Mixed Model Cheat Sheet

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1	Model statement
2	Likelihood

1) Model statement

A linear Gauss-Markov models is presented as:

$$y = Xb + \varepsilon$$
$$y \sim N(Xb, V)$$
$$\varepsilon \sim N(0, V)$$

And independent random terms other than the residuals:

$$\mathbf{\epsilon} = \sum_{i=1}^{i=1} \mathbf{Z}_i \mathbf{u}_i + \mathbf{e}$$

And the variance-covariance matrix is defined by

$$E[\boldsymbol{\epsilon}\boldsymbol{\epsilon}'] = \boldsymbol{V} = \sum_{i=1}^{i=1} \boldsymbol{Z}_{i}\boldsymbol{K}_{i}\boldsymbol{Z}'_{i}\sigma_{i}^{2} + \boldsymbol{I}\sigma_{e}^{2}$$

Reshape the model the model into:

$$y = Xb + \sum_{i=1}^{i=1} Z_i u_i + e$$

$$u \sim N(0, Z_i K_i Z_i' \sigma_i^2)$$

$$e \sim N(0, I \sigma_e^2)$$

2) Likelihood

Gaussian likelihood is defined as:

$$L = 2\pi^{-0.5n} |\mathbf{V}|^{-0.5} \exp(-0.5(\mathbf{y} - \mathbf{Xb})'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{Xb}))$$

Log likelihood:

$$\ln L = -\frac{1}{2}[\ln|\mathbf{V}| + \mathbf{y}'\mathbf{P}\mathbf{y}]$$

And restricted log likelihood (log REML) is defined as:

$$\ln L = -\frac{1}{2} \left[\ln |\mathbf{V}| + \ln |\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}| + \mathbf{y}'\mathbf{P}\mathbf{y} \right]$$

Where the projection matrix **P** is defined by:

$$P = V^{-1} - V^{-1}X(X'V^{-1}X)^{-1}X'V^{-1}$$

Properties and equalities of projection matrices (S and P):

$$\begin{split} & \textbf{SX} = \textbf{PX} = 0, & \textbf{Py} = \textbf{V}^{-1}(\textbf{y} - \textbf{Xb}), & \textbf{PP} = \textbf{P} \\ & \textbf{PVP} = \textbf{P}, & \textbf{y}' \textbf{Py} = \textbf{y}' \textbf{e} \ \sigma_e^{-2}, & \textbf{P} = \textbf{V}^{-1} \textbf{S} \\ & \textbf{S} = \textbf{I} - \textbf{X} (\textbf{X}'\textbf{X})^{-1} \textbf{X}, & \textbf{Sy} = \textbf{y} - \textbf{Xb}, & \textbf{SS} = \textbf{S} \\ & \textbf{ZAZPy} = \textbf{Zu} \sigma_u^{-2}, & \textbf{Py} = (\textbf{y} - \textbf{Xb} - \textbf{Zu}) \sigma_e^{-2} \end{split}$$

The inverse of V can be expressed as follows

$$V^{-1} = I - Z(Z'Z + \lambda K^{-1})^{-1}Z'$$

3) BLUE and BLUP

The solution for the best linear unbiased estimators (**b**, *BLUE*) and best linear unbiased predictor (**u**, *BLUP*) are commonly defined through the equations:

$$\mathbf{b} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$$

$$\mathbf{u} = \sigma_i^2\mathbf{K}_i\mathbf{Z}_i'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\mathbf{b})$$

Conditional to the random terms, BLUE coefficients can be estimated from solving the system of equation:

$$(\mathbf{X}'\mathbf{X}) \mathbf{b} = \mathbf{X}'(\mathbf{y} - \sum \mathbf{Z}_i \mathbf{u}_i)$$

And coefficients of the ith random effect, conditional to fixed effects and other random effects, as:

$$(\mathbf{Z}_{i}'\mathbf{Z}_{i} + \mathbf{K}_{i}^{-1}\lambda_{i}) \mathbf{u} = \mathbf{Z}_{i}'(\mathbf{y} - \mathbf{X}\mathbf{b} - \sum \mathbf{Z}_{-i}\mathbf{u}_{-i})$$

Where $\lambda_i = \sigma_e^2 \sigma_i^{-2}$.

4) General terms

In a more generalized notation, the mixed model equation is described as:

$$\mathbf{y} = \mathbf{X}\mathbf{b} + \sum_{l}^{i=1} \mathbf{Z}_{l}\mathbf{u}_{i} + \mathbf{e} = \mathbf{W}\mathbf{g} + \mathbf{e}$$

Where the design matrices and coefficients are unified as:

$$\mathbf{W} = \mathbf{X}|\mathbf{Z}_1|\mathbf{Z}_{...}|\mathbf{Z}_{\mathrm{I}}$$
$$\mathbf{g} = \mathbf{b}|\mathbf{u}_1|\mathbf{u}_{...}|\mathbf{u}_{\mathrm{I}}$$

And the mixed model can be written as:

$$y = Wg + e$$

 $y \sim N(Wg, I\sigma_e^2)$

Where the coefficients are, in probabilistic terms, described as a normal distribution with a block diagonal variance:

$$\mathbf{g} \sim \mathsf{MVN} \left(\begin{bmatrix} b \\ 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} \infty & 0 & \cdots & 0 \\ 0 & \mathbf{K}_1 \sigma_1^2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \mathbf{K}_1 \sigma_1^2 \end{bmatrix} \right)$$

Thus, the mixed model equation is generalized as:

$$\mathbf{W}'\mathbf{W} + \mathbf{\Sigma} = \mathbf{W}'\mathbf{y}$$

$$\mathbf{\Sigma} = \begin{bmatrix} 0 & 0 & \cdots & 0 \\ 0 & \mathbf{K}_1^{-1} \lambda_1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \mathbf{K}_I^{-1} \lambda_I \end{bmatrix}$$

And the mixed model equation makes the complex problem a simple "Ax = b" mathematic problem:

$$\label{eq:continuity} \begin{split} C &= gr \\ C &\to \ W'W + \Sigma \\ r &\to \ W'y \end{split}$$

5) Coefficients

The conditional expectation of each regression coefficient, assuming $K_i = I$, is expressed as:

$$g_j = \frac{\mathbf{w}_j'(\mathbf{y} - \mathbf{W}_{-j}\mathbf{g}_{-j})}{\mathbf{w}_i'\mathbf{w}_i + \lambda_i} = \frac{\mathbf{w}_j'\tilde{\mathbf{y}}}{\mathbf{w}_i'\mathbf{w}_i + \lambda_i}$$

Where w_i from fixed effect coefficients assume $\lambda = 0$.

Conditional to all other terms via *Gauss-Seidel residual update* (GSRU), coefficients can be efficiently estimated in two steps:

$$\begin{split} g_j^{t+1} &= \frac{w_j' e^t + w_j' w_j g_j^t}{w_j' w_j + \lambda_j} \\ e^{t+1} &= e^t - w_j (g_j^{t+1} - g_j^t) \end{split}$$

6) Kernel trick

It is also possible to expand this framework to scenarios where $K_i \neq I$, that is through the diagonalization of K_i by spectral decomposition:

$$\mathbf{K}_{i} = \mathbf{U}_{i} \mathbf{D}_{i} \mathbf{U}_{i}^{\prime}$$

Where **U** is a matrix of Eigenvectors and **D** is a diagonal matrix of Eigenvalues. With properties

$$\begin{split} \mathbf{K}^{-1} &= \mathbf{U}\mathbf{D}^{-1}\mathbf{U}'\\ \mathbf{K}'\mathbf{K} &= \mathbf{U}\mathbf{D}^{2}\mathbf{U}'\\ \mathbf{D} &= \mathbf{U}'\mathbf{K}\mathbf{U}\\ \mathbf{I} &= \mathbf{U}'\mathbf{U}\\ \mathbf{U}' &= \mathbf{U}^{-1}\\ \mathrm{tr}(\mathbf{K}) &= \mathrm{tr}(\mathbf{D}) \end{split}$$

Such that:

$$\begin{aligned} &\mathbf{u_i} \sim N(0, \mathbf{Z_i} \mathbf{K_i} \mathbf{Z_i'} \sigma_i^2 \) \\ &\mathbf{u_i} \sim \mathbf{U} \ N(0, \mathbf{Z_i} \mathbf{D_i} \mathbf{Z_i'} \sigma_i^2 \) \\ &\mathbf{a_i} \sim N(0, \mathbf{D_i} \sigma_i^2 \) \\ &\mathbf{E}[\mathbf{u_i}] = \mathbf{U_i} \mathbf{a_i} \\ &\mathbf{Z_i} \mathbf{u_i} = \mathbf{Z_i} \mathbf{U_i} \mathbf{a_i} \end{aligned}$$

Assume the notation:

$$V_i = Z_i U_i$$

Then, for ith random effect and jth Eigen-pair, the regression coefficient has expectation (GSRU):

$$g_{ij} = \frac{v'\tilde{y}}{v'v + \lambda_i d_i^{-1}}, \qquad v = Z_i U_{ij}$$

7) Information matrix

$$\begin{split} \textbf{\textit{V}}_{(Zu)} &= \frac{\partial \textbf{\textit{V}}}{\partial \sigma_{i}^{2}} = \textbf{\textit{Z}}_{i}\textbf{\textit{K}}_{i}\textbf{\textit{Z}}_{i} \quad \text{and} \quad \textbf{\textit{V}}_{(e)} = \frac{\partial \textbf{\textit{V}}}{\partial \sigma_{e}^{2}} = \textbf{\textit{I}} \\ &\text{Information} = P\frac{\partial \textbf{\textit{V}}}{\partial \sigma_{i}^{2}}P\frac{\partial \textbf{\textit{V}}}{\partial \sigma_{i}^{2}} = \ P\textbf{\textit{V}}_{i}P\textbf{\textit{V}}_{j} \end{split}$$

8) Variance components

First derivative 1: Expectation-maximization

$$\begin{split} \sigma_e^2 &= \frac{y'Py}{\text{tr}(P)} = \frac{e'e + \text{tr}(WC^{-1}W')\sigma_e^{2^{t-1}}}{n} = \frac{y'e}{n-r} \\ \sigma_i^2 &= \frac{u_i'K_i^{-1}u_i + \text{tr}\big(K_i^{-1}C^{ii}\big)\sigma_e^2}{q_i} = \frac{u_i'K_i^{-1}u_i}{q_i - \lambda \text{tr}(K_i^{-1}C^{ii})} \end{split}$$

$$\begin{split} n &= \text{number of observations} \\ r &= \text{rank of } \boldsymbol{X} \text{ (number of columns of } \boldsymbol{X} \text{)} \\ q_i &= \text{number of columns of } \boldsymbol{Z}_i \end{split}$$

First derivative 2: Gibbs sampling (flat priors: S = 0, v = -2)

$$\sigma_e^2 \sim \frac{e'e + S_e \nu_e}{\chi_{n+\nu_e}^2} \text{ and } \sigma_i^2 \sim \frac{u_i' K_i^{-1} u_i + S_i \nu_i}{\chi_{q_i+\nu_i}^2}$$

First derivative 3: Pseudo-expectation

$$\sigma_e^2 = \frac{\mathbf{y}'\mathbf{e}}{\mathbf{n} - \mathbf{r}}$$
 and $\sigma_i^2 = \frac{(\mathbf{y} - \mathbf{X}\mathbf{b})'\mathbf{Z}_i\mathbf{u}_i}{\operatorname{tr}(\mathbf{S}\mathbf{Z}_i\mathbf{K}_i\mathbf{Z}_i)}$

Second derivative 1: MINQUE

$$\begin{bmatrix} tr(\textbf{PZKZ'PZKZ'}) & tr(\textbf{PZKZ'P}) \\ tr(\textbf{PZKZ'P}) & tr(\textbf{P}) \end{bmatrix} \begin{bmatrix} \sigma_u^2 \\ \sigma_e^2 \end{bmatrix} = \begin{bmatrix} y' \textbf{PZKZ'Py} \\ y' \textbf{Py} \end{bmatrix}$$

Second derivative 2: Fisher scoring

$$\begin{split} \boldsymbol{v}^{t+1} &= \boldsymbol{v}^t - \boldsymbol{I}^{-1}\boldsymbol{s}, & \boldsymbol{v} = \begin{bmatrix} \sigma_u^2 \\ \sigma_e^2 \end{bmatrix} \\ \boldsymbol{I} &= \begin{bmatrix} \operatorname{tr}(\boldsymbol{PZKZ'PZKZ'}) & \operatorname{tr}(\boldsymbol{PZKZ'P}) \\ \operatorname{tr}(\boldsymbol{PZKZ'P}) & \operatorname{tr}(\boldsymbol{P}) \end{bmatrix} \\ \boldsymbol{s} &= \begin{bmatrix} \operatorname{tr}(\boldsymbol{PZKZ'}) - \boldsymbol{y'PZKZ'Py} \\ \operatorname{tr}(\boldsymbol{P}) - \boldsymbol{y'Py} \end{bmatrix} \end{split}$$

Second derivative 3: Average information

$$\begin{aligned} \mathbf{v}^{t+1} &= \mathbf{v}^t - \mathbf{A}\mathbf{I}^{-1}\mathbf{s}, & \mathbf{v} &= \begin{bmatrix} \sigma_u^2 \\ \sigma_e^2 \end{bmatrix} \\ \mathbf{A}\mathbf{I} &= 0.5 \times \begin{bmatrix} \mathrm{tr}(\mathbf{y}'\mathbf{PZKZ}'\mathbf{PZKZ}'\mathbf{Py}) & \mathrm{tr}(\mathbf{y}'\mathbf{PZKZ}'\mathbf{P}) \\ \mathrm{tr}(\mathbf{y}'\mathbf{PZKZ}'\mathbf{P}) & \mathrm{tr}(\mathbf{y}'\mathbf{Py}) \end{bmatrix} \\ \mathbf{s} &= \begin{bmatrix} \mathrm{tr}(\mathbf{PZKZ}') - \mathbf{y}'\mathbf{PZKZ}'\mathbf{Py} \\ \mathrm{tr}(\mathbf{P}) - \mathbf{y}'\mathbf{Py} \end{bmatrix} \end{aligned}$$

Derivation-free method: EMMA

$$\begin{split} \mathbf{Z}\mathbf{K}\mathbf{Z}' &= \mathbf{U}\mathbf{D}\mathbf{U}' \\ \mathbf{b} &= (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y} \\ \sigma_e^2 &= \frac{(\mathbf{y} - \mathbf{X}\mathbf{b})'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\mathbf{b})}{\mathbf{n} - \mathbf{r}} \\ \\ \mathrm{argmax} \ L(\lambda) &= -0.5 \ [\ \ln|\mathbf{V}| + \ln|\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}| + \mathbf{y}'\mathbf{P}\mathbf{y}] \\ \ln|\mathbf{V}| &= \sum \ln \left(\mathbf{d}_j \lambda^{-1} + 1\right) \\ \ln|\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}| &= \ln|\mathbf{X}'\mathbf{U}(\mathbf{D}^{-1}\lambda^{-1} + \mathbf{I})\mathbf{U}'\mathbf{X} \ | \\ \mathbf{y}'\mathbf{P}\mathbf{y} &= \mathbf{y}'\mathbf{e} \end{split}$$