

1. Given

$$f(x) = \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)},$$

where $x, \mu \in \mathbb{R}^k$, Σ is a k -by- k positive definite matrix and $|\Sigma|$ is its determinant.
Show that $\int_{\mathbb{R}^k} f(x) dx = 1$.

$$\text{Let } y = x - \mu$$

$$\Rightarrow I = \int_{\mathbb{R}^k} f(x) dx = \int_{\mathbb{R}^k} \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} e^{-\frac{1}{2} y^T \Sigma^{-1} y} dy$$

$\because \Sigma$ is positive definite

$\therefore \exists P$ is orthogonal matrix, Λ is diagonal matrix s.t. $\Sigma = P \Lambda P^T$

$$\Rightarrow \Sigma^{-1} = (P \Lambda P^T)^{-1} = (P^T)^{-1} \Lambda^{-1} P^{-1} = P \Lambda^{-1} P^T \quad (\because P^T = P^{-1})$$

$$\Rightarrow y^T \Sigma^{-1} y = y^T (P \Lambda^{-1} P^T) y$$

Let $z = P^T y$ and know that $|\det(P^T)| = 1$ (i.e., $dz = dy$)

$$\begin{aligned} \Rightarrow y^T \Sigma^{-1} y &= (Pz)^T (P \Lambda^{-1} P^T) (Pz) = z^T P^T P \Lambda^{-1} P^T P z \\ &= z^T \Lambda^{-1} z \end{aligned}$$

$$\Rightarrow z^T \Lambda^{-1} z = \sum_{i=1}^k \frac{z_i^2}{\lambda_i} \quad \text{where } \lambda_i \text{ is } \Lambda \text{'s eigenvalues } i=1, \dots, k$$

$$\text{Since } |\Sigma| = |P \Lambda P^T| = |P| |\Lambda| |P^T| = 1 \cdot |\Lambda| \cdot 1 = \prod_{i=1}^k \lambda_i,$$

$$I = \int_{\mathbb{R}^k} \frac{1}{\sqrt{(2\pi)^k \prod_{i=1}^k \lambda_i}} e^{-\frac{1}{2} \sum_{i=1}^k \frac{z_i^2}{\lambda_i}} dz$$

$$= \frac{1}{\prod_{i=1}^k \sqrt{2\pi \lambda_i}} \int_{\mathbb{R}^k} \left(\prod_{i=1}^k e^{-\frac{z_i^2}{2\lambda_i}} \right) dz$$

$$= \left(\prod_{i=1}^k \frac{1}{\sqrt{2\pi \lambda_i}} \right) \left(\prod_{i=1}^k \int_{-\infty}^{\infty} e^{-\frac{z_i^2}{2\lambda_i}} dz_i \right)$$

Note $\int_{-\infty}^{\infty} e^{-ax^2} dx = \sqrt{\frac{\pi}{a}}$

Then, $\int_{-\infty}^{\infty} e^{-\frac{z_i^2}{2\lambda_i}} dz_i = \sqrt{\frac{\pi}{(\frac{1}{2\lambda_i})}} \quad (\text{Let } a = \frac{1}{2\lambda_i})$

$$= \sqrt{2\pi\lambda_i}$$

$$\Rightarrow I = \prod_{i=1}^k \left(\frac{\sqrt{2\pi\lambda_i}}{\sqrt{2\pi\lambda_i}} \right) = 1 \quad \#$$

2. Let A, B be n -by- n matrices and x be a n -by-1 vector.

(a) Show that $\frac{\partial}{\partial A} \text{trace}(AB) = B^T$.

(b) Show that $x^T A x = \text{trace}(x x^T A)$.

(b) Derive the maximum likelihood estimators for a multivariate Gaussian.

(a) Note that $\text{trace}(AB) = \sum_{i=1}^n (AB)_{ii} = \sum_{i=1}^n \sum_{j=1}^n a_{ij} b_{ji}$

Suppose we partial a_{kl}

$$\begin{aligned} \left(\frac{\partial}{\partial A} \text{trace}(AB) \right)_{kl} &= \frac{\partial}{\partial a_{kl}} \left(\sum_{i=1}^n \sum_{j=1}^n a_{ij} b_{ji} \right) \\ &= \frac{\partial}{\partial a_{kl}} (a_{kl} b_{lk}) \\ &= b_{lk} \end{aligned}$$

Thus, $\frac{\partial}{\partial A} \text{trace}(AB) = B^T \quad \#$

(b) Since $x^T A x$ is scalar,

$$x^T A x = \text{trace}(x^T A x)$$

Note that $\text{trace}(AB) = \text{trace}(BA)$

$$\Rightarrow \begin{cases} \text{tr}(x^T (A x)) = \text{tr}((A x) x^T) \\ \text{tr}((x x^T) A) = \text{tr}(A (x x^T)) \end{cases}$$

Thus, $x^T A x = \text{tr}(x x^T A) \quad \#$

$$(c) \text{ Note } L(\mu, \Sigma; x) = \prod_{i=1}^N \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} \exp\left(-\frac{1}{2}(x_i - \mu)^T \Sigma^{-1}(x_i - \mu)\right)$$

$$\text{where } x = \{x_1, x_2, \dots, x_N\}$$

$$\text{Let } \ell = \ln L$$

$$\begin{aligned} \ell(\mu, \Sigma) &= \ln \left[(2\pi)^k |\Sigma|^{-\frac{N}{2}} \exp\left(-\frac{1}{2}(x_i - \mu)^T \Sigma^{-1}(x_i - \mu)\right) \right] \\ &= -\frac{Nk}{2} \ln(2\pi) - \frac{N}{2} \ln |\Sigma| - \frac{1}{2}(x_i - \mu)^T \Sigma^{-1}(x_i - \mu) \end{aligned}$$

$$\begin{aligned} \text{Then, } \nabla_{\mu} \ell &= -\frac{1}{2} \sum_{i=1}^N \nabla_{\mu} ((x_i - \mu)^T \Sigma^{-1}(x_i - \mu)) \\ &= -\frac{1}{2} \sum_{i=1}^N (-2 \Sigma^{-1}(x_i - \mu)) \\ &= \Sigma^{-1} \sum_{i=1}^N (x_i - \mu) \end{aligned}$$

$$\text{Let } \nabla_{\mu} \ell = 0$$

$$\because \Sigma \text{ is positive definite } \therefore \exists \Sigma^{-1} \text{ s.t. } \sum_{i=1}^N (x_i - \mu) = 0$$

$$\Rightarrow \sum_{i=1}^N x_i - N\mu = 0$$

$$\Rightarrow N\mu = \sum_{i=1}^N x_i$$

$$\text{Then, } \hat{\mu}_{MLE} = \frac{1}{N} \sum_{i=1}^N x_i$$

$$\begin{aligned} \ln(\mu, \Sigma) &= -\frac{Nk}{2} \ln(2\pi) - \frac{N}{2} \ln |\Sigma| - \frac{1}{2} \sum_{i=1}^N \text{tr}((x_i - \mu)(x_i - \mu)^T \Sigma^{-1}) \text{ by (b)} \\ &= -\frac{Nk}{2} \ln(2\pi) - \frac{N}{2} \ln |\Sigma| - \frac{1}{2} \text{tr}\left(\left(\sum_{i=1}^N (x_i - \mu)(x_i - \mu)^T\right) \Sigma^{-1}\right) \end{aligned}$$

$$\text{Let } S_{\mu} = \sum_{i=1}^N (x_i - \mu)(x_i - \mu)^T \text{ and } C = -\frac{Nk}{2} \ln(2\pi)$$

$$\ln(\mu, \Sigma) = C - \frac{N}{2} \ln |\Sigma| - \frac{1}{2} \text{tr}(S_{\mu} \Sigma^{-1})$$

$$\text{Note that } \frac{\partial}{\partial A} \ln |A| = (A^{-1})^T, \quad \frac{\partial}{\partial A} \text{tr}(BA) = B^T$$

$$\text{Let } U = \Sigma^{-1}$$

$$\ell = C + \frac{N}{2} \ln |U| - \frac{1}{2} \text{tr}(S_{\mu} U)$$

$$\Rightarrow \frac{\partial}{\partial U} = \frac{N}{2} (U^{-1})^T - \frac{1}{2} (S_{\mu})^T$$

$$\because U \text{ and } S_{\mu} \text{ are symmetric} \quad \therefore (U^{-1})^T = \Sigma^T = \Sigma, \quad (S_{\mu})^T = S_{\mu}$$

$$\Rightarrow \frac{\partial}{\partial U} = \frac{N}{2} \Sigma - \frac{1}{2} S_{\mu}$$

$$\text{Let } \frac{\partial}{\partial U} = 0 \quad \Rightarrow \frac{N}{2} \Sigma - \frac{1}{2} S_{\mu} = 0$$

$$\Rightarrow N \Sigma = S_{\mu}$$

$$\begin{aligned} \text{Thus, } \hat{\Sigma}_{MLE} &= \frac{S_{\mu}}{N} = \frac{1}{N} \sum_{i=1}^N (x_i - \mu)(x_i - \mu)^T \\ &= \frac{1}{N} \sum_{i=1}^N (x_i - \hat{\mu}_{MLE})(x_i - \hat{\mu}_{MLE})^T \quad \# \end{aligned}$$

3. Questions

Compared to Softmax regression, which natively handles multi-class problems, in what situations does One-vs-Rest perform poorly?

Does it have any advantages that Softmax does not?