User's Guide

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1. Overview

This MATLAB toolbox "envelope" implements six models in the envelope family, including envelope model, heteroscedastic envelope model, inner envelope model, partial envelope model, scaled envelope model and envelope model for the reduction on X.

1.1. Technical Support

We provides a support website, on which the users can download the package and check recent updates http://code.google.com/p/envlp/ . For further help, the users can contact the authors of the toolbox dennis@stat.umn.edu , suzhihua@stat.umn.edu , and yiyang@umn.edu .

1.2. About this Toolbox

The toolbox "envelope" contains the following utilities:

- Envelope Module "env".
- Heteroscedastic Envelope Module "henv".
- Inner Envelope Module "ienv".
- Partial Envelope Module "penv".
- Scaled Envelope Module "senv".
- Envelope Model in the Predictor Space Module "xenv".
- Examples on Using the Toolbox.
- Data Published in Papers on Envelope Models.

The applicability of this toolbox is described in the following table:

Model	Model Selection Tools	Inference Tools	Section	
Envelope Model	AIC	Estimation and Prediction	3.2	
	BIC	Bootstrap for Estimating Standard Errors		
	LRT	Hypothesis Test on Coefficients		
Heteroscedastic Envelope Model	AIC	Estimation and Prediction	3.3	
	BIC	Bootstrap for Estimating Standard Errors		
	LRT	Hypothesis Test on Coefficients		
Inner Envelope Model	AIC	Estimation and Prediction	3.4	
	BIC	Bootstrap for Estimating Standard Errors		
	LRT	Hypothesis Test on Coefficients		
Partial Envelope Model	AIC	Estimation and Prediction	3.5	
	BIC	Bootstrap for Estimating Standard Errors		
	LRT	Hypothesis Test on Coefficients		
Scaled Envelope Model	AIC	Estimation and Prediction	3.6	
	BIC	Bootstrap for Estimating Standard Errors		
		Hypothesis Test on Coefficients		
Envelope Model in the Predictor Space	AIC	Estimation and Prediction		
	BIC	Bootstrap for Estimating Standard Errors	3.7	
	LRT	Hypothesis Test on Coefficients		

Table 1: Applicability of toolbox "envelope"

1.3. Document Organization

This user's guide is organized as follows:

- **Chapter 1** Overview. Introduces the content and applicability of the toolbox.
- **Chapter 2** Quick start. Describes the installation of the toolbox, and provides simple examples on using the main functions.
- **Chapter 3** Toolbox structure. Discusses the structure of the toolbox, and demonstrates the applicability with more detailed examples.
- **Chapter 4** Optional arguments. Explains how to control the convergence speed, input user-specified starting values, and display iteration process.

2. Quick Start

2.1. Installation

To install the toolbox, direct your MATLAB working directory to the folder "envlp", and type

If the users agree with our license statements, the installation is completed, and all the utilities of the toolbox are added to the MATLAB path. Once installed, no further action is needed to call the functions in the toolbox, even if the users change a working directory, or MATLAB is relaunched.

2.2. A Guided Tour

This "envelope" toolbox consists tools in the following three classes:

- 1. Model selection: Select the dimension of the envelope subspace.
- 2. Model Fitting: Fit the model.
- 3. Post processing: Inference based on the model fitting.

We will present a tour of our toolbox through all three classes.

Model Selection

The model selection tools available in this toolbox are Akaike information criterion (AIC), Bayesian information criterion (BIC) and Likelihood ratio testing (LRT). Each model selection tool can be applied to all models, except that LRT cannot be applied to scaled envelope model, as indicated in Table 1. AIC and BIC generally require longer computing time than LRT, because of the nature of the method. The syntax of AIC, BIC and LRT are

```
u = modelselectaic(X, Y, modelType)
u = modelselectbic(X, Y, modelType)
u = modelselectlrt(X, Y, alpha, modelType)
```

For the inputs: X is a matrix containing the predictors, except that with "penv", it is a list having X.X1 and X.X2; Y is a matrix containing the response; 'alpha' is the significance level; and 'modelType' is a string indicting the model and its choices can be "env", "henv", "ienv", "penv", "senv" and "xenv". The output is the dimension of the envelope subspace selected by the tool. Examples will be provided in Section 2.3.

Model Fitting

The functions to fit the six models in the envelope family are the main drive of this toolbox. Their syntax are

```
ModelOutput = env(X, Y, u)
ModelOutput = henv(X, Y, u)
ModelOutput = ienv(X, Y, u)
ModelOutput = penv(X, Y, u)
ModelOutput = senv(X, Y, u)
ModelOutput = xenv(X, Y, u)
```

The input X and Y are the same as the input in the model selection tools, u is the dimension of the envelope subspace, the user can input any dimension if it is known, and use model selection tools described in model selection part to choose u. The output ModelOutput is a list, containing the maximum likelihood estimators (MLE) under the model, as well as some keys statistics for inference, including the standard error ratios of elements in the regression coefficients for the standard model versus the envelope model, maximized value of the log likelihood, number of parameters, etc.

Post processing

This toolbox provides functions for the following inference:

- Compute bootstrap standard errors for the regression coefficients $\widehat{\beta}$, which gives an estimator for the actual standard errors of $\widehat{\beta}$.
- For a given data point, report its fitted value or predicted value, with associated standard errors.
- Test hypothesis

$$H_0: \mathbf{L}\widehat{\boldsymbol{\beta}}\mathbf{R} = \mathbf{A}$$

 $H_\alpha: \mathbf{L}\widehat{\boldsymbol{\beta}}\mathbf{R} \neq \mathbf{A}$

where L, R and A are given matrices.

The syntax of functions for the inferences are

```
bootse = bootstrapse(X, Y, u, B, modelType)
PredictOutput = prediction(ModelOutput, Xnew, infType, modelType)
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

respectively. For the inputs, modelType in all three functions is the same as in model section part; ModelOutput in prediction and testcoefficient is a list returned from the functions for model fitting; X, Y, and u are discussed in the model selection part; B is the number of bootstrap sample; Xnew is a value of X with which to estimate or predict Y; infType is a string of characters indicting the inference type, the choices can be "estimation" or "prediction"; and TestInput is a list that specifies the null hypothesis, including L, R, and A. TestInput is an option input argument, is missing, testcoefficient will perform the usual F test, i.e. testing if $\beta = 0$. The output of bootstrapse is a matrix the same size of $\hat{\beta}$, with each elements as the bootstrap standard error of its corresponding element in $\hat{\beta}$. The output of prediction is a list containing the fitted or predicted value, its covariance matrix and standard errors. The function testcoefficient will print out a form displaying the test statistic, degrees of freedom of the reference distribution and the p-value, all of which are also summarized in the list TestOutput.

2.3. Two examples

To demonstrate their usage, we take the envelope model and partial envelope model as examples. Other methods can be applied similarly, and are also discussed in details in Section 3.

Wheat Protein Data

For the envelope model, the data we use is the wheatprotein data, which is used as a data analysis example in Cook et al. [2010]. First we load the data and assign column 1 to 6 as the response Y and column 8 as the predictor X. The description of the data is available under the folder /envlp/data.

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1:6);
```

Then we apply the model selection tools and get the following results.

```
modelType = 'env';
u = modelselectaic(X, Y, modelType)
u =
    1
u = modelselectbic(X, Y, modelType)
u =
    1
```

```
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, modelType)
u =
1
```

In this example, all three model selection tools agree that the dimension of the envelope should be 1, which is consistent with the results in Cook et al. [2010]. Then we fit the envelope model with dimension 1.

```
ModelOutput = env(X, Y, u)
ModelOutput =
         beta: [6x1 double]
        Sigma: [6x6 double]
        Gamma: [6x1 double]
       Gamma0: [6x5 double]
          eta: 8.5647
        Omega: 7.8762
       Omega0: [5x5 double]
        alpha: [6x1 double]
            1: -850.7592
    covMatrix: [6x6 double]
       asyEnv: [6x1 double]
        ratio: [6x1 double]
           np: 28
            n: 50
```

We notice that ModelOutput is a list that includes the MLEs and the statistics relevant to the influence of the envelope model. Fore more details of the components in ModelOutput, please refer to the Reference of the toolbox. If we would want to see the regression coefficients, as well as the standard error ratios, we can type

```
ModelOutput.beta

ans =

-1.0644
4.4730
3.6839
-5.9770
0.6013
-1.5986

ModelOutput.ratio
ans =

28.0945
18.4326
23.6384
16.3211
```

```
65.8245
6.4668
```

and get the results. We can also look at the eigenvalues of Σ_1 and Σ_2 by the following commands

```
ModelOutput.Omega

ans =
7.8762
eig(ModelOutput.Omega0)
ans =
1.0e+03 *
6.5166
0.2083
0.0201
0.0004
0.0003
```

These results are consistent with those published in Cook et al. [2010]. With the model fitting results, if we would like to check the fitted value and its standard errors for the second observation, we can use the prediction function.

We can get the standard errors of the fitted value, and also we can compare the fitted value with its true value

```
PredictOutput.SE

ans =

4.8892
4.0227
4.3237
4.7470
6.8186
2.6948
[PredictOutput.value, Y(2,:)']

ans =

474.7135 458.0000
127.4740 112.0000
251.2044 236.0000
380.8280 368.0000
```

```
380.9473 383.0000
-6.3287 -15.0000
```

Now suppose we want to test if $\beta = 0$, we run the testcoefficient as

```
TestOutput = testcoefficient(ModelOutput, modelType)

Test Hypothesis Chisq Statistic DF P-value

L * beta * R = A 116.230 6 0.0000
```

Notice that we did not have TestInput, because we are performing the usual F test on if $\beta = 0$. The test results indicate that we have strong evidence to reject $\beta = 0$.

Fiber and Paper Data

Now we look at another example for the partial envelope model. The data we load is the fiber and paper data used in Su and Cook [2011].

```
load fiberpaper.dat
Y = fiberpaper(:, 1 : 4);
X.X1 = fiberpaper(:, 7);
X.X2 = fiberpaper(:, 5 : 6);
```

We notice that for partial envelope model, X is a list, with X.X1 containing the main predictors and X.X2 containing covariates. The model selection process is parallel to that for the envelope model:

In this example, AIC picks the dimension 3, while BIC and LRT both gives dimension as 1. In Su and Cook [2011], u=1 is used for model fitting and inference. So we fit the partial envelope model with u=1.

```
u = 1;
ModelOutput = penv(X, Y, u)
ModelOutput =
        betal: [4x1 double]
        beta2: [4x2 double]
        alpha: [4x1 double]
        Gamma: [4x1 double]
          eta: 0.0047
        Omega: 0.0149
       Omega0: [3x3 double]
        Sigma: [4x4 double]
            1: -35.6323
           np: 23
    covMatrix: [12x12 double]
      asyPenv: [4x1 double]
        ratio: [4x1 double]
            n: 62
```

Again we get the output as a list, which contains the MLE for the parameters and statistics that are relevant to the inference for the partial envelope model. Detailed description of the elements in the list is in the Reference of the toolbox. To look at the standard error ratios, and the eigenvalues of Σ_1 and Σ_2 , we key in

```
ModelOutput.ratio

ans =
66.0742
6.9326
10.5048
9.6279
ModelOutput.Omega

ans =
0.0149
eig(ModelOutput.Omega0)
ans =
4.9819
0.0999
0.0050
```

The results are almost the same the those in Su and Cook [2011]. They are slightly different, because the toolbox used a different starting value algorithm, which is less likely to be trapped in local minimum and thus leads to a more reliable result. Next we look at the standard errors for elements in $\widehat{\beta}$. Since ModelOutput.asyPenv contains the asymptotic standard errors, for actual standard errors, we need to divide the asymptotic standard errors by \sqrt{n} , where n is the sample size and is returned in ModelOutput.n.

```
ModelOutput.asyPenv / sqrt(ModelOutput.n)
ans =
    1.0e-03 *
    0.3181
    0.8704
    0.9719
    0.4689
```

The standard errors above are computed assuming the partial envelope model. They can also be estimated using bootstrap the residuals.

The uses may get a slightly different results from the above, as a different random seed is used.

3. Toolbox Structure

3.1. Multivariate Linear Regression

The envelope model is originally developed under the framework of multivariate linear regression, and its performance is frequently compared to the standard multivariate linear regression model. In this package, we offer some simple functions on that makes the comparison easier.

A standard multivariate linear regression model is formulated as

$$\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\beta} \mathbf{X} + \boldsymbol{\varepsilon},$$

where $\mathbf{Y} \in \mathbb{R}^r$ is the multivariate response, $\mathbf{X} \in \mathbb{R}^p$ is non-stochastic predictor, and $\boldsymbol{\varepsilon} \in \mathbb{R}^r$ follows a distribution with mean 0, and covariance matrix $\mathbf{\Sigma} \in \mathbb{R}^{r \times r}$, $\boldsymbol{\alpha} \in \mathbb{R}^r$ and $\boldsymbol{\beta} \in \mathbb{R}^{r \times p}$ are the unknown intercept and coefficients. The goal of the envelope model is to reduce the standard errors in estimating $\boldsymbol{\beta}$.

This function to fit this model is

```
ModelOutput = fit_OLS(X, Y)
```

where the input X is an $n \times p$ matrix, with the ith row as the transpose of the ith observation of X, and Y is an $n \times r$ matrix with the ith row as the transpose of the ith observation of Y. The output ModelOutput is a list, which contains the ordinary least squares (OLS) estimators of β , Σ and α .

The standard errors of the OLS estimator $\hat{\beta}$ is often compared to those of the envelope model, the function

```
bootse = bootstrapse OLS(X, Y, B)
```

computes the standard errors of $\widehat{\beta}$ estimated by bootstrapping the residuals. The input B is the number of bootstrap samples, and bootse returns a matrix with the same size as β , with each element as the standard error of the corresponding element in β .

3.2. Envelope Model

The envelope model is a general method in the envelope family to reduce the standard errors in estimating β . It can be applied when the number of the responses is strictly greater than the number of the predictors, and the responses are continuous variables. It has a potential to obtain efficiency gains in estimating β , compared to the OLS estimators.

In the multivariate linear regression context in Section 3.1, a coordinate form of the envelope model is

$$\mathbf{Y} = \boldsymbol{lpha} + \mathbf{\Gamma} \boldsymbol{\eta} \mathbf{X} + \boldsymbol{arepsilon}, \quad \mathbf{\Sigma} = \mathbf{\Gamma} \mathbf{\Omega} \mathbf{\Gamma}^T + \mathbf{\Gamma}_0 \mathbf{\Omega}_0 \mathbf{\Gamma}_0^T,$$

where the regression coefficients $\boldsymbol{\beta} = \boldsymbol{\Gamma} \boldsymbol{\eta}, \boldsymbol{\mathcal{B}} = \operatorname{span}(\boldsymbol{\beta}), \boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans $\mathcal{E}_{\boldsymbol{\Sigma}}(\boldsymbol{\mathcal{B}})$ – the envelope subspace, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\boldsymbol{\Sigma}}(\boldsymbol{\mathcal{B}}), \boldsymbol{\eta} \in \mathbb{R}^{u \times p}, \boldsymbol{\Omega} \in \mathbb{R}^{u \times u}, \boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ are coordinates, and u is the dimension of $\mathcal{E}_{\boldsymbol{\Sigma}}(\boldsymbol{\mathcal{B}})$.

To select the dimension of the envelope subspace u, we can use the following functions

```
u = modelselectaic(X, Y, 'env')
u = modelselectbic(X, Y, 'env')
alpha = 0.01; # Users can specify other significance level
u = modelselectlrt(X, Y, alpha, 'env')
```

The possible values of u can be any integer from 0 to r. When u = r, the envelope model is essentially the standard multivariate linear model. And when u = 0, it means that $\beta = 0$, then the changes in \mathbf{Y} does not depend on \mathbf{X} . After obtaining u, we can fit the envelope model by

```
ModelOutput = env(X, Y, u)
```

The output ModelOutput is a list, which contains the MLE of β , Σ , Γ , Γ_0 , η , Ω , Ω_0 and α , and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\beta)$, the asymptotic standard errors of β , the ratios of the asymptotic standard errors of the standard model versus the envelope model for elements in β , the number of parameters in the model, and the number of observations in the data. After fitting the data and get ModelOutput, we can perform post processing inference as computing bootstrap standard errors of $\hat{\beta}$ by

```
bootse = bootstrapse(X, Y, u, B, 'env')
```

or computing the fitted value or predicted value and associated standard errors given a new \mathbf{X} by

```
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'env')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'env')
```

or testing if a linear combination of β is equal to a particular matrix, i.e. given L, R, and A, testing if $L\beta R = A$,

```
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

The inputs and outputs of these post processing functions are discussed in details in Section 2.2.

3.3. Heteroscedastic Envelope Model

The heteroscedastic envelope model [Su and Cook, a] is used when the data has none constant error structure, and it is developed under the framework of estimating multivariate means for different populations. To use the heteroscedastic envelope model, the number of response should be greater than or equal to the number of populations.

The standard model for estimating multivariate means for p populations can be formulated as $\mathbf{Y}_{(i)j} = \boldsymbol{\mu} + \boldsymbol{\beta}_{(i)} + \boldsymbol{\varepsilon}_{(i)j}, \quad i = 1, \cdots, p, \quad j = 1, \cdots, n_{(i)}.$ We use the subscript (i) to represent the ith group, and we j without the parenthesis to represent the jth observation. Then $\mathbf{Y}_{(i)j} \in \mathbb{R}^r$ is the jth observation in the ith group, $n_{(i)}$ is the number of observations in the ith group, $\boldsymbol{\mu} \in \mathbb{R}^r$ is the grand mean, $\boldsymbol{\beta}_{(i)} \in \mathbb{R}^r$ is the main effect for the ith group and satisfies $\sum_{i=1}^p n_{(i)} \boldsymbol{\beta}_{(i)} = 0$, the errors $\boldsymbol{\varepsilon}_{(i)j}$ follows a distribution with mean 0, and covariance matrix $\boldsymbol{\Sigma}_{(i)} \in \mathbb{R}^{r \times r}$. With this formulation, let $\boldsymbol{\mathcal{B}} = \operatorname{span}(\boldsymbol{\beta}_{(1)} \cdots, \boldsymbol{\beta}_{(p)})$, and $\boldsymbol{\mathcal{M}} = \{\boldsymbol{\Sigma}_{(i)}: i = 1, \cdots, p\}$, the coordinate form of the heteroscedastic envelope model is displayed as follows:

$$\mathbf{Y}_{(i)j} = \boldsymbol{\mu} + \boldsymbol{\Gamma} \boldsymbol{\eta}_{(i)} + \boldsymbol{\varepsilon}_{(i)j}, \quad \boldsymbol{\Sigma}_{(i)} = \boldsymbol{\Gamma} \boldsymbol{\Omega}_{1(i)} \boldsymbol{\Gamma}^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T, \quad i = 1, \dots, p, \ \ j = 1, \dots, n_{(i)},$$

where $\boldsymbol{\beta}_{(i)} = \boldsymbol{\Gamma} \boldsymbol{\eta}_{(i)}$, $\boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans $\mathcal{E}_{\mathcal{M}}(\mathcal{B})$ – the \mathcal{M} envelope of \mathcal{B} , $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\mathcal{M}}(\mathcal{B})$, $\boldsymbol{\eta}_{(i)} \in \mathbb{R}^{u \times 1}$, $\boldsymbol{\Omega}_{1(i)} \in \mathbb{R}^{u \times u}$, $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$, $i = 1, \dots, p$, are coordinates, $\sum_{i=1}^p n_{(i)} \boldsymbol{\eta}_{(i)} = 0$, and u is the dimension of $\mathcal{E}_{\mathcal{M}}(\mathcal{B})$.

To use the heteroscedastic envelope model, first we check if the data has heteroscedastic error structure by the Box's M test [Johnson and Wichern, 2007]. Take the water strider example [Su and Cook, a], the following codes performs the Box's M test.

```
load waterstrider.mat
alpha = 0.01;  # Users can specify other significance level
TestOutput = mtest(X, Y, alpha);
```

The input X is a group indicator, there is no constraint on the number of columns of X, as long as X takes p unique values for p populations. For example, if there are three groups, X can take three unique values as 1, 2, and 3 to indicate the groups, or X can take (0, 1), (1, 0), and (0, 0) to indicate the groups. The output is displayed below

MBox	Chi-sqr.	df	P
157.5977	137.3361	72	0.0000

Covariance matrices are significantly different.

As the data has none constant covariance structure, we can apply the heteroscedastic envelope model. To select the dimension of the envelope subspace, we use

```
u = modelselectaic(X, Y, 'henv')
u = modelselectbic(X, Y, 'henv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'henv')
```

The input X can have the same form as in the function mtest. After obtaining u, the heteroscedastic envelope model can be fit by

```
ModelOutput = henv(X, Y, u)
```

The output ModelOutput is a list containing the MLE of the grand mean μ , the group mean $\mu + \beta_{(i)}$, the fitted values $\widehat{\mathbf{Y}}, \Gamma, \Gamma_0, \beta_{(i)}$, the unique values in $\mathbf{X}, \Sigma_{(i)}, \eta_{(i)}, \Omega_{(i)}$, and Ω_0 , and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $(\mu^T, \beta_{(1)}^T, \cdots, \beta_{(p)}^T)^T$, the asymptotic standard errors of $\beta_{(i)}$, the ratios of the asymptotic standard errors of the standard model versus the inner envelope model for elements in $\beta_{(i)}$, the number of parameters in the model, and the number of observations in the data. After model fitting, the functions for post processing inference are

```
bootse = bootstrapse(X, Y, u, B, 'henv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'henv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'henv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

Their inputs and outputs are similar to those for the envelope model.

3.4. Inner Envelope Model

Like the envelope model, the inner envelope model Su and Cook is also a general method to reduce the standard errors in estimating β . It has a different mechanism to achieve the efficiency gains than the envelope model, so it may be very useful when the envelope model reduces to the standard model and offers no efficiency gains. The inner envelope model also requires that the number of predictors is strictly less than the number of responses.

In the multivariate linear regression context (Section 3.1), the coordinate form of the inner envelope model can be written as

$$\mathbf{Y} = \boldsymbol{\alpha} + (\boldsymbol{\Gamma}_1 \boldsymbol{\eta}_1^T + \boldsymbol{\Gamma}_0 \mathbf{B} \boldsymbol{\eta}_2^T) \mathbf{X} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \boldsymbol{\Gamma}_1 \boldsymbol{\Omega}_1 \boldsymbol{\Gamma}_1^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T,$$

where $\boldsymbol{\beta} = \boldsymbol{\Gamma}_1 \boldsymbol{\eta}_1^T + \boldsymbol{\Gamma}_0 \mathbf{B} \boldsymbol{\eta}_2^T \in \mathbb{R}^{r \times p}$, $\boldsymbol{\mathcal{B}} = \operatorname{span}(\boldsymbol{\beta})$, $\boldsymbol{\Gamma}_1 \in \mathbb{R}^{r \times u}$ spans the inner envelope subspace $\mathcal{IE}_{\boldsymbol{\Sigma}}(\boldsymbol{\mathcal{B}})$, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{IE}_{\boldsymbol{\Sigma}}(\boldsymbol{\mathcal{B}})$, $\mathbf{B} \in \mathbb{R}^{(r-u) \times (p-u)}$ is a semi-orthogonal matrix such that $(\boldsymbol{\Gamma}_1, \ \boldsymbol{\Gamma}_0 \mathbf{B})$ spans $\boldsymbol{\mathcal{B}}, \ \boldsymbol{\eta}_1 \in \mathbb{R}^{p \times u}, \ \boldsymbol{\eta}_2 \in \mathbb{R}^{p \times (p-u)}$, $\boldsymbol{\Omega}_1 \in \mathbb{R}^{u \times u}$, and $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ contains the coordinates, and u is the dimension of $\mathcal{IE}_{\boldsymbol{\Sigma}}(\boldsymbol{\mathcal{B}})$.

To select the dimension of $\mathcal{IE}_{\Sigma}(\mathcal{B})$, we use the same functions as for the envelope model, except that the modelType is 'ienv'.

```
u = modelselectaic(X, Y, 'ienv')
u = modelselectbic(X, Y, 'ienv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'ienv')
```

The possible values of u are integers from 0 to p. When u=0, the inner envelope model reduces to the standard model, and offers no efficiency gains. When u=p, the inner envelope model is equivalent to an envelope model with dimension p.

Given the dimension u, the inner envelope model can be fitted by

```
ModelOutput = ienv(X, Y, u)
```

The output ModelOutput is a list containing the MLE of β , Σ , Γ_1 , Γ_0 , η_1 , B, η_2 , Ω_1 , Ω_0 and α , and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\beta)$, the asymptotic standard errors of β , the ratios of the asymptotic standard errors of the standard model versus the inner envelope model for elements in β , the number of parameters in the model, and the number of observations in the data. After model fitting, the functions for post processing inference are

```
bootse = bootstrapse(X, Y, u, B, 'ienv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'ienv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'ienv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

Their inputs and outputs are similar to those for the envelope model.

3.5. Partial Envelope Model

The partial envelope model [Su and Cook, 2011] can be applied when part of the predictors are of main interest. It only requires that the number of the main predictors is strictly less than the number of the responses. This is particular useful when the number of predictors p is large, but there are only a small number of predictors are of main interest. Suppose that $\mathbf{X}^T = (\mathbf{X}_1^T, \mathbf{X}_2^T)$, where $\mathbf{X}_1 \in \mathbb{R}^{p_1}$ are predictors of main interest, and $\mathbf{X}_2 \in \mathbb{R}^{p_2}$ are covariates, $p_1 + p_2 = p$. Then the standard model is formulated as $\mathbf{Y} = \alpha + \beta_1 \mathbf{X}_1 + \beta_2 \mathbf{X}_2 + \varepsilon$, and the coordinate form of the partial envelope model is

$$\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\Gamma} \boldsymbol{\eta} \mathbf{X}_1 + \boldsymbol{\beta}_2 \mathbf{X}_2 + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \boldsymbol{\Gamma} \boldsymbol{\Omega} \boldsymbol{\Gamma}^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T,$$

where $\boldsymbol{\beta}_1 = \boldsymbol{\Gamma} \boldsymbol{\eta} \in \mathbb{R}^{r \times p_1}$, $\boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans the partial envelope $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B}_1)$ – the partial envelope subspace, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times u}$ spans the orthogonal complement of $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B}_1)$, $\boldsymbol{\eta} \in \mathbb{R}^{u \times p_1}$, $\boldsymbol{\Omega} \in \mathbb{R}^{u \times u}$, $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ are coordinates, $\boldsymbol{\beta}_2 \in \mathbb{R}^{r \times p_2}$ are the coefficients for \mathbf{X}_2 and u is the dimension of $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B}_1)$.

To use the functions related to partial envelope model, the input X is a list, which has two elements X.X1 and X.X2. The element X.X1 is an $n \times p_1$ matrix, with the ith row as the transpose of the ith observation of \mathbf{X}_1 , and X.X2 is an $n \times p_2$ matrix, with the ith row as the transpose of the ith observation of \mathbf{X}_2 . The form of X as a list is unique to the partial envelope model.

The following functions can be applied to select u,

```
u = modelselectaic(X, Y, 'penv')
u = modelselectbic(X, Y, 'penv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'penv')
```

The possible values of u are any integer from 0 to r, when u = r, the partial envelope model reduces the standard model, and when u = 0, the changes in \mathbf{Y} does not depend on the changes in \mathbf{X}_1 .

After obtaining u, the partial envelope model can be fit by

```
ModelOutput = penv(X, Y, u)
```

The output ModelOutput is a list containing the MLE of $\beta_1, \beta_2, \Sigma, \Gamma, \Gamma_0, \eta, \Omega, \Omega_0$ and α , and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $(\operatorname{vec}(\beta_2)^T, \operatorname{vec}(\beta_1)^T)^T$, the asymptotic standard errors of β_1 , the ratios of the asymptotic standard errors of the standard model versus the partial envelope model for elements in β_1 , the number of parameters in the model, and the number of observations in the data. The functions for post processing inference are

```
bootse = bootstrapse(X, Y, u, B, 'penv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'penv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'penv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

The usage of functions are introduced in Section 2.2, and are similar to those for the envelope models as discussed in Section 3.2. The only difference is that in the function prediction, the input Xnew is a list, which has Xnew.X1 and Xnew.X2. Xnew.X1 is a p_1 dimensional column vector containing the new value of \mathbf{X}_1 , and Xnew.X2 is a p_2 dimensional column vector containing the new value of \mathbf{X}_2 .

3.6. Scaled Envelope Model

Scaled envelope model (Cook and Su) is used when the user hopes to have a scale-invariant version of the envelope model. It is more efficient or at least as efficient as the standard model.

The scaled envelope model normally takes longer to compute because of the nature of its iterative computing algorithm. But users can print out the model fitting process, which will be discussed in Chapter 4.

In the multivariate linear regression context (3.1), the coordinate form of the scaled envelope model can be written as

$$\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\Lambda} \boldsymbol{\Gamma} \boldsymbol{\eta} \mathbf{X} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \boldsymbol{\Lambda} \boldsymbol{\Gamma} \boldsymbol{\Omega} \boldsymbol{\Gamma}^T \boldsymbol{\Lambda} + \boldsymbol{\Lambda} \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T \boldsymbol{\Lambda},$$

where $\beta = \Lambda \Gamma \eta \in \mathbb{R}^{r \times p}$, $\Lambda \in \mathbb{R}^{r \times r}$ is the scaling matrix, it is a diagonal matrix with the first element as 1, and the other diagonal elements as positive real numbers, $\Gamma \in \mathbb{R}^{r \times p}$ spans the envelope subspace $\mathcal{E}_{\Lambda^{-1}\Sigma\Lambda^{-1}}(\Lambda^{-1}\mathcal{B})$, $\Gamma_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\Lambda^{-1}\Sigma\Lambda^{-1}}(\Lambda^{-1}\mathcal{B})$, $\eta \in \mathbb{R}^{u \times p}$, $\Omega \in \mathbb{R}^{u \times u}$, and $\Omega_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ carry the coordinates, and u is the dimension of the envelope subspace $\mathcal{E}_{\Lambda^{-1}\Sigma\Lambda^{-1}}(\Lambda^{-1}\mathcal{B})$. If $\Lambda = \mathbf{I}_r$, then the scaled envelope model is equivalent to the envelope model.

To select the dimension of $\mathcal{E}_{\Lambda^{-1}\Sigma\Lambda^{-1}}(\Lambda^{-1}\mathcal{B})$, we can apply the following two functions

```
u = modelselectaic(X, Y, 'senv')
u = modelselectbic(X, Y, 'senv')
```

Notice that likelihood ratio testing can not be applied to the model selection of the scaled envelope model, only AIC and BIC can be used to select u. The possible values of u can be any integer from 0 to r. When u=r, the scaled envelope model is essentially the standard multivariate linear model. And when u=0, it means that $\beta=0$, then the changes in \mathbf{Y} does not depend on \mathbf{X} . After obtaining u, we can fit the scaled envelope model by

```
ModelOutput = senv(X, Y, u)
```

The output ModelOutput is a list, which contains the MLE of β , Σ , Λ , Γ , Γ_0 , η , Ω , Ω_0 and α , and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\beta)$, the asymptotic standard errors of β , the ratios of the asymptotic standard errors of the standard model versus the scaled envelope model for elements in β , the number of parameters in the model, and the number of observations in the data. After model fitting, the users can perform the following post processing inference:

```
bootse = bootstrapse(X, Y, u, B, 'senv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'senv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'senv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

These functions are used similarly as for the envelope model.

3.7. Envelope Model in the Predictor Space

The envelope model in the predictor space is used when the number of predictors is strictly larger than the number of the responses. It has the potential to have better prediction performance compared to the standard model, or even the partial least squares. In fact, in the population version, the envelope estimator in this case is equivalent to the partial least squares estimator, but in the sample version, its performance is normally more stable than the partial least squares estimator.

We slightly change the formulation of the standard model as $\mathbf{Y} = \boldsymbol{\mu} + \boldsymbol{\beta}^T \mathbf{X} + \boldsymbol{\varepsilon}$, to be consistent with the notations in Cook et al. [Submitted]. Then the coordinate form of envelope model in the predictor space is as follows:

$$\mathbf{Y} = \boldsymbol{\mu} + \boldsymbol{\eta}^T \mathbf{\Omega}^{-1} \mathbf{\Gamma}^T \mathbf{X} + \boldsymbol{\varepsilon}, \quad \mathbf{\Sigma}_{\mathbf{X}} = \mathbf{\Gamma} \mathbf{\Omega} \mathbf{\Gamma}^T + \mathbf{\Gamma}_0 \mathbf{\Omega}_0 \mathbf{\Gamma}_0^T,$$

where $\mu \in \mathbb{R}^r$ is the intercept, $\beta = \Gamma \Omega^{-1} \eta \in \mathbb{R}^{p \times r}$, $\Gamma \in \mathbb{R}^{r \times u}$ spans the envelope subspace $\mathcal{E}_{\Sigma_{\mathbf{X}}}(\mathcal{B})$, and $\mathcal{B} = \operatorname{span}(\beta^T)$, $\Gamma_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\Sigma_{\mathbf{X}}}(\mathcal{B})$, $\eta \in \mathbb{R}^{u \times r}$, $\Omega \in \mathbb{R}^{u \times u}$, and $\Omega_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ carry coordinates, and u is the dimension of the envelope $\mathcal{E}_{\Sigma_{\mathbf{X}}}(\mathcal{B})$.

To select the dimension of $\mathcal{E}_{\Sigma_{\mathbf{X}}}(\mathcal{B})$, we apply the following three functions

```
u = modelselectaic(X, Y, 'xenv')
u = modelselectbic(X, Y, 'xenv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'xenv')
```

The possible values for u can be any integer from 0 to p, when u = p, the envelope model reduces to the standard model, and when u = 0, the changes in \mathbf{Y} does not depend on \mathbf{X} . After estimating u, we can fit the model by

```
ModelOutput = xenv(X, Y, u)
```

The output ModelOutput is a list, which contains the MLE of β , $\Sigma_{\mathbf{X}}$, Γ , Γ_0 , η , Ω , Ω_0 and μ , and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\operatorname{vec}(\beta)$, the asymptotic standard errors of β , the ratios of the asymptotic standard errors of the standard model versus the envelope model for elements in β , the number of parameters in the model, and the number of observations in the data. After model fitting, the following post processing inference can be performed:

```
bootse = bootstrapse(X, Y, u, B, 'xenv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'xenv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'xenv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

These functions are used similarly as for the envelope model in Section 3.2.

4. Optional Arguments

The computing of the envelope models involve Grassmann manifold optimization. The package sg_min 2.4.1 by Ross Lippert (http://web.mit.edu/~ripper/www.sgmin.html) uses the analytical

first derivative and numerical second derivative of the objective function, and we find it very stable. The Grassmann manifold optimization in this package is then based on sg_min 2.4.1. Because of the iterative nature of the optimization, we offer some optional arguments so that the users can control the convergence tolerance and speed, as well as monitoring the iteration process. In functions modelselectaic, modelselectbic, modelselectlrt, env, henv, ienv, penv, senv, xenv and bootstrapse, there is an optional input argument Opts. Opts is a list, which has five elements:

1. Opts.maxIter: This controls the maximum number of iterations. The default value is 300. The iterations will terminate once the maximum number of iterations is reached, then the results are based on the last iteration. At the same time, a warning

WARNING: reached maximum number of iterations without convergence for specified tolerances

is printed out.

- 2. Opts.ftol: This controls the tolerance parameter of the objective function. The default value is 1e-10. The iteration will terminate once the tolerance condition for both the objective function and its derivative are reached.
- 3. Opts.gradtol: This controls the tolerance parameter of the derivative of the objective function. The default value is 1e-7.
- 4. Opts.verbose: This is a flag for printing out the computing process. It is logical 0 or 1, with 0 for no print out and 1 for print out. The default value is 0. The print out will depend on the nature of the functions: with functions for model selection (modelselectaic, modelselectbic and modelselectlrt), the current dimension will be printed out; with functions for bootstrap (bootstrapse), the current number of bootstrap sample will be printed out; with the scaled envelope model (senv), the current number of iteration for the alternating algorithm between the scales and Grassmann manifold optimization is printed out; and with all the other functions henv, ienv, penv, senv, and xenv, the current number of iterations of the Grassmann manifold optimization is printed out.
- 5. Opts.init: This argument allows the users to input their starting values. If not specified, the starting values are generated by the functions get_Init or get_Init4henv. This argument is only applicable to env, henv, ienv, penv, senv and xenv.

The users can choose to define some or none of the five elements, if not defined, the default value will be used. If the users choose to define some of the elements, it is added as the last input of the function. For example, suppose the users define Opts for env, then the syntax of env will be

```
ModelOutput = env(X, Y, u, Opts)
```

We will conclude this chapter by a couple of examples.

Control Convergence Arguments

We use the wheat protein data and the function env as an example. First we set

```
Opts.verbose = 1
```

so that we can monitor the iteration process. We leave the other arguments as default values for now

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1:6);
u = 1;
ModelOutput = env(X, Y, u, Opts);
iter grad F(Y)
0 3.109420e+00 -1.166907e+00
1 1.448779e-01 -1.193292e+00
2 2.566091e-01 -1.193338e+00
3 2.434851e-02 -1.193379e+00
4 1.035130e-02 -1.193385e+00
5 1.259475e-02 -1.193385e+00
6 9.240358e-03 -1.193385e+00
7 1.436949e-02 -1.193386e+00
8 1.096367e-02 -1.193389e+00
9 1.643753e-04 -1.193391e+00
10 2.813965e-04 -1.193391e+00
11 8.046247e-05 -1.193391e+00
12 4.414260e-07 -1.193391e+00
13 1.774275e-07 -1.193391e+00
```

We notice that it takes 13 iterations till convergence. Suppose we set the maximum number of iterations as 10, then the algorithm stops at the tenth iteration and a warning is printed out at the end.

```
Opts.maxIter = 10;

ModelOutput = env(X, Y, u, Opts);

iter grad F(Y)

0 3.109420e+00 -1.166907e+00

1 1.448779e-01 -1.193292e+00

2 2.566091e-01 -1.193338e+00

3 2.434851e-02 -1.193379e+00

4 1.035130e-02 -1.193385e+00

5 1.259475e-02 -1.193385e+00

6 9.240358e-03 -1.193385e+00

7 1.436949e-02 -1.193386e+00
```

```
9 1.643753e-04 -1.193391e+00
10 2.813965e-04 -1.193391e+00
WARNING: reached maximum number of iterations without convergence for specia
```

Normally Grassmann manifold optimization takes hundreds of iterations to converge at the default tolerance parameters. We use this example because that we can print its complete iteration process. Now we change the tolerance level of convergence.

```
Opts.ftol = 1e-5;

Opts.gradtol = 1e-3;

ModelOutput = env(X, Y, u, Opts);

iter grad F(Y)

0 3.109420e+00 -1.166907e+00

1 1.448779e-01 -1.193292e+00

2 2.566091e-01 -1.193338e+00

3 2.434851e-02 -1.193379e+00

4 1.035130e-02 -1.193385e+00

5 1.259475e-02 -1.193385e+00

6 9.240357e-03 -1.193385e+00

7 1.436947e-02 -1.193386e+00

8 1.096356e-02 -1.193389e+00

9 1.573313e-04 -1.193391e+00
```

8 1.096367e-02 -1.193389e+00

As we loosen the tolerance level, it requires less number of iterations till convergence. If the users want to specify their starting value, it can be done by

```
Opts.init = [1 0 0 0 0 0]';

ModelOutput = env(X, Y, u, Opts);

iter grad F(Y)

0 3.510184e+00 7.406482e+00

1 1.005063e+01 5.523686e+00

2 2.643888e+00 4.138687e+00

...

74 1.465095e-02 -1.193389e+00

75 1.066872e-04 -1.193391e+00

76 7.018152e-07 -1.193391e+00
```

From the result, we notice that a random starting value normally takes longer to convergence. Furthermore, according to our experience, it is more likely to be trapped in the local minimums, and therefore, should be avoided.

The Display Argument

In the previous example, the print out for env is the Grassmann manifold optimization process. In this example, we will show different print out types for different functions. We will still use the wheat protein data as the background for the first two cases.

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1:6);
```

1. With functions on model selection, the print out is the current dimension:

```
Opts.verbose = 1;
u = modelselectaic(X, Y, 'env', Opts);
Current dimension 0
Current dimension 1
Current dimension 2
Current dimension 3
Current dimension 4
Current dimension 5
```

Suppose the number of responses is r, the print out will start with dimension 0 and end with dimension r-1. The case of u=r is computed but not printed, because it takes little time, compared to the Grasmann manifold optimization process.

2. With function bootstrapse, the print out is the current number of bootstrap sample.

```
u = 1;
B = 15;
Opts.verbose = 1;
bootse = bootstrapse(X, Y, u, B, 'env', Opts);
Current number of bootstrap sample 1
Current number of bootstrap sample 2
Current number of bootstrap sample 3
Current number of bootstrap sample 4
Current number of bootstrap sample 5
Current number of bootstrap sample 6
Current number of bootstrap sample 7
Current number of bootstrap sample 8
Current number of bootstrap sample 9
Current number of bootstrap sample 10
Current number of bootstrap sample 11
Current number of bootstrap sample 12
Current number of bootstrap sample 13
Current number of bootstrap sample 14
Current number of bootstrap sample 15
```

The print out will start from the first bootstrap sample till the last bootstrap sample.

3. With function senv, the print out is the alternating algorithm between the optimization of the scaling parameters and Grassmann manifold.

```
load('sales.txt')
Y = sales(:, 4 : 7);
X = sales(:, 1 : 3);
u = 1;
Opts.verbose = 1;
ModelOutput = senv(X, Y, u, Opts);
Current number of iterations 1
Current number of iterations 2
Current number of iterations 3
Current number of iterations 4
Current number of iterations 5
Current number of iterations 6
Current number of iterations 7
Current number of iterations 8
Current number of iterations 9
Current number of iterations 10
Current number of iterations 11
Current number of iterations 12
Current number of iterations 13
Current number of iterations 14
Current number of iterations 15
Current number of iterations 16
Current number of iterations 17
Current number of iterations 18
Current number of iterations 19
Current number of iterations 20
Current number of iterations 21
Current number of iterations 22
Current number of iterations 23
Current number of iterations 24
Current number of iterations 25
Current number of iterations 26
```

The maximum number of iteration of the alternating algorithm is 1000. So the print out can run from the first iteration to the thousandth iteration.

4. With other functions, the print out is the Grassmann manifold optimization process as in env.

```
load irisf.mat
```

```
d = 1;
Opts.verbose = 1;
ModelOutput = ienv(X, Y, d, Opts)
iter grad F(Y)
0 2.043243e-01 1.808874e-02
1 1.102662e-01 1.399590e-02
2 3.082806e-02 7.353045e-03
3 1.755330e-03 7.285612e-03
4 1.312911e-03 7.285253e-03
5 3.840448e-04 7.284642e-03
6 1.174194e-06 7.284631e-03
7 1.462019e-06 7.284631e-03
8 9.358198e-08 7.284631e-03
9 1.478274e-08 7.284631e-03
10 1.952908e-08 7.284631e-03
```

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