

MCMC_PDF

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1 MCMC (Metropolis-Hesting) on Supernova-Ia dataset

1.1 Cosmology with Supernovae Ia: theory

- The flux of supernova of luminosity L is given by $f = \frac{L}{4\pi D_L^2}$, Where D_L is the *Luminosity distance*.
- D_L as a function of *redshift* Z $D_L = \frac{(1+z)c}{H_0 \sqrt{|1-\Omega|}} S_k(r)$
- Where $r(z) = \sqrt{|1-\Omega|} \int_0^z \frac{dz'}{\sqrt{\Omega_m(1+z')^3 + \Omega_v + (1-\Omega)(1+z')^2}}$, $S_k(r) = \sin r$, r , $\sinh r$ depending upon $\Omega > 1, =1, <1$, here $\Omega \equiv \Omega_m + \Omega_v$
- For flat universe $\Omega = 1$, $D_L(z) = 3000h^{-1}(1+z) \int_0^z \frac{dz'}{\sqrt{\Omega_m(1+z')^3 + 1 - \Omega_m}} \text{Mpc}$, Where $H_0 = 100.h \text{KmS}^{-1} \text{Mpc}^{-1}$
- For flat only universe, another way is to use fitting formula without any integration $D_L(z) = \frac{c}{H_0}(1+z)[\eta(1, \Omega_m) - \eta(\frac{1}{1+z}, \Omega_m)]$
- Where $\eta(a, \Omega_m) = 2\sqrt{s^3 + 1}[\frac{1}{a^4} - 0.1540\frac{s}{a^3} + 0.4304\frac{s^2}{a^2} + 0.19097\frac{s^3}{a} + 0.066941\frac{s^4}{a^0}]^{-1/8}$, and $s^3 \equiv \frac{(1-\Omega_m)}{\Omega_m}$, has accuracy of 0.4% for $0.2 \leq \Omega_m \leq 1$
- Fluxes are usually expressed in magnitudes, where $m = -2.5 \log_{10} F + \text{constant}$. The distance modulus is $\mu = m - M$, where M is the absolute magnitude, which is the value of m if the source is at a distance 10pc:
- $\mu = 25 - 5 \log_{10} h + 5 \log_{10}(\frac{D_L^*}{\text{Mpc}})$, The Hubble constant has been factored out of D_L : $D_L^* \equiv D_L(h = 1)$. If we have measurements of μ , then we can use Bayesian arguments to estimate the parameters Ω_m, Ω_v, h .

Model

defining η : $\eta(a, \Omega_m) = 2\sqrt{s^3 + 1}[\frac{1}{a^4} - 0.1540\frac{s}{a^3} + 0.4304\frac{s^2}{a^2} + 0.19097\frac{s^3}{a} + 0.066941\frac{s^4}{a^0}]^{-1/8}$, where $s^3 \equiv \frac{(1-\Omega_m)}{\Omega_m}$

defining the function for luminosity distance D_L : $D_L(z) = \frac{c}{H_0}(1+z)[\eta(1, \Omega_m) - \eta(\frac{1}{1+z}, \Omega_m)]$

cleverly we have modified this function coefficient as $3000.(1+z)$ as we are choosing $D_L^*, h \equiv 1$
 $D_L^*(z) = 3000.0.(1+z)[\eta(1, \Omega_m) - \eta(\frac{1}{1+z}, \Omega_m)]$

function for distance modulus μ : $\mu = 25 - 5\log_{10}h + 5\log_{10}(\frac{D_L^*}{Mpc})$

1.1.1 Above was the model for which we have to find out the best fit parameters values for Ω_m and h using bayesian methods.

2 Baye's theorem : $P(\theta|D) = \frac{P(\theta).P(D|\theta)}{\Sigma P(D)}$

where θ 's are the parameter , D is the data

* $P(\theta|D)$ is called 'Posterior Probability'.

* $P(\theta)$ is called 'Prior'.

* $P(D|\theta)$ is the 'Likelihood'.

* $\Sigma P(D)$ is our 'Normalization constant'.

- Here, in this problem we have chosen a "flat prior", which means that the distribution of posterior does not depends on the prior's distribution.
- Also, We have a stationary model. So We don't need to bother about the "Normalizing term" as well.
- In this case, the baye's formula can replaced into : $P(\theta|D) \propto P(D|\theta)$

3 MH algorithm:

- looping from (1,N=10000)

Method in Marcov's chain:

- Next values of parameters are chosen from a normal distribution centered around the previous values of parameter with some standard deviation σ (here 0.01).
- Calculate the log likelihood for the new randomly generated parameter values.
- If next value of likelihood is grater than previous 'or' ratio of likelihoods is greater than random uniform (0,1). Then accept the parameter values, else repeat the previous step.