

MATLAB Toolbox Envelope: User's Guide

June 2, 2012

Contents	1
1 Overview	2
1.1 Technical Support	3
1.2 Directory Structure	3
1.3 Document Organization	4
2 Quick Start	5
2.1 Installation	5
2.2 A Guided Tour	5
2.3 Two examples	7
3 Envelope Models	12
3.1 Multivariate Linear Regression	12
3.2 Envelope Model	12
3.3 Envelope Model Using Sequential Algorithm	13
3.4 Heteroscedastic Envelope Model	15
3.5 Inner Envelope Model	16
3.6 Partial Envelope Model	17
3.7 Scaled Envelope Model	18
3.8 Envelope Model in the Predictor Space	19
3.9 Envelope Model in the Predictor Space Using Partial Least Squares Algorithm	20
4 Optional Arguments	22
Bibliography	27

Overview

The envelope model is a new area in multivariate analysis. It uses dimension reduction techniques to achieve efficient estimation of parameters, for example, the regression coefficients in multivariate linear regression (MLR).

This MATLAB toolbox “envelope” currently implements six models in the envelope family, including envelope model (using Grassmann manifold optimization algorithm and sequential algorithm), heteroscedastic envelope model, inner envelope model, partial envelope model, scaled envelope model and envelope model in the predictor space (using Grassmann manifold optimization algorithm and partial least squares algorithm):

- env** Envelope Module. The envelope model is a general tool for efficient estimation in the context of MLR, and it has the potential to achieve substantial efficiency gains when part of the response variables or their linear combination is invariant to the changes of the predictors [Cook et al., 2010].
- envpls** Envelope Model Using Sequential Algorithm Module. This module also implements the envelope model, but uses a sequential computing algorithm. This algorithm makes the envelope model applicable for small sample size cases [Cook et al., 2012].
- henv** Heteroscedastic Envelope Module. The heteroscedastic envelope model is used when the data has none constant error structure [Su and Cook, 2012].
- ienv** Inner Envelope Module. The inner envelope model is also a general tool for efficient estimation in the context of MLR, but has a different mechanism from the envelope model. Therefore it may provide efficiency gains in the cases when the envelope models fail to offer any gains [Su and Cook, to appear].
- penv** Partial Envelope Module. The partial envelope model can be applied when part of the predictors are of main interest. It often gives more efficiency gains than the envelope model in estimating the coefficients of the main predictors [Su and Cook, 2011].
- senv** Scaled Envelope Module. The scaled envelope model is used when the user hopes to have a scale-invariant version of the envelope model [Cook and Su, 2012].
- xenv** Envelope Model in the Predictor Space Module. The envelope model in the predictor space is used when some of the predictors or their linear combinations does not contribute in the change of the responses. It can potentially bring a better prediction performance than the standard model, or even partial least squares [Cook et al., 2012].
- xenvpls** Envelope Model in the Predictor Space Using Partial Least Squares Algorithm Module. This module implements the envelope model in the predictor space, but uses the partial least squares algorithm. This algorithm makes the envelope model in the predictor space applicable for small sample size cases [Cook, 2012].

The complete applicability of this toolbox is described in Table 1.1.

Model	Model Selection	Inference	Section
Envelope Model	AIC	Estimation and Prediction	3.2
	BIC	Bootstrap for Estimating Standard Errors	
	LRT	Hypothesis Test on Coefficients	
Envelope Model Using Sequential Algorithm	m-fold cross validation	Bootstrap for Estimating Standard Errors	3.3
Heteroscedastic Envelope Model	AIC	Estimation and Prediction	3.4
	BIC	Bootstrap for Estimating Standard Errors	
	LRT	Hypothesis Test on Coefficients	
Inner Envelope Model	AIC	Estimation and Prediction	3.5
	BIC	Bootstrap for Estimating Standard Errors	
	LRT	Hypothesis Test on Coefficients	
Partial Envelope Model	AIC	Estimation and Prediction	3.6
	BIC	Bootstrap for Estimating Standard Errors	
	LRT	Hypothesis Test on Coefficients	
Scaled Envelope Model	AIC	Estimation and Prediction	3.7
	BIC	Bootstrap for Estimating Standard Errors	
		Hypothesis Test on Coefficients	
Envelope Model in the Predictor Space	AIC	Estimation and Prediction	3.8
	BIC	Bootstrap for Estimating Standard Errors	
	LRT	Hypothesis Test on Coefficients	
Envelope Model in the Predictor Space Using Partial Least Squares Algorithm	m-fold cross validation	Bootstrap for Estimating Standard Errors	3.9

Table 1.1: Applicability of toolbox “envelope”

1.1. Technical Support

We provides a support website, on which the users can download the toolbox, report bugs and check recent updates <http://code.google.com/p/envlp/>. For further help, the users can contact the authors of the toolbox: Dennis Cook (dennis@stat.umn.edu), Zhihua Su (suzhihua@stat.umn.edu), and Yi Yang (yiyang@umn.edu).

1.2. Directory Structure

Toolbox folder contains following files and sub-folders:

./envelope_license.m Envelope Toolbox License.

./install_envelope.m Envelope Toolbox installation script.

./README.txt Envelope Toolbox Quick Start.

./data/ Data published in papers on envelope models and data description manual.

./doc/ User’s guide, function reference manual.

./examples/ Examples for demonstration.

./src/ Envelope toolbox source code.

1.3. Document Organization

This user's guide is organized as follows:

Chapter 1 Overview. Introduces the content and applicability of the toolbox.

Chapter 2 Quick start. Describes the installation of the toolbox, and provides simple examples on using the main functions.

Chapter 3 Envelope models. Discusses each module of the toolbox, and demonstrates the applicability with more detailed examples.

Chapter 4 Optional arguments. Explains how to control the convergence speed, input user-specified starting values, and display iteration process.

Quick Start

2.1. Installation

To install the toolbox, direct your MATLAB working directory to the folder “envlp”, and type

```
install_envelope
```

If the users agree with our license statements, the installation is completed, and all the utilities of the toolbox are added to the MATLAB path. Once installed, no further action is needed to call the functions in the toolbox, even if the users change a working directory, or MATLAB is relaunched.

2.2. A Guided Tour

This “envelope” toolbox consists tools in the following three classes:

1. Dimension selection: Select the dimension of the envelope subspace.
2. Model Fitting with Selected Dimension: Fit the model.
3. Post processing: Inference based on the model fitting.

We will present a tour of our toolbox through all three classes.

Dimension selection

The dimension selection tools available in this toolbox are Akaike information criterion (AIC), Bayesian information criterion (BIC), Likelihood ratio testing (LRT) and m-fold cross validation. The m-fold cross validation can only be applied to “envpls” and “xenvpls”, while AIC, BIC and LRT can be applied to “env”, “henv”, “ienv”, “penv”, “senv” and “xenv”, except that LRT cannot be applied to “senv”, as indicated in Table 1.1. AIC and BIC generally require longer computing time than LRT, because of the nature of the method. The syntax of AIC, BIC, LRT and m-fold cross validation are

```
u = modelselectaic(X, Y, modelType)
u = modelselectbic(X, Y, modelType)
u = modelselectlrt(X, Y, alpha, modelType)
u = mfoldcv(X, Y, m, modelType)
```

For the inputs: X is a matrix containing the predictors, except that with “penv”, it is a list having X.X1 and X.X2; Y is a matrix containing the response; ‘alpha’ is the significance level; ‘m’ is an integer indicating m-fold cross validation and ‘modelType’ is a string indicting the model. The

choices for `modelselectaic`, `modelselectbic` and `modelselectlrt` can be “env”, “henv”, “ienv”, “penv”, “senv” and “xenv”, and the choices for `mfoldcv` are “envpls” and “xenvpls”. The output is the dimension of the envelope subspace selected by the tool. Examples will be provided in Section 2.3.

Model Fitting with Selected Dimension

The functions to fit the six models in the envelope family are the main drive of this toolbox. Their syntax are

```
ModelOutput = env(X, Y, u)
ModelOutput = envpls(X, Y, u)
ModelOutput = henv(X, Y, u)
ModelOutput = ienv(X, Y, u)
ModelOutput = penv(X, Y, u)
ModelOutput = senv(X, Y, u)
ModelOutput = xenv(X, Y, u)
ModelOutput = xenvpls(X, Y, u)
```

The input X and Y are the same as the input in the dimension selection tools, u is the dimension of the envelope subspace, the user can input any dimension if it is known, and use dimension selection tools described in dimension selection part to choose u . The output `ModelOutput` is a list, containing the maximum likelihood estimators (MLE) under the model, as well as some key statistics for inference, including the standard error ratios of elements in the regression coefficients for the standard model versus the envelope model, maximized value of the log likelihood, number of parameters, etc.

`envpls` and `xenvpls` fit the same model as `env` and `xenv`, respectively, but use a different computing algorithm, making them capable to accommodate small sample size cases. With large sample size cases, `envpls` and `xenvpls` can also be applied, and usually give different results from `env` and `xenv`. [Cook et al., 2012] indicates that `env` and `xenv` normally give better results, and therefore are recommended.

Post-processing

This toolbox provides functions for the following inference:

- Compute bootstrap standard errors for the regression coefficients $\hat{\beta}$, which gives an estimator for the actual standard errors of $\hat{\beta}$.
- For a given data point, report its fitted value or predicted value, with associated standard errors.
- Test hypothesis

$$H_0: \mathbf{L}\hat{\beta}\mathbf{R} = \mathbf{A}$$

$$H_a: \mathbf{L}\hat{\beta}\mathbf{R} \neq \mathbf{A}$$

where \mathbf{L} , \mathbf{R} and \mathbf{A} are given matrices.

The syntax of functions for the inferences are

```
bootse = bootstrapse(X, Y, u, B, modelType)
PredictOutput = prediction(ModelOutput, Xnew, infType, modelType)
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

respectively. For the inputs, `modelType` in all three functions can be “env”, “henv”, “ienv”, “penv”, “senv” and “xenv”, while `modelType` in `bootstrapse` can also be “envpls” and “xenvpls”; `ModelOutput` in `prediction` and `testcoefficient` is a list returned from the functions for model fitting; `X`, `Y`, and `u` are discussed in the dimension selection part; `B` is the number of bootstrap sample; `Xnew` is a value of `X` with which to estimate or predict `Y`; `infType` is a string of characters indicating the inference type, the choices can be “estimation” or “prediction”; and `TestInput` is a list that specifies the null hypothesis, including `L`, `R`, and `A`. `TestInput` is an option input argument, is missing, `testcoefficient` will perform the usual F test, i.e. testing if $\beta = 0$. The output of `bootstrapse` is a matrix the same size of $\hat{\beta}$, with each elements as the bootstrap standard error of its corresponding element in $\hat{\beta}$. The output of `prediction` is a list containing the fitted or predicted value, its covariance matrix and standard errors. The function `testcoefficient` will print out a form displaying the test statistic, degrees of freedom of the reference distribution and the p-value, all of which are also summarized in the list `TestOutput`.

2.3. Two examples

To demonstrate their usage, we take the envelope model and partial envelope model as examples. Other methods can be applied similarly, and are also discussed in details in Section 3.

Wheat Protein Data

For the envelope model, the data we use is the wheatprotein data, which is used as a data analysis example in Cook et al. [2010]. First we load the data and assign column 1 to 6 as the response `Y` and column 8 as the predictor `X`. The description of the data is available under the folder `/envlp/data`.

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1:6);
```

Then we apply the dimension selection tools and get the following results.

```
modelType = 'env';
u = modelselectaic(X, Y, modelType)
u =
    1
u = modelselectbic(X, Y, modelType)
u =
    1
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, modelType)
u =
    1
```

In this example, all three dimension selection tools agree that the dimension of the envelope should be 1, which is consistent with the results in Cook et al. [2010]. Then we fit the envelope model with dimension 1.

```
ModelOutput = env(X, Y, u)
ModelOutput =
```

```

    beta: [6x1 double]
    Sigma: [6x6 double]
    Gamma: [6x1 double]
    Gamma0: [6x5 double]
    eta: 8.5647
    Omega: 7.8762
    Omega0: [5x5 double]
    alpha: [6x1 double]
    l: -850.7592
    covMatrix: [6x6 double]
    asyEnv: [6x1 double]
    ratio: [6x1 double]
    np: 28
    n: 50

```

We notice that `ModelOutput` is a list that includes the MLEs and the statistics relevant to the influence of the envelope model. For more details of the components in `ModelOutput`, please refer to the Reference of the toolbox. If we would want to see the regression coefficients, as well as the standard error ratios, we can type

```

ModelOutput.beta
ans =
    -1.0644
     4.4730
     3.6839
    -5.9770
     0.6013
    -1.5986

ModelOutput.ratio
ans =
    28.0945
    18.4326
    23.6384
    16.3211
    65.8245
     6.4668

```

and get the results. We can also look at the eigenvalues of Σ_1 and Σ_2 by the following commands

```

ModelOutput.Omega
ans =
     7.8762
eig(ModelOutput.Omega0)
ans =
    1.0e+03 *
     6.5166
     0.2083
     0.0201
     0.0004
     0.0003

```

These results are consistent with those published in Cook et al. [2010]. With the model fitting results, if we would like to check the fitted value and its standard errors for the second observation, we can use the `prediction` function.

```
Xnew = X(2, :)';
PredictOutput = predict_env(ModelOutput, Xnew, 'estimation')
PredictOutput =
      value: [6x1 double]
   covMatrix: [6x6 double]
         SE: [6x1 double]
```

We can get the standard errors of the fitted value, and also we can compare the fitted value with its true value

```
PredictOutput.SE
ans =
    4.8892
    4.0227
    4.3237
    4.7470
    6.8186
    2.6948

[PredictOutput.value, Y(2,:)]
ans =
    474.7135    458.0000
    127.4740    112.0000
    251.2044    236.0000
    380.8280    368.0000
    380.9473    383.0000
    -6.3287   -15.0000
```

Now suppose we want to test if $\beta = 0$, we run the `testcoefficient` as

```
TestOutput = testcoefficient(ModelOutput, modelType)
Test Hypothesis      Chisq Statistic      DF      P-value
-----
L * beta * R = A      116.230           6      0.0000
-----
```

Notice that we did not have `TestInput`, because we are performing the usual F test on if $\beta = 0$. The test results indicate that we have strong evidence to reject $\beta = 0$.

Fiber and Paper Data

Now we look at another example for the partial envelope model. The data we load is the fiber and paper data used in Su and Cook [2011].

```
load fiberpaper.dat
Y = fiberpaper(:, 1 : 4);
```

```
X.X1 = fiberpaper(:, 7);
X.X2 = fiberpaper(:, 5 : 6);
```

We notice that for partial envelope model, X is a list, with $X.X1$ containing the main predictors and $X.X2$ containing covariates. The dimension selection process is parallel to that for the envelope model:

```
modelType = 'penv';
u = modelselectaic(X, Y, modelType)
u =

    3

u = modelselectbic(X, Y, modelType)
u =

    1

alpha = 0.01;
u = modelselectlrt(X, Y, alpha, modelType)
u =

    1
```

In this example, AIC picks the dimension 3, while BIC and LRT both gives dimension as 1. In Su and Cook [2011], $u = 1$ is used for model fitting and inference. So we fit the partial envelope model with $u = 1$.

```
u = 1;
ModelOutput = penv(X, Y, u)
ModelOutput =
    beta1: [4x1 double]
    beta2: [4x2 double]
    alpha: [4x1 double]
    Gamma: [4x1 double]
    eta: 0.0047
    Omega: 0.0149
    Omega0: [3x3 double]
    Sigma: [4x4 double]
    l: -35.6323
    np: 23
    covMatrix: [12x12 double]
    asyPenv: [4x1 double]
    ratio: [4x1 double]
    n: 62
```

Again we get the output as a list, which contains the MLE for the parameters and statistics that are relevant to the inference for the partial envelope model. Detailed description of the elements in the list is in the Reference of the toolbox. To look at the standard error ratios, and the eigenvalues of Σ_1 and Σ_2 , we key in

```

ModelOutput.ratio
ans =
    66.0742
     6.9326
    10.5048
     9.6279

ModelOutput.Omega
ans =
    0.0149

eig(ModelOutput.Omega0)
ans =
    4.9819
    0.0999
    0.0050

```

The results are almost the same the those in Su and Cook [2011]. They are slightly different, because the toolbox used a different starting value algorithm, which is less likely to be trapped in local minimum and thus leads to a more reliable result. Next we look at the standard errors for elements in $\hat{\beta}$. Since `ModelOutput.asyPenv` contains the asymptotic standard errors, for actual standard errors, we need to divide the asymptotic standard errors by \sqrt{n} , where n is the sample size and is returned in `ModelOutput.n`.

```

ModelOutput.asyPenv / sqrt(ModelOutput.n)
ans =
    1.0e-03 *
    0.3181
    0.8704
    0.9719
    0.4689

```

The standard errors above are computed assuming the partial envelope model. They can also be estimated using bootstrap the residuals.

```

B = 5000;
bootse = bootstrapse(X, Y, u, B, modelType)
bootse =
    0.0052
    0.0020
    0.0029
    0.0013

```

The uses may get a slightly different results from the above, as a different random seed is used.

Envelope Models

3.1. Multivariate Linear Regression

The envelope model is originally developed under the framework of multivariate linear regression, and its performance is frequently compared to the standard multivariate linear regression model. In this toolbox, we offer some simple functions for the comparison between two models. A standard multivariate linear regression model is formulated as

$$\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\beta}\mathbf{X} + \boldsymbol{\varepsilon},$$

where $\mathbf{Y} \in \mathbb{R}^r$ is the multivariate response, $\mathbf{X} \in \mathbb{R}^p$ is non-stochastic predictor, and $\boldsymbol{\varepsilon} \in \mathbb{R}^r$ follows a distribution with mean 0, and covariance matrix $\boldsymbol{\Sigma} \in \mathbb{R}^{r \times r}$, $\boldsymbol{\alpha} \in \mathbb{R}^r$ and $\boldsymbol{\beta} \in \mathbb{R}^{r \times p}$ are the unknown intercept and coefficients. The goal of the envelope model is to reduce the standard errors in estimating $\boldsymbol{\beta}$.

This function to fit this model is

```
ModelOutput = fit_OLS(X, Y)
```

where the input \mathbf{X} is an $n \times p$ matrix, with the i th row as the transpose of the i th observation of \mathbf{X} , and \mathbf{Y} is an $n \times r$ matrix with the i th row as the transpose of the i th observation of \mathbf{Y} . The output `ModelOutput` is a list, which contains the ordinary least squares (OLS) estimators of $\boldsymbol{\beta}$, $\boldsymbol{\Sigma}$ and $\boldsymbol{\alpha}$.

The standard errors of the OLS estimator $\hat{\boldsymbol{\beta}}$ is often compared to those of the envelope model, the function

```
bootse = bootstrapse_OLS(X, Y, B)
```

computes the standard errors of $\hat{\boldsymbol{\beta}}$ estimated by bootstrapping the residuals. The input B is the number of bootstrap samples, and `bootse` returns a matrix with the same size as $\hat{\boldsymbol{\beta}}$, with each element as the standard error of the corresponding element in $\hat{\boldsymbol{\beta}}$.

3.2. Envelope Model

The envelope model is a general method in the envelope family to reduce the standard errors in estimating $\boldsymbol{\beta}$. It can be applied when the number of the responses is strictly greater than the number of the predictors, and the responses are continuous variables. It has a potential to obtain efficiency gains in estimating $\boldsymbol{\beta}$, compared to the OLS estimators.

In the multivariate linear regression context in Section 3.1, a coordinate form of the envelope model is

$$\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\Gamma}\boldsymbol{\eta}\mathbf{X} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \boldsymbol{\Gamma}\boldsymbol{\Omega}\boldsymbol{\Gamma}^T + \boldsymbol{\Gamma}_0\boldsymbol{\Omega}_0\boldsymbol{\Gamma}_0^T,$$

where the regression coefficients $\boldsymbol{\beta} = \boldsymbol{\Gamma}\boldsymbol{\eta}$, $\mathcal{B} = \text{span}(\boldsymbol{\beta})$, $\boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B})$ – the envelope subspace, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B})$, $\boldsymbol{\eta} \in \mathbb{R}^{u \times p}$, $\boldsymbol{\Omega} \in \mathbb{R}^{u \times u}$, $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ are coordinates, and u is the dimension of $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B})$.

To select the dimension of the envelope subspace u , we can use the following functions

```
u = modelselectaic(X, Y, 'env')
u = modelselectbic(X, Y, 'env')
alpha = 0.01; # Users can specify other significance level
u = modelselectlrt(X, Y, alpha, 'env')
```

The possible values of u can be any integer from 0 to r . When $u = r$, the envelope model is essentially the standard multivariate linear model. And when $u = 0$, it means that $\boldsymbol{\beta} = \mathbf{0}$, then the changes in \mathbf{Y} do not depend on \mathbf{X} . After obtaining u , we can fit the envelope model by

```
ModelOutput = env(X, Y, u)
```

The output `ModelOutput` is a list, which contains the MLE of $\boldsymbol{\beta}$, $\boldsymbol{\Sigma}$, $\boldsymbol{\Gamma}$, $\boldsymbol{\Gamma}_0$, $\boldsymbol{\eta}$, $\boldsymbol{\Omega}$, $\boldsymbol{\Omega}_0$ and $\boldsymbol{\alpha}$, and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\boldsymbol{\beta})$, the asymptotic standard errors of $\boldsymbol{\beta}$, the ratios of the asymptotic standard errors of the standard model versus the envelope model for elements in $\boldsymbol{\beta}$, the number of parameters in the model, and the number of observations in the data. After fitting the data and get `ModelOutput`, we can perform post processing inference as computing bootstrap standard errors of $\hat{\boldsymbol{\beta}}$ by

```
bootse = bootstrapse(X, Y, u, B, 'env')
```

or computing the fitted value or predicted value and associated standard errors given a new \mathbf{X} by

```
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'env')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'env')
```

or testing if a linear combination of $\boldsymbol{\beta}$ is equal to a particular matrix, i.e. given \mathbf{L} , \mathbf{R} , and \mathbf{A} , testing if $\mathbf{L}\boldsymbol{\beta}\mathbf{R} = \mathbf{A}$,

```
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

The inputs and outputs of these post processing functions are discussed in details in Section 2.2.

3.3. Envelope Model Using Sequential Algorithm

The envelope model using sequential algorithm implements the envelope model in Section 3.2, but uses a different computing algorithm. The algorithm estimates the envelope subspace sequentially Cook [2012]. It is faster than the Grassmann manifold optimization algorithm used in `env`, and it is applicable when the sample size n is less than the number of responses r . In large sample cases, the performance of this algorithm is not as good as the Grassmann manifold optimization algorithm, but as it is faster, it can provide a starting value for the Grassmann manifold optimization.

To select the dimension of the envelope subspace u , we can use the m -fold cross validation. Common choices are 5 or 10.

```
m = 5;
u = mfoldcv(X, Y, m, 'envpls')
```

The possible values of u can be any integer from 0 to the minimum of the floor of $(m-1)n/m-1$ and r . When $u = r$, the envelope model is essentially the standard multivariate linear model, and when $u = 0$, it means that $\beta = 0$, and the changes in Y do not depend on the changes in X . If the sample size is large, we can also use the tools for the “env” to select u , as “env” and “envpls” implement the same model.

```
u = modelselectaic(X, Y, 'env')
u = modelselectbic(X, Y, 'env')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'env')
```

After getting u , to fit the envelope model using the sequential algorithm, we use

```
ModelOutput = envpls(X, Y, u)
```

The output `ModelOutput` is a list, which contains the estimators of β , Σ , Γ , Γ_0 , η , Ω , Ω_0 and α , and also the number of parameters in the model and number of observations in the data. It does not contain any information based on the likelihood function, for example, the maximized log-likelihood, the asymptotic covariance matrix of $\text{vec}(\beta)$, the asymptotic standard errors of elements in β , etc. This is because `envpls` is mainly implemented for small sample size cases, under which the likelihood cannot be estimated. Then any inference that is based on the covariance matrix of $\text{vec}(\beta)$ such as estimation, prediction and test coefficients cannot be performed with “envpls”. But we can still estimate the bootstrap standard errors for $\hat{\beta}$ by

```
bootse = bootstrapse(X, Y, u, B, 'envpls')
```

To Dennis: I am not sure if we would like to include this in. Using the sequential algorithm as the starting value normally gives a different result, which is often the local maximum.

Another utility of `envpls` is that it can provide a starting value for `env` when r is large. `env` can be very slow if r is large, while `envpls` is faster than `env`. Using the starting value from `envpls` can save much computing time, and at the same time, may still have the good performance from `env`. An example is shown as follows.

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1 : 6);
u = 1;
temp = envpls(X, Y, u);
Opts.init = temp.Gamma;
ModelOutput = env(X, Y, u, Opts);
```

The usage of `Opts` will be discussed in details in Chapter 4.

3.4. Heteroscedastic Envelope Model

The heteroscedastic envelope model [Su and Cook, 2012] is used when the data has none constant error structure, and it is developed under the framework of estimating multivariate means for different populations. To use the heteroscedastic envelope model, the number of response should be greater than or equal to the number of populations.

The standard model for estimating multivariate means for p populations can be formulated as $\mathbf{Y}_{(i)j} = \boldsymbol{\mu} + \boldsymbol{\beta}_{(i)} + \boldsymbol{\varepsilon}_{(i)j}$, $i = 1, \dots, p$, $j = 1, \dots, n_{(i)}$. We use the subscript (i) to represent the i th group, and we j without the parenthesis to represent the j th observation. Then $\mathbf{Y}_{(i)j} \in \mathbb{R}^r$ is the j th observation in the i th group, $n_{(i)}$ is the number of observations in the i th group, $\boldsymbol{\mu} \in \mathbb{R}^r$ is the grand mean, $\boldsymbol{\beta}_{(i)} \in \mathbb{R}^r$ is the main effect for the i th group and satisfies $\sum_{i=1}^p n_{(i)} \boldsymbol{\beta}_{(i)} = \mathbf{0}$, the errors $\boldsymbol{\varepsilon}_{(i)j}$ follows a distribution with mean 0, and covariance matrix $\boldsymbol{\Sigma}_{(i)} \in \mathbb{R}^{r \times r}$. With this formulation, let $\mathcal{B} = \text{span}(\boldsymbol{\beta}_{(1)} \cdots, \boldsymbol{\beta}_{(p)})$, and $\mathcal{M} = \{\boldsymbol{\Sigma}_{(i)} : i = 1, \dots, p\}$, the coordinate form of the heteroscedastic envelope model is displayed as follows:

$$\mathbf{Y}_{(i)j} = \boldsymbol{\mu} + \boldsymbol{\Gamma} \boldsymbol{\eta}_{(i)} + \boldsymbol{\varepsilon}_{(i)j}, \quad \boldsymbol{\Sigma}_{(i)} = \boldsymbol{\Gamma} \boldsymbol{\Omega}_{1(i)} \boldsymbol{\Gamma}^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T, \quad i = 1, \dots, p, \quad j = 1, \dots, n_{(i)},$$

where $\boldsymbol{\beta}_{(i)} = \boldsymbol{\Gamma} \boldsymbol{\eta}_{(i)}$, $\boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans $\mathcal{E}_{\mathcal{M}}(\mathcal{B})$ – the \mathcal{M} envelope of \mathcal{B} , $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\mathcal{M}}(\mathcal{B})$, $\boldsymbol{\eta}_{(i)} \in \mathbb{R}^{u \times 1}$, $\boldsymbol{\Omega}_{1(i)} \in \mathbb{R}^{u \times u}$, $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$, $i = 1, \dots, p$, are coordinates, $\sum_{i=1}^p n_{(i)} \boldsymbol{\eta}_{(i)} = \mathbf{0}$, and u is the dimension of $\mathcal{E}_{\mathcal{M}}(\mathcal{B})$.

To use the heteroscedastic envelope model, first we check if the data has heteroscedastic error structure by the Box's M test [Johnson and Wichern, 2007]. Take the water strider example [Su and Cook, 2012], the following codes performs the Box's M test.

```
load waterstrider.mat
alpha = 0.01; # Users can specify other significance level
TestOutput = mtest(X, Y, alpha);
```

The input X is a group indicator, there is no constraint on the number of columns of X, as long as X takes p unique values for p populations. For example, if there are three groups, X can take three unique values as 1, 2, and 3 to indicate the groups, or X can take (0, 1), (1, 0), and (0, 0) to indicate the groups. The output is displayed below

```
-----
MBox      Chi-sqr.      df      P
-----
157.5977   137.3361      72      0.0000
-----
Covariance matrices are significantly different.
```

As the data has none constant covariance structure, we can apply the heteroscedastic envelope model. To select the dimension of the envelope subspace, we use

```
u = modelselectaic(X, Y, 'henv')
u = modelselectbic(X, Y, 'henv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'henv')
```

The input X can have the same form as in the function `mtest`. After obtaining u , the heteroscedastic envelope model can be fit by

```
ModelOutput = henv(X, Y, u)
```

The output `ModelOutput` is a list containing the MLE of the grand mean $\boldsymbol{\mu}$, the group mean $\boldsymbol{\mu} + \boldsymbol{\beta}_{(i)}$, the fitted values $\hat{\mathbf{Y}}$, $\boldsymbol{\Gamma}$, $\boldsymbol{\Gamma}_0$, $\boldsymbol{\beta}_{(i)}$, the unique values in \mathbf{X} , $\boldsymbol{\Sigma}_{(i)}$, $\boldsymbol{\eta}_{(i)}$, $\boldsymbol{\Omega}_{(i)}$, and $\boldsymbol{\Omega}_0$, and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $(\boldsymbol{\mu}^T, \boldsymbol{\beta}_{(1)}^T, \dots, \boldsymbol{\beta}_{(p)}^T)^T$, the asymptotic standard errors of $\boldsymbol{\beta}_{(i)}$, the ratios of the asymptotic standard errors of the standard model versus the inner envelope model for elements in $\boldsymbol{\beta}_{(i)}$, the number of parameters in the model, and the number of observations in the data. After model fitting, the functions for post processing inference are

```
bootse = bootstrapse(X, Y, u, B, 'henv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'henv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'henv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

Their inputs and outputs are similar to those for the envelope model.

3.5. Inner Envelope Model

Like the envelope model, the inner envelope model Su and Cook is also a general method to reduce the standard errors in estimating $\boldsymbol{\beta}$. It has a different mechanism to achieve the efficiency gains than the envelope model, so it may be very useful when the envelope model reduces to the standard model and offers no efficiency gains. The inner envelope model also requires that the number of predictors is strictly less than the number of responses.

In the multivariate linear regression context (Section 3.1), the coordinate form of the inner envelope model can be written as

$$\mathbf{Y} = \boldsymbol{\alpha} + (\boldsymbol{\Gamma}_1 \boldsymbol{\eta}_1^T + \boldsymbol{\Gamma}_0 \mathbf{B} \boldsymbol{\eta}_2^T) \mathbf{X} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \boldsymbol{\Gamma}_1 \boldsymbol{\Omega}_1 \boldsymbol{\Gamma}_1^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T,$$

where $\boldsymbol{\beta} = \boldsymbol{\Gamma}_1 \boldsymbol{\eta}_1^T + \boldsymbol{\Gamma}_0 \mathbf{B} \boldsymbol{\eta}_2^T \in \mathbb{R}^{r \times p}$, $\mathcal{B} = \text{span}(\boldsymbol{\beta})$, $\boldsymbol{\Gamma}_1 \in \mathbb{R}^{r \times u}$ spans the inner envelope subspace $\mathcal{SE}_{\boldsymbol{\Sigma}}(\mathcal{B})$, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{SE}_{\boldsymbol{\Sigma}}(\mathcal{B})$, $\mathbf{B} \in \mathbb{R}^{(r-u) \times (p-u)}$ is a semi-orthogonal matrix such that $(\boldsymbol{\Gamma}_1, \boldsymbol{\Gamma}_0 \mathbf{B})$ spans \mathcal{B} , $\boldsymbol{\eta}_1 \in \mathbb{R}^{p \times u}$, $\boldsymbol{\eta}_2 \in \mathbb{R}^{p \times (p-u)}$, $\boldsymbol{\Omega}_1 \in \mathbb{R}^{u \times u}$, and $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ contains the coordinates, and u is the dimension of $\mathcal{SE}_{\boldsymbol{\Sigma}}(\mathcal{B})$.

To select the dimension of $\mathcal{SE}_{\boldsymbol{\Sigma}}(\mathcal{B})$, we use the same functions as for the envelope model, except that the `modelType` is `'ienv'`.

```
u = modelselectaic(X, Y, 'ienv')
u = modelselectbic(X, Y, 'ienv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'ienv')
```

The possible values of u are integers from 0 to p . When $u = 0$, the inner envelope model reduces to the standard model, and offers no efficiency gains. When $u = p$, the inner envelope model is equivalent to an envelope model with dimension p .

Given the dimension u , the inner envelope model can be fitted by

```
ModelOutput = ienv(X, Y, u)
```

The output `ModelOutput` is a list containing the MLE of $\boldsymbol{\beta}$, $\boldsymbol{\Sigma}$, $\boldsymbol{\Gamma}_1$, $\boldsymbol{\Gamma}_0$, $\boldsymbol{\eta}_1$, \mathbf{B} , $\boldsymbol{\eta}_2$, $\boldsymbol{\Omega}_1$, $\boldsymbol{\Omega}_0$ and $\boldsymbol{\alpha}$, and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\boldsymbol{\beta})$, the asymptotic standard errors of $\boldsymbol{\beta}$, the ratios of the asymptotic standard errors of the standard model versus the inner envelope model for elements in $\boldsymbol{\beta}$, the number of parameters in the model, and the number of observations in the data. After model fitting, the functions for post processing inference are

```
bootse = bootstrapse(X, Y, u, B, 'ienv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'ienv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'ienv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

Their inputs and outputs are similar to those for the envelope model.

3.6. Partial Envelope Model

The partial envelope model [Su and Cook, 2011] can be applied when part of the predictors are of main interest. It only requires that the number of the main predictors is strictly less than the number of the responses. This is particular useful when the number of predictors p is large, but there are only a small number of predictors are of main interest. Suppose that $\mathbf{X}^T = (\mathbf{X}_1^T, \mathbf{X}_2^T)$, where $\mathbf{X}_1 \in \mathbb{R}^{p_1}$ are predictors of main interest, and $\mathbf{X}_2 \in \mathbb{R}^{p_2}$ are covariates, $p_1 + p_2 = p$. Then the standard model is formulated as $\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\beta}_1 \mathbf{X}_1 + \boldsymbol{\beta}_2 \mathbf{X}_2 + \boldsymbol{\varepsilon}$, and the coordinate form of the partial envelope model is

$$\mathbf{Y} = \boldsymbol{\alpha} + \boldsymbol{\Gamma} \boldsymbol{\eta} \mathbf{X}_1 + \boldsymbol{\beta}_2 \mathbf{X}_2 + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \boldsymbol{\Gamma} \boldsymbol{\Omega} \boldsymbol{\Gamma}^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T,$$

where $\boldsymbol{\beta}_1 = \boldsymbol{\Gamma} \boldsymbol{\eta} \in \mathbb{R}^{r \times p_1}$, $\boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans the partial envelope $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B}_1)$ – the partial envelope subspace, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B}_1)$, $\boldsymbol{\eta} \in \mathbb{R}^{u \times p_1}$, $\boldsymbol{\Omega} \in \mathbb{R}^{u \times u}$, $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ are coordinates, $\boldsymbol{\beta}_2 \in \mathbb{R}^{r \times p_2}$ are the coefficients for \mathbf{X}_2 and u is the dimension of $\mathcal{E}_{\boldsymbol{\Sigma}}(\mathcal{B}_1)$.

To use the functions related to partial envelope model, the input \mathbf{X} is a list, which has two elements `X.X1` and `X.X2`. The element `X.X1` is an $n \times p_1$ matrix, with the i th row as the transpose of the i th observation of \mathbf{X}_1 , and `X.X2` is an $n \times p_2$ matrix, with the i th row as the transpose of the i th observation of \mathbf{X}_2 . The form of \mathbf{X} as a list is unique to the partial envelope model.

The following functions can be applied to select u ,

```
u = modelselectaic(X, Y, 'penv')
u = modelselectbic(X, Y, 'penv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'penv')
```

The possible values of u are any integer from 0 to r , when $u = r$, the partial envelope model reduces the standard model, and when $u = 0$, the changes in \mathbf{Y} do not depend on the changes in \mathbf{X}_1 . After obtaining u , the partial envelope model can be fit by

```
ModelOutput = penv(X, Y, u)
```

The output `ModelOutput` is a list containing the MLE of $\boldsymbol{\beta}_1$, $\boldsymbol{\beta}_2$, $\boldsymbol{\Sigma}$, $\boldsymbol{\Gamma}$, $\boldsymbol{\Gamma}_0$, $\boldsymbol{\eta}$, $\boldsymbol{\Omega}$, $\boldsymbol{\Omega}_0$ and $\boldsymbol{\alpha}$, and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $(\text{vec}(\boldsymbol{\beta}_2)^T, \text{vec}(\boldsymbol{\beta}_1)^T)^T$, the asymptotic standard errors of $\boldsymbol{\beta}_1$, the ratios of the

asymptotic standard errors of the standard model versus the partial envelope model for elements in β_1 , the number of parameters in the model, and the number of observations in the data. The functions for post processing inference are

```
bootse = bootstrapse(X, Y, u, B, 'penv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'penv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'penv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

The usage of functions are introduced in Section 2.2, and are similar to those for the envelope models as discussed in Section 3.2. The only difference is that in the function `prediction`, the input `Xnew` is a list, which has `Xnew.X1` and `Xnew.X2`. `Xnew.X1` is a p_1 dimensional column vector containing the new value of \mathbf{X}_1 , and `Xnew.X2` is a p_2 dimensional column vector containing the new value of \mathbf{X}_2 .

3.7. Scaled Envelope Model

Scaled envelope model (Cook and Su) is used when the user hopes to have a scale-invariant version of the envelope model. It is more efficient or at least as efficient as the standard model. The scaled envelope model normally takes longer to compute because of the nature of its iterative computing algorithm. But users can print out the model fitting process, which will be discussed in Chapter 4.

In the multivariate linear regression context (3.1), the coordinate form of the scaled envelope model can be written as

$$\mathbf{Y} = \boldsymbol{\alpha} + \Lambda \Gamma \boldsymbol{\eta} \mathbf{X} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma} = \Lambda \Gamma \boldsymbol{\Omega} \Gamma^T \Lambda + \Lambda \Gamma_0 \boldsymbol{\Omega}_0 \Gamma_0^T \Lambda,$$

where $\boldsymbol{\beta} = \Lambda \Gamma \boldsymbol{\eta} \in \mathbb{R}^{r \times p}$, $\Lambda \in \mathbb{R}^{r \times r}$ is the scaling matrix, it is a diagonal matrix with the first element as 1, and the other diagonal elements as positive real numbers, $\Gamma \in \mathbb{R}^{r \times p}$ spans the envelope subspace $\mathcal{E}_{\Lambda^{-1} \boldsymbol{\Sigma} \Lambda^{-1}}(\Lambda^{-1} \mathcal{B})$, $\Gamma_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\Lambda^{-1} \boldsymbol{\Sigma} \Lambda^{-1}}(\Lambda^{-1} \mathcal{B})$, $\boldsymbol{\eta} \in \mathbb{R}^{u \times p}$, $\boldsymbol{\Omega} \in \mathbb{R}^{u \times u}$, and $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ carry the coordinates, and u is the dimension of the envelope subspace $\mathcal{E}_{\Lambda^{-1} \boldsymbol{\Sigma} \Lambda^{-1}}(\Lambda^{-1} \mathcal{B})$. If $\Lambda = \mathbf{I}_r$, then the scaled envelope model is equivalent to the envelope model.

To select the dimension of $\mathcal{E}_{\Lambda^{-1} \boldsymbol{\Sigma} \Lambda^{-1}}(\Lambda^{-1} \mathcal{B})$, we can apply the following two functions

```
u = modelselectaic(X, Y, 'senv')
u = modelselectbic(X, Y, 'senv')
```

Notice that likelihood ratio testing can not be applied to the model selection of the scaled envelope model, only AIC and BIC can be used to select u . The possible values of u can be any integer from 0 to r . When $u = r$, the scaled envelope model is essentially the standard multivariate linear model. And when $u = 0$, it means that $\boldsymbol{\beta} = \mathbf{0}$, then the changes in \mathbf{Y} do not depend on \mathbf{X} . After obtaining u , we can fit the scaled envelope model by

```
ModelOutput = senv(X, Y, u)
```

The output `ModelOutput` is a list, which contains the MLE of $\boldsymbol{\beta}$, $\boldsymbol{\Sigma}$, Λ , Γ , Γ_0 , $\boldsymbol{\eta}$, $\boldsymbol{\Omega}$, $\boldsymbol{\Omega}_0$ and $\boldsymbol{\alpha}$, and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\boldsymbol{\beta})$, the asymptotic standard errors of $\boldsymbol{\beta}$, the ratios of the asymptotic

standard errors of the standard model versus the scaled envelope model for elements in $\boldsymbol{\beta}$, the number of parameters in the model, and the number of observations in the data. After model fitting, the users can perform the following post processing inference:

```
bootse = bootstrapse(X, Y, u, B, 'senv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'senv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'senv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)
```

These functions are used similarly as for the envelope model.

3.8. Envelope Model in the Predictor Space

The envelope model in the predictor space is used when the number of predictors is strictly larger than the number of the responses. It has the potential to have better prediction performance compared to the standard model, or even the partial least squares. In fact, in the population version, the envelope estimator in this case is equivalent to the partial least squares estimator, but in the sample version, its performance is normally more stable than the partial least squares estimator. We slightly change the formulation of the standard model as $\mathbf{Y} = \boldsymbol{\mu} + \boldsymbol{\beta}^T \mathbf{X} + \boldsymbol{\varepsilon}$, to be consistent with the notations in Cook et al. [2012]. Then the coordinate form of envelope model in the predictor space is as follows:

$$\mathbf{Y} = \boldsymbol{\mu} + \boldsymbol{\eta}^T \boldsymbol{\Omega}^{-1} \boldsymbol{\Gamma}^T \mathbf{X} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\Sigma}_X = \boldsymbol{\Gamma} \boldsymbol{\Omega} \boldsymbol{\Gamma}^T + \boldsymbol{\Gamma}_0 \boldsymbol{\Omega}_0 \boldsymbol{\Gamma}_0^T,$$

where $\boldsymbol{\mu} \in \mathbb{R}^r$ is the intercept, $\boldsymbol{\beta} = \boldsymbol{\Gamma} \boldsymbol{\Omega}^{-1} \boldsymbol{\eta} \in \mathbb{R}^{p \times r}$, $\boldsymbol{\Gamma} \in \mathbb{R}^{r \times u}$ spans the envelope subspace $\mathcal{E}_{\boldsymbol{\Sigma}_X}(\mathcal{B})$, and $\mathcal{B} = \text{span}(\boldsymbol{\beta}^T)$, $\boldsymbol{\Gamma}_0 \in \mathbb{R}^{r \times (r-u)}$ spans the orthogonal complement of $\mathcal{E}_{\boldsymbol{\Sigma}_X}(\mathcal{B})$, $\boldsymbol{\eta} \in \mathbb{R}^{u \times r}$, $\boldsymbol{\Omega} \in \mathbb{R}^{u \times u}$, and $\boldsymbol{\Omega}_0 \in \mathbb{R}^{(r-u) \times (r-u)}$ carry coordinates, and u is the dimension of the envelope $\mathcal{E}_{\boldsymbol{\Sigma}_X}(\mathcal{B})$. To select the dimension of $\mathcal{E}_{\boldsymbol{\Sigma}_X}(\mathcal{B})$, we apply the following three functions

```
u = modelselectaic(X, Y, 'xenv')
u = modelselectbic(X, Y, 'xenv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'xenv')
```

The possible values for u can be any integer from 0 to p , when $u = p$, the envelope model reduces to the standard model, and when $u = 0$, the changes in \mathbf{Y} do not depend on \mathbf{X} . After estimating u , we can fit the model by

```
ModelOutput = xenv(X, Y, u)
```

The output `ModelOutput` is a list, which contains the MLE of $\boldsymbol{\beta}$, $\boldsymbol{\Sigma}_X$, $\boldsymbol{\Gamma}$, $\boldsymbol{\Gamma}_0$, $\boldsymbol{\eta}$, $\boldsymbol{\Omega}$, $\boldsymbol{\Omega}_0$ and $\boldsymbol{\mu}$, and also statistics computed from the model, including the maximized log likelihood, the asymptotic covariance matrix of $\text{vec}(\boldsymbol{\beta})$, the asymptotic standard errors of $\boldsymbol{\beta}$, the ratios of the asymptotic standard errors of the standard model versus the envelope model for elements in $\boldsymbol{\beta}$, the number of parameters in the model, and the number of observations in the data. After model fitting, the following post processing inference can be performed:

```

bootse = bootstrapse(X, Y, u, B, 'xenv')
PredictOutput = prediction(ModelOutput, Xnew, 'estimation', 'xenv')
PredictOutput = prediction(ModelOutput, Xnew, 'prediction', 'xenv')
TestOutput = testcoefficient(ModelOutput, modelType, TestInput)

```

These functions are used similarly as for the envelope model in Section 3.2.

3.9. Envelope Model in the Predictor Space Using Partial Least Squares Algorithm

The envelope model in the predictor space using partial least squares algorithm module estimates exactly the same model as in section 3.8, but it estimates the envelope subspace $\mathcal{E}_{\Sigma_X}(\mathcal{B})$ using partial least squares. Then this case make the envelope model in the predictor space applicable to small sample cases, where n is less than p .

To select u , the dimension of $\mathcal{E}_{\Sigma_X}(\mathcal{B})$, we can use m-fold cross validation, for example, 5-fold cross validation:

```
u = mfoldcv(X, Y, m, 'xenvpls')
```

The possible values of u can be any integer from 0 to the minimum of $(m-1)n/m-1$ and p . When $u = p$, the envelope model degenerates to the standard model, and when $u = 0$, then $\beta = 0$, and the changes in \mathbf{Y} do not depend the changes in \mathbf{X} . If the sample size is large, we can also use the dimension selection tools for “xenv”, as “xenv” and “xenvpls” implement the same model.

```

u = modelselectaic(X, Y, 'xenv')
u = modelselectbic(X, Y, 'xenv')
alpha = 0.01;
u = modelselectlrt(X, Y, alpha, 'xenv')

```

After obtaining u , to fit the envelope model in the predictor space using partial least squares algorithm, we use

```
ModelOutput = xenvpls(X, Y, u)
```

The output ModelOutput is a list, which contains the estimators of β , Σ_X , Γ , Γ_0 , η , Ω , Ω_0 and μ , and also the number of parameters in the model and number of observations in the data. It does not contain any information based on the likelihood function, for example, the maximized log-likelihood, the asymptotic covariance matrix of $\text{vec}(\beta)$, the asymptotic standard errors of elements in β , etc. This is because `envpls` is mainly implemented for small sample size cases, under which the likelihood cannot be estimated. The estimator is obtained by partial least squares algorithm. We can check this by

```

ModelOutput = xenvpls(X, Y, u);
[XL, YL, XS, YS, BETA, PCTVAR, MSE, stats] = plsregress(X, Y, u);
subspace(ModelOutput.Gamma, stats.W)

```

The function `plsregress` is built in the Statistics toolbox in MATLAB, it does the partial least squares regression to the data. The function `subspace` computes the angles between the two subspaces, if the angle is 0, then the two subspaces are the same.

For inference, any inference that is based on the covariance matrix of $\text{vec}(\boldsymbol{\beta})$ such as estimation, prediction and test coefficients cannot be performed with “`xenvpls`”. But we can still estimate the bootstrap standard errors for $\hat{\boldsymbol{\beta}}$ by

```
bootse = bootstrapse(X, Y, u, B, 'xenvpls')
```

Optional Arguments

The computing of the envelope models involve Grassmann manifold optimization. The package `sg_min` 2.4.1 by Ross Lippert <http://web.mit.edu/~ripper/> uses the analytical first derivative and numerical second derivative of the objective function, and we find it very stable. The Grassmann manifold optimization in this toolbox is then based on `sg_min` 2.4.1. Because of the iterative nature of the optimization, we offer some optional arguments so that the users can control the convergence tolerance and speed, as well as monitoring the iteration process. In functions `modelselectaic`, `modelselectbic`, `modelselectlrt`, `env`, `henv`, `ienv`, `penv`, `senv`, `xenv` and `bootstrapse`, there is an optional input argument `Opts`. `Opts` is a list, which has five elements:

1. `Opts.maxIter`: This controls the maximum number of iterations. The default value is 300. The iterations will terminate once the maximum number of iterations is reached, then the results are based on the last iteration. At the same time, a warning

WARNING: reached maximum number of iterations without convergence for specified tolerances

is printed out.

2. `Opts.ftol`: This controls the tolerance parameter of the objective function. The default value is $1e-10$. The iteration will terminate once the tolerance condition for both the objective function and its derivative are reached.
3. `Opts.gradtol`: This controls the tolerance parameter of the derivative of the objective function. The default value is $1e-7$.
4. `Opts.verbose`: This is a flag for printing out the computing process. It is logical 0 or 1, with 0 for no print out and 1 for print out. The default value is 0. The print out will depend on the nature of the functions: with functions for dimension selection (`modelselectaic`, `modelselectbic` and `modelselectlrt`), the current dimension will be printed out; with functions for bootstrap (`bootstrapse`), the current number of bootstrap sample will be printed out; with the scaled envelope model (`senv`), the current number of iteration for the alternating algorithm between the scales and Grassmann manifold optimization is printed out; and with all the other functions `henv`, `ienv`, `penv`, `senv`, and `xenv`, the current number of iterations of the Grassmann manifold optimization is printed out.
5. `Opts.init`: This argument allows the users to input their starting values. If not specified, the starting values are generated by the functions `get_Init` or `get_Init4henv`. This argument is only applicable to `env`, `henv`, `ienv`, `penv`, `senv` and `xenv`.

The users can choose to define some or none of the five elements, if not defined, the default value will be used. If the users choose to define some of the elements, it is added as the last input of the function. For example, suppose the users define `Opts` for `env`, then the syntax of `env` will be

```
ModelOutput = env(X, Y, u, Opts)
```

We will conclude this chapter by a couple of examples.

Control Convergence Arguments

We use the wheat protein data and the function `env` as an example. First we set

```
Opts.verbose = 1
```

so that we can monitor the iteration process. We leave the other arguments as default values for now

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1:6);
u = 1;
ModelOutput = env(X, Y, u, Opts);
```

<i>iter</i>	<i>grad</i>	<i>F(Y)</i>
0	3.109420e+00	-1.166907e+00
1	1.448779e-01	-1.193292e+00
2	2.566091e-01	-1.193338e+00
3	2.434851e-02	-1.193379e+00
4	1.035130e-02	-1.193385e+00
5	1.259475e-02	-1.193385e+00
6	9.240358e-03	-1.193385e+00
7	1.436949e-02	-1.193386e+00
8	1.096367e-02	-1.193389e+00
9	1.643753e-04	-1.193391e+00
10	2.813965e-04	-1.193391e+00
11	8.046247e-05	-1.193391e+00
12	4.414260e-07	-1.193391e+00
13	1.774275e-07	-1.193391e+00

We notice that it takes 13 iterations till convergence. Suppose we set the maximum number of iterations as 10, then the algorithm stops at the tenth iteration and a warning is printed out at the end.

```
Opts.maxIter = 10;
ModelOutput = env(X, Y, u, Opts);
```

<i>iter</i>	<i>grad</i>	<i>F(Y)</i>
0	3.109420e+00	-1.166907e+00
1	1.448779e-01	-1.193292e+00
2	2.566091e-01	-1.193338e+00
3	2.434851e-02	-1.193379e+00
4	1.035130e-02	-1.193385e+00
5	1.259475e-02	-1.193385e+00
6	9.240358e-03	-1.193385e+00
7	1.436949e-02	-1.193386e+00
8	1.096367e-02	-1.193389e+00
9	1.643753e-04	-1.193391e+00

```
10          2.813965e-04          -1.193391e+00
WARNING: reached maximum number of iterations without convergence for
specified tolerances
```

Normally Grassmann manifold optimization takes hundreds of iterations to converge at the default tolerance parameters. We use this example because that we can print its complete iteration process. Now we change the tolerance level of convergence.

```
Opts.ftol = 1e-5;
Opts.gradtol = 1e-3;
ModelOutput = env(X, Y, u, Opts);
iter      grad      F(Y)
0         3.109420e+00    -1.166907e+00
1         1.448779e-01    -1.193292e+00
2         2.566091e-01    -1.193338e+00
3         2.434851e-02    -1.193379e+00
4         1.035130e-02    -1.193385e+00
5         1.259475e-02    -1.193385e+00
6         9.240357e-03    -1.193385e+00
7         1.436947e-02    -1.193386e+00
8         1.096356e-02    -1.193389e+00
9         1.573313e-04    -1.193391e+00
```

As we loosen the tolerance level, it requires less number of iterations till convergence. If the users want to specify their starting value, it can be done by

```
Opts.init = [1 0 0 0 0 0]';
ModelOutput = env(X, Y, u, Opts);
iter      grad      F(Y)
0         3.510184e+00    7.406482e+00
1         1.005063e+01    5.523686e+00
2         2.643888e+00    4.138687e+00
...
74        1.465095e-02    -1.193389e+00
75        1.066872e-04    -1.193391e+00
76        7.018152e-07    -1.193391e+00
```

From the result, we notice that a random starting value normally takes longer to convergence. Furthermore, according to our experience, it is more likely to be trapped in the local minimums, and therefore, should be avoided.

The Display Argument

In the previous example, the print out for `env` is the Grassmann manifold optimization process. In this example, we will show different print out types for different functions. We will still use the wheat protein data as the background for the first two cases.

```
load wheatprotein.txt
X = wheatprotein(:, 8);
Y = wheatprotein(:, 1:6);
```

1. With functions on dimension selection, the print out is the current dimension:

```

Opts.verbose = 1;
u = modelselectaic(X, Y, 'env', Opts);
Current dimension 0
Current dimension 1
Current dimension 2
Current dimension 3
Current dimension 4
Current dimension 5

```

Suppose the number of responses is r , the print out will start with dimension 0 and end with dimension $r - 1$. The case of $u = r$ is computed but not printed, because it takes little time, compared to the Grassmann manifold optimization process.

2. With function `bootstrapse`, the print out is the current number of bootstrap sample.

```

u = 1;
B = 15;
Opts.verbose = 1;
bootse = bootstrapse(X, Y, u, B, 'env', Opts);
Current number of bootstrap sample 1
Current number of bootstrap sample 2
Current number of bootstrap sample 3
Current number of bootstrap sample 4
Current number of bootstrap sample 5
Current number of bootstrap sample 6
Current number of bootstrap sample 7
Current number of bootstrap sample 8
Current number of bootstrap sample 9
Current number of bootstrap sample 10
Current number of bootstrap sample 11
Current number of bootstrap sample 12
Current number of bootstrap sample 13
Current number of bootstrap sample 14
Current number of bootstrap sample 15

```

The print out will start from the first bootstrap sample till the last bootstrap sample.

3. With function `senv`, the print out is the alternating algorithm between the optimization of the scaling parameters and Grassmann manifold.

```

load('sales.txt')
Y = sales(:, 4 : 7);
X = sales(:, 1 : 3);
u = 1;
Opts.verbose = 1;
ModelOutput = senv(X, Y, u, Opts);
Current number of iterations 1
Current number of iterations 2
Current number of iterations 3
Current number of iterations 4

```

```

Current number of iterations 5
Current number of iterations 6
Current number of iterations 7
Current number of iterations 8
Current number of iterations 9
Current number of iterations 10
Current number of iterations 11
Current number of iterations 12
Current number of iterations 13
Current number of iterations 14
Current number of iterations 15
Current number of iterations 16
Current number of iterations 17
Current number of iterations 18
Current number of iterations 19
Current number of iterations 20
Current number of iterations 21
Current number of iterations 22
Current number of iterations 23
Current number of iterations 24
Current number of iterations 25
Current number of iterations 26

```

The maximum number of iteration of the alternating algorithm is 1000. So the print out can run from the first iteration to the thousandth iteration.

4. With other functions, the print out is the Grassmann manifold optimization process as in `env`.

```

load irisf.mat
d = 1;
Opts.verbose = 1;
ModelOutput = ienv(X, Y, d, Opts)

```

<i>iter</i>	<i>grad</i>	<i>F(Y)</i>
0	2.043243e-01	1.808874e-02
1	1.102662e-01	1.399590e-02
2	3.082806e-02	7.353045e-03
3	1.755330e-03	7.285612e-03
4	1.312911e-03	7.285253e-03
5	3.840448e-04	7.284642e-03
6	1.174194e-06	7.284631e-03
7	1.462019e-06	7.284631e-03
8	9.358198e-08	7.284631e-03
9	1.478274e-08	7.284631e-03
10	1.952908e-08	7.284631e-03

Bibliography

- R.D. Cook. Lecture Notes on Dimension Reduction. School of Statistics, University of Minnesota, Minneapolis, 2012.
- R.D. Cook and Z. Su. Scaled envelopes: Scale invariant and efficient estimation in multivariate linear regression. To be submitted shortly, 2012.
- R.D. Cook, B. Li, and F. Chiaromonte. Envelope models for parsimonious and efficient multivariate linear regression (with discussion). *Statist. Sinica*, 20:927–1010, 2010.
- R.D. Cook, I. Helland, and Z. Su. Envelopes and partial least squares regression. Submitted, 2012.
- R.A. Johnson and D.W. Wichern. *Applied Multivariate Statistical Analysis*. Upper Saddle River, NJ: Prentice Hall, 2007.
- Z. Su and R.D. Cook. Partial envelopes for efficient estimation in multivariate linear regression. *Biometrika*, 98(1):133–146, 2011. ISSN 0006-3444.
- Z. Su and R.D. Cook. Estimation of Multivariate Means with Heteroscedastic Errors Using Envelope Models. *Statistica Sinica*, to appear, 2012.
- Z. Su and R.D. Cook. Inner Envelopes: Efficient Estimation in Multivariate Linear Regression. *Biometrika*, to appear.