Programming Assignment Unit 3

September 19, 2025

1 Multiple Linear Regression Exercises

1.1 a) Calculate the parameter estimates (0, 1, 2 and 2), in addition find the usual 95% confidence intervals for 0, 1, and 2.

```
[]: # Load the data
     D <- data.frame(</pre>
       x1=c(0.58, 0.86, 0.29, 0.20, 0.56, 0.28, 0.08, 0.41, 0.22, 0.35,
            0.59, 0.22, 0.26, 0.12, 0.65, 0.70, 0.30, 0.70, 0.39, 0.72,
            0.45, 0.81, 0.04, 0.20, 0.95),
       x2=c(0.71, 0.13, 0.79, 0.20, 0.56, 0.92, 0.01, 0.60, 0.70, 0.73,
            0.13, 0.96, 0.27, 0.21, 0.88, 0.30, 0.15, 0.09, 0.17, 0.25,
            0.30, 0.32, 0.82, 0.98, 0.00),
       y=c(1.45, 1.93, 0.81, 0.61, 1.55, 0.95, 0.45, 1.14, 0.74, 0.98,
            1.41, 0.81, 0.89, 0.68, 1.39, 1.53, 0.91, 1.49, 1.38, 1.73,
            1.11, 1.68, 0.66, 0.69, 1.98)
     )
     fit_full \leftarrow lm(y \sim x1 + x2, data = D)
     summary(fit full)
                            # parameter estimates & tests
     confint(fit_full, 0.95) # 95% CIs
     summary(fit full)$sigma^2
                                  # 2 estimate
    Call:
    lm(formula = y \sim x1 + x2, data = D)
    Residuals:
                    1Q
                         Median
                                      3Q
                                               Max
    -0.15493 -0.07801 -0.02004 0.04999 0.30112
    Coefficients:
                 Estimate Std. Error t value Pr(>|t|)
                                       6.571 1.31e-06 ***
    (Intercept) 0.433547
                            0.065983
                            0.095245 17.355 2.53e-14 ***
                 1.652993
    x2
                 0.003945
                            0.074854
                                       0.053
                                                 0.958
```

Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1127 on 22 degrees of freedom

Multiple R-squared: 0.9399, Adjusted R-squared: 0.9344

F-statistic: 172 on 2 and 22 DF, p-value: 3.699e-14

A matrix: 0×2 of type dbl 2.5 % 97.5 %

0.0127052276665322

1.2 b) Reduce the model if appropriate (=0.05)

Call:

 $lm(formula = y \sim x1, data = D)$

Residuals:

Min 1Q Median 3Q Max -0.15633 -0.07633 -0.02145 0.05157 0.29994

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.43609 0.04399 9.913 9.02e-10 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1102 on 23 degrees of freedom

Multiple R-squared: 0.9399, Adjusted R-squared: 0.9373

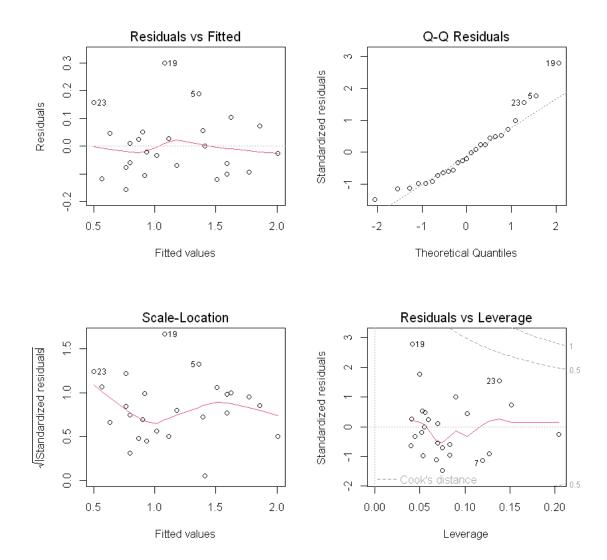
F-statistic: 359.6 on 1 and 23 DF, p-value: 1.538e-15

1.3 c) Residual analysis

```
[]: par(mfrow = c(2, 2))
plot(fit)  # diagnostic plots
shapiro.test(residuals(fit))  # normality check
```

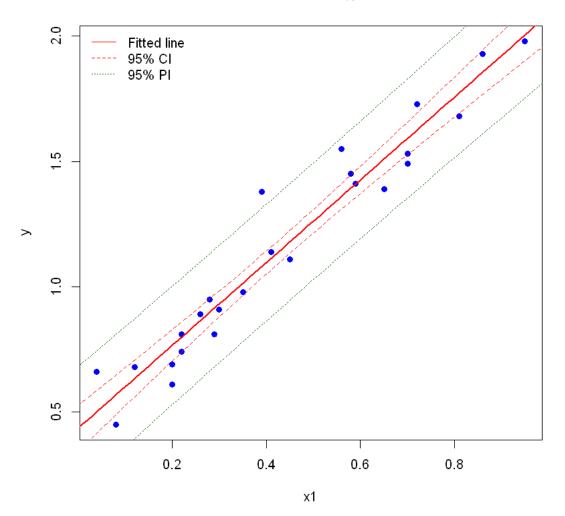
Shapiro-Wilk normality test

data: residuals(fit)
W = 0.93532, p-value = 0.1154



1.4 d) Fitted line with 95% confidence & prediction intervals

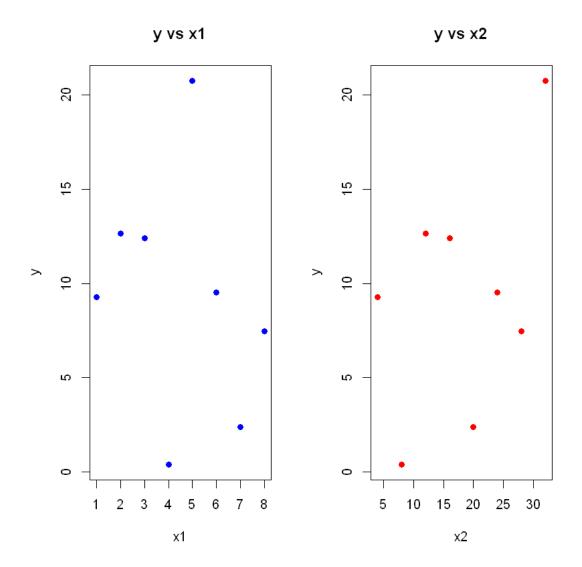
Fitted line with 95% CI & PI



2 MLR simulation exercise

```
[]: ## Load data
D <- data.frame(
    y = c(9.29,12.67,12.42,0.38,20.77,9.52,2.38,7.46),
    x1 = c(1,2,3,4,5,6,7,8),
    x2 = c(4,12,16,8,32,24,20,28)
)</pre>
```

2.1 a) Plots of y vs x1 and y vs x2



The plot of y vs x1 shows no obvious upward or downward trend—the points are scattered. The plot of y vs x2 is also diffuse. Neither predictor shows a clear linear relationship with y.

2.2 b) Fit the two simple linear models

```
[]: fit_x1 <- lm(y ~ x1, data = D)
fit_x2 <- lm(y ~ x2, data = D)

# Summaries with parameter estimates and p-values
summary(fit_x1)
summary(fit_x2)</pre>
```

```
# 95% confidence intervals for coefficients
confint(fit_x1, level = 0.95)
confint(fit_x2, level = 0.95)
```

Call:

lm(formula = y ~ x1, data = D)

Residuals:

Min 1Q Median 3Q Max -9.2942 -3.0504 0.6933 1.8381 11.7217

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 12.1775 5.1984 2.343 0.0576 .
x1 -0.6258 1.0294 -0.608 0.5655

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6.672 on 6 degrees of freedom

Multiple R-squared: 0.05802, Adjusted R-squared: -0.09897

F-statistic: 0.3696 on 1 and 6 DF, p-value: 0.5655

Call:

lm(formula = y ~ x2, data = D)

Residuals:

Min 1Q Median 3Q Max -7.554 -5.104 1.036 4.212 7.397

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 4.2039 4.8069 0.875 0.415
x2 0.2865 0.2380 1.204 0.274

Residual standard error: 6.169 on 6 degrees of freedom

Multiple R-squared: 0.1946, Adjusted R-squared: 0.06035

F-statistic: 1.45 on 1 and 6 DF, p-value: 0.2739

A matrix: 2×2 of type dbl (Intercept) -0.5426374 24.897637 -3.1447959 1.893129A matrix: 2×2 of type dbl (Intercept) -7.5580921 15.9659492 -2.2957889 0.8688246

Linear model with x1

Model: = 0 + 11 +

term	estimate	95% CI	p-value
Intercept x1 slope	12.18 -0.63	-0.54 , 24.90 -3.14 , 1.89	$0.058 \\ 0.566$

- The slope is small and the confidence interval includes 0.
- 2 = 0.06: x1 explains only about 6 % of the variation in y.

Linear model with x2

Model: = 0 + 12 +

term	estimate	95% CI	p-value
Intercept	4.20	-7.56, 15.97	$0.415 \\ 0.274$
x2 slope	0.29	-0.30, 0.87	

- The slope is not significant; its confidence interval also spans 0.
- 2 = 0.19: x2 explains about 19 % of the variation, still weak.

Interpretation

- Neither x1 nor x2 provides a statistically significant linear relationship with y at the 5 % level.
- Both models have wide confidence intervals and low R^2 values, so the data do not support using x1 or x2 alone to predict y.
- The visual scatterplots agree: there is no clear linear trend.

We can conclude that, with this small dataset, neither x1 nor x2 is an effective predictor of y in a simple linear model.