Course project

"Optimization approaches to community detection"

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- 3.1 Semidefinite relaxations
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It is assumed, that latent group labels Z has a multinomial distribution with parameter π and each element of A has a Bernoulli distribution with parameter θ_{ij} :

$$Z_i \sim \text{Poly}(\pi), \quad \pi^T = (\pi_1, \dots, \pi_k),$$

 $a_{ij} \sim \text{Bernoulli}(P_{Z(i)Z(j)}), \quad i, j = \overline{1, n},$

A Bayesian approach is used to estimate the most probable configuration of Z given an adjacency matrix A

$$Z^* = \arg\max_{Z} p(Z|A) = \arg\max_{Z} \iint p(Z, \pi, P|A) d\pi dP$$
 (3.2.1)

It treats parameters π and θ_{ij} as random variables with following prior distributions

$$\pi \sim \text{Dirichlet}(\alpha)$$

$$P_{ii} \sim \text{Beta}(\beta), \quad i = \overline{1, k},$$

where α and β are predefined hyperparameters and P_{ij} , $i \neq j$ are set to equal to a small constant ε . Futhermore, to tackle the intractable integration in 3.2.1 a restriction on the family of factorized distributions is considered

$$\mathcal{Q}=\{q:\,q(Z,\pi,P)=q(\pi)q(P)\prod_i q(Z_i)\},$$

where

$$q(Z_i): Z_i \sim \operatorname{Poly}(\tilde{\pi})$$

 $q(\pi): \pi \sim \operatorname{Dirichlet}(\tilde{\alpha})$
 $q(P): P_{ii} \sim \operatorname{Beta}(\tilde{\beta}_i)$

The optimal distribution $q^* \in \mathcal{Q}$, that approximates the true posterior $p(Z, \pi, P|A)$ minimizes the Kullback-Leibler divergence

$$q^* = \arg\min_{q \in \mathcal{Q}} \text{KL}\left(q \| p(Z, \pi, P | X)\right)$$
(3.2.2)

Define

$$\mathcal{L}(q) = \sum_{Z} \iint q(Z, \pi, P) \log \frac{p(A, Z, \pi, P)}{q(Z, \pi, P)} d\pi dP$$
(3.2.3)

Since

$$\mathcal{L}(q) + \text{KL}\left(q \| p(Z, \pi, P | X)\right) = \log p(A)$$

the minimization problem 3.2.2 can be equivalently solved by maximizing $\mathcal{L}(q)$

$$q^* = \arg\max_{q \in \mathcal{Q}} \mathcal{L}(q) \tag{3.2.4}$$

Denote $\tilde{\Pi} = ||\tilde{\pi}_{ij}||$, $i = \overline{1, n}$, $j = \overline{1, k}$. Given $\tilde{\Pi}$ values of parameters, that maximize $\mathcal{L}(q)$ can be found according to formulas

$$\tilde{\alpha} = \alpha + \tilde{\Pi}^T 1$$

$$\tilde{\beta}_i = \beta + \frac{1}{2} \left(\tilde{\Pi}_{\cdot i}^T A \tilde{\Pi}_{\cdot i}, \tilde{\Pi}_{\cdot i}^T \overline{A} \tilde{\Pi}_{\cdot i} \right)^T, \quad i = \overline{1, k},$$
(3.2.5)

where 1 denotes an all-ones vector, $\overline{A} = 11^T - I - A$, and $\tilde{\Pi}_{i}$ stands for the *i*-th column of the matrix $\tilde{\Pi}$. A corresponding value of log-likelihood is equal to

$$\mathcal{L}(\tilde{\Pi}) = \sum_{i < j} \tilde{\Pi}_{\cdot i}^T A \tilde{\Pi}_{\cdot j} \log \varepsilon + \tilde{\Pi}_{\cdot i}^T \overline{A} \tilde{\Pi}_{\cdot j} \log (1 - \varepsilon) - \sum_{i, j} \tilde{\pi}_{ij} \log \tilde{\pi}_{ij} + \log \frac{\mathcal{B}(\tilde{\alpha})}{\mathcal{B}(\alpha)} + \sum_{i} \log \frac{\mathcal{B}(\tilde{\beta})}{\mathcal{B}(\beta)},$$
(3.2.6)

where $\tilde{\alpha} = \tilde{\alpha}(\tilde{\Pi})$ and $\tilde{\beta} = \tilde{\beta}(\tilde{\Pi})$ can be found from 3.2.5, $\mathcal{B}(\alpha) \triangleq \frac{\Gamma\left(\sum_{i} \alpha_{i}\right)}{\prod_{i} \Gamma(\alpha_{i})}$ and Γ is gamma-function. Now the maximization problem 3.2.4 can be reformulated as follows

$$\begin{cases} \mathcal{L}(\tilde{\Pi}) \longrightarrow \max \\ \sum_{j} \tilde{\pi}_{ij} = 1, \quad i = \overline{1, n} \end{cases}$$
 (3.2.7)

One can use reparametrization

$$\tilde{\pi}_{ij} = e^{\theta_{ij} - \mathcal{A}_i}, \quad i = \overline{1, n}, j = \overline{1, k - 1},$$

$$\tilde{\pi}_{ik} = e^{-\mathcal{A}_i}, \quad i = \overline{1, n},$$

$$\mathcal{A}_i = \log(1 + \sum_{j=1}^{k-1}) e^{\theta_{ij}}, \qquad i = \overline{1, n}$$

$$(3.2.8)$$

and obtain a problem of unconstrained maximization

$$\mathcal{L}(\theta) \longrightarrow \max$$
 (3.2.9)

The problem 3.2.9 can be solved via natural conjugate gradient method. Namely, given an initial value $\theta^{(0)}$, one iteratively finds optimal value of θ as follows

$$\theta^{(t+1)} = \theta^{(t)} + \lambda^{(t)} d^{(t)}, \quad t = 0, 1, 2, \dots$$

Here $d^{(t)}$ is so called natural conjugate gradient. $d^{(t)}$ can be found according to formulas

$$d^{(t)} = \begin{cases} g^{(t)}, & t = 0\\ g^{(t)} + \frac{\|g^{(t)}\|_{\theta}^{2}}{\|g^{(t-1)}\|_{\theta}^{2}} d^{(t-1)}, & t > 0, \end{cases}$$
(3.2.10)

where $g^{(t)}$ is a natural gradient of $\mathcal{L}(\theta)$. $\|\cdot\|$ stands for the norm with respect to Riemannian metrics

$$G(\theta) = \operatorname{diag} (\mathcal{I}(\theta_1), \dots, \mathcal{I}(\theta_N)),$$

where $\theta_i = (\theta_{i1}, \dots, \theta_{ik})^T$, $i = \overline{1, n}$ and (θ_i) is a Fischer information. This method is nothing else but a conjugate gradient method in a Riemannian space with metrics $G(\theta)$. Given θ , g and $||g||_{\theta}$ can be found as follows

$$g = \nabla_{\tilde{\Pi}} \mathcal{L}(\tilde{\Pi}) = \begin{pmatrix} (I, -1) \nabla_{\tilde{\pi}_{1}} \mathcal{L}(\tilde{\Pi}) \\ \vdots \\ (I, -1) \nabla_{\tilde{\pi}_{n}} \mathcal{L}(\tilde{\Pi}) \end{pmatrix}$$

$$||g||_{\theta}^{2} = \sum_{i=1}^{n} \left(\nabla_{\tilde{\pi}_{1}} \mathcal{L}(\tilde{\Pi}) \right)^{T} \left(\operatorname{diag}(\tilde{\pi}_{i}) - \tilde{\pi}_{i}^{T} \tilde{\pi}_{i} \right) \left(\nabla_{\tilde{\pi}_{1}} \mathcal{L}(\tilde{\Pi}) \right),$$

$$(3.2.11)$$

where $\tilde{\Pi} = \tilde{\Pi}(\theta)$ and

$$\frac{\partial \mathcal{L}(\tilde{\Pi})}{\partial \pi_{ij}} = \sum_{l \neq j} \left(A_i \cdot \tilde{\Pi}_{\cdot l} \log \varepsilon + \overline{A}_i \cdot \tilde{\Pi}_{\cdot j} \log(1 - \varepsilon) \right) + \left(A_i \cdot \tilde{\Pi}_{\cdot j}, \overline{A}_i \cdot \tilde{\Pi}_{\cdot j} \right) \nabla \log \mathcal{B}(\tilde{\beta}_j) + \psi(\tilde{\alpha}_j) - \log \tilde{\pi}_{ij} - 1, \quad (3.2.12)$$

where $\psi(\cdot)$ is digamma function.

Algorithm 1: Natural Conjugate Gradient

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Input: adjacency matrix A, maximum number of clusters k, tolerance \eta, maximum
               number of iterations t_{\text{max}}
Output: array of predicted labels Z
Initialize \theta;
\mathcal{L}_{old} \leftarrow -\infty;
\lambda \leftarrow 1;
for t in range (t_{\text{max}}) do
      Calculate \tilde{\Pi} using 3.2.8;
       Calculate \tilde{\alpha}, \tilde{\beta} using 3.2.5;
       Calculate \mathcal{L} using 3.2.6;
      if 0 \le \frac{\mathcal{L} - \mathcal{L}_{old}}{|\mathcal{L}|} < \eta then
        break
      if \eta > 0 then
             Update d using 3.2.10, 3.2.11, 3.2.12;
             \theta \leftarrow \theta_{old} + \lambda d;
             \mathcal{L}_{old} \leftarrow \mathcal{L};
       else
          \begin{array}{l} \lambda \leftarrow \frac{\lambda}{2} \\ \theta \leftarrow \theta_{old} + \lambda |\eta| d \end{array}
end
Z = \left(\arg\max_{1 \le j \le k} \tilde{\pi}_{1j}, \dots, \arg\max_{1 \le j \le k} \tilde{\pi}_{nj}\right)
return Z
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Initial value of θ was generated from a standart normal distribution $\mathcal{N}(0,1)$. Probability of occurrence an inter-cluster edge ε was set to 10^{-10} . The maximal number of iterations and relative tolerance were taken equal to 100 and 10^{-6} respectively. Examples of performance of the natural conjugate gradient method can be found, for example, in [1].

- 3.3 Modularity-based methods
- 3.4 Spectral method
- 4 Data
- 5 Experimental results
- 6 Work split
- 7 References

References

[1] A fast inference algorithm for stochastic blockmodel, Zhiqiang Xu, Yiping Ke, Yi Wang, 2014.