

## Topic 17: False Discovery Rate (FDR) and Knockoffs

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**Key points:** Constructing knockoff variables to control FDR when estimating regression coefficients.

**Disclaimer:** The note is built on Prof. [Jinchi Lv](#)'s lectures of the course at USC, DSO 607, High-Dimensional Statistics and Big Data Problems.

### 17.1 Motivation

Consider the classical linear regression setting

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

where  $\boldsymbol{\beta} \in \mathbb{R}^p$  is the unknown vector of coefficients and  $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$ . In a high-dimensional problem, we would like to just select a subset of all variables  $\hat{S} \subset \{1, \dots, p\}$  s.t. conditional on  $\{\mathbf{X}_j\}_{j \in \hat{S}}$ ,  $\mathbf{y}$  is **independent** of all other variables, we can define the **False Discovery Rate (FDR)** in can be defined as

#### Definition 17.1.1: False Discovery Rate (FDR)

$$\text{FDR} = \mathbb{E}(\text{FDP}) = \mathbb{E} \left[ \frac{|\hat{S} \cap \mathcal{H}_0|}{|\hat{S}|} = \frac{\#\{j : j \in \hat{S} \setminus S\}}{\#\{j : j \in \hat{S}\}} \right]$$

where  $\mathcal{H}_0 \subset \{1, \dots, p\}$  is the set of **null** variables:  $\mathbf{X}_j$  is **null** iff  $\mathbf{Y}$  is independent of  $\mathbf{X}_j$  conditional on the other variables  $\mathbf{X}_{-j} = \{\mathbf{X}_1, \dots, \mathbf{X}_p\} \setminus \{\mathbf{X}_j\}$ .

In this note, we consider a series of knockoff-based methods to control FDR. They all follow a common procedure:

- **Step 1:** Construct Knockoffs
- **Step 2:** Calculate test statistics for both original and knockoff variables
- **Step 3:** Calculate a threshold for the test statistics, controlling for a desired FDR level
- **Step 4:** Select variables that pass the threshold

### 17.2 Barber and Candès (2015)

**Constructing the knockoffs** [Barber and Candès \(2015\)](#) construct the knockoffs by the following procedure

- Calculate the Gram matrix  $\boldsymbol{\Sigma} = \mathbf{X}'\mathbf{X}$  for the normalized original variables, where  $\Sigma_{jj} = \|\mathbf{X}_j\|_2^2 = 1$

- Construct the knockoffs  $\tilde{\mathbf{X}}$  s.t.

$$\tilde{\mathbf{X}}' \tilde{\mathbf{X}} = \Sigma \quad \mathbf{X}' \tilde{\mathbf{X}} = \Sigma - \text{diag}\{\mathbf{s}\}$$

where  $\mathbf{s} \in \mathbb{R}_+^p$  is a  $p$ -dimensional non-negative vector (larger  $s_j$  indicates higher power) and

- $\tilde{\mathbf{X}}$  exhibits the **same** covariance structure as the original design  $\mathbf{X}$
- The correlation between distinct original variables and knockoffs are the same as between the originals:

$$\mathbf{X}'_j \tilde{\mathbf{X}}_k = \mathbf{X}'_j \mathbf{X}_k, \quad \forall j \neq k$$

- The correlation between the original variables and their own knockoffs is **less than 1**

$$\mathbf{X}'_j \tilde{\mathbf{X}}_j = \Sigma_{jj} - s_j = 1 - s_j$$

To construct such knockoffs,

- Given a proper  $\mathbf{s}$ , if  $n \geq 2p$ , then

$$\tilde{\mathbf{X}} = \mathbf{X}(\mathbf{I} - \Sigma^{-1} \text{diag}\{\mathbf{s}\}) + \tilde{\mathbf{U}}\mathbf{C}$$

where  $\tilde{\mathbf{U}} \in \mathbb{R}^{n \times p}$  is an **orthonormal** matrix s.t.  $\tilde{\mathbf{U}}' \mathbf{X} = \mathbf{0}$  and  $\mathbf{C}' \mathbf{C} = 2 \text{diag}\{\mathbf{s}\} - \text{diag}\{\mathbf{s}\} \Sigma^{-1} \text{diag}\{\mathbf{s}\} \geq \mathbf{0}$

- A sufficient and necessary condition for  $\tilde{\mathbf{X}}$  to exist:  $\text{diag}\{\mathbf{s}\} \leq 2\Sigma$

2 types of knockoffs can be constructed, following these procedures

T1 **Equi-correlated** knockoffs: set  $s_j = 2\lambda_{\min}(\Sigma) \wedge 1$  for all  $j$ , then  $\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle = 1 - 2\lambda_{\min}(\Sigma) \wedge 1$  for all  $j$ . This is essentially minimizing  $|\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle|$

T2 **SDP** knockoffs: solve the convex problem

$$\arg \min_{\mathbf{s}} \sum_j (1 - s_j) \quad \text{s.t. } 0 \leq s_j \leq 1, \text{diag}\{\mathbf{s}\} \leq 2\Sigma$$

which is essentially minimizing the average of  $\langle \mathbf{X}_j, \tilde{\mathbf{X}}_j \rangle$

**Calculate test statistics** Define and calculate test statistics  $W_j$  for each  $\beta_j \in \{1, \dots, p\}$  using  $[\mathbf{X} \quad \tilde{\mathbf{X}}]$ :

- the test statistic  $W_j$  should be constructed s.t. large positive values are evidence against the null hypothesis  $\beta_j = 0$ , for example, consider a Lasso on  $[\mathbf{X} \quad \tilde{\mathbf{X}}]$

$$\hat{\beta}(\lambda) = \arg \min_{\mathbf{b}} \left\{ \frac{1}{2} \|\mathbf{y} - [\mathbf{X} \quad \tilde{\mathbf{X}}] \mathbf{b}\|_2^2 + \lambda \|\mathbf{b}\|_1 \right\}$$

where  $\lambda$  is the point on the Lasso path at which the feature enters the model as

$$Z_j = \sup \{ \lambda : \hat{\beta}_j(\lambda) \neq 0 \}$$

$$\text{and set } W_j = (Z_j \vee \tilde{Z}_j) \cdot \begin{cases} +1, & Z_j > \tilde{Z}_j \\ -1, & Z_j < \tilde{Z}_j \end{cases}$$

- In general, the statistics  $W$  should satisfy the **sufficient** property and **anti-symmetry** property:

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<sup>1</sup>Other choices of  $W_j$  are  $W_j = |\mathbf{X}'_j \mathbf{y}| - |\tilde{\mathbf{X}}'_j \mathbf{y}|$ , or  $|\hat{\beta}_j^{\text{LS}}| - |\hat{\beta}_{j+p}^{\text{LS}}|$

**Definition 17.2.1: Property of Test Statistics  $W_j$** 

The test statistic  $W_j$  is said to obey

- the **sufficient** property if  $\mathbf{W}$  depends only on the Gram matrix and on feature-response inner products, that is

$$\mathbf{W} = f\left([\mathbf{X} \quad \tilde{\mathbf{X}}]' [\mathbf{X} \quad \tilde{\mathbf{X}}], [\mathbf{X} \quad \tilde{\mathbf{X}}]' \mathbf{y}\right)$$

- the **antisymmetry** property if swapping the original  $\mathbf{X}_j$  and its knockoff  $\tilde{\mathbf{X}}_j$  has the effect of **switching the sign** of  $W_j$ , that is

$$W_j(Z_j, \tilde{Z}_j) = -W_j(\tilde{Z}_j, Z_j)$$

## References

Rina Foygel Barber and Emmanuel J. Candès. Controlling the false discovery rate via knockoffs. *Annals of Statistics*, 43(5):2055–2085, 2015.