

FE5222 Solutions to Homework 2

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1. (Q2) (**Chooser Option**) A chooser option gives the owner the right to choose at time $t_0 > 0$ to buy a call or a put option with strike K and expiry $T > t_0$. Hence at time t_0 the value of the chooser option is

$$\max \{C(t_0), P(t_0)\}$$

where $C(t_0)$ and $P(t_0)$ are the value of call and put option with strike K and expiry T respectively. Use risk neutral approach to price this option at time $t = 0$ under the Black-Scholes-Merton model. Hint: use option call-put parity.

Solution:

Let $C(t, T, K)$ be the price of a call option with strike K and expiry T at time t , and $P(t, T, K)$ be the price of a put option with strike K and expiry T at time t .

Using risk neutral pricing, we have

$$C(t, T, K) = e^{-r(T-t)} \mathbb{E} [(S(T) - K)^+ | \mathcal{F}_t]$$

and

$$P(t, T, K) = e^{-r(T-t)} \mathbb{E} [(K - S(T))^+ | \mathcal{F}_t]$$

where $\mathbb{E}[\bullet]$ is the expectation w.r.t. risk neutral measure.

From

$$(S(T) - K)^+ - (K - S(T))^+ = S(T) - K$$

and using the fact that discounted stock price is a martingale under risk neutral measure, we can derive the call-put parity

$$C(t, T, K) - P(t, T, K) = S(t) - e^{-r(T-t)} K$$

Since $\max \{x, y\} = \max \{x - y, 0\} + y$, the value of a chooser option at time t_0 is

$$\begin{aligned} \max \{C(t_0, T, K), P(t_0, T, K)\} &= C(t_0, T, K) + \max \{0, P(t_0, T, K) - C(t_0, T, K)\} \\ &= C(t_0, T, K) + \max \{0, e^{-r(T-t_0)} K - S(t_0)\} \end{aligned}$$

Hence the price of the chooser option at time $t = 0$ is

$$\begin{aligned} V &= e^{-rt_0} \mathbb{E} [\max \{C(t_0, T, K), P(t_0, T, K)\}] \\ &= e^{-rt_0} \mathbb{E} [C(t_0, T, K)] + e^{-rt_0} \mathbb{E} [\max \{0, e^{-r(T-t_0)} K - S(t_0)\}] \end{aligned}$$

Note that

$$\begin{aligned} e^{-rt_0} \mathbb{E}[C(t_0, T, K)] &= e^{-rt_0} \mathbb{E} \left[e^{-r(T-t_0)} \mathbb{E}[(S(T) - K)^+ | \mathcal{F}_t] \right] \\ &= e^{-rT} \mathbb{E} \left[\mathbb{E}[(S(T) - K)^+ | \mathcal{F}_t] \right] \\ &= e^{-rT} \mathbb{E}[(S(T) - K)^+] \end{aligned}$$

which is the price of a call option with strike K and expiry T .

$$e^{-rt_0} \mathbb{E} \left[\max \left\{ 0, e^{-r(T-t_0)} K - S(t_0) \right\} \right]$$

is the price of a put option with strike $e^{-r(T-t_0)} K$ and expiry t_0 .

Hence

$$V = C(0, T, K) + P(0, t_0, e^{-r(T-t_0)} K)$$

Q.E.D.

2. (Q3) Let $\alpha(t)$ and $\beta(t)$ be non-random time-dependent functions. $X(0) = 1$ and $X(t)$ satisfies the following SDE

$$dX(t) = \alpha(t)X(t)dt + \beta(t)X(t)dW(t)$$

Solution:

Note that $\alpha(t)$ and $\beta(t)$ are non-random, but time-dependent. A common mistake in the homework is to assume these two are constant.

Applying Ito's formula to $\ln X(t)$ we have

$$\begin{aligned} d \ln X(t) &= \frac{dX(t)}{X(t)} - \frac{1}{2} \frac{dX(t)}{X(t)} \frac{dX(t)}{X(t)} \\ &= \alpha(t)dt + \beta(t)dW(t) - \frac{1}{2}\beta^2(t)dt \\ &= \left(\alpha(t) - \frac{1}{2}\beta^2(t) \right) dt + \beta(t)dW(t) \end{aligned}$$

Integrating it from 0 to t and using the fact that $\ln X(0) = 0$, we have

$$\ln X(t) = \int_0^t \left(\alpha(s) - \frac{1}{2}\beta^2(s) \right) ds + \int_0^t \beta(s)dW(s)$$

Hence

$$X(t) = e^{\int_0^t (\alpha(s) - \frac{1}{2}\beta^2(s)) ds + \int_0^t \beta(s)dW(s)}$$

Since β is deterministic (i.e., non-random), we have proven in class

$$\int_0^t \beta(s)dW(s) \sim \mathcal{N} \left(0, \int_0^t \beta^2(s)ds \right)$$

Hence

$$\ln X(t) \sim \mathcal{N} \left(\int_0^t \left(\alpha(s) - \frac{1}{2}\beta^2(s) \right) ds, \int_0^t \beta^2(s)ds \right)$$

Q.E.D.

3. (Q5) Consider the multidimensional model with m stocks,

$$\frac{dS_i}{S_i} = \alpha_i(t)dt + \sum_{j=1}^d \sigma_{i,j}(t)dW_j(t)$$

for $i = 1, \dots, m$. Assume the riskless rate is a constant r .

Suppose there exists a solution $\Theta_j(t), j = 1, \dots, d$, for the market price of risk equations (see lecture notes) and let $\tilde{\mathbb{P}}$ be the risk neutral measure. Then

$$d\tilde{W}_j(t) = dW_j(t) + \Theta_j(t)dt$$

is a Brownian motion under $\tilde{\mathbb{P}}$.

(a) (8 Points) Let

$$\sigma_i(t) = \sqrt{\sum_{j=1}^d \sigma_{i,j}^2(t)}$$

and

$$dB_i(t) = \frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t) dW_j(t)$$

Prove that $B_i(t), i = 1, \dots, m$, is a Brownian motion under \mathbb{P} .

(b) (5 Points) Derive the instantaneous correlation of B_i and B_j for $i \neq j$.

(c) (8 Points) Define $\gamma_i(t) = \frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t) \Theta_j(t)$. Show that

$$\tilde{B}_i(t) = B_i(t) + \int_0^t \gamma_i(u) du$$

is a Brownian motion under $\tilde{\mathbb{P}}$.

(d) (4 Points) Show that

$$\frac{dS_i(t)}{S_i(t)} = rdt + \sigma_i(t) d\tilde{B}_i(t)$$

(e) (5 Points) Derive the instantaneous correlation of $\tilde{B}_i(t)$ and $\tilde{B}_j(t)$ for $i \neq j$. Compare the result with part (b).

Solution:

(a) We will use Levy's Theorem to show that $B_i(t)$ is a Brownian motion. We will have to prove that $B_i(t)$ is continuous, a martingale and $[B_i, B_i](t) = t$

Note that

$$B_i(t) = \int_0^t \frac{1}{\sigma_i(s)} \sum_{j=1}^d \sigma_{i,j}(s) dW_j(s) = \sum_{j=1}^d \int_0^t \frac{1}{\sigma_i(s)} \sigma_{i,j}(s) dW_j(s)$$

Since each $\int_0^t \frac{1}{\sigma_i(s)} \sigma_{i,j}(s) dW_j(s)$ is an Ito's integral, it is continuous and a martingale. Hence $B_i(t)$, as the sum of continuous martingales, is also a continuous martingale.

$$\begin{aligned} dB_i(t)dB_i(t) &= \left(\frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t) dW_j(t) \right) \left(\frac{1}{\sigma_i(t)} \sum_{k=1}^d \sigma_{i,k}(t) dW_k(t) \right) \\ &= \frac{1}{\sigma_i^2(t)} \sum_{j,k} \sigma_{i,j}(t) \sigma_{i,k}(t) dW_j(t) dW_k(t) \\ &= \frac{1}{\sigma_i^2(t)} \sum_j \sigma_{i,j}^2(t) dt \end{aligned}$$

where the last equality follows from the fact that $dW_j(t)dW_k(t) = 0, j \neq k$ and $dW_j(t)dW_k(t) = dt, j = k$.

By definition $\frac{1}{\sigma_i^2(t)} \sum_j \sigma_{i,j}^2(t) dt = dt$, hence $dB_i(t)dB_i(t) = dt$. That is

$$[B_i, B_i](t) = t.$$

We can conclude that $B_i(t)$ is a Brownian motion.

(b) The instantaneous correlation is defined as

$$\frac{\mathbb{E}[dX(t)dY(t)|\mathbb{F}_t]}{\sqrt{\text{Var}(dX(t)|\mathbb{F}_t)\text{Var}(dY(t)|\mathbb{F}_t)}}$$

where $dX(t)$ and $dY(t)$ shall be interpreted as change of $X(t)$ and $Y(t)$ respectively at an infinitesimal time interval $(t, t+dt)$ and $\mathbb{E}[\bullet|\mathcal{F}_t]$ is the conditional expectation and $\text{Var}(\bullet|\mathcal{F}_t)$ is the conditional variance. A less formal definition is

$$\frac{dX(t)dY(t)}{\sqrt{dX(t)dX(t)}\sqrt{dY(t)dY(t)}}$$

Since

$$\begin{aligned} dB_i(t)dB_j(t) &= \left(\frac{1}{\sigma_i(t)} \sum_{k=1}^d \sigma_{i,k}(t)dW_k(t) \right) \left(\frac{1}{\sigma_j(t)} \sum_{l=1}^d \sigma_{j,l}(t)dW_l(t) \right) \\ &= \frac{1}{\sigma_i(t)\sigma_j(t)} \sum_{k,l} \sigma_{i,k}(t)\sigma_{j,l}(t)dW_k(t)dW_l(t) \\ &= \frac{1}{\sigma_i(t)\sigma_j(t)} \sum_k \sigma_{i,k}(t)\sigma_{j,k}(t)dt \end{aligned}$$

$dB_i(t)dB_i(t) = dt$ and $dB_j(t)dB_j(t) = dt$, the instantaneous correlation is

$$\frac{dB_i(t)dB_j(t)}{\sqrt{dB_i(t)dB_i(t)}\sqrt{dB_j(t)dB_j(t)}} = \frac{1}{\sigma_i(t)\sigma_j(t)} \sum_k \sigma_{i,k}(t)\sigma_{j,k}(t)$$

(c) Note that

$$\begin{aligned} d\tilde{B}_i(t) &= dB_i(t) + \gamma_i(t)dt \\ &= \frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t)dW_j(t) + \frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t)\Theta_j(t)dt \\ &= \frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t) (dW_j(t) + \Theta_j(t)dt) \\ &= \frac{1}{\sigma_i(t)} \sum_{j=1}^d \sigma_{i,j}(t)d\tilde{W}_j(t) \end{aligned}$$

Hence the proof is the same as that in (a) with \tilde{W}_j in place of W_j .