

# MCMC\_PDF

April 21, 2020

## 1 MCMC (Metropolis-Hesting) on Supernova-Ia dataset

### 1.1 Cosmology with Supernovae Ia: theory

- The flux of supernova of luminosity  $L$  is given by  $f = \frac{L}{4\pi D_L^2}$ , Where  $D_L$  is the *Luminosity distance*.
- $D_L$  as a function of *redshift*  $z$   $D_L = \frac{(1+z)c}{H_0 \sqrt{|1-\Omega|}} S_k(r)$
- Where  $r(z) = \sqrt{|1-\Omega|} \int_0^z \frac{dz'}{\sqrt{\Omega_m(1+z')^3 + \Omega_v + (1-\Omega)(1+z')^2}}$ ,  $S_k(r) = \sin r$ ,  $r$ ,  $\sinh r$  depending upon  $\Omega > 1, =1, <1$ , here  $\Omega \equiv \Omega_m + \Omega_v$
- For flat universe  $\Omega = 1$ ,  $D_L(z) = 3000h^{-1}(1+z) \int_0^z \frac{dz'}{\sqrt{\Omega_m(1+z')^3 + 1 - \Omega_m}} \text{Mpc}$ , Where  $H_0 = 100 h \text{KmS}^{-1} \text{Mpc}^{-1}$
- For flat only universe, another way is to use fitting formula without any integration  $D_L(z) = \frac{c}{H_0}(1+z)[\eta(1, \Omega_m) - \eta(\frac{1}{1+z}, \Omega_m)]$
- Where  $\eta(a, \Omega_m) = 2\sqrt{s^3 + 1}[\frac{1}{a^4} - 0.1540\frac{s}{a^3} + 0.4304\frac{s^2}{a^2} + 0.19097\frac{s^3}{a} + 0.066941\frac{s^4}{a^0}]^{-1/8}$ , and  $s^3 \equiv \frac{(1-\Omega_m)}{\Omega_m}$ , has accuracy of 0.4% for  $0.2 \leq \Omega_m \leq 1$
- Fluxes are usually expressed in magnitudes, where  $m = -2.5 \log_{10} F + \text{constant}$ . The distance modulus is  $\mu = m - M$ , where  $M$  is the absolute magnitude, which is the value of  $m$  if the source is at a distance 10pc:
- $\mu = 25 - 5 \log_{10} h + 5 \log_{10}(\frac{D_L^*}{\text{Mpc}})$ , The Hubble constant has been factored out of  $D_L$ :  $D_L^* \equiv D_L(h = 1)$ . If we have measurements of  $\mu$ , then we can use Bayesian arguments to estimate the parameters  $\Omega_m, \Omega_v, h$ .

# Model

**defining  $\eta$ :**  $\eta(a, \Omega_m) = 2\sqrt{s^3 + 1}[\frac{1}{a^4} - 0.1540\frac{s}{a^3} + 0.4304\frac{s^2}{a^2} + 0.19097\frac{s^3}{a} + 0.066941\frac{s^4}{a^0}]^{-1/8}$ , where  $s^3 \equiv \frac{(1-\Omega_m)}{\Omega_m}$

**defining the function for luminosity distance  $D_L$ :**  $D_L(z) = \frac{c}{H_0}(1+z)[\eta(1, \Omega_m) - \eta(\frac{1}{1+z}, \Omega_m)]$

cleverly we have modified this function coefficient as  $3000.(1+z)$  as we are choosing  $D_L^*, h \equiv 1$   
 $D_L^*(z) = 3000.0.(1+z)[\eta(1, \Omega_m) - \eta(\frac{1}{1+z}, \Omega_m)]$

function for distance modulus  $\mu$ :  $\mu = 25 - 5\log_{10}h + 5\log_{10}(\frac{D_L^*}{Mpc})$

**1.1.1** Above was the model for which we have to find out the best fit parameters values for  $\Omega_m$  and  $h$  using bayesian methods.

**2 Baye's theorem :**  $P(\theta|D) = \frac{P(\theta).P(D|\theta)}{\Sigma P(D)}$

where  $\theta$ 's are the parameter ,  $D$  is the data

\*  $P(\theta|D)$  is called 'Posterior Probability'.

\*  $P(\theta)$  is called 'Prior'.

\*  $P(D|\theta)$  is the 'Likelihood'.

\*  $\Sigma P(D)$  is our 'Normalization constant'.

- Here, in this problem we have chosen a "flat prior", which means that the distribution of posterior does not depends on the prior's distribution.
- Also, We have a stationary model. So We don't need to bother about the "Normalizing term" as well.
- In this case, the baye's formula can replaced into :  $P(\theta|D) \propto P(D|\theta)$

### 3 MH algorithm:

- looping from (1,N=10000)

**In Marcov's chain parts we basically do like this:**

- Next values of parameters are chosen from a normal distribution centered around the previous values of parameter with some standard deviation  $\sigma$ (here 0.01).

next,

- we calculate the log likelyhood for the new randomly generated parameter values.
- if next value of likelyhood is grater than previous 'or' ratio of likelyhoods is greater than random uniform (0,1). Then accept the parameter values, else repeat the previous step.