, Least Angle Regression (LARS).

#### **Forward Selection**

## 11.2 Penalized Least Square Estimation

Lasso is one special class of Penalized Least Square (PLS) Estimation. For the linear regression model  $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$ , if  $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_n)$ , we have PLS as

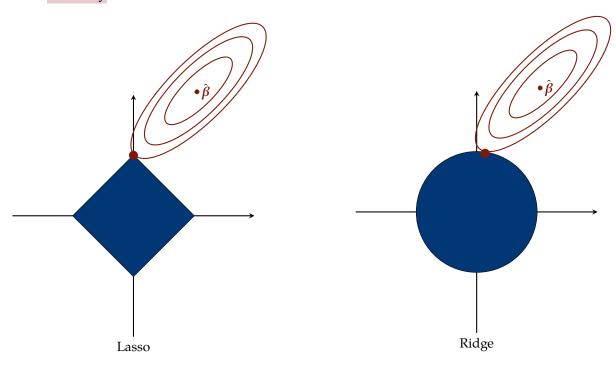
$$\min_{\boldsymbol{\beta} \in \mathbb{R}^p} \left\{ \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \sum_{j=1}^p p_{\lambda} (|\beta_j|) \right\}$$

where  $p_{\lambda}(\cdot)$  is a penalty function indexed by the regularization parameter  $\lambda \geq 0$ . Antoniadis and Fan (2001) showed that the PLS estimator  $\hat{\beta}$  has the following properties:

- sparsity: if  $\min_{t\geq 0} \left\{ t + p'_{\lambda}(t) \right\} > 0$
- approximate unbiasedness: if  $p'_{\lambda}(t) = 0$  for t large enough
- **continuity**: iff  $\arg\min_{t\geq 0} \{t + p'_{\lambda}(t)\} = 0$

In general

- the **sigularity** of penalty function at the origin,  $p'_{\lambda}(0_{+}) > 0$  is needed for generating **sparsity** in variable selection
- the **concavity** is needed to reduce the bias



# References

Anestis Antoniadis and Jianqing Fan. Regularization of wavelet approximations. *Journal of the American Statistical Association*, 96(455):939–967, 2001.

Bradley Efron, Trevor Hastie, Iain Johnstone, and Robert Tibshirani. Least angle regression. *The Annals of Statistics*, 32(2):407 – 499, 2004. doi: 10.1214/009053604000000067. URL https://doi.org/10.1214/009053604000000067.

Robert Tibshirani. Regression shrinkage and selection via the lasso. *Journal of the Royal Statistical Society: Series B (Methodological)*, 58(1):267–288, 1996.