

## Week 8

# Large Deviations for Random Walks

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### Nash Inequality (continued)

**Theorem 8.0.1.** *Let  $(\Gamma, \mu)$  be a weighted graph, and let  $\alpha \geq 1$ . TFAE.*

(a) (Nash Inequality)  $(\Gamma, \mu)$  satisfies  $(N_\alpha)$

(b) (On Diagonal Bounds) There exists  $C_H > 0$  such that for every  $x \in V$  and  $n \geq 0$

$$p_n(x, x) \leq \frac{C_H}{(n \vee 1)^{\alpha/2}}$$

(c) (Off Diagonal Bounds) There exists  $C'_H > 0$  such that for every  $x, y \in V$  and  $n \geq 0$

$$p_n(x, y) \leq \frac{C'_H}{(n \vee 1)^{\alpha/2}}$$

*Proof.* We provide only a sketch of the proof. From Worksheet 2, (a)  $\implies$  (b) holds, and (c)  $\implies$  (b) is trivial. First, we show (b)  $\implies$  (c). So assume (b). Let  $m \geq 0$ . If  $n$  is even, with  $n = 2m$ , then for any  $x, y \in V$ , we have

$$p_{2m}(x, x) \leq \frac{C_H}{(2m \vee 1)^{\alpha/2}} \quad \text{and} \quad p_{2m}(y, y) \leq \frac{C_H}{(2m \vee 1)^{\alpha/2}}$$

As an exercise, show that  $p_{2m}(x, y) \leq \sqrt{p_{2m}(x, x)p_{2m}(y, y)}$ , and using this, we get (b)  $\implies$  (c) with  $C'_H = C_H$ . If  $n = 2m + 1$ , then, since  $p_{2m+1}(x, y) \leq \sqrt{p_{2m}(x, x)p_{2m+2}(y, y)}$  (by a similar exercise), we get

$$p_{2m+1}(x, y) \leq \sqrt{\frac{C_H^2}{(2m \vee 1)^{\alpha/2}(2m+2 \vee 1)^{\alpha/2}}} \leq \frac{C'}{(2m+1 \vee 1)^{\alpha/2}}$$

for some  $C' > 0$ . To show the last inequality above, use the fact that there exists  $C_\alpha > 0$  such that  $(2m)^{\alpha/2}(2m+2)^{\alpha/2} \leq C_\alpha(2m+1)^{\alpha/2}$  (details left as exercises). Thus (b)  $\implies$  (c).

Now, we show (c)  $\implies$  (a). Assuming (c), observe that (by taking supremum over  $x \in V$ )

$$|P_n f(x)| \leq \sum_{y \in V} p_n(x, y) |f(y)| \mu_y \implies \|P_n f\|_\infty \leq \frac{C_H}{(n \vee 1)^{\alpha/2}} \|f\|_1$$

$$\text{and } \|P_n f\|_2^2 = \langle P_n f, P_n f \rangle = \langle P_{2n} f, f \rangle \leq \|P_{2n} f\|_\infty \|f\|_1 \leq \frac{C_H}{(2n \vee 1)^{\alpha/2}} \|f\|_1^2 \quad (8.1)$$

Now, we make use of the following inequality - (verify!)

$$\mathcal{E}(f, f) \geq \frac{1}{2n} [\|f\|_2^2 - \|P_n f\|_2^2]$$

Using this, and (8.1), we get

$$\mathcal{E}(f, f) \geq \frac{1}{2n} \left[ \|f\|_2^2 - \frac{C_H}{(2n \vee 1)^{\alpha/2}} \|f\|_1^2 \right]$$

WLOG, assume  $\|f\|_1 = 1$ , and choose smallest possible  $k$  such that

$$\frac{C_H}{(2n \vee 1)^{\alpha/2}} \leq \frac{\|f\|_2^2}{2} \quad \text{so that} \quad \mathcal{E}(f, f) \geq \frac{1}{4k} \|f\|_2^2$$

Since  $k \geq 1$ , we have  $k^{-\alpha/2} \leq C^2 \|f\|_2^2$  for some  $C > 0$ , and hence  $k^{-\alpha/2} \leq C \|f\|_2$ . Therefore,

$$\mathcal{E}(f, f) \geq \frac{C_2 \|f\|_2^2}{\|f\|_2^{\frac{4}{\alpha}}} = C_2 \|f\|_2^{2-4/\alpha} \implies (N_\alpha)$$

□

## Carne-Varopoulos Bound

We begin with a few lemmas and some results involving Chebyshev polynomials.

**Lemma 8.0.1.** *Let  $\{S_n\}_{n \geq 0}$  denote the simple symmetric random walk on  $\mathbb{Z}$  with  $S_0 = 0$ . Then*

(a)

$$\mathbf{P}(S_n \geq D) \leq \exp\left(-\frac{D^2}{2n}\right)$$

(b)

$$\mathbf{E}[\lambda^{S_n}] = \sum_{r \in \mathbb{Z}} \lambda^r \mathbf{P}(S_n = r) = 2^{-n} \sum_{r=0}^n \binom{n}{r} \left(\frac{1}{\lambda}\right)^{2n-r}$$

*Proof.* (a) was given in Worksheet 2, and (b) is trivial using results from Week 1. □

**Definition 8.0.1.** *(Chebyshev Polynomials) For  $-1 \leq t \leq 1$ , define*

$$H_k(t) := \frac{1}{2}(t + i\sqrt{1-t^2})^k + \frac{1}{2}(t - i\sqrt{1-t^2})^k$$

**Lemma 8.0.2.** *For each  $k \geq 0$ , we have*

(a)  $H_k$  is a real polynomial of degree  $k$ .

(b)  $t^n = \sum_{k \in \mathbb{Z}} \mathbf{P}(S_n = k) H_{|k|}(t)$

*Proof.* To show (a), fix  $t \in [-1, 1]$  and set  $s = \sqrt{1 - t^2}$ . Observe that

$$H_k(t) = \frac{1}{2} \sum_{r=0}^k \binom{k}{r} t^{k-r} [(is)^r + (-is)^r] = \frac{1}{2} \sum_{r=0}^{k/2} \binom{k}{2r} t^{k-2r} \psi(s)$$

where  $\psi$  is some real function of  $s$ .

To show (b) set  $z_1 = t + is$  and  $z_2 = t - is$  so that  $|z_1| = |z_2| = 1$  and  $z_1 z_2 = 1$ . Then,

$$H_k(t) = \frac{1}{2} (z_1^k + z_2^k) = H_{-k}(t) \implies |H_k(t)| \leq 1$$

Now, observe that  $t = (z_1 + z_2)/2$ , so that

$$t^n = \sum_{k=0}^n \frac{1}{2^n} \binom{n}{k} z_1^k z_2^{n-k} = \sum_{k=0}^n \frac{1}{2^n} \binom{n}{k} z_1^{2k-n} = \frac{1}{2^n} \sum_{r \in \mathbb{Z}} \mathbf{P}(S_n = r) z_1^r$$

Repeating the same arguments above, we get

$$\begin{aligned} t^n &= \frac{1}{2^n} \sum_{r \in \mathbb{Z}} \mathbf{P}(S_n = r) z_1^r = \frac{1}{2^n} \sum_{r \in \mathbb{Z}} \mathbf{P}(S_n = r) z_2^r \\ \implies t^n &= \frac{1}{2^n} \sum_{r \in \mathbb{Z}} \mathbf{P}(S_n = r) \left( \frac{z_1^r + z_2^r}{2} \right) = \sum_{r \in \mathbb{Z}} \mathbf{P}(S_n = r) H_{|r|}(t) \end{aligned}$$

□

**Theorem 8.0.2.** (*Carne-Varopoulos bound*) Let  $(\Gamma, \mu)$  be a weighted graph. Then, for every  $x, y \in V$  and  $n \geq 1$

$$p_n(x, y) \leq \frac{2}{\sqrt{\mu_x \mu_y}} \exp \left( - \frac{d(x, y)^2}{2n} \right)$$

*Proof.* Proved in Worksheet 2. □

## Large Deviations for Random Walks

Let  $\{\xi_i\}_{i \geq 1}$  be IID  $\mathbb{Z}$  valued random variables such that  $\mathbf{E}[\xi_1] = \mu$  and  $\text{Var}[\xi_1] < \infty$ . Define  $S_0 = 0$  and  $S_n = \sum_{i=1}^n \xi_i$ . Then, the strong law of large numbers (SLLN) and the central limit theorem (CLT) respectively state that

$$\mathbf{P} \left( \lim_{n \rightarrow \infty} \frac{S_n}{n} = \mu \right) = 1 \quad \text{and} \quad \frac{S_n - n\mu}{\sqrt{n}} \xrightarrow{d} \mathcal{N}(0, 1)$$

Thus, the CLT loosely states that  $S_n \approx n\mu + \sqrt{n}Z$ , where  $Z \sim \mathcal{N}(0, 1)$ .

As an exercise, show that for every  $\epsilon > 0$ ,  $\mathbf{P}(A_n^\epsilon) \rightarrow 0$  as  $n \rightarrow \infty$ , where  $A_n^\epsilon = \{S_n \geq n(\mu + \epsilon)\}$ . What is the rate of decay of  $\mathbf{P}(A_n^\epsilon)$  (as  $n \rightarrow \infty$ )?

(Hint:  $\mathbf{P}(S_n \geq n(\mu + \epsilon)) \approx \mathbf{P}(\xi_i > \mu + \epsilon \ \forall \ 1 \leq i \leq n) = [\mathbf{P}(\xi_1 > \mu + \epsilon)]^n \approx e^{-Cn}$  for some  $C > 0$ )

**Theorem 8.0.3.** Let  $\{\xi_i\}_{i \geq 0}$  be IID random variables with  $\mathbf{P}(\xi_1 = 0) = \mathbf{P}(\xi_1 = 1) = 1/2$ . Then, for every  $a > 1/2$ ,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log[\mathbf{P}(S_n \geq an)] = -I(a)$$

where

$$I(z) = \begin{cases} \log 2 + a \log a + (1-a) \log a & \text{if } 0 \leq z \leq 1 \\ \infty & \text{otherwise} \end{cases}$$

**Observations:**

- (1) Minima of  $I(z)$  is achieved at  $z = 1/2$ , and the graph increases from  $[1/2, 1]$ . This implies rate of exponential decay increases as  $1/2 \rightarrow a \rightarrow 1$ .
- (2) Symmetry of the function  $I(\cdot)$  around  $1/2$  suggests that for  $a < 1/2$ , (Requires a proof)

$$\frac{1}{n} \log[\mathbf{P}(S_n \geq an)] \rightarrow -I(a)$$

- (3) The theorem implies SLLN. The idea of the proof makes use of the following inequality

$$\mathbf{P}(S_n > (1/2 + \delta)n) \leq \exp\{-I(n(1/2 + \delta))\}$$

*Proof.*

If,  $a > 1$  then, since  $S_n$  can be at most  $n$ ,  $\mathbf{P}(S_n > an) = 0$  so the result follows. Now, consider  $\frac{1}{2} < a \leq 1$ , then

$$\mathbf{P}(S_n > an) = \sum_{an < k \leq n} \mathbf{P}(S_n = k) = \sum_{an < k \leq n} \binom{n}{k} \frac{1}{2^n} = \frac{1}{2^n} \sum_{an < k \leq n} \binom{n}{k}$$

Let,  $Q_n(a) = \max_{an < k \leq n} \binom{n}{k}$ . So, we have,

$$2^{-n} Q_n(a) \leq \mathbf{P}(S_n > an) \leq 2^{-n} Q_n(a) (n+1) \quad (8.2)$$

First equality follows from the fact that one summand in the  $\sum_{an < k \leq n} \binom{n}{k}$  attains maximum and the second equality follows since, each summand of  $\sum_{0 \leq k \leq n} \binom{n}{k}$  is  $\leq Q_n(a)$ .

**Claim:**

For,  $\frac{1}{2} < a < 1$ ,

$$\frac{1}{n} \log Q_n(a) \xrightarrow{n \rightarrow \infty} -a \log a - (1-a) \log(1-a)$$

Now, from (8.2),

$$-\log 2 + \frac{1}{n} \log Q_n(a) \leq \frac{1}{n} \log \mathbf{P}(S_n > an) \leq -\log 2 + \frac{1}{n} \log Q_n(a) + \frac{1}{n} \log(n+1) \quad (8.3)$$

assuming the claim as LHS and RHS of (8.3) goes to  $-I(a)$ , the result follows. We now prove the claim.

**Proof of claim:**

Since,  $a > \frac{1}{2}$ ,  $\max_{an < k \leq n} \binom{n}{k} = \binom{n}{\lceil an \rceil}$ . Now, from stirling's approximation

$$\binom{n}{\lceil an \rceil} = \frac{n!}{\lceil an \rceil! (n - \lceil an \rceil)!} \sim \frac{n^n e^{-n} \sqrt{2\pi n}}{\lceil an \rceil^{\lceil an \rceil} e^{-\lceil an \rceil} \sqrt{2\pi \lceil an \rceil}} \cdot \frac{1}{(n - \lceil an \rceil)^{n - \lceil an \rceil} e^{n - \lceil an \rceil} \sqrt{2\pi (n - \lceil an \rceil)}}$$

For,  $a > \frac{1}{2}, a < 1; \lceil an \rceil \rightarrow \infty$  and  $n - \lceil an \rceil \rightarrow \infty$  as  $n \rightarrow \infty$  (Check!) and

$$\begin{aligned} \frac{1}{n} \log Q_n(a) &\sim \frac{1}{n} \left[ \left(n + \frac{1}{2}\right) \log n - \left(\lceil an \rceil + \frac{1}{2}\right) \log \lceil an \rceil - \left(n - \lceil an \rceil + \frac{1}{2}\right) \log (n - \lceil an \rceil) - \log(\sqrt{2\pi}) \right] \\ &= \log n + \frac{1}{2n} \log n - \frac{\lceil an \rceil}{n} \log \lceil an \rceil - \frac{1}{2n} \log \lceil an \rceil - \frac{1}{n} \log \sqrt{2\pi} - \frac{n - \lceil an \rceil}{n} \log (n - \lceil an \rceil) - \frac{1}{2} \log (n - \lceil an \rceil) \end{aligned}$$

the second, fourth, fifth and seventh summand of the above equation tends to 0 as  $n$  tends to  $\infty$  and from the exercise (?) we have that

$$\frac{\lceil an \rceil}{n} \log \frac{\lceil an \rceil}{n} \xrightarrow{n \rightarrow \infty} a \log a \quad \text{and} \quad \frac{n - \lceil an \rceil}{n} \log \frac{n - \lceil an \rceil}{n} \xrightarrow{n \rightarrow \infty} (1 - a) \log(1 - a)$$

which proves the claim.  $\square$

### Cramer, 1930's

$\{\xi_i\}_{i \geq 1}$  i.i.d random variables with  $\mathbf{E}[\xi_i] = \mu < \infty$ ,  $\mathbf{E}[e^{r\xi_i}] < \infty$ ,  $\forall r \in \mathbb{R}$ . For any  $a > \mu$ ,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbf{P}(S_n > an) = -I(a)$$

where,  $I(a) = \sup_{z \in \mathbb{R}} [za - \mathbf{E}[e^{z\xi}]]$

### Sanov, 1961 (Level 2 of LDP)

$\mathbf{P}(S_n > an) = \mathbf{P} \circ S_n^{-1}((an, \infty)) := \mu_n((an, \infty))$

$$-\frac{1}{n} \log \mu_n((an, \infty)) \xrightarrow{n \rightarrow \infty} \infty$$

### Varadhan's LDP setup, 1960's

Let,  $X_n : \Omega \rightarrow \mathbb{R}$  be a random variable of  $(\Omega, \mathcal{F}, \mathbf{P})$ .  $A$  be an event,  $\mathbf{P}_n(A) := \mathbf{P}(S_n \in A)$ , then  $\mathbf{P}(\cdot)$  is a probability on  $\mathbb{R}$ .

A sequence  $\{\mathcal{P}_n\}_{n \geq 1}$  of probability measures on  $\mathbb{R}$  (can be any metric space  $(X, d)$ ) is said to satisfy large deviation principle with rate  $n$  and rate function  $I : \mathbb{R} \rightarrow [0, \infty) \cup \{\infty\}$ , if

1.  $I \not\equiv \infty$ ,  $I$  is lower-semi continuous and has compact level sets.
2.  $\overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \log \mathbf{P}_n(\mathcal{C}) \leq -I(\mathcal{C}) \forall$  closed sets  $\mathcal{C}$
3.  $\underline{\lim}_{n \rightarrow \infty} \frac{1}{n} \log \mathbf{P}_n(\mathcal{O}) \geq -I(\mathcal{O}) \forall$  open sets  $\mathcal{O}$

where,  $A \subseteq \mathbb{R}$ ,  $I(A) = \inf_{y \in A} I(y)$ .

**Theorem 8.0.4.**  $\{\mathcal{P}_n\}_{n \geq 1}$  satisfied LDP with rate  $n$  then ,  $I(\cdot)$  is unique.

**Theorem 8.0.5** (Varadhan's lemma). *If,  $\{\mathcal{P}_n\}_{n \geq 1}$  satisfies LDP with rate  $n$  and rate function  $I(\cdot)$ , let  $F_n(x) = \mathbf{P}_n((-\infty, x])$  for some continuous and bounded above function  $F : \mathbb{R} \rightarrow \mathbb{R}$ , we have*

$$\int e^{nF(x)} dF_n(x) \xrightarrow[n \rightarrow \infty]{} \sup_{x \in \mathbb{R}} [F(x) - I(x)]$$

## Applications

For,  $\theta \in S^1$ ,  $t \in \mathbb{R}$ ,  $u : S^1 \times \mathbb{R}_+ \rightarrow \mathbb{R}$ ,

$$\begin{aligned} \frac{\partial u}{\partial t} &= \frac{1}{2} \frac{\partial^2 u}{\partial \theta^2} + V(\theta)u \\ u(0, \theta) &= 1 \end{aligned}$$

then,

$$\frac{1}{t} \log u(t, \theta) \xrightarrow[t \rightarrow \infty]{} \lambda_1 = \sup_{f \in \dots} \left\{ \int V(\theta) f(\theta) d\theta - \frac{1}{8} \int \frac{(f'(\theta))^2}{f(\theta)} d\theta \right\}$$

we can represent this as follows,

$$u(t, \theta) = \mathbf{E} e^{\int_0^t V(\theta_s) ds}, \quad \{\theta_s\} - \text{brownian motion on } S^1$$

## Exercises

1. For any  $a \in \mathbb{R}$ , show that,

$$\frac{[an]}{n} \xrightarrow[n \rightarrow \infty]{} a \quad \text{and} \quad \frac{n - [an]}{n} \xrightarrow[n \rightarrow \infty]{} 1 - a$$