

## Week 9

# Random walk in trap environment

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## 9.1 Continuous time random walk

In this model the random walker waits an exponential amount of time to perform a jump like a discrete time random walk. Consider  $\{V_i : i \geq 1\}$  to be a collection of independent  $\text{Exponential}(\lambda)$  random variables. Let  $\lambda = 1$ . Define  $T_k$  to be the sum of the first  $k$   $V_i$ 's. Also, define  $N_t$  to be the number of  $T_k$  less than  $t$ . Hence,

- $\mathbf{P}(V_i \leq t) = 1 - e^{-t}$
- $T_k = \sum_{i=1}^k V_i$
- $N_t = \sum_{k=1}^{\infty} \mathbb{1}(T_k \leq t)$
- $\{N_t = k\} = \{T_k \leq T_{k+1}\}$

**Theorem 9.1.1.** 1.  $N_t \sim \text{Poisson}(t)$

2.  $N_t - N_s$  is independent of  $N_r$  where  $r \leq s \leq t$ .

3. For  $0 \leq t_0 \leq t_1 \leq \dots \leq t_n$

$\{N_{t_{i+1}} - N_{t_i} : i = 0, \dots, n-1\}$  are independent

So,  $N_{t_{i+1}} - N_{t_i} \sim \text{Poisson}(t_{i+1} - t_i)$

**Definition 9.1.1.** Let  $U_n : n \geq 0$  be a random walk on  $(\Gamma, \mu)$ . Define, a continuous time random walk on  $(\Gamma, \mu)$  with rate 1 to be:

$$Y_t = U_{N_t} \forall t \geq 0$$

.

*Remark.* The random variable  $Y_t$  is a random step function which is right continuous with left limits.

## 9.2 Random walk in trap environment

### Continuous Time set-up

Consider the graph  $\mathbb{Z}^d$  with natural weights. Let  $\{X_t\}_{t \geq 0}$  be a continuous time random walk on  $\mathbb{Z}^d$  starting at 0, with rate  $\kappa$ . Now, let us set up the traps, i.e., for each  $y \in \mathbb{Z}^d$  let  $N_y \sim \text{Poisson}(\rho)$ . This  $N_y$  denote the number of *traps* at  $y$ . Each trap  $(Y^{j,y})$  perform a continuous time random walk  $(\{Y_t^{j,y}\}_{t \geq 0})$  with rate  $\nu$ ; where  $1 \leq j \leq N_y$ . The random walk gets killed if it meets a trap. There are two ways of killing viz,

Hard The walk gets killed upon intersection with any  $Y^{j,y}$ .

Soft At each site  $x$  at time  $t \geq 0$ , define

$$\xi(t, x) := \sum_{y \in \mathbb{Z}^d, 1 \leq j \leq N_y} \#\{Y^{j,y} \text{ at } x\}.$$

Now  $X_t$  gets killed at rate  $\gamma \xi(t, x)$  where  $\gamma \in \mathbb{R}$ .

*Remark.* Hard killing infact corresponds to  $\gamma = \infty$  case of soft killing.

The probability of survival is given by

$$Z_{\gamma, t} = \mathbf{E}^X[\exp(-\gamma \int_0^t \xi(s, X(s)) ds)]$$

### Discrete Time set-up

Let  $\{X_t\}_{t \geq 0}$  be a random walk on  $\mathbb{Z}^d$  with natural weights starting at 0. For each  $y \in \mathbb{Z}^d$  let  $N_y \sim \text{Poisson}(\rho)$  denotes the number of traps at  $y$ . Each trap  $(Y^{j,y})$  perform a lazy random walk  $(\{Y_t^{j,y}\}_{t \geq 0})$  on  $\mathbb{Z}^d$ ; where  $1 \leq j \leq N_y$ . The trap kills the random walk with probability  $q$  if it meets the random walk;  $q \in (0, 1)$ . Let  $\xi(n, x)$  denote the number of traps at location  $x$ , i.e.

$$\xi(n, x) = \sum_{y \in \mathbb{Z}^d, 1 \leq j \leq N_y} \delta_x(Y_n^{j,y}).$$

Assume  $X_k$  has survived till  $k \leq n$ . Given  $X_n$  the probability that  $X_n$  will survive at time  $n$  is  $(1 - q)^{\xi(n, X_n)}$ . Hence,

$$\begin{aligned} \sigma^X(n, \xi) &= \mathbf{P}(X \text{ has survived till time } n \text{ given } \{Y_m^{j,y}\}_{1 \leq j \leq m, y \in \mathbb{Z}^d} \text{ where } m \leq n) \\ &= (1 - q)^{\sum_{i=1}^n \xi(i, X_i)}. \end{aligned} \tag{9.1}$$

## 9.3 Pascal's Theorem

The average survival probability of a given trajectory  $X$  is given by  $\sigma^X(n) = \mathbf{E}^\xi[(1 - q)^{\sum_{i=1}^n \xi(i, X_i)}]$ .

**Theorem 9.3.1** (Pascal). *The survival probability is maximized by the trajectory  $\underline{0}$  where  $\underline{0}_k = 0$  for every  $k \in \mathbb{N} \cup 0$ , i.e.,*

$$\sigma^X(n) \leq \sigma^{\underline{0}}(n).$$

**Lemma 9.3.1.**  $\sigma^X(n) = \exp(-\lambda \sum_{y \in \mathbb{Z}^d} W_X(n, y))$  where  $W_X(n, y) = 1 - \mathbf{E}^y[1 - (1 - q)^{\sum_{i=1}^n \delta(Y_i^y)}]$ . The  $Y_i^y$  is a random variable with ditribution same as i.i.d.  $Y_i^{j,y}$ .

*Proof.* Let  $X : \mathbb{N} \cup 0 \rightarrow \mathbb{Z}^d$  with  $X_0 = 0$  be the trajectory. Now,

$$\begin{aligned}
\sigma^X(n) &= \mathbf{E}^\xi[(1 - q)^{\sum_{i=1}^n \xi(i, X_i)}] \\
&= \mathbf{E}^\xi[(1 - q)^{\sum_{i=1}^n \sum_{y \in \mathbb{Z}^d} \sum_{1 \leq j \leq N_y} \delta_{X_i}(Y_n^{j,y})}] \\
&= \prod_{y \in \mathbb{Z}^d} \mathbf{E}^\xi[\prod_{1 \leq j \leq N_y} (1 - q)^{\sum_{i=1}^n \delta_{X_i}(Y_n^{j,y})}] \\
&= \prod_{y \in \mathbb{Z}^d} \mathbf{E}^y \mathbf{E}^{N_y}[\prod_{1 \leq j \leq N_y} (1 - q)^{\sum_{i=1}^n \delta_{X_i}(Y_n^{j,y})}] \\
&= \prod_{y \in \mathbb{Z}^d} \sum_{k=0}^{\infty} \frac{e^{-\lambda} \lambda^k}{k!} \mathbf{E}^y[\prod_{1 \leq j \leq k} (1 - q)^{\sum_{i=1}^n \delta_{X_i}(Y_n^{j,y})}] \\
&= \prod_{y \in \mathbb{Z}^d} \sum_{k=0}^{\infty} \frac{e^{-\lambda} \lambda^k}{k!} (\prod_{1 \leq j \leq k} (1 - q)^{\sum_{i=1}^n \delta_{X_i}(Y_n^{j,y})})^k \\
&= \prod_{y \in \mathbb{Z}^d} e^{-\lambda(1 - \mathbf{E}^y((1 - q)^{\sum_{i=1}^n \delta_{X_i}(Y_n^{j,y})}))} \\
&= e^{-\lambda \sum_{y \in \mathbb{Z}^d} W_X(n, y)}.
\end{aligned}$$

□

**Lemma 9.3.2.**  $W_X(n, y) = 1 - \mathbf{E}^y[1 - (1 - q)^{\sum_{i=1}^n \delta(Y_i^y)}] = \mathbf{P}_y^X(\tau \leq n)$ , where  $\tau = \min\{i \geq 0 | X_i = Y_i, Z_i = 1\}$ .

**Lemma 9.3.3.**  $\sum_{y \in \mathbb{Z}^d} \mathbf{P}_y^X(\tau \leq n) \geq \sum_{y \in \mathbb{Z}^d} \mathbf{P}_y^0(\tau \leq n)$ .

*Proof.* Let

$$\begin{aligned}
q &= \mathbf{P}(Z_n = 1) \\
&= \mathbf{P}_{X_n}(\bigcup_{y \in \mathbb{Z}^d} \{Z_n = 1, Y_n = y\}) \\
&= \sum_{y \in \mathbb{Z}^d} \mathbf{P}_{X_n}(Z_n = 1, Y_n = y) \\
&= \sum_{y \in \mathbb{Z}^d} \mathbf{P}_{X_n}(Z_n = 1, Y_n = X_n) \\
&= \sum_{y \in \mathbb{Z}^d} [\mathbf{P}^X(\tau = n) + \sum_{k=0}^{n-1} \mathbf{P}_y^X(\tau = k) p_{n-k}^y(X_n - X_k) q]
\end{aligned}$$

**Lemma 9.3.4.** For a lazy symmetric random walk on  $\mathbb{Z}^d$ .

$$\begin{aligned}
p_n^Y(0) &\geq p_n^Y(y), \forall y \in \mathbb{Z}^d \\
p_n^Y(0) &\geq p_{n+1}^Y(0).
\end{aligned}$$

Therefore using the above lemma, we get:

$$q \leq \sum_{y \in \mathbb{Z}^d} [\mathbf{P}^X(\tau = n) + \sum_{k=0}^{n-1} \mathbf{P}_y^X(\tau = k) p_{n-k}^y(0) q].$$

Also, replacing  $X = \underline{0}$  in  $\sum_{y \in \mathbb{Z}^d} [\mathbf{P}^X(\tau = n) + \sum_{k=0}^{n-1} \mathbf{P}_y^X(\tau = k) p_{n-k}^y(X_n - X_k) q]$ , we get:

$$q = \sum_{y \in \mathbb{Z}^d} [\mathbf{P}^0(\tau = n) + \sum_{k=0}^{n-1} \mathbf{P}_y^0(\tau = k) p_{n-k}^y(0) q]$$

Let

$$\begin{aligned} S_n^X &= \sum_{y \in \mathbb{Z}^d} \mathbf{P}_y^X(\tau \leq n) \\ S_n^0 &= \sum_{y \in \mathbb{Z}^d} \mathbf{P}_y^0(\tau \leq n) \\ S_n^X - S_{n-1}^X &= \sum_{y \in \mathbb{Z}^d} \mathbf{P}_y^X(\tau = n) \end{aligned}$$

We define  $S_{-1}^X = S_{-1}^0 = 0$ .

We have

$$\begin{aligned} q &= \sum_{y \in \mathbb{Z}^d} \left[ \mathbf{P}^X(\tau = n) + \sum_{k=0}^{n-1} \mathbf{P}_y^X(\tau = k) p_{n-k}^y(X_n - X_k) q \right] \\ \implies (S_n^X - S_n^0) &\geq (1 - qp_i^Y(0))(S_{n-1}^X - S_{n-1}^0) + q \sum_{k=0}^{n-2} (S_k^X - S_k^0)(p_{n-k-1}^Y(0) - p_{n-k}^Y(0)) \end{aligned}$$

Now using induction, we get  $S_n^X \geq S_n^0$ .

□

*Remark.* The continuous case has a similar proof and can be found here [?].

## 9.4 Strategy

The strategy is to find an event such that  $\mathbf{P}_\epsilon^X(\text{event}) \approx \sigma^0(x)$  We first define the following events:

$$\begin{aligned} G_n &= \{X_n \text{ stays inside } B(0, R_n)\} \\ E_n &= \{\text{no traps in } B(0, R_n)\} \\ F_n &= \{\text{traps outside } B(0, R_n) \text{ by time } n\}. \end{aligned}$$

We get the probability of the given events as:

$$\mathbf{P}(E_n) = e^{-cR_n^d} \tag{9.2}$$

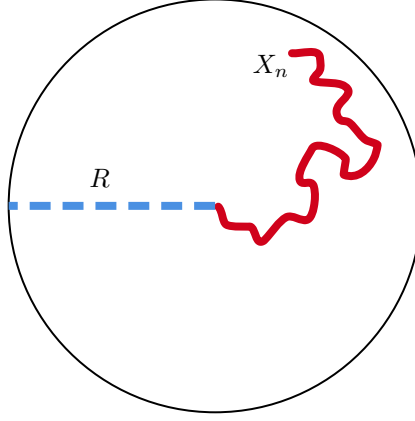


Figure 9.1: Strategy for avoiding traps by not allowing the walk to move outside the ball of radius  $R$

$$\begin{aligned}
\mathbf{P}(G_n) &= \mathbf{P}(\sup_{0 \leq k \leq n} |X_k| \leq R_n) \\
&= \mathbf{P}(\tau_{B(0, R_n)} \geq n) \\
&\geq e^{-cn/R_n^2}
\end{aligned} \tag{9.3}$$

and

$$\mathbf{P}(F_n) \geq e^{-c\sqrt{n}} \tag{9.4}$$

Choose  $R_n$ , s.t.

$$c_1 R_n^d = c_2 n / R_n^2 \text{ i.e. } R_n = c_1 n^{1/d+2}$$

This shows that  $P(G_n \cap E_n \cap F_n)$  is of the same order as  $\sigma_0(n)$ .