

1 day

naive bayes:  $\hat{y} = \underset{y}{\operatorname{argmax}} p(y|x) = \underset{y}{\operatorname{argmax}} \frac{p(x|y) p(y)}{p(x)} \xrightarrow[p(x) \text{ is Constant}]{p(y=1)=p(y=2)=p(y=3)} \underset{y}{\operatorname{argmax}} p(x|y) \quad \textcircled{\text{I}}$

$$(x|y) \sim \mathcal{N}(\mu, \Sigma) \Rightarrow p(x|y) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)\right]$$

$$\begin{aligned} \Rightarrow \ln p(x|y) &= -\ln(2\pi)^{D/2} |\Sigma|^{1/2} - \frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu) \\ &= -\frac{D}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma| - \frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu) \quad \textcircled{\text{II}} \end{aligned}$$

$$\xrightarrow{\textcircled{\text{I}}, \textcircled{\text{II}}} \hat{y} = \underset{y}{\operatorname{argmax}} p(x|y) = \underset{y}{\operatorname{argmax}} \ln p(x|y) = \underset{y}{\operatorname{argmax}} \left\{ -\frac{D}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma| - \frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu) \right\}$$

$$\xrightarrow{\text{leave out Constants}} = \underset{y}{\operatorname{argmax}} \left\{ -\frac{1}{2} \ln |\Sigma| - \frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu) \right\} = \underset{y}{\operatorname{argmin}} \left\{ \ln |\Sigma| + (x-\mu)^T \Sigma^{-1} (x-\mu) \right\}$$

$$|\Sigma_1| = 0.49, \quad \Sigma_1^{-1} = \begin{bmatrix} 10/7 & 0 \\ 0 & 10/7 \end{bmatrix}, \quad \ln |\Sigma_1| = -0.713$$

$$|\Sigma_2| = 0.16 - 0.09 = 0.07, \quad \Sigma_2^{-1} = \begin{bmatrix} \frac{20}{7} & -\frac{30}{7} \\ -\frac{30}{7} & \frac{80}{7} \end{bmatrix}, \quad \ln |\Sigma_2| = -2.660$$

$$|\Sigma_3| = 0.56 - 0.04 = 0.52, \quad \Sigma_3^{-1} = \begin{bmatrix} \frac{20}{13} & -\frac{5}{13} \\ -\frac{5}{13} & \frac{35}{26} \end{bmatrix}, \quad \ln |\Sigma_3| = -0.654$$

$$x = [50, 0.5] \xrightarrow{y=1} (50, 0.5) \begin{pmatrix} 1/7 & 0 \\ 0 & 1/7 \end{pmatrix} \begin{pmatrix} 50 \\ 0.5 \end{pmatrix} = \frac{5000.5}{14} = 3571.786$$

$$\xrightarrow{y=2} (49, -0.5) \begin{pmatrix} 2/7 & -3/7 \\ -3/7 & 8/7 \end{pmatrix} \begin{pmatrix} 49 \\ -0.5 \end{pmatrix} = \frac{49510}{7} = 7072.857$$

$$\xrightarrow{y=3} (49, -0.5) \begin{pmatrix} 2/13 & -5/13 \\ -3/13 & 35/26 \end{pmatrix} \begin{pmatrix} 49 \\ -0.5 \end{pmatrix} = \frac{386155}{104} = 3713.029$$

$$\arg\min \{ 3571.786 - 0.713, 7072.857 - 2.660, 3713.029 - 0.654 \} = \boxed{1 = \hat{y}}$$

$$x = [0.5, 0.5] \xrightarrow{y=1} (0.5, 0.5) \begin{pmatrix} 1/7 & 0 \\ 0 & 1/7 \end{pmatrix} \begin{pmatrix} 0.5 \\ 0.5 \end{pmatrix} = \frac{5}{7} \rightarrow \frac{5}{7} - 0.713 = 0.001286$$

$$\xrightarrow{y=2} (-0.5, 0.5) \begin{pmatrix} 2/7 & -3/7 \\ -3/7 & 8/7 \end{pmatrix} \begin{pmatrix} -0.5 \\ 0.5 \end{pmatrix} = \frac{10}{7} \rightarrow \frac{10}{7} - 2.660 = -1.2314$$

$$\xrightarrow{y=3} (-0.5, -0.5) \begin{pmatrix} 2/13 & -5/13 \\ -5/13 & 35/26 \end{pmatrix} \begin{pmatrix} -0.5 \\ -0.5 \end{pmatrix} = \frac{55}{104} \rightarrow \frac{55}{104} - 0.654 = -0.125$$

$$\arg\min \rightarrow \boxed{\hat{y} = 2}$$

$$y(x_n, \omega) = \omega_0 + \sum_{i=1}^D \omega_i x_{ni} \quad \rightarrow E_D(\omega) = \frac{1}{2} \sum_{n=1}^N [y(x_n, \omega) - y_n]^2$$

$$\tilde{y}_n = \omega_0 + \sum_{i=1}^D \omega_i (x_{ni} + \epsilon_{ni}) = \omega_0 + \sum_{i=1}^D \omega_i x_{ni} + \sum_{i=1}^D \omega_i \epsilon_{ni} = y(x_n, \omega) + \sum_{i=1}^D \omega_i \epsilon_{ni}$$

$$\begin{aligned} E_D(\omega) &= \frac{1}{2} \sum_{n=1}^N \{ \tilde{y}_n - y_n \}^2 = \frac{1}{2} \sum_{n=1}^N \{ \tilde{y}_n^2 - 2\tilde{y}_n y_n + y_n^2 \} \\ &= \frac{1}{2} \sum_{n=1}^N \left\{ y(x_n, \omega)^2 + \underbrace{2y(x_n, \omega) \sum_{i=1}^D \omega_i \epsilon_{ni}}_{(2)} + \underbrace{\left( \sum_{i=1}^D \omega_i \epsilon_{ni} \right)^2}_{(3)} - 2y(x_n, \omega)y_n - 2 \underbrace{\sum_{i=1}^D \omega_i \epsilon_{ni} y_n}_{(5)} + y_n^2 \right\} \end{aligned} \quad (I)$$

حالا باید از عبارت فوقاً  $(E_D(\omega))$  استفاده کنیم.

باقی به این  $\epsilon_i \sim N(0, \sigma^2 I)$  است،  $E(\epsilon_i) = 0$  می باشد. حساب (2) و (3) باقی به حالت خطای بودن می باشد. صفر شود. حال برابر صفر سوم داریم:

$$\begin{aligned} (3) \quad E \left[ \left( \sum_{i=1}^D \omega_i \epsilon_{ni} \right)^2 \right] &= E \left[ \sum_{i=1}^D \sum_{j=1, j \neq i}^D \omega_i \epsilon_{ni} \omega_j \epsilon_{nj} + \sum_{i=1}^D \omega_i^2 \epsilon_{ni}^2 \right] \\ &= \sum_{i=1}^D \sum_{j=1, j \neq i}^D \omega_i \omega_j E(\epsilon_{ni} \epsilon_{nj}) + \sum_{i=1}^D \omega_i^2 E(\epsilon_{ni}^2) \end{aligned}$$

$$\begin{aligned} \text{با } \epsilon_i \text{ iid} \quad \left. \begin{aligned} E(\epsilon_{ni} \epsilon_{nj}) &= E(\epsilon_{ni}) E(\epsilon_{nj}) = 0 \\ E((\epsilon_{ni} - 0)^2) &= \text{var}(\epsilon_{ni}) = \sigma^2 \end{aligned} \right\} \rightarrow = 0 + \sum_{i=1}^D \omega_i^2 \sigma^2 \quad (II) \end{aligned}$$

$$\begin{aligned} (I), (II) \rightarrow E(E_D(\omega)) &= \frac{1}{2} \sum_{n=1}^N \left\{ y(x_n, \omega)^2 + 0 + \sum_{i=1}^D \omega_i^2 \sigma^2 - 2y(x_n, \omega)y_n - 0 + y_n^2 \right\} \\ &= \frac{1}{2} \sum_{n=1}^N \left\{ y(x_n, \omega)^2 - 2y(x_n, \omega)y_n + y_n^2 \right\} + \frac{N}{2} \sum_{i=1}^D \omega_i^2 \sigma^2 \\ &= \frac{1}{2} \sum_{n=1}^N [y(x_n, \omega) - y_n]^2 + \frac{N \sigma^2}{2} \sum_{i=1}^D \omega_i^2 \\ &= \boxed{E_D(\omega) + \frac{N \sigma^2}{2} \sum_{i=1}^D \omega_i^2} \end{aligned}$$

مسئله 3  
الف) به جای اینکه  $p(c_i | x)$  را تعیین کنیم، فرضیات که  $p(c_i | x)$  را به صورت زیر در نظر می‌گیریم:

$$p(c_i | x) = y_i(x) = \frac{\exp(-w_i^T x)}{\sum_{j=1}^K \exp(-w_j^T x)}$$

حال اگر ما تریس در نظر بگیریم که هر سورا آن بردار one-hot انکود شده  $t^{(i)}$  است و در دیتا پوینت را به صورت  $x^{(i)}, t^{(i)}$  در نظر می‌گیریم. تابع loss برابر این مدل به صورت زیر خواهد بود:

$$p(t^{(i)} | x^{(i)}, w) = y_{t^{(i)}}(x^{(i)})$$

$$L(w) = - \sum_{i=1}^N \left( \sum_{k=1}^K T_{ik} \log(y_k(x^{(i)})) \right)$$

ب) ابتدا  $y^{(i)}$  را به عنوان برآورد one-hot برابر  $x^{(i)}$  در نظر می گیریم. حال  $o^{(i)}$  به صورت زیر خواهد بود:

$$o^{(i)} = \begin{bmatrix} v_1^T x^{(i)} \\ v_2^T x^{(i)} \\ \vdots \\ v_{k-1}^T x^{(i)} \\ 0 \end{bmatrix}$$

حال اگر بخواهیم  $x^{(i)}$  را برآورد  $\hat{y}$ ؛ آنگاه مفادیم دات:

$$p(y^{(i)} | x^{(i)}) = \text{log} \frac{e^{o_k^{(i)}}}{\sum_j e^{o_j^{(i)}}} = o_k^{(i)} - \text{log} \left( \sum_j e^{o_j^{(i)}} \right) = (y^{(i)})^T o^{(i)} - \text{log} \left( \sum_j e^{o_j^{(i)}} \right)$$

$$l(w_1, w_2, \dots, w_{k-1}) = \sum_{i=1}^n \left[ (y^{(i)})^T o^{(i)} - \text{log} \left( \sum_j e^{o_j^{(i)}} \right) \right] \quad \text{حال برابر مفادیم دات:}$$

ج) گزاین  $\lambda$  را به یک  $k$  ها به دست می آوریم:

$$\begin{aligned} \frac{\partial L}{\partial w_{nL}} &= \sum_{i=1}^n \left[ y_n^{(i)} x_L^{(i)} - \frac{e^{o_n^{(i)}}}{\sum_j e^{o_j^{(i)}}} x_L^{(i)} \right] \\ &= \sum_{i=1}^n \left[ y_n^{(i)} - p_n^{(i)} \right] x_L^{(i)} \end{aligned}$$

حال  $x^{(i)}$  عضو کلاس  $m$  است:  $p_m^{(i)} = p(y=m | x=x^{(i)})$  برقرار است

$$\frac{\partial L}{\partial y_j} = \sum_{i=1}^n (y_j^{(i)} - p_j^{(i)}) x^{(i)} \Rightarrow \frac{\partial L}{\partial w} = (Y - P)^T X$$

که در آن:  $x = [x^{(1)}, x^{(2)}, \dots, x^{(n)}]^T$

$Y = [y^{(1)}, y^{(2)}, \dots, y^{(n)}]^T$

$P = [p^{(1)}, p^{(2)}, \dots, p^{(n)}]^T$

$w = [w_1, \dots, w_{k-1}]^T$

$$\frac{\partial f}{\partial w} = (Y - P)^T X - \lambda w$$

د) گزاین  $\lambda$  به صورت متغیر می باشد:

سوال 4

الف) در این حالت ماتریس  $X$  به صورت  $X$  خلاصه می شود

$$\begin{aligned} \hat{y} = w^T x &\rightarrow \hat{y} = w_j^T x_j \rightarrow E(w_j) = (y - w_j^T x_j)^2 = \|y - w_j^T x_j\|^2 = (y - w_j^T x_j)^T (y - w_j^T x_j) \\ &= (y^T - x_j^T w_j)(y - w_j^T x_j) = y^T y - 2y^T x_j w_j + \underbrace{w_j^T x_j^T x_j w_j}_{= w_j^T x_j^T x_j w_j} \end{aligned}$$

$$w_j^* \rightarrow \frac{\partial E(w_j)}{\partial (w_j)} = 0 \Rightarrow \frac{\partial E(w_j)}{\partial (w_j)} = -2y^T x_j + 2w_j^T x_j^T x_j = 0$$

$$\Rightarrow y^T x_j = w_j^T x_j^T x_j \Rightarrow w_j = (x_j^T x_j)^{-1} y^T x_j \Rightarrow \boxed{w_j^* = \frac{y^T x_j}{x_j^T x_j}}$$

ب) از آنجایی که  $x_j$  ها متعامد هستند برابر هر  $j \neq i$  داریم:  $x_j \cdot x_i = 0$ . بنابراین  $X^T X$  ماتریس قطری است.

$$X^T X = \text{diag}(x_1^T x_1, x_2^T x_2, \dots, x_n^T x_n) \Rightarrow (X^T X)^{-1} = \text{diag}((x_1^T x_1)^{-1}, \dots, (x_n^T x_n)^{-1})$$

می دانیم در حالت کلی که از همه ویژگی ها استفاده می کنیم به صورت زیر محاسبه می شود:

$$\begin{aligned} y = w^T X &\rightarrow E(w) = \sum_{i=1}^n (y^{(i)} - w^T x_i)^2 = \|y - Xw\|^2 = (y - Xw)^T (y - Xw) \\ &= w^T X^T X w - 2y^T X w + y^T y \end{aligned}$$

$$\frac{\partial E(w)}{\partial (w)} = 0 \rightarrow w X^T X = y^T X \Rightarrow w = (X^T X)^{-1} X^T y$$

بنابراین

$$w = (X^T X)^{-1} X^T y = \text{diag}((x_1^T x_1)^{-1}, \dots, (x_n^T x_n)^{-1}) X^T y$$

$$\Rightarrow w_j = (\text{diag}((x_1^T x_1)^{-1}, \dots, (x_n^T x_n)^{-1}) X^T y)_j = (\text{diag}((x_1^T x_1)^{-1}, \dots, (x_n^T x_n)^{-1}))_j (X^T y)_j$$



$$= (x_j^T x_j)^{-1} (x_j^T y)_j = \frac{(x_j^T y)_j}{x_j^T x_j} = \frac{(x_j^T)_j y}{x_j^T x_j} = \frac{x_j^T y}{x_j^T x_j}$$

که همانطور که در بین ما دیده می شود این که حاصل از  $\min$  کردن در هر انتخاب می شود برابر است و این نتیجه را می توانیم اثبات کرد.

(2) توجه داریم که در نرم ماتریس، در حالت کلی تعاریف از ویژگی ها داده شده است و از آنجا که خواص ماتریس:

$$X = \begin{bmatrix} x_j & \vdots \\ x_j & \vdots \end{bmatrix}_{n \times 2} \rightarrow X^T X = \begin{bmatrix} x_j^T x_j & \text{sum}(x_j) \\ \text{sum}(x_j) & n \end{bmatrix}_{2 \times 2}$$

مستند از  $\text{sum}(x_j)$  مجموع  $n$  عضو ستون  $x_j$  است.

$$\Rightarrow (X^T X)^{-1} = \frac{1}{n \|x_j\|^2 - \text{sum}(x_j)^2} \begin{bmatrix} n & -\text{sum}(x_j) \\ -\text{sum}(x_j) & x_j^T x_j \end{bmatrix}$$

$$X^T y = \left( \begin{bmatrix} x_j & \vdots \\ x_j & \vdots \end{bmatrix} \right)^T y = \begin{bmatrix} x_j^T y \\ \text{sum}(y) \end{bmatrix}$$

از طرف:

$$\Rightarrow \begin{cases} w = [w_j, w_0] \\ w = (X^T X)^{-1} X^T y = \frac{1}{n \|x_j\|^2 - \text{sum}(x_j)^2} \begin{bmatrix} n & -\text{sum}(x_j) \\ -\text{sum}(x_j) & x_j^T x_j \end{bmatrix} \begin{bmatrix} x_j^T y \\ \text{sum}(y) \end{bmatrix} \end{cases}$$

$$\textcircled{1} w_j = \frac{n x_j^T y - \text{sum}(x_j) \text{sum}(y)}{n \|x_j\|^2 - \text{sum}(x_j)^2} = \frac{\frac{x_j^T y}{n} - \frac{\text{sum}(x_j)}{n} \frac{\text{sum}(y)}{n}}{\frac{\|x_j\|^2}{n} - \left( \frac{\text{sum}(x_j)}{n} \right)^2} = \frac{E(x_j y) - E(x_j) E(y)}{E(x_j^2) - E(x_j)^2} = \frac{\text{Cov}(x_j, y)}{\text{Var}(x_j)}$$

$$\begin{aligned} \textcircled{2} w_0 &= \frac{\text{sum}(y) \|x_j\|^2 - \text{sum}(x_j) (x_j^T y)}{n \|x_j\|^2 - \text{sum}(x_j)^2} = \frac{\text{sum}(y) \|x_j\|^2 - \text{sum}(x_j) (x_j^T y)}{n \|x_j\|^2 - \text{sum}(x_j)^2} \\ &= \frac{n^2 E(y) E(x_j^2) - n^2 E(x_j) E(x_j y)}{n^2 \text{Var}(x_j)} = \frac{E(y) E(x_j^2) - E(x_j) E(x_j y)}{\text{Var}(x_j)} \\ &= \frac{E(y) (E(x_j^2) - E(x_j)^2) + E(y) E(x_j)^2 - E(x_j) E(x_j y)}{\text{Var}(x_j)} = E(y) + \frac{E(y) E(x_j)^2 - E(x_j) E(x_j y)}{\text{Var}(x_j)} \\ &= E(y) + E(x_j) \frac{E(y) E(x_j) - E(x_j y)}{\text{Var}(x_j)} \stackrel{\textcircled{1}}{=} E(y) - w_j E(x_j) \end{aligned}$$

$$E(x) = \sum_x x p(x), \quad x \geq 0$$

(أ)

$$E(x) = \sum_{x \geq a} x p(x) \geq \sum_{x \geq a} a p(x) \geq \sum_{x \geq a} a p(x) = a \sum_{x \geq a} p(x) = a P(x \geq a) \Rightarrow E(x) \geq a P(x \geq a)$$

$$\Rightarrow \frac{E(x)}{a} \geq P(x \geq a)$$

(ب)

$$Z = \frac{(x - \mu)^2}{\sigma^2} \xrightarrow{Z \geq \varepsilon^2} \frac{E(Z)}{\varepsilon^2} \geq P(Z \geq \varepsilon^2)$$

$$E(Z) = E((x - \mu)^2) = \text{var}(x) = \sigma^2$$

$$\Rightarrow P(|x - \mu| \geq \varepsilon) = P((x - \mu)^2 \geq \varepsilon^2) = P(Z \geq \varepsilon^2) \leq \frac{E(Z)}{\varepsilon^2} = \frac{\sigma^2}{\varepsilon^2} \Rightarrow P(|x - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{\varepsilon^2}$$

$$P(|x - \bar{x}_4| \geq 0.01 \times \bar{x}_4) \leq \frac{\sigma^2}{n} \frac{1}{(\bar{x}_4 \times 0.01)^2} = \frac{\bar{x}_4 (1 - \bar{x}_4)}{n} \times \frac{1}{(\bar{x}_4 \times 0.01)^2} \leq 0.05 \quad (2)$$

$$\Rightarrow \frac{\frac{3.1415}{4} (1 - \frac{3.1415}{4})}{n} \times \frac{1}{(\frac{3.1415}{4} \times 0.01)^2} \leq 0.05$$

منه نحتاج 6 متغيرين

$$\Rightarrow \frac{2732.33}{n} \leq 0.05 \Rightarrow 54647.9 \leq n \Rightarrow n = 54648$$

(أ)

$$\Sigma = \begin{bmatrix} \sigma_1 & & 0 \\ & \sigma_2 & \\ 0 & & \ddots \\ & & & \sigma_n \end{bmatrix}, \quad \sigma_1 \geq \sigma_2 \geq \sigma_3 \geq \dots \geq \sigma_n > 0$$

$$\Sigma \Sigma^T = I \Rightarrow \Sigma^{-1} = \begin{bmatrix} 1/\sigma_1 & & 0 \\ & 1/\sigma_2 & \\ 0 & & \ddots \\ & & & 1/\sigma_n \end{bmatrix} \quad \sigma_1 \geq \sigma_2 \geq \sigma_3 \geq \dots \geq \sigma_n$$

$$G(A^{-1}) : (A^{-1})^T A^{-1} x = G^2(A^{-1}) x \rightarrow (A^{-1})^T A^{-1} x = U (\Sigma^{-1})^T \underbrace{V^T V}_I \Sigma^{-1} U^T = U (\Sigma^{-1})^2 U^T$$

$$\xrightarrow{\text{نضرب في } U^T} U^T U (\Sigma^{-1})^2 U^T = (\Sigma^{-1})^2 U^T = G^2(A^{-1})$$

$$\rightarrow \underbrace{((\Sigma^{-1})^2 - G^2(A^{-1}) I)}_{=0} U^T x = 0 \rightarrow (\Sigma^{-1})^2 = G^2(A^{-1}) I$$

$$\textcircled{1} \rightarrow \sigma_i^2(A^{-1}) = \frac{1}{\sigma_i^2(A)} \xrightarrow{\sigma_i > 0} \sigma_i(A^{-1}) = \frac{1}{\sigma_i(A)}$$

$$G_{\max}(A) G_{\max}(A^{-1}) = G_{\max}(A) \frac{1}{G_{\min}(A)} = \frac{G_{\max}(A)}{G_{\min}(A)} \xrightarrow{G_{\max}(A) \geq G_{\min}(A)} \geq 1 \rightarrow G_{\max}(A) G_{\max}(A^{-1}) \geq 1$$



$$A = U \Sigma V^T = U \begin{bmatrix} \sigma_1 & & & \\ & \sigma_2 & & \\ & & \ddots & \\ & & & \sigma_r \\ & & & & 0 \end{bmatrix} V^T, \quad r = \text{rank}(A), \quad \sigma_1, \sigma_2, \dots, \sigma_r > 0$$

(ب)

$$\|A\|_2 = \max_{\|v\|_2=1} \|Av\| = \max(\sigma_i) = \sigma_1$$

$$\|A\|_F^2 = \sum_i \sum_j A(i,j)^2 = \text{tr}(A^T A) = \sum_{i=1}^r \lambda_i = \sum_{i=1}^r \sigma_i^2$$

↓  
eigenvalue of  $A^T A$

$$\|A\|_2 = \sigma_1 = \sqrt{\sigma_1^2} \leq \|A\|_F = \sqrt{\sigma_1^2 + \dots + \sigma_r^2} \leq \sqrt{r \sigma_1^2} = \sqrt{r} \cdot \sigma_1 = \sqrt{r} \|A\|_2$$

↓  
 $\sigma_1^2, \sigma_2^2, \dots, \sigma_r^2 > 0$

$$\Rightarrow \|A\|_2 \leq \|A\|_F \leq \sqrt{\text{rank}(A)} \|A\|_2$$

$$\tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}} = \frac{e^a(1 - e^{-2a})}{e^a(1 + e^{-2a})} = \frac{1 - e^{-2a}}{1 + e^{-2a}} = \frac{2}{1 + e^{-2a}} - 1 = 2\sigma(2a) - 1 \quad (*)$$

sigmoid function  $\sigma$

$$a_j = \frac{(x - \mu_j)}{s}$$

جواب السؤال

$$y(x, w) = w_0 + \sum_{j=1}^n [w_j \sigma(2a_j)]$$

$$= w_0 + \sum_{j=1}^n \left[ \frac{w_j}{2} (2\sigma(2a_j) - 1 + 1) \right] = w_0 + \sum_{j=1}^n \frac{w_j}{2} (2\sigma(2a_j) - 1) + \sum_{j=1}^n \frac{w_j}{2}$$

$$= w_0 + \sum_{j=1}^n \frac{w_j}{2} \tanh(a_j) + \sum_{j=1}^n \frac{w_j}{2}$$

$$\begin{aligned} w_0 &\triangleq w_0 + \sum_{j=1}^n \frac{w_j}{2} \\ w_j &\triangleq \frac{w_j}{2} \end{aligned}$$

$$= \left( w_0 + \sum_{j=1}^n \frac{w_j}{2} \right) + \sum_{j=1}^n \frac{w_j}{2} \tanh(a_j) = w_0 + \sum_{j=1}^n w_j \tanh\left(\frac{x - \mu_j}{s}\right) \quad \square$$