

Noise-Induced Randomization in Regression Discontinuity Designs

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Outline

- 1 Introduction
- 2 Key Argument
- 3 Estimation
- 4 Confidence Intervals
- 5 Applications
- 6 Discussion

Introduction

RD Identification

Z_i
running variable

RD Identification

Z_i
running variable

W_i
treatment

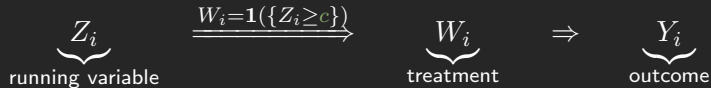
RD Identification

$$\underbrace{Z_i}_{\text{running variable}} \xrightarrow{W_i = \mathbf{1}(\{Z_i \geq c\})} \underbrace{W_i}_{\text{treatment}}$$

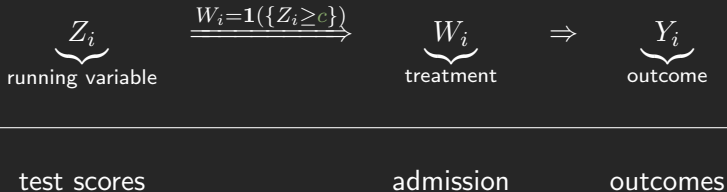
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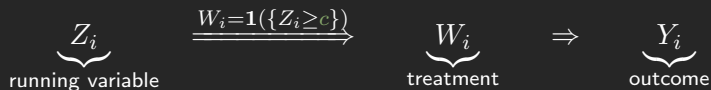
RD Identification



RD Identification



RD Identification



test scores
test results

admission
medication

outcomes
outcomes

RD Identification: Continuity Argument

For potential outcomes $\{Y_i(0), Y_i(1)\}$: $Y_i = Y_i(W_i)$, a weighted **causal effect** can be identified as

$$\tau_c = \mathbb{E}[Y_i(1) - Y_i(0) \mid Z_i = c]$$

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assuming

- the conditional response functions $\mu_w(z) = \mathbb{E}[Y(w) \mid Z = z]$ are continuous

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assuming

- the conditional response functions $\mu_w(z) = \mathbb{E}[Y(w) \mid Z = z]$ are continuous
- $\mu_w(z)$ to have a uniformly bounded 2nd derivative for CIs (Armstrong and Kolesár, 2018, 2020)

RD Identification: Problems of Continuity Argument

Assumption: continuous $\mu_w(z) = \mathbb{E}[Y(w) \mid Z = z]$

$$\tau_c = \lim_{z \downarrow c} \mathbb{E}[Y \mid Z = z] - \lim_{z \uparrow c} \mathbb{E}[Y \mid Z = z]$$

Where does this continuity come from?

RD Identification: Problems of Continuity Argument

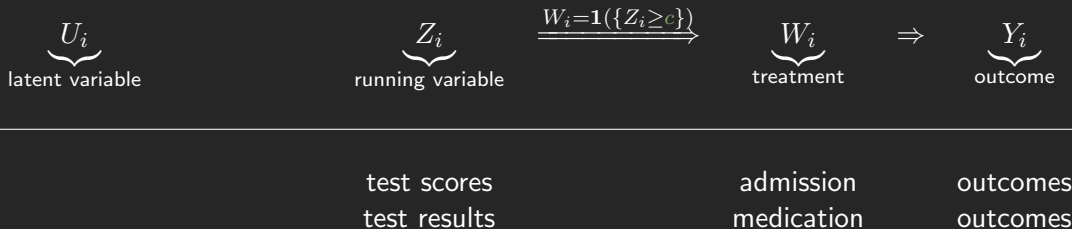
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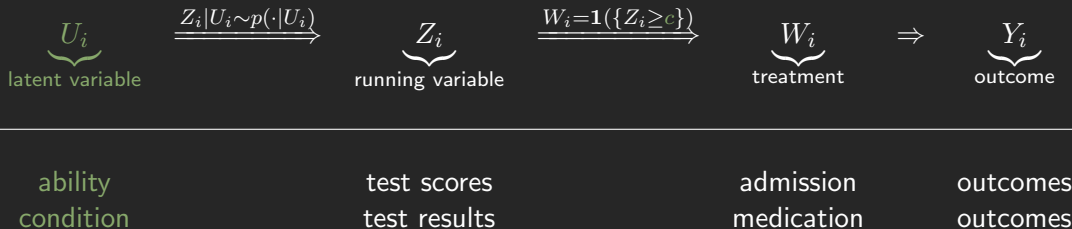
Where does this continuity come from?

Lee (2008): continuous measurement error in the running variable by units

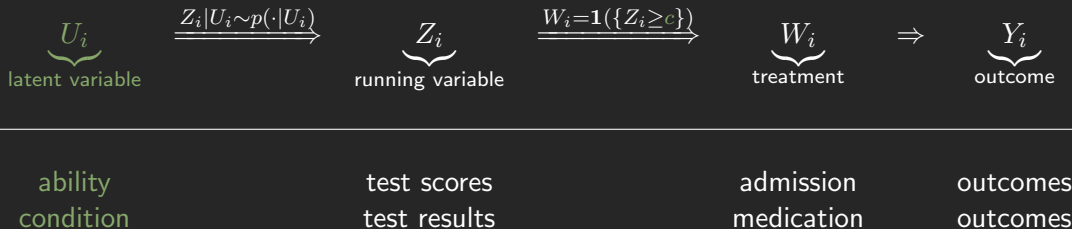
RD Identification: Measurement Error



RD Identification: Measurement Error

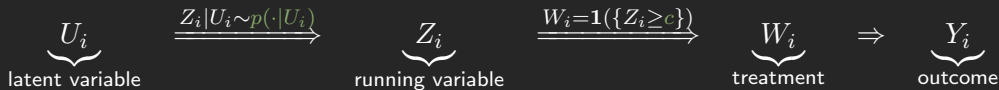


RD Identification: Measurement Error



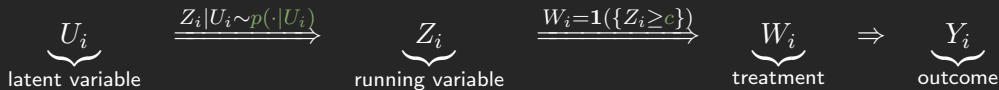
Why don't we take advantage of the measurement error itself for inference?

This Paper



Weighted treatment effects can be estimated if the measurement error in Z_i

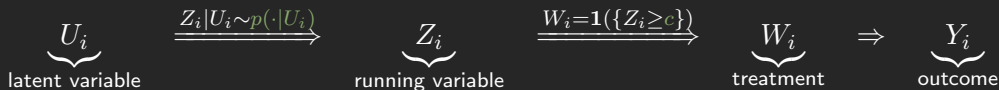
This Paper



Weighted treatment effects can be estimated if the measurement error in Z_i

- has a known distribution

This Paper

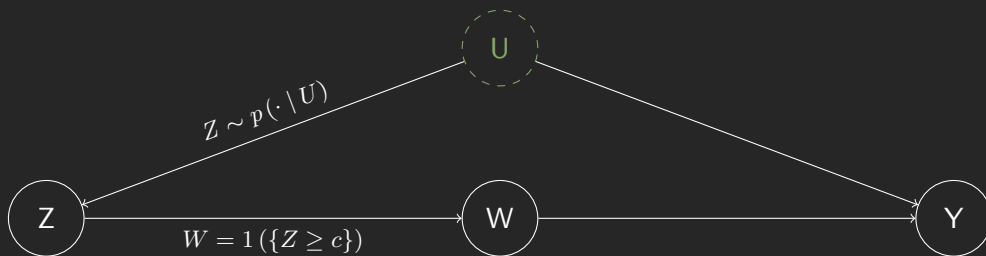


Weighted treatment effects can be estimated if the measurement error in Z_i

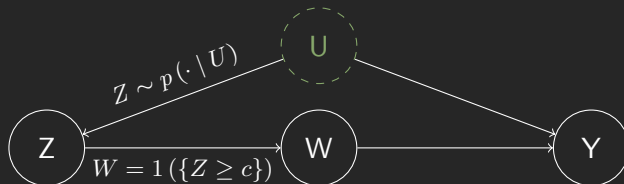
- has a known distribution
- is conditionally (on U_i) independent of potential outcomes

Key Argument

Sharp RD Design with A Noisy Running Variable



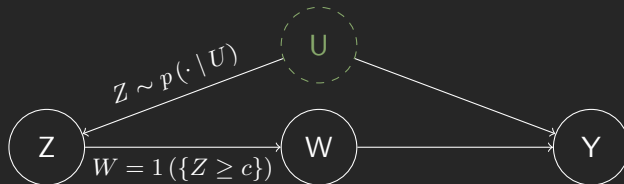
Sharp RD Design with A Noisy Running Variable



Assumption 1: Sharp RD design

- I.I.D. samples $\{Y_i(0), Y_i(1), Z_i\} \in \mathbb{R}^3, i = 1, \dots, n$
- treatment assignment: $W_i = 1(\{Z_i \geq c\})$, where $c \in \mathbb{R}$ is the cutoff
- observation: $\{Y_i, Z_i\}$ where $Y_i = Y_i(W_i)$

Sharp RD Design with A Noisy Running Variable

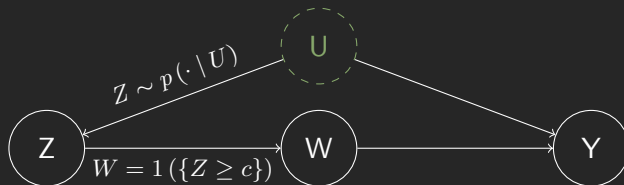


Assumption 2: Noisy running variable

$$Z_i | U_i \sim p(\cdot | U_i)$$

where $p(\cdot | \cdot)$ is a **known** conditional density w.r.t. to a measure λ , the latent variable U_i has an **unknown** distribution G

Sharp RD Design with A Noisy Running Variable

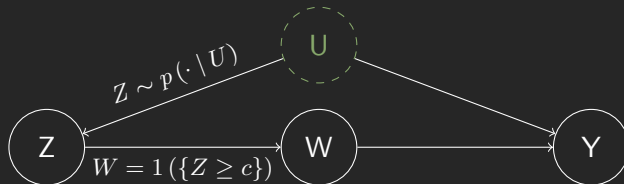


Assumption 2: Noisy running variable

$$Z_i | U_i \sim \mathcal{N}(U_i, \nu^2), \nu > 0$$

where $p(\cdot | \cdot)$ is a **known** conditional density w.r.t. to a measure λ , the latent variable U_i has an **unknown** distribution G

Sharp RD Design with A Noisy Running Variable

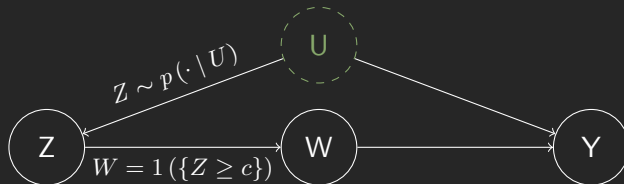


Assumption 2: Noisy running variable

$$Z_i | U_i \sim \text{Binomial}(K, U_i), K \in \mathbb{N}$$

where $p(\cdot | \cdot)$ is a **known** conditional density w.r.t. to a measure λ , the latent variable U_i has an **unknown** distribution G

Sharp RD Design with A Noisy Running Variable

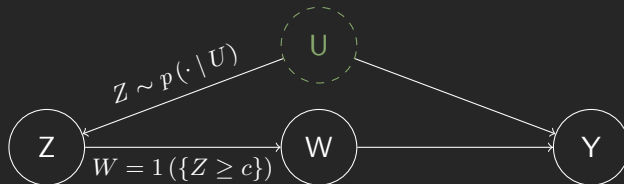


Assumption 3: Exogeneity

$$[\{Y_i(0), Y_i(1)\} \perp Z_i] \mid U_i$$

which implies $\mathbb{E}[Y_i \mid U_i, Z_i] = \alpha_{(W_i)}(u)$

Sharp RD Design with A Noisy Running Variable

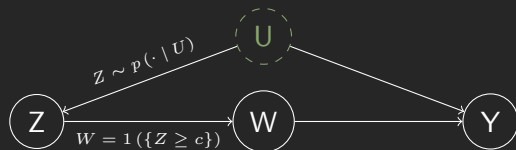


Assumption 3: Exogeneity

$$[\{Y_i(0), Y_i(1)\} \perp Z_i] \mid U_i$$

which implies $\mathbb{E}[Y_i \mid U_i, Z_i] = \alpha_{(W_i)}(u)$, where $\alpha_{(w)}(u) = \mathbb{E}[Y_i(w) \mid U_i = u]$ is the **response functions** for the potential outcomes conditional on the latent variable u

Sharp RD Design with A Noisy Running Variable

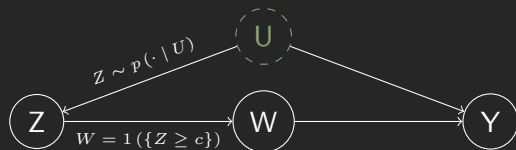


A1 Sharp RD

A2 Noisy Z_i : $Z_i | U_i \sim p(\cdot | U_i)$

A3 Exogeneity:
 $[\{Y_i(0), Y_i(1)\} \perp Z_i | U_i]$

Sharp RD Design with A Noisy Running Variable



A1 **Sharp** RD

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Proposition 1

Let $\gamma_+(\cdot), \gamma_-(\cdot)$ be measurable functions of Z , then under A1-A3:

$$\mathbb{E}[\gamma_+(Z)Y] = \mathbb{E}[\alpha_{(1)}(U)h(U, \gamma_+)], \quad \mathbb{E}[\gamma_-(Z)Y] = \mathbb{E}[\alpha_{(0)}(U)h(U, \gamma_-)]$$

where $h(u, \gamma) := \int \gamma(z)p(z \mid u)d\lambda(z)$, $\alpha_{(w)}(u) = \mathbb{E}[Y_i(w) \mid U_i = u]$

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$$\blacksquare \mathbb{E}[Y^2], \mathbb{E}[\gamma_-(Z)^2], \mathbb{E}[\gamma_+(Z)^2] < \infty$$

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- $\gamma_+(\cdot), \gamma_-(\cdot)$ are weighting functions s.t.

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 - $\gamma_-(z) = 0$ for $z \geq c$: assign non-zero weights only to **control** units

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Proof:

$$\mathbb{E}[\gamma_+(Z)Y | U]$$

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Proof:

$$\mathbb{E}[\gamma_+(Z)Y | U] = \mathbb{E}[\gamma_+(Z)Y \cdot \mathbf{1}(\{Z \geq c\}) | U]$$

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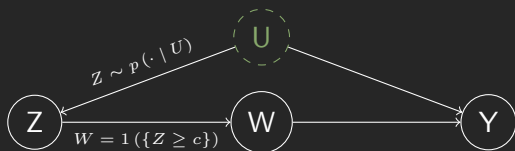
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Proof:

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Sharp RD Design with A Noisy Running Variable



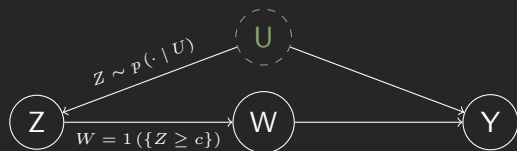
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Sharp RD Design with A Noisy Running Variable



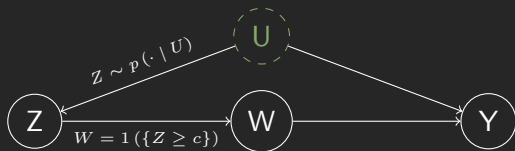
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- No need to know G (distribution of U)
- Need to know $p(z | u)$ (conditional distribution of the noise)

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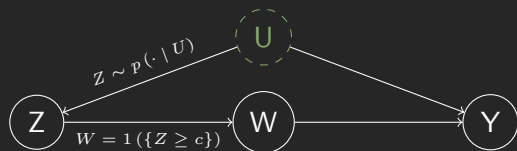
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■ No need to know G (distribution of U)

■ Need to know $p(z | u)$ (conditional distribution of the noise)

- test-retest data, prior modelling of responses to tests, physical model of the measurement device, biomedical knowledge, etc.

Sharp RD Design with A Noisy Running Variable



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■ No need to know G (distribution of U)

■ Need to know $p(z | u)$ (conditional distribution of the noise)

- test-retest data, prior modelling of responses to tests, physical model of the measurement device, biomedical knowledge, etc.
- still valid when **underestimating** the true noise level

Estimation

Estimation of Weighted Treatment Effects

Proposition: The Key Argument

Let $\gamma_+(\cdot), \gamma_-(\cdot)$ be measurable functions of Z , then under A1-A3:

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where $h (u, \gamma) := \int \gamma (z) p (z | u) d\lambda (z)$, $\alpha_{(w)} (u) = \mathbb{E} [Y_i (w) | U_i = u]$

ratio-form estimators:

$$\begin{aligned} \hat{\tau}_\gamma &= \hat{\mu}_{\gamma,+} \\ &= \frac{\sum_i \gamma_+ (Z_i) Y_i}{\sum_i \underbrace{\gamma_+ (Z_i)}} \end{aligned}$$

$$\begin{aligned} & - \hat{\mu}_{\gamma,-} \\ &= \frac{\sum_i \gamma_- (Z_i) Y_i}{\sum_i \underbrace{\gamma_- (Z_i)}} \end{aligned}$$

Estimation of Weighted Treatment Effects

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$$\begin{aligned} & - \hat{\mu}_{\gamma,-} \\ &= \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \underbrace{\gamma_-(Z_i)}_{\gamma_-(z)=0, z \geq c}} \end{aligned}$$

Estimation of Weighted Treatment Effects

Ratio-form estimators:

$$\hat{\tau}_{\gamma} = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)}$$

What's the weighted treatment effects to conduct inference for?

Estimation of Weighted Treatment Effects

Ratio-form estimators:

$$\hat{\tau}_{\gamma} = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)}$$

What's the weighted treatment effects to conduct inference for?

$$\tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u), w(\cdot) \geq 0$$

Estimation of Weighted Treatment Effects

Ratio-form estimators:

$$\hat{\tau}_{\gamma} = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)}$$

What's the weighted treatment effects to conduct inference for?

$$\tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u), w(\cdot) \geq 0$$

where $\tau(u)$ (**Conditional Average Treatment Effects**) is

$$\tau(u) = \mathbb{E}[Y_i(1) - Y_i(0) \mid U_i = u] = \alpha_{(1)}(u) - \alpha_{(0)}(u)$$

Weighted Treatment Effects: Example

$$\tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u), w(\cdot) \geq 0$$

where $\tau(u)$ (**CATE**) is $\tau(u) = \mathbb{E}[Y_i(1) - Y_i(0) \mid U_i = u] = \alpha_{(1)}(u) - \alpha_{(0)}(u)$

■ RD paramater:

$$\begin{aligned} \tau_c &= \mathbb{E}[Y_i(1) - Y_i(0) \mid Z_i = c] = \mathbb{E}[\tau(U_i) \mid Z_i = c] \\ &= \int \frac{p(c \mid u)}{\int p(c \mid u) dG(u)} \tau(u) dG(u) \end{aligned}$$

Weighted Treatment Effects: Example

$$\tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u), w(\cdot) \geq 0$$

where $\tau(u)$ (**CATE**) is $\tau(u) = \mathbb{E}[Y_i(1) - Y_i(0) \mid U_i = u] = \alpha_{(1)}(u) - \alpha_{(0)}(u)$

■ RD paramater:

$$\begin{aligned} \tau_c &= \mathbb{E}[Y_i(1) - Y_i(0) \mid Z_i = c] = \mathbb{E}[\tau(U_i) \mid Z_i = c] \\ &= \int \frac{p(c \mid u)}{\underbrace{\int p(c \mid u) dG(u)}_{=f_G(c)}} \tau(u) dG(u) \end{aligned}$$

Asymptotic Bias

Theorem: Asymptotic Limit of $\hat{\tau}_\gamma$

$$\hat{\tau}_\gamma = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)}$$

Asymptotic Bias

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Asymptotic Bias

Theorem: Asymptotic Limit of $\hat{\tau}_\gamma$

$$\begin{aligned}
 \hat{\tau}_\gamma = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} &= \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)} \\
 &\xrightarrow{p} \frac{\mathbb{E}[\gamma_+(Z)Y]}{\mathbb{E}[\gamma_+(Z)]} - \frac{\mathbb{E}[\gamma_-(Z)Y]}{\mathbb{E}[\gamma_-(Z)]} \\
 \text{(Prop.1)} &= \frac{\mathbb{E}[\alpha_{(1)}(U) h(U, \gamma_+)]}{\mathbb{E}[h(U, \gamma_+)]} - \frac{\mathbb{E}[\alpha_{(0)}(U) h(U, \gamma_-)]}{\mathbb{E}[h(U, \gamma_-)]} = \mu_{\gamma,+} - \mu_{\gamma,-} \equiv \theta_\gamma
 \end{aligned}$$

where

$$h(u, \gamma) := \int \gamma(z) p(z | u) d\lambda(z), \quad \alpha_{(w)}(u) = \mathbb{E}[Y_i(w) | U_i = u]$$

Asymptotic Bias

Theorem: Asymptotic Limit of $\hat{\tau}_\gamma$

$$\hat{\tau}_\gamma = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)} \xrightarrow{p} \frac{\mathbb{E} [\alpha_{(1)}(U) h(U, \gamma_+)]}{\mathbb{E} [h(U, \gamma_+)]} - \frac{\mathbb{E} [\alpha_{(0)}(U) h(U, \gamma_-)]}{\mathbb{E} [h(U, \gamma_-)]} = \mu_{\gamma,+} - \mu_{\gamma,-} \equiv \theta_\gamma$$

How biased is this asymptotic limit? Comparing to

$$\tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u), w(\cdot) \geq 0$$

Asymptotic Bias: Decomposition

$$a\text{Bias} \left[\gamma_{\pm}, \tau_w; \alpha_{(0)}(\cdot), \tau(\cdot), G \right] = \theta_{\gamma} - \tau_w$$

Asymptotic Bias: Decomposition

$$a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)} (\cdot), \tau (\cdot), G] = \theta_{\gamma} - \tau_w$$

$$= \frac{\mathbb{E} [\alpha_{(1)} (U) h (U, \gamma_+)]}{\mathbb{E} [h (U, \gamma_+)]} - \frac{\mathbb{E} [\alpha_{(0)} (U) h (U, \gamma_-)]}{\mathbb{E} [h (U, \gamma_-)]} - \int \frac{w (u)}{\mathbb{E}_G [w (U)]} \tau (u) \, dG (u)$$

Asymptotic Bias: Decomposition

$$\begin{aligned}
 a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)}(\cdot), \tau(\cdot), G] &= \theta_{\gamma} - \tau_w \\
 &= \frac{\mathbb{E} [\alpha_{(1)}(U) h(U, \gamma_+)]}{\mathbb{E} [h(U, \gamma_+)]} - \frac{\mathbb{E} [\alpha_{(0)}(U) h(U, \gamma_-)]}{\mathbb{E} [h(U, \gamma_-)]} - \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u) \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} \right) \alpha_{(1)}(u) dG(u) - \int \left(\frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) \\
 &\quad - \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u)
 \end{aligned}$$

Asymptotic Bias: Decomposition

$$\begin{aligned}
 a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)} (\cdot), \tau (\cdot), G] &= \theta_{\gamma} - \tau_w \\
 &= \frac{\mathbb{E} [\alpha_{(1)} (U) h (U, \gamma_+)]}{\mathbb{E} [h (U, \gamma_+)]} - \frac{\mathbb{E} [\alpha_{(0)} (U) h (U, \gamma_-)]}{\mathbb{E} [h (U, \gamma_-)]} - \int \frac{w (u)}{\mathbb{E}_G [w (U)]} \tau (u) dG (u) \\
 &= \int \left(\frac{h (u, \gamma_+)}{\mathbb{E}_G [h (U, \gamma_+)]} \right) \alpha_{(1)} (u) dG (u) - \int \left(\frac{h (u, \gamma_-)}{\mathbb{E}_G [h (U, \gamma_-)]} \right) \alpha_{(0)} (u) dG (u) \\
 &\quad - \int \frac{w (u)}{\mathbb{E}_G [w (U)]} \tau (u) dG (u)
 \end{aligned}$$

Remember? $\tau(u)$ (**Conditional Average Treatment Effects**) is

$$\tau (u) = \mathbb{E} [Y_i (1) - Y_i (0) \mid U_i = u] = \alpha_{(1)} (u) - \alpha_{(0)} (u) \Rightarrow \boxed{\alpha_{(1)} (u) = \tau (u) + \alpha_{(0)} (u)}$$

Asymptotic Bias: Decomposition

$$\begin{aligned}
 a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)}(\cdot), \tau(\cdot), G] &= \theta_{\gamma} - \tau_w \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} \right) \underbrace{\alpha_{(1)}(u)}_{=\tau(u) + \alpha_{(0)}(u)} dG(u) \\
 &\quad - \int \left(\frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) - \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u)
 \end{aligned}$$

Asymptotic Bias: Decomposition

$$\begin{aligned}
 a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)} (\cdot), \tau (\cdot), G] &= \theta_{\gamma} - \tau_w \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} \right) \underbrace{\alpha_{(1)}(u)}_{=\tau(u) + \alpha_{(0)}(u)} dG(u) \\
 &\quad - \int \left(\frac{h(u, \gamma_-)}{\mathbb{E}_G [h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) - \int \frac{w(u)}{\mathbb{E}_G [w(U)]} \tau(u) dG(u) \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G [h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) \\
 &\quad + \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G [w(U)]} \right) \tau(u) dG(u)
 \end{aligned}$$

Asymptotic Bias: Decomposition

$$\begin{aligned}
 a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)} (\cdot), \tau (\cdot), G] &= \theta_{\gamma} - \tau_w \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} \right) \underbrace{\alpha_{(1)}(u)}_{=\tau(u) + \alpha_{(0)}(u)} dG(u) \\
 &\quad - \int \left(\frac{h(u, \gamma_-)}{\mathbb{E}_G [h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) - \int \frac{w(u)}{\mathbb{E}_G [w(U)]} \tau(u) dG(u) \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G [h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) \\
 &\quad + \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G [w(U)]} \right) \tau(u) dG(u)
 \end{aligned}$$

Confounding bias

Asymptotic Bias: Decomposition

$$\begin{aligned}
 a\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_{(0)} (\cdot), \tau (\cdot), G] &= \theta_{\gamma} - \tau_w \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} \right) \underbrace{\alpha_{(1)}(u)}_{=\tau(u) + \alpha_{(0)}(u)} dG(u) \\
 &\quad - \int \left(\frac{h(u, \gamma_-)}{\mathbb{E}_G [h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) - \int \frac{w(u)}{\mathbb{E}_G [w(U)]} \tau(u) dG(u) \\
 &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G [h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) \\
 &\quad + \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G [h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G [w(U)]} \right) \tau(u) dG(u)
 \end{aligned}$$

Confounding bias

CATE heterogeneity bias

Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

Confounding bias

CATE heterogeneity bias

How to minimize them?

Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

Confounding bias

CATE heterogeneity bias

How to minimize them?

- **Confounding bias:** $h(\cdot, \gamma_+) \approx h(\cdot, \gamma_-)$

Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

Confounding bias

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

CATE heterogeneity bias

How to minimize them?

- **Confounding bias:** $h(\cdot, \gamma_+) \approx h(\cdot, \gamma_-)$ where $h(u, \gamma) := \int \gamma(z) p(z | u) d\lambda(z)$
How well the units are **balanced** via the latent variable u

Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

Confounding bias

CATE heterogeneity bias

How to minimize them?

- **Confounding bias:** $h(\cdot, \gamma_+) \approx h(\cdot, \gamma_-)$

How well the units are **balanced** via the latent variable u

- **CATE heterogeneity bias:**

Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

Confounding bias

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Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

Confounding bias

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

CATE heterogeneity bias

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How well the units are **balanced** via the latent variable u

- **CATE heterogeneity bias:**

- $\tau(u)$ being constant w.r.t. u , a constant conditional treatment effect

Asymptotic Bias: Decomposition

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)$$

Confounding bias

$$\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)$$

CATE heterogeneity bias

How to minimize them?

- **Confounding bias:** $h(\cdot, \gamma_+) \approx h(\cdot, \gamma_-)$

How well the units are **balanced** via the latent variable u

- **CATE heterogeneity bias:**

- $\tau(u)$ being constant w.r.t. u , a constant conditional treatment effect
- $h(u, \gamma_+) = w(u), \forall u$, an absolutely “correct” weighting function

Confidence Intervals

Asymptotic Normality

$$\hat{\tau} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)}$$

$$\hat{\tau}_\gamma \xrightarrow{p} \theta_\gamma = \frac{\mathbb{E} [\alpha_{(1)}(U) h(U, \gamma_+)]}{\mathbb{E} [h(U, \gamma_+)]} - \frac{\mathbb{E} [\alpha_{(0)}(U) h(U, \gamma_-)]}{\mathbb{E} [h(U, \gamma_-)]}$$

Asymptotic Normality

Theorem: Asymptotic Normality of $\hat{\tau}$

Suppose the sequence of weighting kernels $\gamma_+^{(n)}$ and $\gamma_-^{(n)}$ is deterministic, and $\exists \beta \in (0, \frac{1}{2})$, $C, C' > 0$ s.t. $\forall n$ large enough:

$$\sup_z \left| \gamma_\diamond^{(n)}(z) \right| < C n^\beta \mathbb{E} \left[\gamma_\diamond^{(n)}(Z_i) \right] \quad \sup_u \left| h(u, \gamma_\diamond^{(n)}) \right| < C' \mathbb{E} \left[\gamma_\diamond^{(n)}(Z_i) \right], \quad \diamond = \{+, -\}$$

Then

$$\frac{\sqrt{n} (\hat{\tau}_\gamma - \theta_\gamma)}{\sqrt{V_\gamma}} \xrightarrow{d} \mathcal{N}(0, 1)$$

where

$$V_\gamma = \frac{\mathbb{E} \left[\gamma_+^2(Z_i) (Y_i - \mu_{\gamma,+})^2 \right]}{\mathbb{E} [\gamma_+(Z_i)]^2} + \frac{\mathbb{E} \left[\gamma_-^2(Z_i) (Y_i - \mu_{\gamma,-})^2 \right]}{\mathbb{E} [\gamma_-(Z_i)]^2}$$

Asymptotic Normality

Theorem: Asymptotic Normality of $\hat{\tau}$

Suppose the sequence of weighting kernels $\gamma_+^{(n)}$ and $\gamma_-^{(n)}$ is deterministic, and $\exists \beta \in (0, \frac{1}{2})$, $C, C' > 0$ s.t. $\forall n$ large enough: $\sup_z |\gamma_\diamond^{(n)}(z)| < C n^\beta \mathbb{E}[\gamma_\diamond^{(n)}(Z_i)]$, $\sup_u |h(u, \gamma_\diamond^{(n)})| < C' \mathbb{E}[\gamma_\diamond^{(n)}(Z_i)]$ where $\diamond = \{+, -\}$. Then

$$\frac{\sqrt{n}(\hat{\tau}_\gamma - \theta_\gamma)}{\sqrt{V_\gamma}} \xrightarrow{d} \mathcal{N}(0, 1)$$

where $V_\gamma = \frac{\mathbb{E}[\gamma_+^2(Z_i)(Y_i - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2} + \frac{\mathbb{E}[\gamma_-^2(Z_i)(Y_i - \mu_{\gamma,-})^2]}{\mathbb{E}[\gamma_-(Z_i)]^2}$

Assumption:

- The response Y_i is bounded: $Y_i \in [0, 1]$
- $\inf_z \text{Var}[Y_i | Z_i = z] > 0$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\sqrt{n} [(\hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-}) - (\mu_{\gamma,+} - \mu_{\gamma,-})]}{\sqrt{\frac{\mathbb{E}[\gamma_+^2(Z_i)(Y_i - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2} + \frac{\mathbb{E}[\gamma_-^2(Z_i)(Y_i - \mu_{\gamma,-})^2]}{\mathbb{E}[\gamma_-(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\sqrt{n} [(\hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-}) - (\mu_{\gamma,+} - \mu_{\gamma,-})]}{\sqrt{\frac{\mathbb{E}[\gamma_+^2(Z_i)(Y_i - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2} + \frac{\mathbb{E}[\gamma_-^2(Z_i)(Y_i - \mu_{\gamma,-})^2]}{\mathbb{E}[\gamma_-(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

$$\frac{\sqrt{n} (\hat{\mu}_{\gamma,+} - \mu_{\gamma,+})}{\sqrt{\frac{\mathbb{E}[\gamma_+(Z_i)^2(Y_i(1) - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\sqrt{n} [(\hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-}) - (\mu_{\gamma,+} - \mu_{\gamma,-})]}{\sqrt{\frac{\mathbb{E}[\gamma_+^2(Z_i)(Y_i - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2} + \frac{\mathbb{E}[\gamma_-^2(Z_i)(Y_i - \mu_{\gamma,-})^2]}{\mathbb{E}[\gamma_-(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

$$\frac{\sqrt{n} (\hat{\mu}_{\gamma,+} - \mu_{\gamma,+})}{\sqrt{\frac{\mathbb{E}[\gamma_+(Z_i)^2(Y_i(1) - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

$$\frac{\sqrt{n} \left(\frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \mu_{\gamma,+} \right)}{\sqrt{\frac{\mathbb{E}[\gamma_+(Z_i)^2(Y_i(1) - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\sqrt{n} \left(\frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \mu_{\gamma,+} \right)}{\sqrt{\frac{\mathbb{E}[\gamma_+(Z_i)^2 (Y_i(1) - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\sqrt{n} \left(\frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \mu_{\gamma,+} \right)}{\sqrt{\frac{\mathbb{E}[\gamma_+(Z_i)^2 (Y_i(1) - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

$$\frac{\frac{\sum_i \gamma_+(Z_i) (Y_i(1) - \mu_{\gamma,+})}{\sum_i \gamma_+(Z_i)}}{\sqrt{\frac{n \mathbb{E}[\gamma_+(Z_i)^2 (Y_i(1) - \mu_{\gamma,+})^2]}{n \mathbb{E}[\gamma_+(Z_i)]}}} = \underbrace{\frac{\sum_i \gamma_+(Z_i) (Y_i(1) - \mu_{\gamma,+})}{\sqrt{n \mathbb{E}[\gamma_+(Z_i)^2 (Y_i(1) - \mu_{\gamma,+})^2]}}}_{\xrightarrow{d} \mathcal{N}(0,1)} \cdot \underbrace{\frac{1}{\frac{\frac{1}{n} \sum_i \gamma_+(Z_i)}{\mathbb{E}[\gamma_+(Z_i)]}}}_{=1+o_p(1)} \xrightarrow{d} \mathcal{N}(0, 1)$$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\sum_i \gamma_+(Z_i) (Y_i(1) - \mu_{\gamma,+})}{\sqrt{n \mathbb{E} [\gamma_+(Z_i)^2 (Y_i(1) - \mu_{\gamma,+})^2]}} \xrightarrow{d} \mathcal{N}(0, 1)$$

- $\underline{\sigma}^2 = \inf_z \text{Var} [Y_i \mid Z_i = z] > 0$: $\text{Var} [\gamma_+(Z_i) (Y_i(1) - \mu_{\gamma,+})] \geq \underline{\sigma}^2 \mathbb{E} [\gamma_+(Z_i)^2]$
- $\sup_u \left| h(u, \gamma_\diamond^{(n)}) \right| < C' \mathbb{E} [\gamma_\diamond^{(n)}(Z_i)]$ and $Y_i \in [0, 1]$: $|\mu_{\gamma,+}| = \left| \frac{\mathbb{E}[\alpha_{(1)}(U)h(U, \gamma_+)]}{\mathbb{E}[\gamma_+(Z_i)]} \right| \leq C'$
- $\sup_z \left| \gamma_\diamond^{(n)}(z) \right| < C n^\beta \mathbb{E} [\gamma_\diamond^{(n)}(Z_i)]$ and Lyapunov CLT condition:

$$\frac{n \mathbb{E} [|\gamma_+(Z_i) (Y_i(1) - \mu_{\gamma,+})|^{2+q}]}{(n \text{Var} [\gamma_+(Z_i) (Y_i(1) - \mu_{\gamma,+})])^{\frac{2+q}{2}}} \leq \left(\frac{C' + 1}{\underline{\sigma}} \right)^{2+q} \left(C n^{\beta - \frac{1}{2}} \right)^q \xrightarrow{0 \rightarrow \infty} 0$$

Asymptotic Normality: A Sketch of the Proof

$$\frac{\frac{1}{n} \sum_i \gamma_+(Z_i)}{\mathbb{E}[\gamma_+(Z_i)]} \xrightarrow{p} 1$$

■ $\left[\sup_z \left| \gamma_{\diamond}^{(n)}(z) \right| < C n^{\beta} \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right] \right]$ and Chebyshev's inequality:

$$\begin{aligned} \mathbf{P} \{ |\mathbb{E}_n[\gamma_+(Z_i)] - \mathbb{E}[\gamma_+(Z_i)]| \geq \epsilon \mathbb{E}[\gamma_+(Z_i)] \} &\leq \frac{\text{Var}[\gamma_+(Z_i)]}{n \epsilon^2 \mathbb{E}[\gamma_+(Z_i)]^2} \\ &\leq \left(\frac{C}{\epsilon} \cdot n^{\beta - \frac{1}{2}} \right)^2 \xrightarrow{n \rightarrow \infty} 0 \end{aligned}$$

Asymptotic Normality: A Sketch of the Proof

$$\begin{aligned}
 \frac{\frac{\sum_i \gamma_+(Z_i)(Y_i(1) - \mu_{\gamma,+})}{\sum_i \gamma_+(Z_i)}}{\sqrt{\frac{n\mathbb{E}[\gamma_+(Z_i)^2(Y_i(1) - \mu_{\gamma,+})^2]}{n\mathbb{E}[\gamma_+(Z_i)]}}} &= \underbrace{\frac{\sum_i \gamma_+(Z_i)(Y_i(1) - \mu_{\gamma,+})}{\sqrt{n\mathbb{E}[\gamma_+(Z_i)^2(Y_i(1) - \mu_{\gamma,+})^2]}}}_{\xrightarrow{d} \mathcal{N}(0,1)} \cdot \underbrace{\frac{1}{\frac{\frac{1}{n} \sum_i \gamma_+(Z_i)}{\mathbb{E}[\gamma_+(Z_i)]}}}_{=1+o_p(1)} \xrightarrow{d} \mathcal{N}(0,1) \\
 \Rightarrow \frac{\sqrt{n} \left(\frac{\sum_i \gamma_+(Z_i)Y_i}{\sum_i \gamma_+(Z_i)} - \mu_{\gamma,+} \right)}{\sqrt{\frac{\mathbb{E}[\gamma_+(Z_i)^2(Y_i(1) - \mu_{\gamma,+})^2]}{\mathbb{E}[\gamma_+(Z_i)]^2}}} &\xrightarrow{d} \mathcal{N}(0,1) \Rightarrow \frac{\sqrt{n}(\hat{\tau}_\gamma - \theta_\gamma)}{\sqrt{V_\gamma}} \xrightarrow{d} \mathcal{N}(0,1)
 \end{aligned}$$

Plug-in Estimator for V_γ

$$V_\gamma = \frac{\mathbb{E} \left[\gamma_+^2 (Z_i) (Y_i - \mu_{\gamma,+})^2 \right]}{\mathbb{E} [\gamma_+ (Z_i)]^2} + \frac{\mathbb{E} \left[\gamma_-^2 (Z_i) (Y_i - \mu_{\gamma,-})^2 \right]}{\mathbb{E} [\gamma_- (Z_i)]^2}$$

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Proposition: Plug-in Estimator \hat{V}_γ

Under the same assumptions, we have $\frac{\hat{V}_\gamma}{V_\gamma} = 1 + o_p(1)$ where

$$\hat{V}_\gamma = \frac{\frac{1}{n} \sum_{i=1}^n \gamma_+^2 (Z_i) (Y_i - \hat{\mu}_{\gamma,+})^2}{\left[\frac{1}{n} \sum_{i=1}^n \gamma_+ (Z_i) \right]^2} + \frac{\frac{1}{n} \sum_{i=1}^n \gamma_-^2 (Z_i) (Y_i - \hat{\mu}_{\gamma,-})^2}{\left[\frac{1}{n} \sum_{i=1}^n \gamma_- (Z_i) \right]^2}$$

Upper Bound for the Potential Bias $|b_\gamma| = |\theta_\gamma - \tau_w|$

$$\begin{aligned}
 \text{aBias} = \theta_\gamma - \tau_w = & \underbrace{\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u)}_{\text{Confounding bias}} \\
 & + \underbrace{\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)}_{\text{CATE heterogeneity bias}}
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 \end{aligned}$$

The distribution of $G(\cdot)$ is **unknown**

Upper Bound for the Potential Bias $|b_\gamma| = |\theta_\gamma - \tau_w|$

Bound the **worst-case** bias:

Upper Bound for the Potential Bias $|b_\gamma| = |\theta_\gamma - \tau_w|$

Bound the **worst-case** bias:

- Back out the class of latent variable distribution from $\hat{F}_n(t) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}(Z_i \leq t)$, the empirical distribution of Z_i (Massart, 1990):

$$\mathcal{G}_n = \left\{ G(\cdot) : \sup_{t \in \mathbb{R}} |F_G(t) - \hat{F}_n(t)| \leq \sqrt{\frac{\log(2/\alpha_n)}{2n}} \right\}, \quad \alpha_n = \min \{0.05, n^{-1/4}\}$$

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- Take treatment effect heterogeneity into consideration:

$$\mathcal{T}_M = \{\tau(\cdot) \mid \tau(u) = \bar{\tau} + \Delta(u), \bar{\tau} \in \mathbb{R}, |\Delta(u)| \leq M\}, \quad M \in [0, 1]$$

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- \mathcal{T}_0 ($M = 0$): constant CATE
- \mathcal{T}_1 ($M = 1$): no assumptions
- $\mathcal{T}_{1/2}$ ($M = 0$): a conservative choice for a monotonicity restriction

Upper Bound for the Potential Bias $|b_\gamma| = |\theta_\gamma - \tau_w|$

Proposition: Upper Bound for the Potential Bias $\hat{B}_{\gamma,M}$

Under asymptotic normality and all necessary assumptions, for $\tau(\cdot) \in \mathcal{T}_M$, the upper bound of bias is

$$\hat{B}_{\gamma,M} = \sup \left\{ |\text{Bias}[\gamma_\pm, \tau_w; \alpha_0(\cdot), \tau(\cdot), G]| : G \in \mathcal{G}_n, \alpha_{(0)}(\cdot) \in [0, 1], \tau(\cdot) \in \mathcal{T}_M \right\}$$

then $\mathbf{P} \left(|b_\gamma| \leq \hat{B}_{\gamma,M} \right) \xrightarrow{n \rightarrow \infty} 1$

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$$\{G \in \mathcal{G}_n\} \subset \{|b_\gamma| \leq \hat{B}_{\gamma,M}\} \Rightarrow \mathbf{P}(|b_\gamma| \leq \hat{B}_{\gamma,M}) \geq \mathbf{P}(G \in \mathcal{G}_n)$$

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and for $G \in \mathcal{G}_n$, Dvoretzky-Kiefer-Wolfowitz (DKW) inequality gives:

$$\mathbf{P}(G \in \mathcal{G}_n) \geq \mathbf{P} \left[\sup_{t \in \mathbb{R}} |F_G(t) - \hat{F}_n(t)| \leq \sqrt{\frac{\log(2/\alpha_n)}{2n}} \right] \geq 1 - \alpha_n \xrightarrow{n \rightarrow \infty} 1$$

Recap: What We Have Derived

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- Asymptotic limit:

$$\hat{\tau}_{\gamma} \xrightarrow{p} \theta_{\gamma} = \frac{\mathbb{E} [\alpha_{(1)} (U) h (U, \gamma_+)]}{\mathbb{E} [h (U, \gamma_+)]} - \frac{\mathbb{E} [\alpha_{(0)} (U) h (U, \gamma_-)]}{\mathbb{E} [h (U, \gamma_-)]}$$

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■ Asymptotic normality:

$$\frac{\sqrt{n} (\hat{\tau}_{\gamma} - \theta_{\gamma})}{\sqrt{\hat{V}_{\gamma}}} \xrightarrow{d} \mathcal{N} (0, 1)$$

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- Upper bound of the asymptotic bias:

$$\hat{B}_{\gamma, M} = \sup \left\{ |\text{Bias} [\gamma_{\pm}, \tau_w; \alpha_0 (\cdot), \tau (\cdot), G]| : G \in \mathcal{G}_n, \alpha_{(0)} (\cdot) \in [0, 1], \tau (\cdot) \in \mathcal{T}_M \right\}$$

Bias-aware Confidence Intervals

Corollary: Valid Confidence Intervals

Under asymptotic normality and all necessary assumptions, for $\tau(\cdot) \in \mathcal{T}_M$, consider the CIs

$$\hat{\tau}_\gamma \pm l_\alpha, \quad l_\alpha = \min \left\{ l : \mathbf{P} \left[\left| b + n^{-\frac{1}{2}} \hat{V}_\gamma^{\frac{1}{2}} \tilde{Z} \right| \leq l \right] \geq 1 - \alpha, \forall |b| \leq \hat{B}_{\gamma, M} \right\}$$

where

- \tilde{Z} is a standard Guassian random variable
- $\alpha \in (0, 1)$ is the significant level
- \hat{V}_γ is an estimate of the sampling variance V_γ

then

$$\liminf_{n \rightarrow \infty} \mathbf{P} [\tau_w \in \hat{\tau}_\gamma \pm l_\alpha] \geq 1 - \alpha$$

Bias-aware Confidence Intervals

CLT (with bias considered) is

$$\frac{\sqrt{n}(\hat{\tau}_\gamma - \theta_\gamma)}{\hat{V}_\gamma^{1/2}} = \frac{\sqrt{n}(\hat{\tau}_\gamma - \tau_w - b_\gamma)}{\hat{V}_\gamma^{1/2}} \xrightarrow{d} \mathcal{N}(0, 1)$$

where $b_\gamma = \theta_\gamma - \tau_w$, then let $\tilde{Z} \sim \mathcal{N}(0, 1)$ we have

$$\begin{aligned} \mathbf{P}[\tau_w \in \hat{\tau}_\gamma \pm l_\alpha] &= \mathbf{P}[-l_\alpha - b_\gamma \leq \hat{\tau}_\gamma - \tau_w - b_\gamma \leq l_\alpha - b_\gamma] \\ &= \mathbf{P}\left[-\sqrt{n}\hat{V}_\gamma^{-1/2}(l_\alpha + b_\gamma) \leq \sqrt{n}\hat{V}_\gamma^{-1/2}(\hat{\tau}_\gamma - \tau_w - b_\gamma) \leq \sqrt{n}\hat{V}_\gamma^{-1/2}(l_\alpha - b_\gamma)\right] \\ &= \mathbb{E}\left(\mathbf{P}\left[-\sqrt{n}\hat{V}_\gamma^{-1/2}(l_\alpha + b_\gamma) \leq \tilde{Z} \leq \sqrt{n}\hat{V}_\gamma^{-1/2}(l_\alpha - b_\gamma)\right] \mid \hat{V}_\gamma, \hat{B}_{\gamma, M}, \hat{\tau}_\gamma\right) + o(1) \\ &= \mathbb{E}\left[\mathbf{P}\left(-l_\alpha \leq n^{-1/2}\hat{V}_\gamma^{1/2}\tilde{Z} + b_\gamma \leq l_\alpha\right) \mid \hat{V}_\gamma, \hat{B}_{\gamma, M}, \hat{\tau}_\gamma\right] + o(1) \\ &= 1 - \alpha + o(1) \end{aligned}$$

Bias-aware Confidence Intervals

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$$\frac{\sqrt{n}(\hat{\tau}_\gamma - \theta_\gamma)}{\hat{V}_\gamma^{1/2}} = \frac{\sqrt{n}(\hat{\tau}_\gamma - \tau_w - b_\gamma)}{\hat{V}_\gamma^{1/2}} \xrightarrow{d} \mathcal{N}(0, 1)$$

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Robustness to CATE Heterogeneity Misspecification

$$\hat{\tau}_{\gamma} \pm l_{\alpha}, l_{\alpha} = \min \left\{ l : \mathbf{P} \left[\left| b + n^{-\frac{1}{2}} \hat{V}_{\gamma}^{\frac{1}{2}} \tilde{Z} \right| \leq l \right] \geq 1 - \alpha, \forall |b| \leq \hat{B}_{\gamma, M} \right\}$$

where

$$\mathcal{T}_M = \{ \tau(\cdot) \mid \tau(u) = \bar{\tau} + \Delta(u), \bar{\tau} \in \mathbb{R}, |\Delta(u)| \leq M \}, \quad M \in [0, 1]$$

Consider an extreme misspecification of CATE heterogeneity: $M = 0$, are the CIs robust?

Robustness to CATE Heterogeneity Misspecification

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Corollary: Robustness to CATE Heterogeneity Misspecification

The CIs under the misspecification of $M = 0$ is still valid, but only for the convenience-weighted treatment effect:

$$\tau_{h,+} := \int \frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} \tau(u) dG(u)$$

Robustness to CATE Heterogeneity Misspecification

Corollary: Robustness to CATE Heterogeneity Misspecification

The CIs under the misspecification of $M = 0$ is still valid, but only for:

$$\tau_{h,+} := \int \frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} \tau(u) dG(u)$$

$$\begin{aligned} a\text{Bias} = \theta_\gamma - \tau_w &= \int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{h(u, \gamma_-)}{\mathbb{E}_G[h(U, \gamma_-)]} \right) \alpha_{(0)}(u) dG(u) \\ &\quad + \underbrace{\int \left(\frac{h(u, \gamma_+)}{\mathbb{E}_G[h(U, \gamma_+)]} - \frac{w(u)}{\mathbb{E}_G[w(U)]} \right) \tau(u) dG(u)}_{\text{CATE heterogeneity bias}} \end{aligned}$$

Applications

Design Estimators

The goal: Make the confidence intervals **shorter**

$$\hat{\tau}_{\gamma} \pm l_{\alpha}, \quad l_{\alpha} = \min \left\{ l : \mathbf{P} \left[\left| b + n^{-\frac{1}{2}} \hat{V}_{\gamma}^{\frac{1}{2}} \tilde{Z} \right| \leq l \right] \geq 1 - \alpha, \forall |b| \leq \hat{B}_{\gamma, M} \right\}$$

by minimizing the worst-case MSE of

$$\hat{\tau} = \hat{\mu}_{\gamma,+} - \hat{\mu}_{\gamma,-} = \frac{\sum_i \gamma_+(Z_i) Y_i}{\sum_i \gamma_+(Z_i)} - \frac{\sum_i \gamma_-(Z_i) Y_i}{\sum_i \gamma_-(Z_i)}$$

Design Estimators: Quadratic Programming

Solve

$$\min_{\gamma_{\pm}(\cdot)} \frac{1}{n} \left(\int \gamma_-^2(z) d\bar{F}(z) + \int \gamma_+^2(z) d\bar{F}(z) \right) + (t_1 + t_2)^2$$

s.t.

$$|h(u, \gamma_+) - h(u, \gamma_-)| \leq t_1, \quad \forall u$$

$$M |h(u, \gamma_{\diamond}) - \bar{w}(u)| \leq t_2, \quad \forall u, \diamond \in \{\pm\}$$

$$\int \gamma_+(z) d\bar{F}(z) = \int \gamma_-(z) d\bar{F}(z) = 1$$

$$\gamma_-(z) = 0, \quad z \geq c$$

$$\gamma_+(z) = 0, \quad z < c$$

$$|\gamma_{\diamond}(z)| \leq Cn^{\beta}, \quad \forall z, \diamond \in \{\pm\}$$

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$$\int \gamma_{+}(z) d\bar{F}(z) = \int \gamma_{-}(z) d\bar{F}(z) = 1$$

$$\gamma_{-}(z) = 0, \quad z \geq c$$

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confounding bias

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normalization constraint

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Sharp RD

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$$|h(u, \gamma_+) - h(u, \gamma_-)| \leq t_1, \quad \forall u$$

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$$\gamma_-(z) = 0, \quad z \geq c$$

Sharp RD

$$\gamma_+(z) = 0, \quad z < c$$

$$|\gamma_{\diamond}(z)| \leq Cn^{\beta}, \quad \forall z, \diamond \in \{\pm\} \quad \text{no observation is given excessive influence}$$

Design Estimators: Quadratic Programming

Solve

$$\min_{\gamma_{\pm}(\cdot)} \frac{1}{n} \left(\int \gamma_{-}^2(z) d\bar{F}(z) + \int \gamma_{+}^2(z) d\bar{F}(z) \right) + (t_1 + t_2)^2$$

s.t.

$$M |h(u, \gamma_{\diamond}) - \bar{w}(u)| \leq t_2, \quad \forall u, \diamond \in \{\pm\}$$

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normalization constraint

$$\bar{F}(\cdot) : \quad F_G(t) = \int \mathbf{1}(\{z \leq t\}) \int p(z | u) dG(u) d\lambda(z)$$

$$\bar{w}(\cdot) : \quad \tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u)$$

Design Estimators: Quadratic Programming

$$\bar{F}(\cdot) : \quad F_G(t) = \int \mathbf{1}(\{z \leq t\}) \int p(z | u) dG(u) d\lambda(z)$$

$$\bar{w}(\cdot) : \quad \tau_w = \int \frac{w(u)}{\mathbb{E}_G[w(U)]} \tau(u) dG(u)$$

- $\bar{F}(\cdot)$ assigns non-trivial mass to $[c, \infty)$ and $\bar{w}(\cdot)$ is bounded: $\exists k > 1$ s.t.

$$\mathbb{P} \left[\frac{1}{k} < \bar{F}([c, \infty)) < 1 - \frac{1}{k}, \sup_u |\bar{w}(u)| < k \right] \xrightarrow{n \rightarrow \infty} 1$$

- $\int \gamma_{\diamond}^{(n)}(z) dF(z)$ is asymptotically lower bounded by a strictly positive number:

$$\exists \delta > 0 \text{ s.t. } \mathbb{P} \left[\int \gamma_{\diamond}^{(n)}(z) dF(z) > \delta \right] \xrightarrow{n \rightarrow \infty} 1$$

Design Estimators: Quadratic Programming

$$\frac{1}{k} < \bar{F}([c, \infty)) < 1 - \frac{1}{k}, \sup_u |\bar{w}(u)| < k \quad \Rightarrow \quad \sup_z \left| \gamma_{\diamond}^{(n)}(z) \right| < C n^{\beta} \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right]$$

$$\int \gamma_{\diamond}^{(n)}(z) dF(z) > \delta \quad \Rightarrow \quad \sup_u \left| h(u, \gamma_{\diamond}^{(n)}) \right| < C' \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right]$$

Design Estimators: Quadratic Programming

$$\frac{1}{k} < \bar{F}([c, \infty)) < 1 - \frac{1}{k}, \sup_u |\bar{w}(u)| < k \quad \Rightarrow \quad \sup_z \left| \gamma_{\diamond}^{(n)}(z) \right| < C n^{\beta} \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right] \\ \int \gamma_{\diamond}^{(n)}(z) dF(z) > \delta \quad \Rightarrow \quad \sup_u \left| h(u, \gamma_{\diamond}^{(n)}) \right| < C' \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right] \Rightarrow$$

Theorem: Asymptotic Normality of $\hat{\tau}$

Suppose the sequence of weighting kernels $\gamma_+^{(n)}$ and $\gamma_-^{(n)}$ is deterministic, and $\exists \beta \in (0, \frac{1}{2})$, $C, C' > 0$ s.t. $\forall n$ large enough: $\sup_z \left| \gamma_{\diamond}^{(n)}(z) \right| < C n^{\beta} \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right]$, $\sup_u \left| h(u, \gamma_{\diamond}^{(n)}) \right| < C' \mathbb{E} \left[\gamma_{\diamond}^{(n)}(Z_i) \right]$ where $\diamond = \{+, -\}$. Then

$$\frac{\sqrt{n}(\hat{\tau}_{\gamma} - \theta_{\gamma})}{\sqrt{V_{\gamma}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

Design Estimators: Procedure

■ Input:

- samples $\{Z_i, Y_i, W_i\}$ and cutoff c
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Design Estimators: Procedure

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■ Procedure:

S1 guess/estimate $\bar{F}(\cdot)$ and $\bar{w}(\cdot)$ via nonparametric maximum likelihood

S2 solve the minimax program, get γ_+, γ_-

S3 form the point estimate $\hat{\tau}_\gamma$ and its variance \hat{V}_γ

S4 estimate the worst-case bias

$$\hat{B}_\gamma = \sup \left\{ |\text{Bias}[\gamma_\pm, \tau_w; \alpha_0(\cdot), \tau(\cdot), G]| : G \in \mathcal{G}_n, \alpha_{(0)}(\cdot) \in [0, 1], \tau(\cdot) \in \mathcal{T}_M \right\}$$

S5 form the bias-aware CIs at level α as

$$\hat{\tau}_\gamma \pm l_\alpha, l_\alpha = \min \left\{ l : \mathbf{P} \left[\left| b + n^{-\frac{1}{2}} \hat{V}_\gamma^{\frac{1}{2}} \tilde{Z} \right| \leq l \right] \geq 1 - \alpha, \forall |b| \leq \hat{B}_{\gamma, M} \right\}$$

Discussion

Literature: Continuity-Based RD

Most popular: local linear regression (Hahn et al., 2001; G. W. Imbens and Lemieux, 2008)

$$\hat{\tau}_c = \arg \min_{\tau} \left\{ \sum_{i=1}^n \underbrace{K}_{\text{weighting}} \left(\underbrace{\frac{|Z_i - c|}{h_n}}_{\text{bandwidth}} \right) (Y_i - a - \tau W_i - \beta_- (Z_i - c)_- - \beta_+ (Z_i - c)_+)^2 \right\}$$

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Robust CIs (Armstrong and Kolesár, 2020; Calonico et al., 2014; Kolesár and Rothe, 2018); Data-adaptive bandwidths (G. Imbens and Kalyanaraman, 2012)

Literature: Continuity-Based RD extended

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If further assume **convexity** of $\mu_{(w)}(z)$, e.g.:

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Optimization-based RD: the treatment effect τ_c can be estimated (minimax linear estimation) via **numerical convex optimization** (Armstrong and Kolesár, 2018; G. Imbens and Wager, 2019)

Link Noise-Induced RD and Continuity-Based RD

$$\begin{aligned}\mu_{(w)}(z) &= \mathbb{E}[Y_i(w) \mid Z_i = z] \\ &= \frac{\int \mathbb{E}[Y_i(w) \mid U_i = u, Z_i = z] p(z \mid u) dG(u)}{f_G(z)} = \frac{\int \alpha_{(w)}(u) p(z \mid u) dG(u)}{\int p(z \mid u) dG(u)}\end{aligned}$$

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Then the worst-case possible curvature is:

$$\text{Curv}(z, \rho, p) = \sup \left\{ \left| \frac{d^2 \mu_{(w)}(z)}{dz^2} \right| : f_G(z) = \int p(z \mid u) dG(u) \geq \rho > 0, \alpha_{(w)}(\cdot) \in [0, 1] \right\}$$

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Armstrong and Kolesár (2020): fit 4th-degree polynomials to $\mu_{(0)}(z)$ and $\mu_{(1)}(z)$, and take the largest estimated curvature obtained anywhere

Literature: Randomization Inference RD

Posit a non-trivial interval \mathcal{I} with $c \in \mathcal{I}$ s.t.

$$\{Y_i(0), Y_i(1)\} \perp Z_i \mid \{Z_i \in \mathcal{I}\}$$

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- Design-based approach (Rubin, 2008)

- Strong assumption

No **data-driven way** of choosing \mathcal{I}

If the interval \mathcal{I} is known a-priori, the problem collapses to a **RCT**

Measurement Error Induced RD

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- A stronger assumption: observing at least 3 noisy measurements of the latent variable U_i , $\{Z_i, Z'_i, Z''_i\}$
 - (U_i, Z_i, Z'_i, Z''_i) is **joint normal**
 - $\alpha_{(w)}(u) = \mathbb{E}[Y_i(w) | U_i = u]$ is **linear** w.r.t. u

RD with Ordinal Running Variables

Similarly, ordinal Z_i (bond rating, custody security score, etc.) can be seen as a noisy measurement of a latent variable U_i (Li et al., 2021)

Li et al. (2021) assume

$$U_i = \mathbf{X}_i\beta + \epsilon_i$$

then use **inverse-propensity weighting** with estimated propensities $e(u) = \mathbb{P}[Z_i \geq c \mid U_i = u]$ for inference.

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Assuming: U_i can be well predicted by observable \mathbf{X}_i

Measurement Errors

- The **running variable** is unobserved, only a noisy measurement is observed
Bartalotti et al. (2021), Davezies and Le Barbanchon (2017), Dong and Kolesár (2021), and Pei and Shen (2017)
- Measurement error in causal inference beyond RD
Jiang and Ding (2020), Kuroki and Pearl (2014), and Pearl (2012)

A Comparison

RD designs Assumptions

Noise-induced RD
Noise-induced RD (Rokkanen, 2015)

a **known distribution** of the measurement error $p(\cdot \mid u)$
multiple joint-normal noisy measurements (U_i, Z_i, Z'_i, Z''_i)
linear $\alpha_{(w)}(u) = \mathbb{E}[Y_i(w) \mid U_i = u]$

Continuity-based RD
OPTimization-based RD
Randomization inference RD
RD with ordinal Z_i

$\mu_{(w)} = \mathbb{E}[Y(w) \mid Z = z]$ is **smooth**
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an **"RCT"** interval \mathcal{I} : $\{Y_i(0), Y_i(1)\} \perp Z_i \mid \{Z_i \in \mathcal{I}\}$
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Thank you!