## Raised Incidence Model for Hamburg Case

This script implements the raised incidence model by Diggle (1990) and the extensions to accommodate multiple potential sources by Diggle & Rowlingson (1994) on Data of the region of Hamburg.

- Population Data as the population at risk  $[\lambda_0(x)]$
- Retailer location data as sources  $[x_k]$
- Outbreak case locations as outbreak  $[\lambda_1(x)]$

All spatial data is stored in the coordinate system ETRS89-LAEA Europe - EPSG:3035.

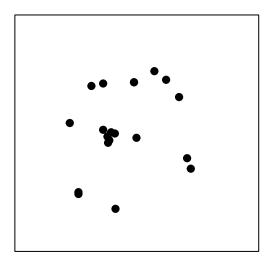
### 1 Load Libraries

```
library(spatstat)
library(sf)
```

### 2 Load Outbreak from Shapefile

```
sf outbreak <- st read("./Data/Outbreak/Aldi Outbreak.shp")</pre>
## Reading layer 'Aldi_Outbreak' from data source
     'C:\Users\srude\Documents\Dev Kram\Traceback_Model\Data\Outbreak\Aldi_Outbreak.shp'
##
##
     using driver 'ESRI Shapefile'
## Simple feature collection with 20 features and 1 field
## Geometry type: POINT
## Dimension:
## Bounding box: xmin: 4312050 ymin: 3372150 xmax: 4331650 ymax: 3394450
## Projected CRS: ETRS89-extended / LAEA Europe
Convert into spatstat ppp object
ppp_outbreak <- as.ppp(X=sf_outbreak$geometry, W=owin(c(4303150,4342650), c(3365250,3403550)))
ppp_outbreak
## Planar point pattern: 20 points
## window: rectangle = [4303150, 4342650] x [3365250, 3403550] units
Plot
plot(ppp_outbreak, pch = 19, main = "Artificial Outbreak")
```

### **Artificial Outbreak**



## 3 Load Population Data

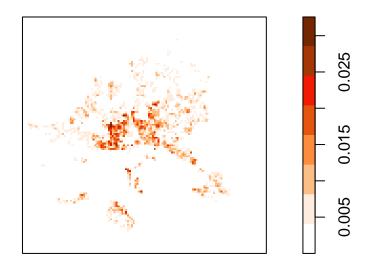
We applied kernel smoothing (Diggle, 1985) to the population data to obtain an unnormalised estimate of the spatially varying population density of susceptibles  $(\lambda_0(x))$ .

```
im_population <- readRDS("./Data/im_population.rds")</pre>
```

#### Plot

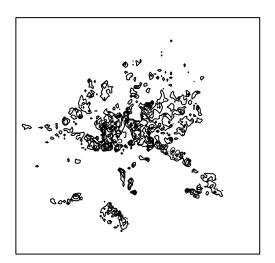
```
colors <- c("#ffffff", "#FEEDDE", "#FDBE85", "#FD8D3C", "#E6550D", "#F01B00", "#A63603", "#732600")
plot(im_population, main = "Kernel Smoothed Population Density Hamburg 2011", col =colors)</pre>
```

# **Kernel Smoothed Population Density Hamburg 2011**



contour(im\_population, main = "Contour of Population Density Hamburg 2011")

## **Contour of Population Density Hamburg 2011**



### 4 Load Potential Pattern

For now, we only use the one point source model, so we read only the Shapefile of one store of the chain.

```
sf_stores <- st_read("./Data/Potential Pattern/Aldi_store_6.shp")</pre>
## Reading layer 'Aldi_store_6' from data source
    'C:\Users\srude\Documents\Dev Kram\Traceback_Model\Data\Potential Pattern\Aldi_store_6.shp'
##
## using driver 'ESRI Shapefile'
## Simple feature collection with 1 feature and 33 fields
## Geometry type: POINT
## Dimension:
                  XΥ
## Bounding box: xmin: 4316914 ymin: 3385953 xmax: 4316914 ymax: 3385953
## Projected CRS: ETRS89-extended / LAEA Europe
Convert into spatstat ppp object
ppp_stores <- as.ppp(X=sf_stores$geometry, W=owin(c(4303150,4342650), c(3365250,3403550)))
ppp_stores
## Planar point pattern: 1 point
## window: rectangle = [4303150, 4342650] x [3365250, 3403550] units
Plot
plot(ppp_stores, pch = 7, main = "Source")
```

### Source

×

```
print(ppp_stores$x[1])
```

## [1] 4316914

```
print(ppp_stores$y[1])
## [1] 3385953
5
     Fit Null Model
H_0: \lambda_1(x) = \rho \lambda_0(x)
fit0 <- ppm(ppp_outbreak ~ offset(log(im_population)))</pre>
## Warning: glm.fit: algorithm did not converge
## Warning: glm.fit: algorithm did not converge
## Warning in glm(fmla, family = quasi(link = "log", variance = "mu"), weights =
## .mpl.W, : fitting to calculate the null deviance did not converge -- increase
## 'maxit'?
fit0
## Nonstationary Poisson process
##
## Log intensity: ~offset(log(im_population))
## Fitted trend coefficient: (Intercept) = 320.0382
##
##
                                  S.E. CI95.lo CI95.hi Ztest
                Estimate
                                                                          7.val
## (Intercept) 320.0382 2.294984e-73 320.0382 320.0382
                                                              *** 1.394511e+75
## *** Fitting algorithm for 'glm' did not converge ***
     Fit Alternative Model
6
H_0: \lambda_1(x) = \rho \lambda_0(x) (1 + \alpha e^{-\beta x'x})
We define the squared distance to the source:
d2source <- function(x, y, xPOS=ppp_stores$x[1], yPOS = ppp_stores$y[1]) {
  (x - xPOS)^2 + (y - yPOS)^2
We define the functional form of the raised incidence part:
raisin <- function(x, y, alpha, beta) {</pre>
  1 + alpha * exp(-beta * d2source(x, y))
We fit the model to our outbreak
fit1 <- ippm(ppp_outbreak ~ offset(log(im_population) + log(raisin)),</pre>
  start = list(alpha = 5, beta = 1))
```

## Warning: glm.fit: algorithm did not converge

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```
## Warning in glm(fmla, family = quasi(link = "log", variance = "mu"), weights =
## .mpl.W, : fitting to calculate the null deviance did not converge -- increase
## 'maxit'?
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## .mpl.W, : fitting to calculate the null deviance did not converge -- increase
## 'maxit'?
fit1
## Nonstationary Poisson process
##
## Log intensity: ~offset(log(im_population) + log(raisin))
## Fitted trend coefficient: (Intercept) = 320.0382
## Irregular parameters (covfunargs) fitted by 'ippm':
## alpha = 5
## beta = 1
                                S.E. CI95.lo CI95.hi Ztest
               Estimate
## (Intercept) 320.0382 2.294984e-73 320.0382 320.0382
                                                         *** 1.394511e+75
```

## \*\*\* Fitting algorithm for 'glm' did not converge \*\*\*

## 6.1 Statistical Inference comparing the two models

```
anova(fit0, fit1, test="LRT")

## Analysis of Deviance Table
##

## Model 1: ~offset(log(im_population)) Poisson
## Model 2: ~offset(log(im_population) + log(raisin)) Poisson
## Npar Df Deviance Pr(>Chi)
## 1  1
## 2  3  2  0  1
```