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Chapters 5.6-5.10 & 6.4 of Probability, Statistics, and Random Processes by A. Leon-Garcia

- 1. Let X and Y be independent random variables that are uniformly distributed in the interval [0,1].
 - (a) Find the pdf of A = X + Y.

Solution: Let $F_U(u)$ and $f_U(u)$ be the cdf and pdf of the uniform random variable. Let $F_A(a)$ and $f_A(a)$ be the cdf and pdf of A. As such,

$$F_A(a) = P(X + Y \le a)$$

$$= \int_{-\infty}^{\infty} P(X + y \le a | Y = y) f_U(y) dy$$

$$= \int_{-\infty}^{\infty} P(X \le a - y | Y = y) f_U(y) dy$$

$$= \int_{-\infty}^{\infty} F_U(a - y) f_U(y) dy$$

By taking the derivative with respect to a, we get

$$f_A(a) = \frac{dF_A(a)}{da}$$
$$= \int_{-\infty}^{\infty} f_U(a-y) f_U(y) dy$$

which is the convolution of the two probability distributions. We note that

$$f_U(u) = \begin{cases} 1 & 0 \le u \le 1\\ 0 & \text{otherwise} \end{cases}$$

and

$$f_U(a-u) = \begin{cases} 1 & 0 \le a - u \le 1 \\ 0 & \text{otherwise.} \end{cases}$$

Therefore, the only region that $f_U(a-y)f_U(y)$ is non-zero is $\max(a-1,0) \le y \le \min(a,1)$. From here, we get

$$f_A(a) = \begin{cases} a & 0 \le a \le 1\\ 2 - a & 1 \le a \le 2\\ 0 & \text{otherwise.} \end{cases}$$

(b) Find the pdf of B = X - Y.

Solution: We get a similar solution to part (a) except that the pdf of B is

$$f_B(b) = \int_{-\infty}^{\infty} f_U(b+y) f_U(y) dy.$$

As such, we get

$$f_B(b) = \begin{cases} 1 + b & -1 \le b \le 0 \\ 1 - b & 0 \le b \le 1 \\ 0 & \text{otherwise.} \end{cases}$$

(c) Find the pdf of C = XY.

Solution: As in part (a), we calculate the cdf of C to find the pdf. Let $F_C(c)$ and $f_C(c)$ be the cdf and pdf of C.

$$F_C(c) = P(XY \le c)$$

$$= \int_{-\infty}^{\infty} P(Xy \le c | Y = y) f_U(y) dy$$

$$= \int_{-\infty}^{\infty} P(X \le \frac{c}{y} | Y = y) f_U(y) dy$$

$$= \int_{-\infty}^{\infty} F_U(\frac{c}{y}) f_U(y) dy$$

We note that the third step is only valid because Y is non-negative so multiplying both sides by y does not change the sign.

By taking the derivative with respect to c, we get

$$f_C(c) = \frac{dF_C(c)}{dc}$$
$$= \int_{-\infty}^{\infty} f_U(\frac{c}{y}) f_U(y) \frac{1}{y} dy$$

Consider the case when $0 \le c \le 1$ since all other cases must have zero for the pdf. As such,

$$f_C(c) = \int_{-\infty}^{\infty} f_U(\frac{c}{y}) f_U(y) \frac{1}{y} dy$$
$$= \int_c^1 \frac{1}{y} dy$$
$$= \ln(y) |_c^1 = \ln(1) - \ln(c) = -\ln(c)$$

Hence, we get

$$f_C(c) = \begin{cases} -\ln(c) & 0 < c \le 1\\ 0 & \text{otherwise.} \end{cases}$$

Another way to get this result is to calculate the cdf first and then differentiate. Again, consider the case when $0 \le c \le 1$ which results in

$$F_C(c) = \int_{-\infty}^{\infty} F_U(\frac{c}{y}) f_U(y) dy$$
$$= \int_0^c dy + \int_c^1 \frac{c}{y} dy$$
$$= c - c \ln(c).$$

By taking the derivative with respect to c, we get the same result.

(d) Find the covariance of A and B.

Solution:

$$Cov(A, B) = \mathbb{E}[AB] - \mathbb{E}[A]\mathbb{E}[B]$$

$$= \mathbb{E}[(X+Y)(X-Y)] - \mathbb{E}[(X+Y)]\mathbb{E}[(X-Y)]$$

$$= \mathbb{E}[X^2 - Y^2] - \mathbb{E}[(X+Y)]\mathbb{E}[(X-Y)]$$

$$= 0$$

(e) Find the covariance of A and C.

Solution:

$$Cov(A, C) = \mathbb{E}[AC] - \mathbb{E}[A]\mathbb{E}[C]$$

$$= \mathbb{E}[(X+Y)XY] - \mathbb{E}[(X+Y)]\mathbb{E}[XY]$$

$$= \mathbb{E}[X^{2}Y + XY^{2}] - \mathbb{E}[(X+Y)]\mathbb{E}[XY]$$

$$= \mathbb{E}[X^{2}Y] + \mathbb{E}[XY^{2}] - \mathbb{E}[(X+Y)]\mathbb{E}[XY]$$

$$= \mathbb{E}[X^{2}]\mathbb{E}[Y] + \mathbb{E}[X]\mathbb{E}[Y^{2}] - \mathbb{E}[X]^{2}\mathbb{E}[Y] - \mathbb{E}[Y]^{2}\mathbb{E}[X]$$

$$= \mathbb{E}[Y] \cdot (\mathbb{E}[X^{2}] - \mathbb{E}[X]^{2}) + \mathbb{E}[X] \cdot (\mathbb{E}[Y^{2}] - \mathbb{E}[Y]^{2})$$

$$= \mathbb{E}[Y] \cdot Var(X) + \mathbb{E}[X] \cdot Var(Y)$$

$$= \frac{1}{2} \cdot \frac{1}{12} + \frac{1}{2} \cdot \frac{1}{12} = \frac{1}{12}$$

- 2. Let X be a Gaussian random variable with mean 0 and variance $\sigma^2 = 1$. We define three new random variables as Y = aX + b, $Z = X^2$, and $W = X^3$. We require that $a \neq 0$. For reference, $\mathbb{E}[X^3] = 0$, $\mathbb{E}[X^4] = 3\sigma^4$, $\mathbb{E}[X^5] = 0$, and $\mathbb{E}[X^6] = 15\sigma^6$.
 - (a) Find the correlation coefficients for each pair (X, Y), (X, Z), and (X, W). **Solution:** First, we calculate the expectation of each random variable which are $\mathbb{E}[X] = 0 = \mathbb{E}[Y] = \mathbb{E}[W]$ and $\mathbb{E}[Z] = \sigma^2$.

Next, we calculate the variances of each random variable:

$$Var(X) = \sigma^{2} = 1$$

$$Var(Y) = Var(aX + b) = a^{2}Var(X) = a^{2}$$

$$Var(Z) = Var(X^{2}) = \mathbb{E}[X^{4}] - \mathbb{E}[X^{2}]^{2}$$

$$= 3\sigma^{4} - \sigma^{4} = 2\sigma^{4} = 2$$

$$Var(W) = Var(X^{3}) = \mathbb{E}[X^{6}] - \mathbb{E}[X^{3}]^{2}$$

$$= 15\sigma^{6} = 15.$$

Now, we calculate the covariance of X to all the other terms.

$$\begin{split} Cov(X,Y) &= \mathbb{E}[X(aX+b)] - \mathbb{E}[aX+b]\mathbb{E}[X] \\ &= a\mathbb{E}[X^2] - a\mathbb{E}[X]^2 = aVar(X) = a \\ Cov(X,Z) &= \mathbb{E}[XX^2] - \mathbb{E}[X]\mathbb{E}[X^2] \\ &= \mathbb{E}[X^3] - \mathbb{E}[X]\mathbb{E}[X^2] = 0 \\ Cov(X,W) &= \mathbb{E}[XX^3] - \mathbb{E}[X]\mathbb{E}[X^3] \\ &= \mathbb{E}[X^4] - \mathbb{E}[X]\mathbb{E}[X^3] = 3\sigma^4 = 3. \end{split}$$

Finally, we can calculate the correlation coefficients:

$$\rho_{X,Y} = \frac{Cov(X,Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}$$

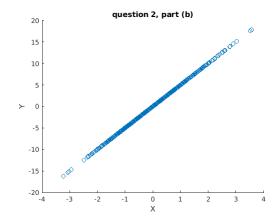
$$= \frac{a}{\sqrt{a^2}} = \frac{a}{|a|} = \begin{cases} 1 & a > 0\\ -1 & a < 0 \end{cases}$$

$$\rho_{X,Z} = \frac{0}{\sqrt{2\sigma^4}} = 0$$

$$\rho_{X,W} = \frac{3}{\sqrt{15}}$$

(b) For the case of a = 5 and b = 0, randomly sample X 1000 times, use each X to get Y, and use a scatter plot to plot (X, Y) where the values of X is along the x-axis and the values of Y are along the y-axis. Describe how the correlation coefficient relates to what you see in the scatter plot.

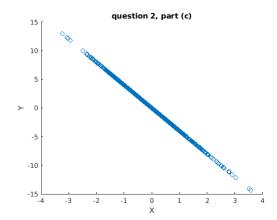
Solution:



We can see that having a positive correlation coefficient implies that X and Y are positively correlated which means that when one increases, the other one is also likely to increase. Even more strongly, when the coefficient is 1, it means that X and Y actually have a linear relationship to each other. That is why you see a straight line in the figure.

(c) For the case of a = -4 and b = 0, randomly sample X 1000 times, use each X to get Y, and use a scatter plot to plot (X, Y) where the values of X is along the x-axis and the values of Y are along the y-axis. Describe how the correlation coefficient relates to what you see in the scatter plot and explain the difference to part (b).

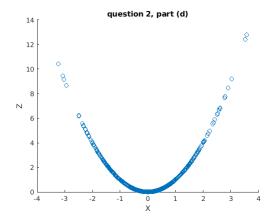
Solution:



Similarly, we can see that having a negative correlation coefficient implies that X and Y are negativel correlated which means that when one increases, the other one is likely to decrease. Additionally,we see that when the coefficient is -1, X and Y also follow a linear relationship to each other only negatively this time.

(d) Now, randomly sample X 1000 times, use each X to get Z, and use a scatter plot to plot (X, Z) where the values of X is along the x-axis and the values of Z are along the y-axis. Describe how the correlation coefficient relates to what you see in the scatter plot and explain the difference to parts (b) and (c).

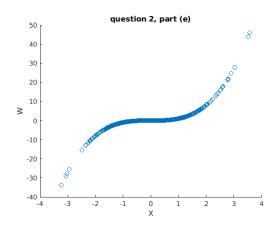
Solution:



Now, we can see that the plot has Z both decreasing and increasing as X increases. This is what is implied by the correlation coefficient being 0. As such, when the coefficient is 0, Z can, on average, either decrease or increase as X increases.

(e) Now, randomly sample X 1000 times, use each X to get W, and use a scatter plot to plot (X, W) where the values of X is along the x-axis and the values of W are along the y-axis. Describe how the correlation coefficient relates to what you see in the scatter plot.

Solution:



Here we see that the correlation coefficient is positive but not 1. As such, we can clearly see a positive trend since W increases as X increases. Yet, the trend is clearly not linear which is why the correlation coefficient is not 1.

- 3. Let X and Y be two random variables with identical distributions. These two random variables are not necessarily independent. Answer the following questions given that C = aX + bY and D = aX bY.
 - (a) Find COV[C, D] in terms of the variances and covariances of X and Y Solution:

$$COV[C, D] = COV[aX + bY, aX - bY]$$

$$= E[(aX + bY)(aX - bY)] - E[aX + bY]E[aX - bY]$$

$$= E[a^{2}X^{2} - b^{2}Y^{2}] - (aE[X] + bE[Y])(aE[X] - bE[Y])$$

$$= a^{2}E[X^{2}] - b^{2}E[Y^{2}] - a^{2}E[X]^{2} + b^{2}E[Y]^{2}$$

$$= a^{2}VAR[X] - b^{2}VAR[Y]$$

$$= (a^{2} - b^{2}) * VAR[X]$$

(b) Find the relation between a and b if the random variables C and D are independent Solution:

Since it is given that C and D are independent, therefore, COV[C, D] = 0.

$$COV[C, D] = (a^2 - b^2) * VAR[X]$$

 $0 = (a^2 - b^2) * VAR[X]$

Since Var(X) > 0 (since X is a non-constant random variable), we must have either a + b = 0 or a - b = 0. If a + b = 0, then a = -b, and if a - b = 0, then a = b. Therefore, $a = \pm b$

- 4. Answer the following. Show all your work.
 - (a) Consider two random variables X and Y. Prove that the correlation coefficient $\rho_{X,Y}$ satisfies $-1 \le \rho_{X,Y} \le 1$.

Solution:

$$0 \leq \mathbb{E}\left[\left(\frac{X - \mathbb{E}[X]}{\sigma_X} \pm \frac{Y - \mathbb{E}[Y]}{\sigma_Y}\right)^2\right]$$

$$= \mathbb{E}\left[\left(\frac{X - \mathbb{E}[X]}{\sigma_X}\right)^2\right] \pm 2\mathbb{E}\left[\frac{(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])}{\sigma_X\sigma_Y}\right] + \mathbb{E}\left[\left(\frac{Y - \mathbb{E}[Y]}{\sigma_Y}\right)^2\right]$$

$$= 1 \pm 2\rho_{X,Y} + 1 = 2(1 \pm \rho_{X,Y})$$

$$\implies -1 < \rho_{X,Y} < 1$$

(b) Let X be a random variable, and Y be another random variable given by Y = aX + b. What is the correlation coefficient between X and Y. Does the answer depend on the sign of a?

Solution:

The correlation coefficient is given by:

$$\rho_{X,Y} = \frac{E[XY] - E[X]E[Y]}{\sigma_X \sigma_Y}.$$

We have that,

$$E[XY] = E[aX^{2} + bX] = aE[X^{2}] + bE[X],$$

$$E[X]E[Y] = E[X](aE[X] + b) = aE[X]^{2} + bE[X],$$

$$COV[X, Y] = E[XY] - E[X]E[Y] = aE[X^{2}] + bE[X] - aE[X]^{2} - bE[X] = aVAR[X],$$

Next, $\sigma_Y = \sqrt{VAR[aX + b]} = \sqrt{a^2VAR[X]} = a\sigma_X$, if a > 0 and $-a\sigma_X$, if a < 0. and,

$$\sigma_X \sigma_Y = \pm a \sigma_X^2 = \pm a VAR[X].$$

Thus, if $VAR[X] \neq 0$,

$$\rho_{X,Y} = \begin{cases} \frac{aVar[X]}{aVar[X]} = 1 \text{ if } a > 0\\ \frac{aVar[X]}{-aVar[X]} = -1 \text{ if } a < 0 \end{cases}.$$

Clearly, the sign of a does matter.

5. Let X and Y be jointly Gaussian random variables with $\mathbb{E}[Y] = 0$, $\sigma_X = 4$, $\sigma_Y = 3$ and $\mathbb{E}[X|Y] = \frac{4Y}{9} + 2$. Find the joint pdf of X and Y.

Solution:

Let m_x and m_y be the means of X and Y. By the definition of marginal distributions for jointly Gaussian random variables, we know that

$$\mathbb{E}[X|Y] = \rho_{X,Y} \frac{\sigma_X}{\sigma_Y} (y - m_y) + m_x.$$

As such,

$$\mathbb{E}[X|Y] = \frac{4Y}{9} + 2 \implies m_x = 2, \rho_{X,Y} = \frac{1}{3}.$$

Now we have all the parameters to find the joint pdf of X and Y which is

$$f_{X,Y}(x,y) = \frac{\exp\left[\frac{1}{2(1-\rho_{X,Y}^2)}\left(\left(\frac{x-m_x}{\sigma_X}\right)^2 - 2\rho_{X,Y}\left(\frac{x-m_x}{\sigma_X}\right)\left(\frac{y-m_y}{\sigma_Y}\right) + \left(\frac{y-m_y}{\sigma_Y}\right)^2\right)\right]}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho_{X,Y}^2}}$$

$$= \frac{\exp\left[\frac{1}{2\left(\frac{8}{9}\right)}\left(\left(\frac{x-2}{4}\right)^2 - \frac{(x-2)(y)}{18} + \left(\frac{y}{3}\right)^2\right)\right]}{2\pi(12)\sqrt{\frac{8}{9}}}$$

$$= \frac{\exp\left[\frac{9}{16}\left(\left(\frac{x-2}{4}\right)^2 - \frac{(x-2)y}{18} + \left(\frac{y}{3}\right)^2\right)\right]}{16\pi\sqrt{2}}$$

6. Q3 from HW6

Let X and Y be two jointly continuous random variables with joint pdf

$$f_{XY}(x,y) = \begin{cases} 6xy, & 0 \le x \le 1, 0 \le y \le \sqrt{x}, \\ 0, & \text{otherwise,} \end{cases}$$

(a) Find $f_X(x)$.

Solution:

For $0 \le x \le 1$,

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) dy$$
$$= \int_{0}^{\sqrt{x}} 6xy dy$$
$$= 3x^2$$

Thus

$$f_X(x) = \begin{cases} 3x^2, & 0 \le x \le 1\\ 0, & \text{otherwise.} \end{cases}$$

(b) Find the conditional pdf of X given Y = y, $f_{X|Y}(x|y)$. Solution:

For $0 \le y \le 1$,

$$f_Y(y) = \int_{-\infty}^{\infty} f_{XY}(x, y) dx$$
$$= \int_{y^2}^{1} 6xy dx$$
$$= 3y(1 - y^4)$$

Thus

$$f_Y(y) = \begin{cases} 3y(1 - y^4), & 0 \le y \le 1 \\ 0, & \text{otherwise,} \end{cases}$$

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

$$= \begin{cases} \frac{2x}{1 - y^4} & y^2 \le x \le 1 \\ 0, & \text{otherwise.} \end{cases}$$

(c) Find E[X|Y=y], for $0 \le y < 1$. What is E[X|Y]? Solution:

For $0 \le y < 1$

$$E[X|Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$
$$= \int_{y^2}^{1} x \frac{2x}{1 - y^4} dx$$
$$= \frac{2(1 - y^6)}{3(1 - y^4)}$$

$$E[X|Y] = \frac{2(1-Y^6)}{3(1-Y^4)}$$

(d) Let A be the event $\{X \geq \frac{1}{2}\}$. Find P[A], $f_{X|A}(x)$, and E[X|A]. Solution:

$$P(A) = \int_{A} f_{X}(x)dx = \int_{\frac{1}{2}}^{1} 3x^{2}dx = \frac{7}{8}$$

$$f_{X|A}(x|A) = \begin{cases} \frac{f_{X}(x)}{P[A]} & \text{if } x \in A\\ 0 & \text{otherwise} \end{cases}$$

$$f_{X|A}(x|A) = \begin{cases} \frac{24x^{2}}{7} & \text{if } \frac{1}{2} \le x \le 1\\ 0 & \text{otherwise} \end{cases}$$

Therefore,

$$E[X|A] = \int_{\infty}^{\infty} x f_{X|A}(x|A) dx = \int_{\frac{1}{2}}^{1} x \frac{24x^{2}}{7} dx = \left[\frac{6x^{4}}{7}\right]_{\frac{1}{2}}^{1} = \frac{45}{56}$$

Appendix: Reference code for Q2

```
(a) % Part (b)
 _{2} X = randn(1000,1);
 a = 5;
 _{4} Y = a*X;
 5 figure ()
 6 scatter(X,Y)
 7 xlabel("X")
   ylabel ("Y")
  title ("question 2, part (b)")
10 % Part (c)
_{11} X = randn(1000,1);
a = -4;
_{13} Y = a*X;
14 figure ()
  scatter(X,Y)
   xlabel("X")
   ylabel ("Y")
   title ("question 2, part (c)")
  % Part (d)
_{21} X = randn(10000,1);
Z = X.^2;
  figure()
  scatter(X,Z)
   xlabel("X")
  ylabel ("Z")
   title ("question 2, part (d)")
  % Part (e)
_{30} X = randn(100,1);
^{31} W = X.^{3};
  figure()
  scatter (X,W)
  xlabel("X")
   ylabel ("W")
   title ("question 2, part (e)")
  %%%
38
```