

## Week 2

# More on random walks

LECTURER: SIVA ATHREYA

SCRIBE: SANCHAYAN BHOWAL, VENKAT TRIVIKRAM

**Theorem 2.0.1.** *Let  $T : \Omega \rightarrow 0, 1, \dots, N$  be a stopping time. Then,*

$$\mathbf{E}[S_T^2] = E[T].$$

*Proof.*

$$\begin{aligned} S_T^2 &= \sum_{k=1}^N S_k^2 \mathbb{1}\{T = k\} \\ &= \sum_{k=1}^N (S_k^2 - S_{k-1}^2) \mathbb{1}\{T \geq k\} \\ &= \sum_{k=1}^N (X_k + S_{k-1})^2 - S_{k-1}^2 \mathbb{1}\{T \geq k\} \\ &= \sum_{k=1}^N (1 + 2X_k S_{k-1}) \mathbb{1}\{T \geq k\}. \end{aligned}$$

Now, consider  $V_k = S_{k-1} \mathbb{1}\{T \geq k\}$ . Note that this is a bet sequence. Hence,

$$\begin{aligned} \mathbf{E}[S_T^2] &= \mathbf{E} \left[ \sum_{k=1}^N \mathbb{1}\{T \geq k\} \right] + 2 \sum_{k=1}^N \mathbf{E}[X_k V_k] \\ &= \sum_{k=1}^N \mathbf{P}(T \geq k) + 0 \\ &= E[T]. \end{aligned}$$

□

## 2.1 Reflection Principle

Assume that  $a \in \mathbb{Z}$  and  $c > 0$ . There is a bijection between the paths that cross  $a + c$  and those that do not. This bijection is obtained by reflecting the part of the path crossing  $a + c$  as shown in the Figure ???. So,

$$|S_n = a + c| = |\sigma_a \leq n \text{ \& } S_n = a + c| = |\sigma_a \leq n \text{ \& } S_n = a - c|$$

Now, we know that all the paths have equal probability. Hence, we get the following lemma.

**Lemma 2.1.1.**  $\mathbf{P}(S_n = a + c) = \mathbf{P}(\sigma_a \leq n \text{ \& } S_n = a - c)$  where  $a \in \mathbb{Z}$  and  $c > 0$ . This is also known as the reflection principle.

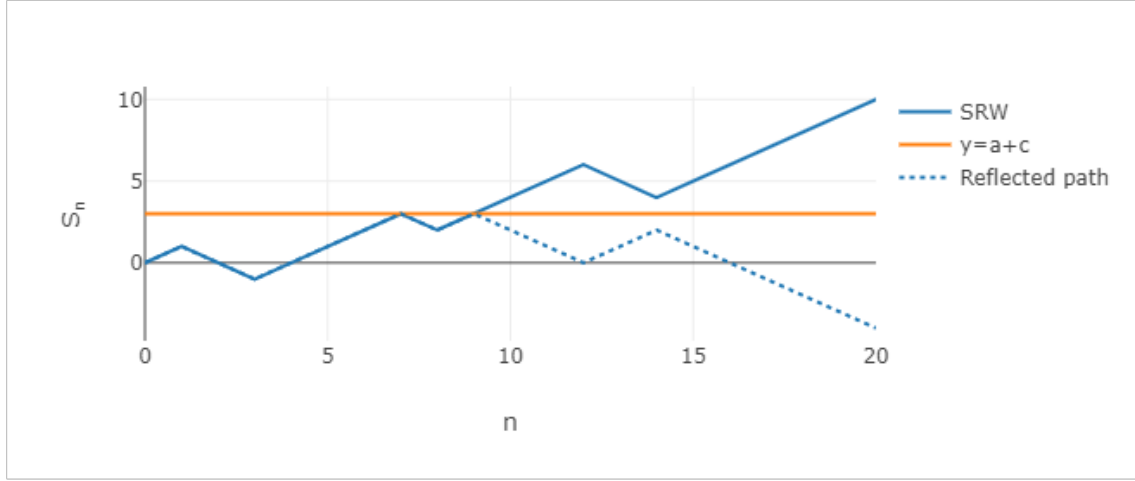


Figure 2.1: The figure shows that the bijection between the paths that cross  $a+c=3$  and those that do not.

**Theorem 2.1.1.**  $\mathbf{P}(\sigma_a \leq n) = \mathbf{P}(S_n \notin [-a, a])$  where  $a \in \mathbb{Z} \setminus \{0\}$ .

*Proof.*

$$\begin{aligned}
 \mathbf{P}(\sigma_a \leq n) &= \mathbf{P}(\sigma_a \leq n, \bigcup_{b \in \mathbb{Z}} S_n = b) \\
 &= \sum_{b \in \mathbb{Z}} \mathbf{P}(\sigma_a \leq n, S_n = b) \\
 &= \sum_{b \in \mathbb{Z}, b \geq a} \mathbf{P}(\sigma_a \leq n, S_n = b) + \sum_{b \in \mathbb{Z}, b < a} \mathbf{P}(\sigma_a \leq n, S_n = b) \\
 &= \sum_{b \in \mathbb{Z}, b \geq a} \mathbf{P}(S_n = b) + \sum_{b \in \mathbb{Z}, b < a} \mathbf{P}(S_n = 2a - b) \\
 &= \mathbf{P}(S_n \geq a) + \mathbf{P}(S_n > a) \\
 &= \mathbf{P}(S_n \geq a) + \mathbf{P}(S_n < -a) \\
 &= \mathbf{P}(S_n \notin [-a, a])
 \end{aligned}$$

□

**Corollary 2.1.1.**  $\mathbf{P}(\sigma_a = n) = \frac{1}{2} [\mathbf{P}(S_{n-1} = a - 1) - \mathbf{P}(S_{n-1} = a + 1)]$  where  $a \in \mathbb{Z}$ .

*Proof.*

□

## 2.2 Arc-Sine Law

Let  $L$  denote the last time the random walk hits 0, i.e.,  $L = \max_{0 \leq n \leq 2N} S_n = 0$ , where  $N$  denotes the length of the walk.

**Theorem 2.2.1.**

$$\mathbf{P}(L = 2n) = \frac{1}{2^{2N}} \binom{2n}{n} \binom{2N-2n}{N-n}.$$

*Remark.* By Stirling's approximation,

$$\begin{aligned} \mathbf{P}(L = 2n) &\sim \frac{1}{\pi N} \frac{1}{\sqrt{\left(\frac{n}{N}\right) \left(1 - \frac{n}{N}\right)}}. \\ \mathbf{P}\left(\frac{L}{2N} \leq x\right) &= \mathbf{P}(L \leq 2Nx) \\ &= \sum_{n=0}^{[2Nx]} \mathbf{P}(L = 2n) \\ &\sim \sum_{n=0}^{[2Nx]} \frac{1}{\pi N} \frac{1}{\sqrt{\left(\frac{n}{N}\right) \left(1 - \frac{n}{N}\right)}} \\ &\sim \int_0^x \frac{dy}{\pi \sqrt{y(1-y)}} \\ &= \frac{2}{\pi} \sin^{-1}(\sqrt{x}). \end{aligned}$$

*Proof of Theorem ??.* Define  $\tilde{\sigma}_0 = \inf\{n : S_n = 0, 0 < n \leq N\}$ . Consider a path of length  $2N$  with  $L = 2n$ . This path can be formed by a path which takes  $S_{2n} = 0$  and followed by a path of length  $2N - 2n$  with  $\sigma_0 > 2N - 2n$ . Hence, number of paths of length  $2N$  with  $L = 2n$  is the product of the number of paths of length  $2n$  with  $S_{2n} = 0$  and the number of paths of length  $2N - 2n$  with  $\sigma_0 > 2N - 2n$ . Hence,

$$\mathbf{P}(L = 2n) = \mathbf{P}(S_{2n} = 0) \mathbf{P}(\tilde{\sigma}_0 > 2N - 2n), \quad (2.1)$$

Now let us compute the distribution of  $\tilde{\sigma}_0$ .

$$\begin{aligned} \mathbf{P}(\tilde{\sigma}_0 > 2k) &= \mathbf{P}(S_1 \neq 0, \dots, S_{2k} \neq 0) \\ &= 2\mathbf{P}(S_1 > 0, \dots, S_{2k} > 0) \\ &= \frac{2}{2^{2k}} \{\text{No. of paths start at 0 and stay above -1 for } 2k - 1 \text{ steps}\} \\ &= \frac{2}{2^{2k}} \{\text{No. of paths start at 0 and stay below 1 for } 2k - 1 \text{ steps}\} \quad (2.2) \\ &= \mathbf{P}(\sigma_1 > 2k - 1) \\ &= 1 - \mathbf{P}(\sigma_1 \geq 2k - 1) \\ &= \mathbf{P}(S_{2k-1} = -1) + \mathbf{P}(S_{2k-1} = 0) \\ &= \mathbf{P}(S_{2k-1} = -1) \end{aligned}$$

Using (??) and (??),

$$\begin{aligned}\mathbf{P}(L = 2n) &= \mathbf{P}(S_{2n} = 0)\mathbf{P}(S_{2N-2n-1} = -1) \\ &= \mathbf{P}(S_{2n} = 0)\mathbf{P}(S_{2N-2n} = 0) \\ &= \frac{1}{2^{2N}} \binom{2n}{n} \binom{2N-2n}{N-n}.\end{aligned}$$

The first step analysis of  $S_{2n}$  shows that,  $\mathbf{P}(S_{2N-2n} = 0) = \frac{1}{2}\mathbf{P}(S_{2N-2n-1} = 1) + \frac{1}{2}\mathbf{P}(S_{2N-2n-1} = -1)$ . Using the symmetry of the walk we know that  $\mathbf{P}(S_{2N-2n-1} = 1) = \mathbf{P}(S_{2N-2n-1} = -1)$ . This gives the second inequality.  $\square$

## 2.3 SRW of length $N$ in $\mathbb{Z}^d$

### 2.3.1 Notations and notions in higher dimension

- $e_i \in \mathbb{Z}^d, \forall i \in \{1, 2, \dots, d\}$ , defined as the vector of length  $d$  with all entries zeroes except  $i^{th}$  being 1.

$$e_i = (0, 0, \dots, \underbrace{1}_{i^{th}}, 0, \dots, 0)$$

- For  $x \in \mathbb{Z}^d$ ,

$$x = \sum_{i=1}^d x_i e_i, \quad x_i \in \mathbb{Z} \quad \|x\| = \left( \sum_{i=1}^d x_i^2 \right)^{\frac{1}{2}}$$

- $\Omega_N = \{(\omega_1, \omega_2, \dots, \omega_N) \mid \omega_i \in \mathbb{Z}^d, \|\omega_i\| = 1 \forall 1 \leq i \leq N\}$
- We have, for  $1 \leq k, n \leq N$

$$X_k : \Omega_N \rightarrow \mathbb{Z}^d, \quad X_k(\omega) = \omega_k \quad S_n : \Omega_N \rightarrow \mathbb{Z}^d, \quad S_n(\omega) = \sum_{k=1}^n X_k(\omega)$$

with  $S_0(\omega) = 0$ . We can consider  $S_n$  as a  $d$ -dimensional vector given by  $S_n = (S_n^{(1)}, S_n^{(2)}, \dots, S_n^{(d)})$ , where each  $S_n^{(i)}$  is a random walk on  $\mathbb{Z}$ .

- The probability function  $\mathbf{P}^N$ , given by,

$$\mathbf{P}^N : \mathcal{P}(\Omega_N) \rightarrow [0, 1], \quad \mathbf{P}(A) = \frac{|A|}{(2d)^N} \forall A \subseteq \Omega_N$$

### 2.3.2 Infinite length random walk

On extending  $N \rightarrow \infty$ , we preserve something called as “consistency”. First, let us define, for  $0 < N < M$ ,

$$\pi_N : \Omega_M \rightarrow \Omega_N, \quad \pi_N(\omega_1, \omega_2, \dots, \omega_M) = (\omega_1, \omega_2, \dots, \omega_N)$$

Under  $(\Omega_N, \mathcal{P}(\Omega_N), \mathbf{P}^N)$  and  $(\Omega_M, \mathcal{P}(\Omega_M), \mathbf{P}^M)$ , if we observe the walk till time  $n < N$  the probability of evenets concerning the walk should be same under  $\mathbf{P}^N$  or  $\mathbf{P}^M$ . For any event  $\{\tilde{\omega} \in \Omega_N\}$ , there exists a corresponding same event namely  $\{\omega \in \Omega_M : \pi_N(\omega) = \tilde{\omega}\}$ . We have,

$$\mathbf{P}^N(\{\tilde{\omega}\}) = \frac{1}{(2d)^N} \quad \mathbf{P}^M(\{\omega \in \Omega_M : \pi_N(\omega) = \tilde{\omega}\}) = \frac{(2d)^{M-N}}{(2d)^M} = \frac{1}{(2d)^N}$$

So, we say the sequence of probability spaces  $(\Omega_1, \mathbf{P}^1), (\Omega_2, \mathbf{P}^2), \dots, (\Omega_N, \mathbf{P}^N)$  satisfies the consistency condition

$$\mathbf{P}^N(\{\tilde{\omega}\}) = \frac{1}{(2d)^N} = \frac{(2d)^{M-N}}{(2d)^M} = \mathbf{P}^M(\{\omega \in \Omega_M : \pi_N(\omega) = \tilde{\omega}\}), \quad 0 < N < M, \quad \tilde{\omega} \in \Omega_N$$

We define the space of infinite sequences,

$$\Omega_\infty = \{\omega = (\omega_k)_{k \geq 1} \mid \omega_k \in \mathbb{Z}^d, \|\omega_k\| = 1\}$$

$\mathcal{A}_\infty (\equiv \mathcal{P}(\Omega_\infty))$  denotes the class of events observable “for ever”

For  $N \in \mathbb{N}$ ,

$$\pi_N : \Omega_\infty \rightarrow \Omega_N, \quad \pi_N(\omega) = (\omega_1, \omega_2, \dots, \omega_N)$$

**Theorem 2.3.1 (Kolmogorov Consistency Theorem).** There exists a unique probability measure on  $(\Omega_\infty, \mathcal{A}_\infty)$  such that  $\forall N \geq 1, \forall \tilde{\omega} \in \Omega_N$ ,

$$\mathbf{P}^N(\{\tilde{\omega}\}) = \mathbf{P}^M(\{\omega \in \Omega_M : \pi_N(\omega) = \tilde{\omega}\}) = \frac{1}{(2d)^N}$$

Now, we can define,

$$X_k : \Omega_\infty \rightarrow \mathbb{Z}^d, \quad X_k(\omega) = \omega_k \quad S_n = \sum_{k=1}^n X_k \quad \forall n \geq 1$$

under  $\mathbf{P}$ ,  $\{S_n\}_{n \geq 1}$  is a simple random walk starting at  $S_0 = 0$ .

**Definition 2.3.1.**  $A \subseteq \Omega_\infty$  is said to be **observable** by time  $n$  if  $A$  is a union of the events of the form

$$\{\omega \in \Omega_\infty : \omega_i = o_i, 1 \leq i \leq N\} \text{ with } o_i \in \mathbb{Z}^d, \|o_i\| = 1$$

For,  $k \in \mathbb{N}_0$ ,  $\mathcal{A}_k$  denotes the set of all events in  $\Omega_\infty$  observable by time  $k$ .

**Definition 2.3.2.**  $T : \Omega_\infty \rightarrow \mathbb{N} \cup \{\infty\} \cup \{0\}$  is a **stopping time** if

$$\text{for any } k \in \mathbb{N}_0, \{T = k\} \in \mathcal{A}_k$$

For example,  $\sigma_a = \min\{n \geq 0 \mid S_n = a\}$  is a stopping time.

### 2.3.3 Speed of the walk

**Definition 2.3.3.** For,  $S_n = \sum_{k=1}^n X_k$ , we define **speed of the walk** as

$$\text{Speed} = \frac{S_n}{n} = \frac{1}{n} \sum_{k=1}^n X_k$$

We have,  $X_k = (X_k^{(1)}, X_k^{(2)}, \dots, X_k^{(d)})$ ,  $\{X_k\}_{k \geq 1}$  which is an i.i.d sequence of random variables with

$$\mathbf{P}(X_k = e_i) = \frac{1}{2d} = \mathbf{P}(X_k = -e_i)$$

$$\Rightarrow \mathbf{E}[X_k] = 0 \text{ and } \mathbf{E}[\|X_k\|] = 1 (\leq \infty)$$

**Theorem 2.3.2 (Strong law of large numbers).** For simple random walk on  $\mathbb{Z}^d$ ,

$$\frac{S_n}{n} \rightarrow 0 \text{ with probability 1 under } (\Omega_\infty, \mathcal{A}_\infty, \mathbf{P})$$

### 2.3.4 Typical position of the walk

For  $d = 1$ ,

$$\begin{aligned} \frac{S_n - (n)(0)}{\sqrt{n}} &\xrightarrow{d} \mathcal{N}(0, 1) \\ \Rightarrow \sqrt{n} \left( \frac{S_n}{n} \right) &\xrightarrow{d} \mathcal{N}(0, 1) \end{aligned}$$

For  $d > 1$ ,  $\mu \in \mathbb{R}^d$  and a positive definite matrix  $\Sigma \in \mathbb{R}^{d \times d}$ , we have  $d$ -dimensional normal distribution as,

$$\Phi_{d,\mu,\Sigma}(y) = \frac{1}{(2\pi)^{d/2}} \frac{1}{\det(\Sigma)^{1/2}} \exp \left( -\frac{1}{2} (x - \mu)^t \Sigma^{-1} (x - \mu) \right)$$

$$\mathbf{P} \left( \frac{S_n}{\sqrt{n}} \in \prod_{i=1}^d [a_i, b_i] \right) \xrightarrow{n \rightarrow \infty} \int_{\prod_{i=1}^d [a_i, b_i]} \Phi_{d,0,\Sigma^d}(y) dy$$

where,  $\mu = 0$ ,  $\Sigma^d = \text{diag}(\frac{1}{d}, \dots, \frac{1}{d})$

### 2.3.5 Large deviation principle

From the CLT, we have that

$$\mathbf{P}(\|S_n\| > a\sqrt{n}) \xrightarrow{n \rightarrow \infty} \int_{\|x\| > a} \Phi_{d,0,\Sigma^d}(y) dy$$

We consider the events of the form  $\{\|S_n\| > an\}$ ,  $a \in [0, \infty)$ , which are “rare” in the sense that their probability tends to 0 as  $n \rightarrow \infty$ . On formal application of CLT shows that probability of these rare events are exponentially small.

**Theorem 2.3.3 (Cramer’s theorem).** For,  $a > 0$ ,

$$\lim_{n \rightarrow \infty} \frac{\log(\mathbf{P}(\|S_n\| > an))}{n} = -I(a)$$

where,

$$I(a) = \begin{cases} \log 2 + \frac{1+a}{2} \log \frac{1+a}{2} + \frac{1-a}{2} \log \frac{1-a}{2}, & \text{for } a \in [-1, 1] \\ \infty, & \text{otherwise} \end{cases}$$

It can be vaguely interpreted as,  $\mathbf{P}(\|S_n\| > na) \sim e^{-nI(a)}$

## 2.4 Exercises

1. Complete the proof of Reflection Principle (Lemma ??).
2. Find the distribution of  $M_k = \max_{1 \leq k \leq n} S_k$ .
3. Show that  $\mathbf{E}[\|X_k\|] = 1$ .