

This document contains solution's to Exercises 6.1 to 6.9 from chapter 6, of Statistical inference by Casella and Berger , focusing on sufficient statistic, minimal sufficient statistic.

Some problem's are still in progress , and the document will be updated with their solution's soon.

6.1

Let  $x$  be one observation from a  $N(0, \sigma^2)$  population. Is  $|x|$  a sufficient statistic?

Soln

$$X \sim N(0, \sigma^2)$$

$$f(x|\sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left(-\frac{1}{2\sigma^2} x^2\right)$$

$$= \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left(-\frac{1}{2\sigma^2} |x|^2\right)$$

$$= \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left(-\frac{1}{2\sigma^2} (\tau(x))^2\right)$$

$$= g(\tau(x)|\sigma^2) \cdot h(x)$$

where  $h(x) = 1$  & the above  
conditional function is  $g(\tau(x)|\sigma^2)$

$\Rightarrow |x|$  is sufficient statistic

(6.2)

Let  $x_1, x_2, \dots, x_n$  be independent r.v.  
with densities

$$f_{X_i}(x|\theta) = \begin{cases} e^{\theta - x} & x > \theta \\ 0 & \text{else} \end{cases}$$

Prove that  $T = \min_i \left( \frac{x_i}{\theta} \right)$

is a sufficient statistic for  $\theta$ .

Sol'n

$$f_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n | \theta)$$

$$= \begin{cases} e^{\theta(\sum x_i) - \sum x_i} & \text{if } \forall i \quad x_i > \theta \\ 0 & \text{else} \end{cases}$$

$$= \begin{cases} e^{\theta \cdot \frac{n(n+1)}{2} - \sum x_i} & \text{if } \forall i \quad x_i > \theta \\ 0 & \text{else} \end{cases}$$

$$\theta, 2\theta, 3\theta, \dots, n\theta$$

$$\Rightarrow x_1 > 0, x_2 > 20, x_3 > 30$$

$$\dots x_n > n0$$

$$\Rightarrow x_1 > 0, \frac{x_2}{2} > 0, x_3 > 0, \dots \frac{x_n}{n} > 0$$

this can be re-expressed as

$$\min_i \left( \frac{x_i}{i} \right) > 0$$

$$\Rightarrow f(x | \theta) = \begin{cases} e^{\frac{\theta \cdot n(n+1)}{2} - \sum x_i} & \text{if } \min_i \left( \frac{x_i}{i} \right) > 0 \\ 0 & \text{otherwise} \end{cases}$$

Take the statistic as

$$T(\bar{x}) = \min_i \left( \frac{x_i}{i} \right)$$

$$g(t | \theta) = \begin{cases} e^{\frac{\theta \cdot n(n+1)}{2}} & \text{if } t > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$h(x) = e^{-\sum x_i}$$

$$\Rightarrow f(x|\theta) = g(\tau(x)|\theta) h(x)$$

$$\Rightarrow \tau(x) = \min_i \left( \frac{x_i}{i} \right) \text{ is}$$

a sufficient statistic.

6.3

Let  $x_1, x_2, \dots, x_n$  be a random sample from the Pdf,

$$f(x|u, \sigma) = \frac{1}{\sigma} e^{-(x-u)/\sigma}$$

$$u < x < \infty$$

$$0 < \sigma < \infty$$

Find a two-dimensional sufficient statistic for  $(u, \sigma)$

Soln

$$x_1, x_2, \dots, x_n \sim f(x|u, \sigma)$$

$$f(x|u, \sigma) = \frac{1}{\sigma} e^{-\frac{(x-u)}{\sigma}}$$

Joint distribution

$$f(x | \mu, \sigma) = \begin{cases} \frac{1}{\sigma^n} \cdot e^{-\frac{1}{\sigma^2}(\sum x_i - n\mu)} & \text{if } x_i > \mu \\ 0 & \text{otherwise} \end{cases}$$

This can be re-expressed as

$$f(x | \mu, \sigma) = \begin{cases} \frac{1}{\sigma^n} \cdot \exp\left(-\frac{1}{\sigma^2}(\sum x_i - n\mu)\right) & \min_i x_i > \mu \\ 0 & \text{otherwise} \end{cases}$$

$$T_1(x) = \min_i x_i$$

$$T_2(x) = \sum x_i$$

$$\Rightarrow g(T_1(x), T_2(x) | \mu, \sigma)$$

$$= \frac{1}{\sigma^n} \exp\left(-\frac{1}{\sigma^2}(T_2(x) - n\mu)\right)$$

$$h(x)=1$$

Therefore  $(\min_i x_i, \sum x_i)$  are

the sufficient statistic's for  $\mu, \sigma$

6.4

Prove theorem 6.2.10

SOL<sup>n</sup> Let  $x_1, \dots, x_n$  iid observations from a pdf or pmf  $f(x|\theta)$  that belongs to an exponential family given by

$$f(x|\theta) = h(x) c(\theta) \exp \left( \sum_{i=1}^k w_i(\theta) t_i(x) \right)$$

where  $\theta = (\theta_1, \theta_2, \dots, \theta_d)$ ,  $d \leq k$

SOL<sup>n</sup>

$$f_{x_1, x_2, \dots, x_n}(x_1, x_2, \dots, x_n | \theta)$$

$$= \prod_{j=1}^n h(x_j) c(\theta) \exp \left( \sum_{i=1}^k w_i(\theta) t_i(x_j) \right)$$

$$= C(\theta)^n \left[ \prod_{j=1}^n h(x_j) \right] \left[ \exp \left( \sum_{j=1}^n \sum_{i=1}^k w_i(\theta) t_i(x_j) \right) \right]$$

Example:  $n=3$ ,  $\rightarrow$

$$w_1(\theta) \cdot t_1(x_1) + w_2(\theta) \cdot t_2(x_2) + w_3(\theta) \cdot t_3(x_3)$$

$$\rightarrow w_1(\theta) \cdot t_1(x_1) + w_2(\theta) \cdot t_1(x_2) + w_3(\theta) \cdot t_1(x_3)$$

$$+ w_1(\theta) \cdot t_2(x_1) + w_2(\theta) \cdot t_2(x_2) + w_3(\theta) \cdot t_2(x_3)$$

$$\Rightarrow w_1(\theta) \cdot (\sum t_1(x_i))$$

$$+ w_2(\theta) (\sum t_2(x_i))$$

$$+ w_3(\theta) (\sum t_3(x_i))$$

$$\Rightarrow \sum_{i=1}^k w_i(\theta) \cdot \left[ \sum_{j=1}^n t_i(x_j) \right]$$

$$f(x|\theta)$$

$$= \left[ \prod_{j=1}^n h(x_j) \right] c(\theta)^n \left[ \exp \left( \sum_{i=1}^k w_i(\theta) \cdot \left[ \sum_{j=1}^n t_i(x_j) \right] \right) \right]$$

$n(\omega)$      $g(T(x)|\theta)$

where  $T(x) = \left( \sum_{j=1}^n t_1(x_j), \sum_{j=1}^n t_2(x_j), \dots, \sum_{j=1}^n t_k(x_j) \right)$

(6.5) Let  $x_1, x_2, \dots, x_n$  be independent random variables with Pdf's.

$$f(x_i|\theta) = \begin{cases} \frac{1}{2\theta} & -i(\theta-1) < x_i < i(\theta+1) \\ 0 & \text{otherwise} \end{cases}$$

where  $\theta > 0$ , Find a two-dimensional sufficient statistic for  $\theta$ .

SOL<sup>n</sup>

$$f(x_i | \theta) = \begin{cases} \frac{1}{2\theta} & -i(\theta-1) < x_i < i(\theta+1) \\ 0 & \text{o.w.} \end{cases}$$

$$\Rightarrow f(X | \theta) = \begin{cases} \frac{1}{(2\theta)^n} \cdot \prod_{i=1}^n \frac{1}{i} & \forall i = \{1, 2, \dots, n\} \\ & -i(\theta-1) < x_i < i(\theta+1) \\ 0 & \text{o.w.} \end{cases}$$

$$-i(\theta-1) < x_i < i(\theta+1) \quad \text{for all } i$$

$$\Rightarrow -(\theta-i) < x_1 < \theta+i, \quad -2(\theta-i) < x_2 < 2(\theta+i)$$

⋮

$$\Rightarrow -(\theta-i) < \frac{x_i}{\theta} < \theta+i \quad \forall i$$

$$\Rightarrow -\theta + 1 < \frac{x_i}{\theta} < \theta + 1 \quad \forall i$$

$$\Rightarrow -\theta < \frac{x_i^o}{i} - 1 < \theta \quad \forall i$$

$$\Rightarrow \left| \frac{x_i^o}{i} - 1 \right| < \theta \quad \forall i \quad \& \theta > 0$$

$$\Rightarrow \max_{i^o} \left| \frac{x_i^o}{i} - 1 \right| < \theta$$

$$\Rightarrow f(x|\theta) = \begin{cases} \frac{1}{(2\theta)^n} \prod_{i=1}^n \frac{1}{i} & \text{if } \max_i |x_i^o - 1| < \theta \\ 0 & \text{o.w.} \end{cases}$$

therefore  $h(\omega) = \begin{cases} \prod_{i=1}^n \frac{1}{i} & \text{for } i=1, 2, \dots, n \\ 0 & \text{o.w.} \end{cases}$

$$T_1(x) = \max_i \left| \frac{x_i^o}{i} - 1 \right|$$

sufficient statistic

as we want 2D sufficient statistic

$$-\theta < \frac{x_i}{\theta} - 1 < \theta$$

$$\Rightarrow \max_i \left( \frac{x_i}{\theta} \right) < \theta + 1$$

$$\& \min_i \left( \frac{x_i}{\theta} \right) > -\theta - 1$$

we can take

$$T(x) = \left( \min_i \left( \frac{x_i}{\theta} \right), \max_i \left( \frac{x_i}{\theta} \right) \right)$$

6.6

Let  $x_1, x_2, \dots, x_n$  be random

sample from a gamma ( $\alpha, \beta$ )

Population. Find a two-dimensional

Sufficient Statistic of  $\alpha, \beta$

Sol

$$f(x_i | \alpha, \beta) = \frac{1}{\Gamma(\alpha) \beta^\alpha} x_i^{\alpha-1} \cdot e^{-\frac{x_i}{\beta}}$$

$x_i \in (0, \infty)$

$$f(x | \alpha, \beta) = \left[ \frac{1}{\Gamma(\alpha) \beta^\alpha} \right]^n \cdot e^{-\frac{\sum x_i}{\beta}} \cdot \prod_{i=1}^n x_i^{\alpha-1}$$

$$= \left[ \frac{1}{\Gamma(\alpha) \beta^\alpha} \right]^n \cdot e^{-\frac{\sum x_i}{\beta}} \cdot \left( \prod_{i=1}^n x_i \right)^{\alpha-1}$$



$$g\left(\sum x_i, \prod_{i=1}^n x_i | \alpha, \beta\right) \propto$$

$$h(x) =$$

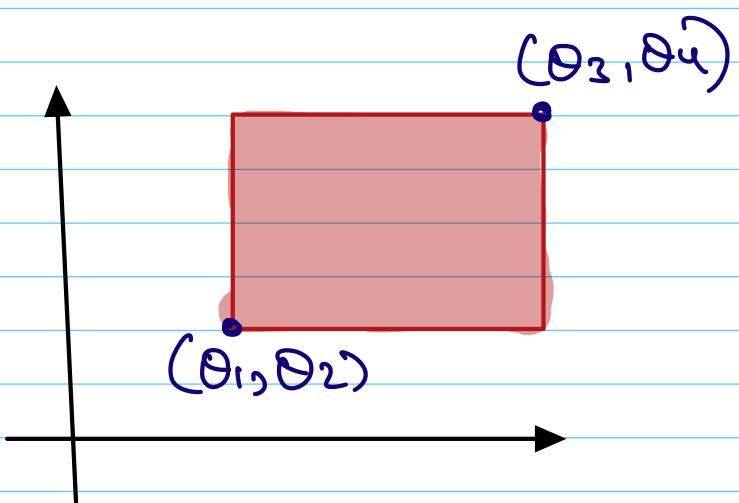
$\Rightarrow$  the 2 Dim sufficient statistic

for  $(\alpha, \beta)$  are  $\left(\sum_{i=1}^n x_i, \prod_{i=1}^n x_i\right)$

6.7

Let  $f(x_1, y_1 | \theta_1, \theta_2, \theta_3, \theta_4)$  be the bivariate Pdf for the uniform distribution on the rectangle with lower left corner  $(\theta_1, \theta_2)$  and upper right corner  $(\theta_3, \theta_4)$  in  $\mathbb{R}^2$ . The parameters satisfy  $\theta_1 < \theta_3$  and  $\theta_2 < \theta_4$ . Let  $(x_1, y_1), \dots, (x_n, y_n)$  be a random sample from this Pdf. find a four-dimensional sufficient statistic for  $(\theta_1, \theta_2, \theta_3, \theta_4)$

Soln



$$f_x(x|\theta_1, \theta_3) = \begin{cases} \frac{1}{\theta_3 - \theta_1} & \text{if } \theta_1 < x < \theta_3 \\ 0 & \text{o.w} \end{cases}$$

$$f_y(y|\theta_2, \theta_4) = \begin{cases} \frac{1}{\theta_4 - \theta_2} & \text{if } \theta_2 < y < \theta_4 \\ 0 & \text{o.w} \end{cases}$$

$$f_{x,y}(x,y|\theta_1, \theta_2, \theta_3, \theta_4)$$

$$= \begin{cases} \left(\frac{1}{\theta_3 - \theta_1}\right) \left(\frac{1}{\theta_4 - \theta_2}\right) & \text{if } \theta_1 < x < \theta_3 \\ 0 & \text{o.w} \end{cases}$$

the Joint Pdf of  $(x_i, y_i) = \{(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)\}$

$$f_{x,y}(x_i, y_i|\theta_1, \theta_2, \theta_3, \theta_4)$$

$$= \begin{cases} \left[\frac{1}{(\theta_3 - \theta_1)(\theta_4 - \theta_2)}\right]^n & \text{if } \forall i \in \{1, 2, \dots, n\} \\ & \theta_1 < x_i < \theta_3 \\ & \theta_2 < y_i < \theta_4 \\ 0 & \text{o.w} \end{cases}$$

$$\forall i, \quad \Theta_1 < x_i < \Theta_3$$

$$\Theta_2 < y_i < \Theta_4$$

This can be expressed as

$$\min_i x_i > \Theta_1$$

$$\max_i x_i < \Theta_3$$

$$\min_i y_i > \Theta_2$$

$$\max_i y_i < \Theta_4$$

$$h(x) = \begin{cases} 1 & x_i \in \mathbb{R}, \text{ for } i=1,2,\dots,n \\ 0 & \text{otherwise} \end{cases}$$

$$T_1(x) = \min_i x_i$$

$$T_2(x) = \min_i y_i$$

$$T_3(x) = \max_i x_i$$

$$T_4(x) = \max_i y_i$$

$$\Rightarrow g(T_1, T_2, T_3, T_4 | \theta_1, \theta_2, \theta_3, \theta_4)$$

$$= \begin{cases} \left[ \frac{1}{(\theta_3 - \theta_1)(\theta_4 - \theta_2)} \right]^n & \text{for } t_1 > \theta_1 \\ & t_2 > \theta_2 \\ & t_3 < \theta_3 \\ & t_4 < \theta_4 \\ 0 & \text{otherwise} \end{cases}$$

Therefore the 4-Dim sufficient statistic

are  $(\min_i x_i, \min_i y_i, \max_i x_i, \max_i y_i)$

for  $(\theta_1, \theta_2, \theta_3, \theta_4)$

6.9

Let  $x_1, x_2, \dots, x_n$  be a random sample from a population with location pdf  $f(x-\theta)$ . Show that the order statistic's

$$T(x_1, \dots, x_n) = (x_{(1)}, x_{(2)}, \dots, x_{(n)})$$

are a sufficient statistic for  $\theta$  and no further reduction is possible

we use Theorem 3.5.6 and work with  $z_1, z_2, \dots, z_n$

$z_1, z_2, \dots, z_n$  iid  $f(x)$  ( $\theta=0$ )

$$x_1 = z_1 + \theta, \quad x_2 = z_2 + \theta, \quad \dots \quad x_n = z_n + \theta$$

where  $-\infty < \theta < \infty$ .

6.9

for each of the following distribution's

let  $x_1, x_2, \dots, x_n$  be a random

sample. Find a minimal

sufficient statistic for  $\theta$

$$(a) f(x|\theta) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\theta)^2}{2}} \quad -\infty < x < \infty$$

$$\quad \quad \quad -\infty < \theta < \infty$$

(normal)

$$x_1, x_2, \dots, x_n \sim N(\theta, 1)$$

$$f(x|\theta) = \frac{1}{(2\pi)^n |I_2|} \cdot e^{-\frac{1}{2} \left[ \sum x_i^2 + n\theta^2 - 2\theta \sum x_i \right]}$$

$$\Rightarrow f(x|\theta) = \frac{1}{(2\pi)^n |I_2|} \cdot e^{-\frac{1}{2} \left[ \sum x_i^2 + n\theta^2 - 2\theta \sum x_i \right]}$$

if

$x, y$  are two random sample  
points then

$$\frac{f(x|\theta)}{f(x_0|\theta)} = e^{-\frac{1}{2} \left[ \sum x_i^2 - \sum x_0^2 - 2\theta(\sum x_i - \sum x_0) \right]}$$

The ratio will be constant as a function

$$\text{of } \theta \iff \sum x_i = \sum x_0$$

Thus  $\sum x_i$  is minimal sufficient statistic

b)  $f(x|\theta) = e^{-(x-\theta)} \cdot \theta < x < \infty$   
 $-\infty < \theta < \infty$

(location Exponent.)

Some  $x_1, x_2, \dots, x_n$  iid  $f(x|\theta)$

$x = (x_1, x_2, \dots, x_n)$  The Joint PDF

$$f(x|\theta) = e^{n\theta - \sum x_i} \quad \forall i = 1, 2, \dots, n$$

$$x_i > \theta$$

$$\Rightarrow f(x|\theta) = e^{\eta\theta - \sum x_i} \quad \min_i x_i > 0$$

$$\frac{f(x|\theta)}{f(y|\theta)} = \frac{e^{\eta\theta} \cdot e^{-\sum x_i}}{e^{\eta\theta} \cdot e^{-\sum y_i}} \frac{I_{(0,\infty)}(\min x_i)}{I_{(0,\infty)}(\min y_i)}$$

this will be independent

$$\Leftrightarrow \min_i x_i = \min_i y_i$$

$$\Leftrightarrow T(x) = T(y)$$

therefore  $\min_i x_i$  is the minimum sufficient statistic.

(c)  $f(x|\theta) = \frac{e^{-(x-\theta)}}{(1+e^{-(x-\theta)})^2}$

$$-\infty < x < \infty$$

$$-\infty < \theta < \infty$$

(Logistic)

Sol'n

Let  $X = (X_1, X_2, \dots, X_n)$  iid

then the Joint distribution

$$f(X|\theta) = \prod_{i=1}^n \frac{e^{-(x_i-\theta)}}{1+e^{-(x_i-\theta)}}^2$$

$$= e^{n\theta} \cdot e^{-\sum x_i} \cdot \prod_{i=1}^n \left( \frac{1}{1+e^{-(x_i-\theta)}} \right)^2$$

$$\frac{f(X|\theta)}{f(x|\theta)} = e^{-(\sum x_i - \sum v_i)} \cdot \prod_{i=1}^n \left[ \frac{1+e^{-(x_i-\theta)}}{1+e^{-(x_i-\theta)}} \right]^2$$

||

this need to be

1

$\Rightarrow$  entire order statistic is minimal sufficient statistic.

(d)  $f(x|\theta) = \frac{1}{\pi [1 + (x-\theta)^2]}$

$-\infty < x < \infty$   
 $-\infty < \theta < \infty$   
 (Cauchy)

Soln  $f(x|\theta) = \frac{1}{\pi^n} \cdot \frac{1}{\prod_{i=1}^n (1 + (x_i - \theta)^2)}$

$$\frac{f(x|\theta)}{f(y|\theta)} = \prod_{i=1}^n \frac{1 + (y_i - \theta)^2}{1 + (x_i - \theta)^2}$$

↓

To make this independent  
of  $\theta$ , then order statistic is  
minimal sufficient statistic

(e)  $f(x|\theta) = \frac{1}{2} e^{-(|x-\theta|)}$

$-\infty < x < \infty$   
 $-\infty < \theta < \infty$   
 (double exponential)

$$\text{so}^n \quad f(x|\theta) = \frac{1}{2^n} \cdot e^{-\sum_{i=1}^n |x_i - \theta|}$$

$$\frac{f(x|\theta)}{f(y|\theta)} = e^{-\left(\sum_{i=1}^n |x_i - \theta| - \sum_{i=1}^n |y_i - \theta|\right)}$$

||

To make this  
independent of  $\theta$

Both  $x, y$  should have same  
number of sample's  $<\theta, >\theta$

i.e # of sample's of  $x < \theta$

= # of sample's of  $y < \theta$