

Student ID \_\_A0197899R\_\_

Student Name \_\_ZHENG HAO\_\_

## Midterm

**Tuesday, 1 October 2019@7-10pm**

1. Do bring your laptop for the midterm.
2. Download
  - i) your individual data (named after your student ID) and
  - ii) d-cdsALL.txt
 from the folder

LumiNUS->files->Midterm

Using wrong data will lead to 0 marks.

3. You are required to fill in the following table to summarize your results and place the table on the front page of your answer sheets.
4. Your answers and R code should be properly named following the format  
StudentID\_studentName.docx or StudentID\_studentName.pdf  
and StudentID\_studentName.R.
5. Your answer sheets and R code should be submitted to the folder  
LumiNUS -> files -> Midterm Student Submission
6. **The submission folder will be closed punctually at 10:00pm** (LumiNUS clock). A 10-30% penalty will apply to late submission. Any late submission will get at least a 10% penalty. A late submission received on the next day, i.e. 2 October, will be given a 20% penalty. A submission later than 2 days, i.e. from 3 October onwards, will get a 30% penalty.
7. This is an open-book test. You are allowed to use any teaching materials and internet resources. *HOWEVER, this is an individual test! You shouldn't collude with any other individual, or plagiarize their work. Suspected collusion or plagiarism will be dealt with according to the NUS' assessment regulations.*
  - Do not let other people see your answers.
  - If other students ask for help, tell them to ask the lecturer.

	<b>Student ID</b>	<b>A0197899R</b>
Q1.	Type of stationarity	<b>Non-stationary</b>
	Fitted model	<b>~arma(1,1)+garch(1,1),con.dist='std'</b>
Q2.	Fitted model	<b>~arma(1,4)+garch(1,1)</b>
	Interval forecasts	<b>Step1 = [89.00989,89.02330] , step2 = [89.00659,89.02004]</b>
Q3	Fitted model	<b>~arma(5,5)+garch(1,1)</b>
	Fitted model with dummy	<b>~arma(0,5)+garch(1,1)</b>

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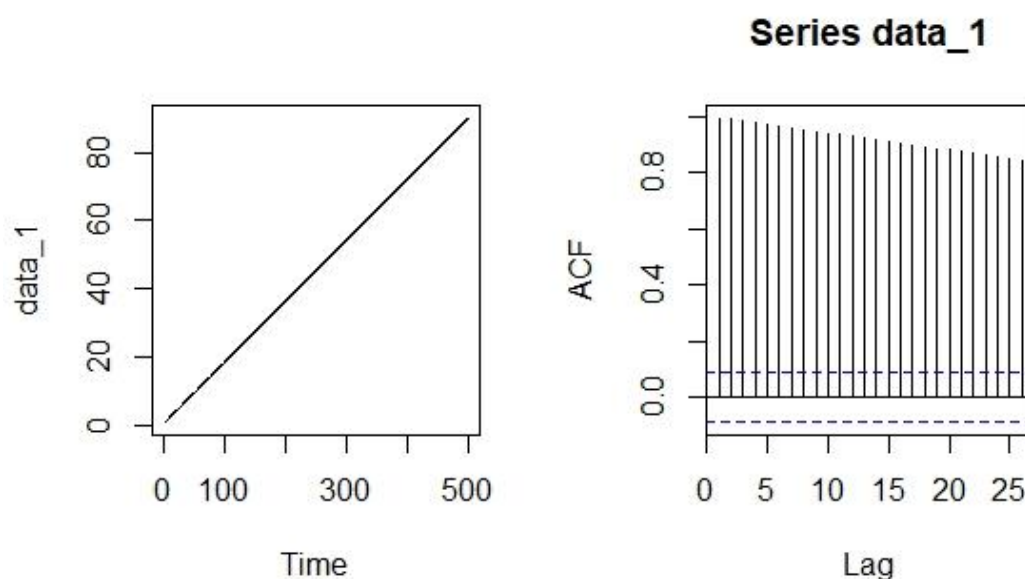
**Analyze data and answer the following 3 questions with proper justifications.**

Load your data using the following R code  
`read.table(filename, sep="\t", header=FALSE)`

**Question 1.** Consider the first time series (the second column in your data), which is monthly sales of SDA – a FinTech company.

1. Plot the time series and ACF plots. Do you think the time series is stationary? Why or why not? What type of non-stationarity can you see in the series and how do you handle this type of non-stationarity? Do you observe any heteroskedasticity in the time series?

**Answer:** the time series and acf plots are as the following. We can see the time series have obvious trend and the acf of it decrease very slowly. So I think it is not stationary. Then I will do the ADF and KPSS test.



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```
Augmented Dickey-Fuller Test

data: data_1
Dickey-Fuller = -7.7034, Lag order = 7, p-value = 0.01
alternative hypothesis: stationary

Warning message:
In adf.test(data_1) : p-value smaller than printed p-value
> kpss.test(data_1)

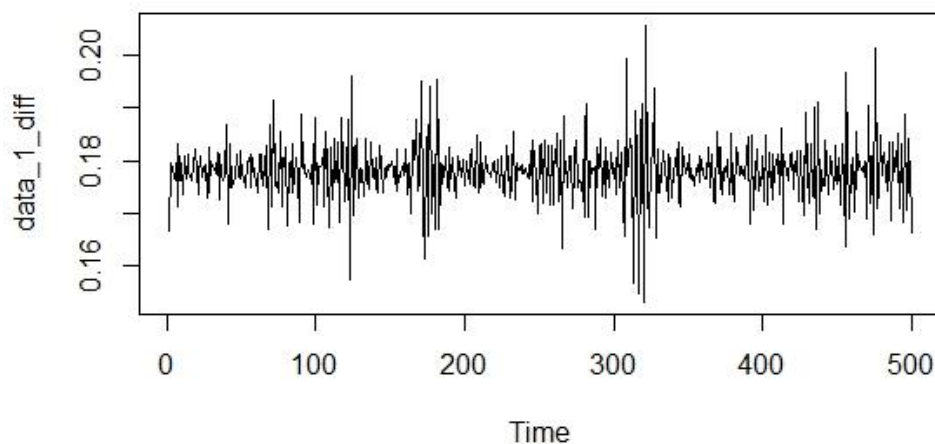
KPSS Test for Level Stationarity

data: data_1
KPSS Level = 8.4317, Truncation lag parameter = 5, p-value = 0.01
```

Both ADF and KPSS reject the null hypothesis, to be conservative, and also with the information of the ts plot, let's assume the original data to be non-stationary.

And it is hard tell

So I choose to do one order difference of the data. And the `diff(data)` is:



It seems mean-reverting, we do the ADF and KPSS test again:

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```
Augmented Dickey-Fuller Test

data: data_1_diff
Dickey-Fuller = -11.577, Lag order = 7, p-value =
0.01
alternative hypothesis: stationary

warning message:
In adf.test(data_1_diff) : p-value smaller than printed p-
> kpss.test(data_1_diff)

KPSS Test for Level Stationarity

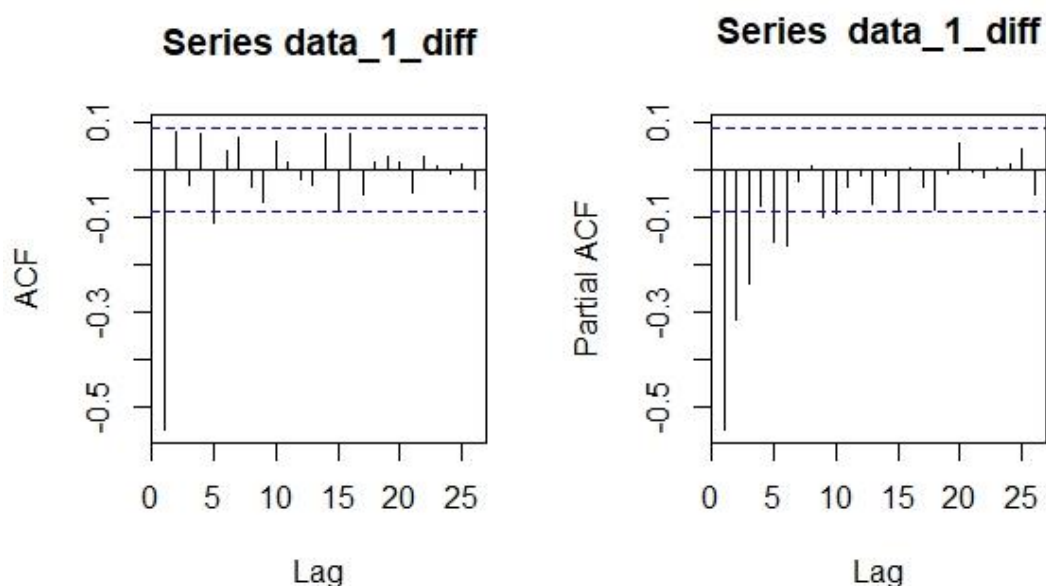
data: data_1_diff
KPSS Level = 0.017571, Truncation lag parameter = 5,
p-value = 0.1
```

We can see that the ADF and KPSS give the same conclusion that the diff\_ts is stationary.

And Heteroskedastic is hard to tell from these two plots.

2. Propose an appropriate model for the sales data and justify your model assumptions with supporting evidence, including but not necessarily limited to autocorrelations, GARCH effect and distributional assumption.

**Answer:** From question 1 we know that we accept the one order difference of original data is stationary, so we plot the acf and pacf of the diff(data):



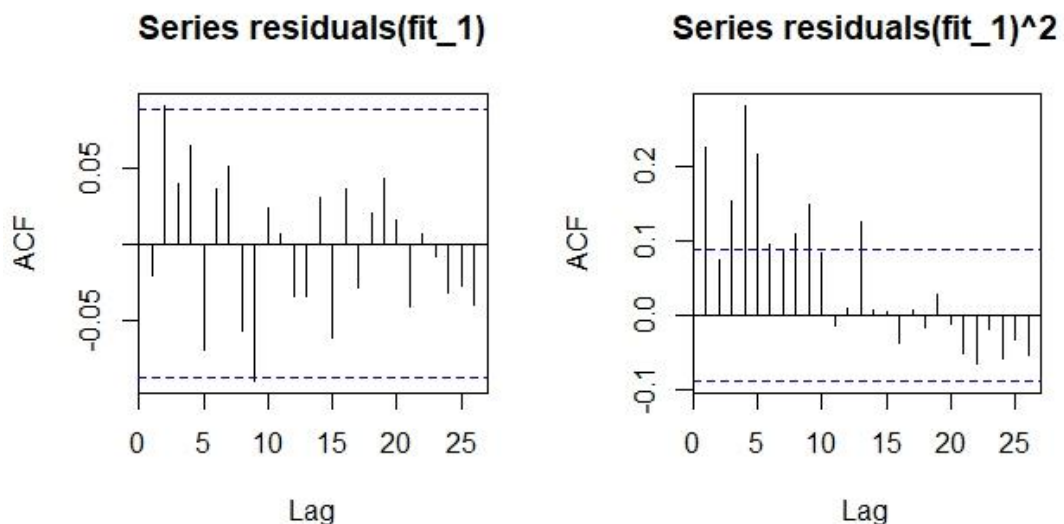
We can see from the ACF that this data may fit MA(1) model, so we use

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arma(0,0,1) for the diff\_data and we plot the ACF of residuals and residuals square.

```
fit_1 = arima(data_1_diff,order=c(0,0,1));  
acf(residuals(fit_1))  
acf(residuals(fit_1)^2)
```



I can see that ARIMA(0,1,1) may fit well for the data, however, the residuals have obvious Garch effect according the acf(residuals^2) plot. Thus I choose use Garch model to fit this data.

I choose three model as the following:

```
gfit_1 = garchFit(~arma(0,1)+garch(1,1),data=data_1_diff)  
gfit_2 = garchFit(~arma(1,0)+garch(1,1),data=data_1_diff)  
gfit_3 = garchFit(~arma(1,1)+garch(1,1),data=data_1_diff)  
according to the AIC and BIC of each model, the gfit_3 model fit the best.
```

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```

Error Analysis:
      Estimate   Std. Error   t value   Pr(>|t|)
mu      1.816e-01         NA         NA         NA
ar1     -2.040e-02    3.596e-05   -567.302   < 2e-16 ***
ma1     -1.000e+00         NA         NA         NA
omega    2.047e-06    6.567e-07     3.116 0.001831 **
alpha1   1.951e-01    5.055e-02     3.859 0.000114 ***
beta1    6.938e-01    6.716e-02    10.330   < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Log Likelihood:
2070.497      normalized:  4.149293

Description:
Tue Oct 01 19:29:14 2019 by user: zhenghaobaby

```

```

Standardised Residuals Tests:

Jarque-Bera Test   R      chi^2   216.3668   0
Shapiro-wilk Test  R      W      0.9638317 9.768234e-10
Ljung-Box Test     R      Q(10)   14.80296  0.1394123
Ljung-Box Test     R      Q(15)   20.92855  0.1391278
Ljung-Box Test     R      Q(20)   22.0783   0.3362772
Ljung-Box Test     R^2    Q(10)    8.146543  0.6145253
Ljung-Box Test     R^2    Q(15)   13.73447  0.5457551
Ljung-Box Test     R^2    Q(20)   22.64489  0.3065655
LM Arch Test       R      TR^2    9.191858  0.6864618

```

```

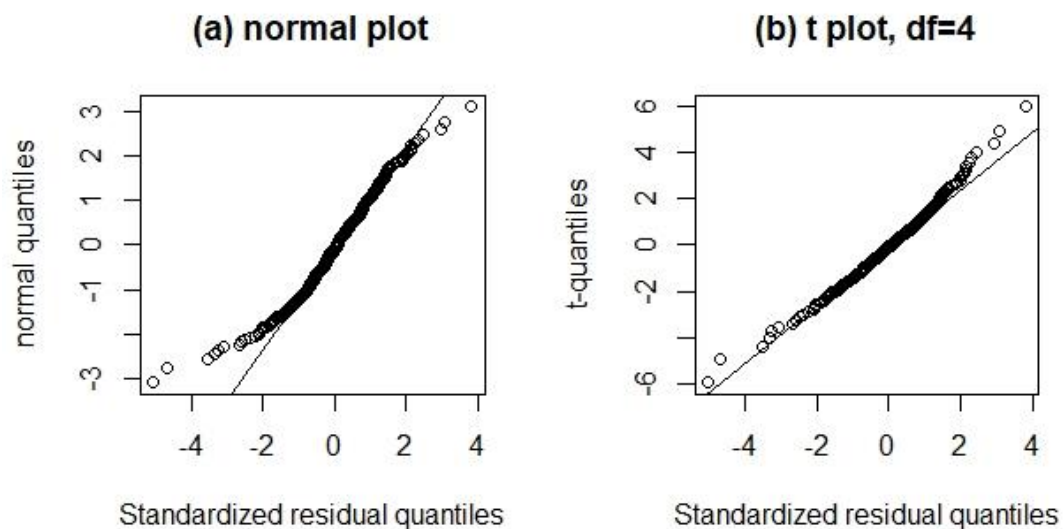
Information Criterion Statistics:
      AIC      BIC      SIC      HQIC
-8.274537 -8.223885 -8.274822 -8.254660

```

Then we should see the residuals of this model.

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We can see that the residuals may more fit t-distribution rather than normal distribution. Thus we should change our assumption about the distribution.

Model:

`gfit_4 = garchFit(~arma(1,1)+garch(1,1),data=data_1_diff,cond.dist = 'std')`

3. Write down the fitted model for sales. Interpret the meaning of each coefficient and discuss significance.

**Answer:**

$$(r_t - 1.8166e - 1) = 2.0562e - 2 (r_{t-1} - 1.8166e - 1) + \varepsilon_t - \varepsilon_{t-1}$$

$$\varepsilon_t = \sqrt{h_t} * z_t \quad z_t \sim t(4.6281)$$

$$h_t = 1.4808e^{-6} + 1.8773e^{-1} \varepsilon_{t-1}^2 + 0.74581 h_{t-1}$$

```
Call:
garchFit(formula = ~arma(1, 1) + garch(1, 1), data = data_1_diff,
cond.dist = "std")

Mean and Variance Equation:
data ~ arma(1, 1) + garch(1, 1)
<environment: 0x00000227f70ec908>
[data = data_1_diff]

Conditional Distribution:
std

Coefficient(s):
      mu      ar1      ma1      omega      alpha1      beta1      shape
1.8166e-01 -2.0562e-02 -1.0000e+00 1.4808e-06 1.8773e-01 7.4581e-01 4.6281e+00
```

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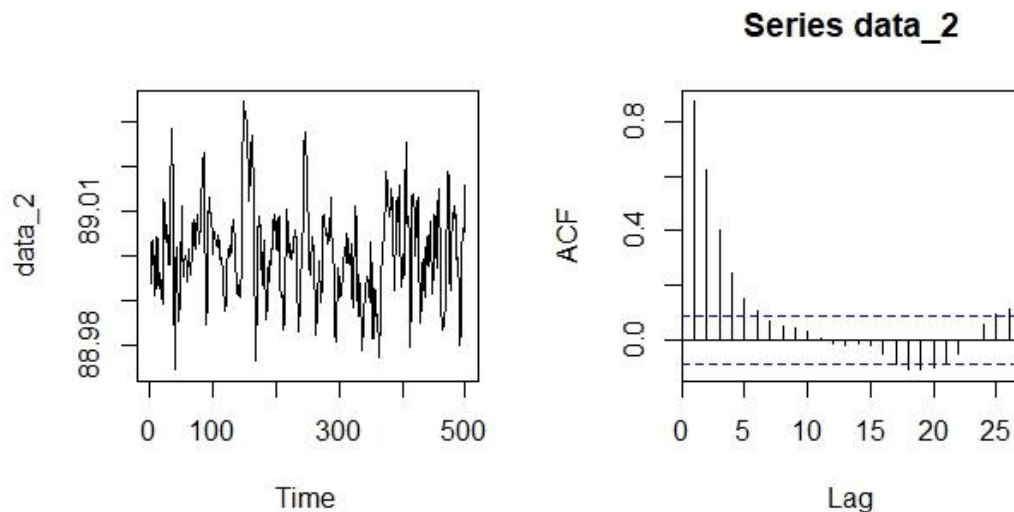
```
Error Analysis:
      Estimate Std. Error  t value Pr(>|t|)
mu      1.817e-01  3.900e-06 46576.823 < 2e-16 ***
ar1     -2.056e-02  3.602e-05 -570.926 < 2e-16 ***
ma1     -1.000e+00  2.116e-02 -47.261 < 2e-16 ***
omega   1.481e-06  6.283e-07   2.357 0.01843 *
alpha1  1.877e-01  5.712e-02   3.287 0.00101 **
beta1   7.458e-01  6.373e-02  11.703 < 2e-16 ***
shape   4.628e+00  9.736e-01   4.753 2e-06 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The mu, ar1, ma1 are the coefficients for conditional mean and the omega, alpha1, beta1 are the coefficients for conditional variance. And the shape is the df of t-distribution. We can see from the picture that all the coefficients are significant according to 5% standard.

**Question 2.** Over years, SDA made gains in many forms of investing in real estate, bonds, etc. as displayed in the third column of your data.

4. Plot the time series and ACF plots. Do you think the time series is stationary? If yes, what type of non-stationarity can you see in the series and how do you handle this type of non-stationarity? Do you observe any heteroskedasticity in the time series?

**Answer:** First I plot the time-series and ACF.



It seems that the time series are mean-reverting, I choose to do the ADF and KPSS test.:



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```

Augmented Dickey-Fuller Test

data: data_2
Dickey-Fuller = -5.9825, Lag order = 7,
p-value = 0.01
alternative hypothesis: stationary

Warning message:
In adf.test(data_2) : p-value smaller than printed p-value
> kpss.test(data_2)

KPSS Test for Level Stationarity

data: data_2
KPSS Level = 0.19497, Truncation lag
parameter = 5, p-value = 0.1

```

Both ADF and KPSS test conclude that the time series are stationary. From the time series plots I think there exists **heteroskedasticity** because some time it has large variance.

- Propose an appropriate model for the gains data. Write down the fitted model and report its AIC. Interpret the meaning of each coefficient and discuss significance.

**Answer:** `fit_2_g = garchFit(~arma(1,4)+garch(1,1),data=data_2)`

**Model** `arma(1,4)+garch(1,1)`

Error Analysis:					
	Estimate	Std. Error	t value	Pr(> t )	
mu	2.425e+01	3.447e-04	70358.954	< 2e-16	***
ar1	7.275e-01	3.405e-05	21366.270	< 2e-16	***
ma1	1.000e+00	5.041e-02	19.836	< 2e-16	***
ma2	2.078e-01	7.272e-02	2.857	0.00428	**
ma3	4.321e-02	7.363e-02	0.587	0.55731	
ma4	3.541e-03	4.721e-02	0.075	0.94022	
omega	1.124e-06	4.192e-07	2.682	0.00733	**
alpha1	5.319e-02	2.222e-02	2.394	0.01668	*
beta1	8.570e-01	4.258e-02	20.124	< 2e-16	***
---					

Standard AIC is

Information Criterion Statistics:			
AIC	BIC	SIC	HQIC
-8.446588	-8.370725	-8.447220	-8.416819

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So the AIC IS

```
> -8.446588*length(data_2)
[1] -4223.294
>
```

The  $\mu$ ,  $\text{ar1}$ ,  $\text{ma1}$ ,  $\text{ma2}$ ,  $\text{ma3}$ ,  $\text{ma4}$  are coefficients for conditional mean, and the  $\omega$ ,  $\alpha_1$ ,  $\beta_1$  are coefficients for conditional variance. The **ma3** and **ma4** can is not significant according to the %5 standard and other coefficients are all significant.

6. Compute 1- and 2-step ahead forecasts and show their confidence interval.

**Answer:**

	meanForecast	meanError	standardDeviation
1	89.01659	0.003419152	0.003419152
2	89.01332	0.006830294	0.003429992

From the picture we can see the 1and 2 step ahead forecast and the interval is as the following:

```
> interval_low
[1] 89.00989 89.00659
> interval_high
[1] 89.02330 89.02004
>
```

Step1 = [89.00989,89.02330] , step2 = [89.00659,89.02004]

**Question 3.** Consider the daily CDS spreads (3-year maturity) of Allstate Insurance from January 01, 2004 to September 19, 2014. The period includes the financial crisis of 2008 so that the CDS spreads vary substantially. The data are in the file d-cdsALL.txt (column 2). Since the spreads are small, we consider the time series  $x_t = 100 \times (\text{spread3y})$ . In addition, sample ACF of  $x_t$  shows strong persistence in the serial dependence. Therefore, we analyze the differenced series  $y_t = (1 - B)x_t$ .

7. Build a time series model for  $y_t$ . Write down the fitted model. [You may start with a model suggested by the auto.arima command in forecast package.]

Answer: `garchFit(~arma(5,5)+garch(1,1),data=data_3_diff)`

**Model:** `arma(5,5)+garch(1,1)`

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### Error Analysis:

	Estimate	Std. Error	t value	Pr(> t )	
mu	-1.055e-06	3.788e-05	-0.028	0.97778	
ar1	2.840e-01	1.147e-01	2.476	0.01330	*
ar2	3.764e-02	8.240e-02	0.457	0.64781	
ar3	-9.285e-03	1.009e-01	-0.092	0.92669	
ar4	-2.556e-01	1.036e-01	-2.467	0.01364	*
ar5	7.360e-01	7.498e-02	9.816	< 2e-16	***
ma1	-2.913e-01	1.176e-01	-2.478	0.01321	*
ma2	-6.066e-04	7.704e-02	-0.008	0.99372	
ma3	-2.999e-02	9.374e-02	-0.320	0.74906	
ma4	3.294e-01	1.020e-01	3.230	0.00124	**
ma5	-7.428e-01	8.069e-02	-9.206	< 2e-16	***
omega	3.270e-07	3.956e-10	826.557	< 2e-16	***
alpha1	1.725e-01	1.457e-02	11.838	< 2e-16	***
beta1	8.629e-01	8.492e-03	101.616	< 2e-16	***
---					

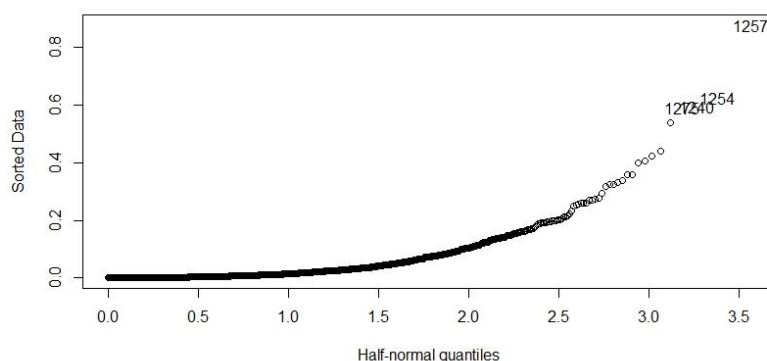
AIC:

```
Information Criterion Statistics:
      AIC      BIC      SIC      HQIC
-5.314800 -5.284810 -5.314851 -5.303968

> -5.3148*length(data_3_diff)
[1] -14700.74
> |
```

8. To improve the fit, identify sequentially the largest four outliers of the fitted model. Write down the fitted model with the four largest outliers included as dummy.

Answer: First, use the halfnorm function to find the outliers



From this picture I can know the index of the largest four outliers is 1257,1254,1240,1275, Thus make this four to be dummy. Then we fit the data again and get the model as the following:

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fit\_3\_g\_mend = garchFit(~arma(0,5)+garch(1,1),data=data\_3\_diff\_mend

**Model:** arma(0,5)+garch(1,1)

```
Information Criterion Statistics:
      AIC      BIC      SIC      HQIC
-5.330328 -5.311048 -5.330349 -5.323364
> -5.330328*length(data_3_diff_mend)
[1] -14743.69
> |
```

And we can see the AIC significantly decrease by setting the largest four as dummy.

9. Let  $a_t$  be the residuals of the model in question (8) and  $\rho_i$  be the lag- $i$  ACF of  $a_t$ . Test  $H_0: \rho_1 = \dots = \rho_{10} = 0$  versus  $H_a: \rho_i \neq 0$  for some  $1 \leq i \leq 10$ . Draw your conclusion.

**Answer:**

```
at = residuals(fit_3_g)
```

```
acf(at)
```

```
Box.test(at,lag=10,type="Ljung-Box")
```

```
at_std = residuals(fit_3_g)/fit_3_g@sigma.t
```

```
acf(at_std)
```

```
Box.test(at_std,lag=10,type="Ljung-Box")
```

```
Box-Ljung test

data:  at
X-squared = 64.827, df = 10, p-value = 4.375e-10
```

We can see for residuals, the p-value is very small and we can reject the  $H_0$ .

```
Box-Ljung test

data:  at_std
X-squared = 8.223, df = 10, p-value = 0.6071
```

However, for the standard residuals, we accept the  $H_0$