

# Stochastic MCMC techniques

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<http://bayesgroup.ru>



# What stochasticity we are talking about?

We want to sample from posterior distribution

$$p(\theta|X) \propto p(\theta) \prod_{i=1}^N p(x_i|\theta)$$


Full dataset!

How to use **minibatches** instead of full dataset?

# Minibatch MCMC techniques

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# Variational Bayes vs. MCMC

	MCMC	IPM	Variational Bayes
Bias	No	??	Strong
Sampling/Ensembling	Inefficient	??	Efficient
Density	No	??	Yes
Likelihood	Needed	??	Needed

# Metropolis-Hastings

$$\alpha(\theta, \theta') = \frac{p(\theta'|X)q(\theta|\theta')}{p(\theta|X)q(\theta'|\theta)}$$

$$\alpha(\theta, \theta') = \frac{p_0(\theta') \prod_{i=1}^N p(x_i|\theta')q(\theta|\theta')}{p_0(\theta) \prod_{i=1}^N p(x_i|\theta)q(\theta'|\theta)}$$

$p_0(\theta)$  – prior distribution

Accept  $\theta'$  if

$$\alpha(\theta, \theta') > u, \quad u \sim \text{Uniform}[0,1]$$



Full dataset!

# An Efficient Minibatch Acceptance Test for Metropolis-Hastings

*Deniel Seita, Xinlei Pan, Haoyu Chen, Jhon Canny*

# Barker lemma

$$\Delta(\theta, \theta') = \log \frac{p_0(\theta') \prod_{i=1}^N p(x_i | \theta') q(\theta | \theta')}{p_0(\theta) \prod_{i=1}^N p(x_i | \theta) q(\theta' | \theta)}$$

For any function  $g(s)$  such that  $g(s) = \exp(s) g(-s)$ ,  
 $\alpha(\theta, \theta') \triangleq g(\Delta(\theta, \theta'))$  satisfies detailed balance.

# What does it mean?

$$\Delta(\theta, \theta') = \log \frac{p_0(\theta') \prod_{i=1}^N p(x_i | \theta') q(\theta | \theta')}{p_0(\theta) \prod_{i=1}^N p(x_i | \theta) q(\theta' | \theta)}$$

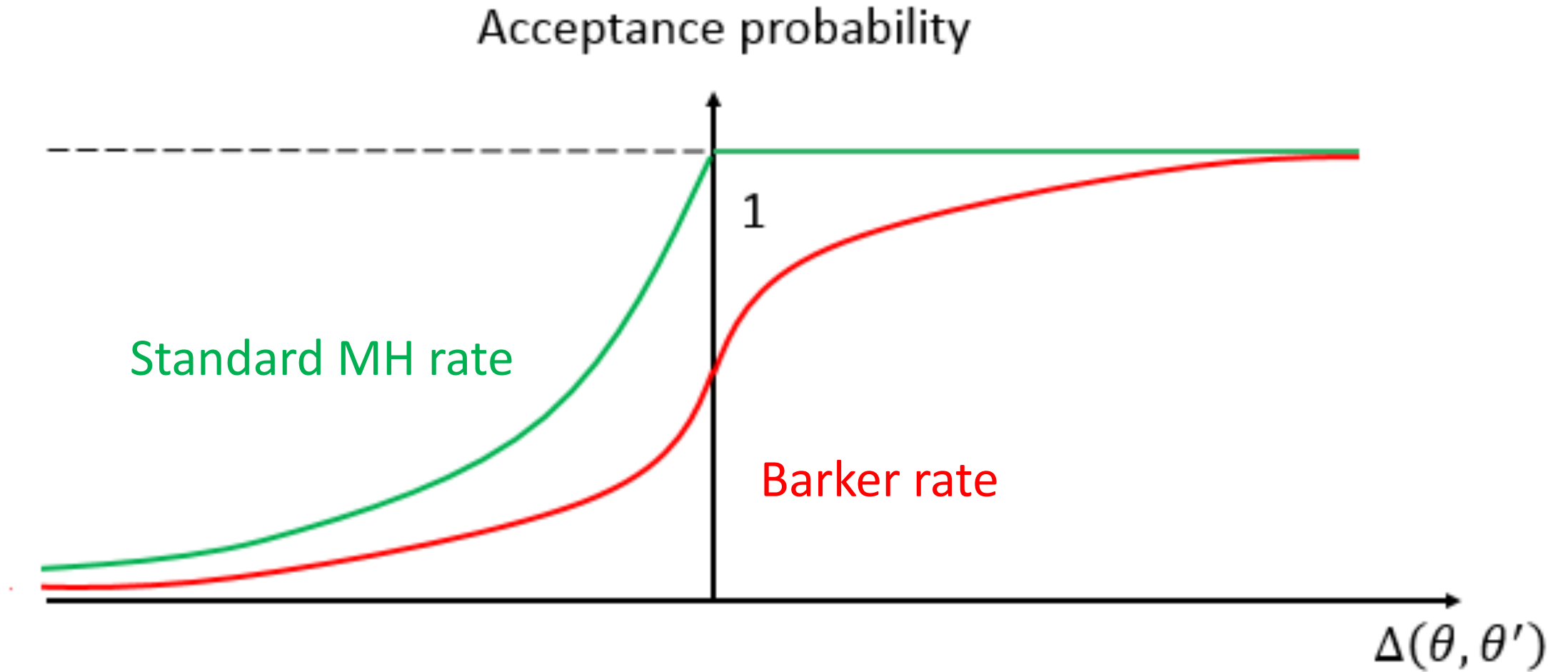
If  $g$  satisfies Barker lemma, then performing the test

$$g(\Delta(\theta, \theta')) > u, \quad u \sim \text{Uniform}[0,1]$$

we sample from true posterior distribution!



# Acceptance rate



# Barker acceptance function

Let  $g(s) = \frac{1}{1+\exp(-s)}$ , then test

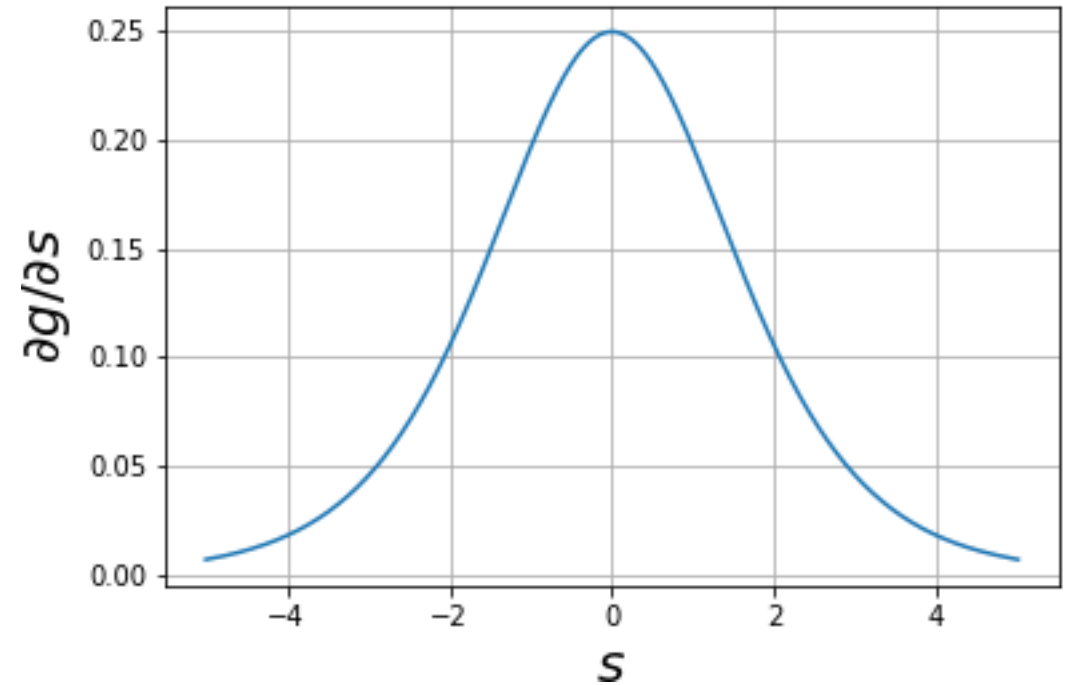
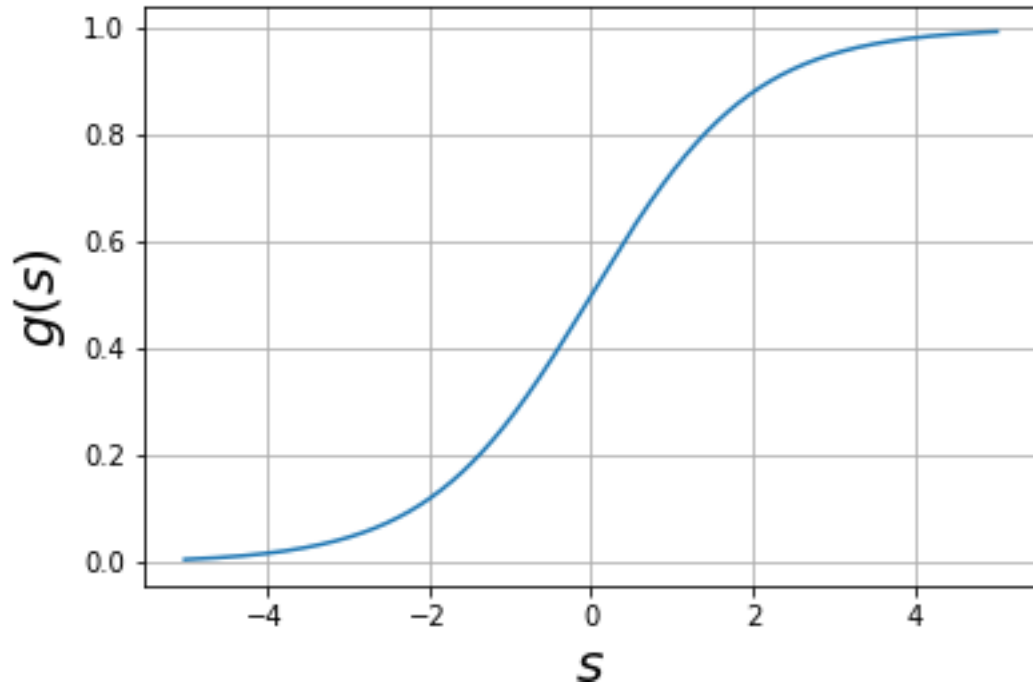
$$g(\Delta(\theta, \theta')) > u, \quad u \sim \text{Uniform}[0,1]$$

satisfies detailed balance and

$$\Delta(\theta, \theta') > X = g^{-1}(u), \quad u \sim \text{Uniform}[0,1]$$

also satisfies detailed balance

$g^{-1}(u)$  – sample from logistic distribution



$$X = X_{log} \sim \text{Logistic}(0,1)$$

$$-X_{log} \sim \text{Logistic}(0,1)$$

# New acceptance test

$$\Delta(\theta, \theta') = \sum_i^N \log \frac{p(x_i|\theta')}{p(x_i|\theta)} + \log \frac{p_0(\theta')q(\theta|\theta')}{p_0(\theta)q(\theta'|\theta)}$$

Accept  $\theta'$  if

$$\Delta(\theta, \theta') > X = g^{-1}(u), \quad u \sim \text{Uniform}[0,1]$$

Or equivalently

$$\Delta(\theta, \theta') + X_{log} > 0, \quad X_{log} \sim \text{Logistic}(0,1)$$

Exact, but we still use full dataset to sample one point

# Minibatch acceptance test

$$\Delta^*(\theta, \theta') = \frac{N}{b} \sum_{i=1}^b \log \frac{p(x_i|\theta')}{p(x_i|\theta)} + \log \frac{p_0(\theta')q(\theta|\theta')}{p_0(\theta)q(\theta'|\theta)}$$

$$\Delta^* = \Delta + X_{\text{norm}}, \quad X_{\text{norm}} \sim \overline{\mathcal{N}}(0, \sigma^2(\Delta^*))$$

$X_{\text{norm}}$  – approximately normal distribution (Central Limit Theorem)

$$\Delta_i = N \log \frac{p(x_i|\theta')}{p(x_i|\theta)} + \log \frac{p_0(\theta')q(\theta|\theta')}{p_0(\theta)q(\theta'|\theta)}$$

$$\sigma^2(\Delta^*) = \sum_{i=1}^b (\Delta_i - \overline{\Delta})^2$$

How to use  $\Delta^*$  instead of  $\Delta$ ?

$$\Delta(\theta, \theta') = \sum_i^N \log \frac{p(x_i|\theta')}{p(x_i|\theta)} + \log \frac{p_0(\theta')q(\theta|\theta')}{p_0(\theta)q(\theta'|\theta)}$$

Accept  $\theta'$  if

$$\Delta(\theta, \theta') + X_{log} > 0, \quad X_{log} \sim \text{Logistic}(0,1)$$

Our  
current  
test

But for minibatches we have value

$$\Delta^* = \Delta + X_{\text{norm}}, \quad X_{\text{norm}} \sim \overline{\mathcal{N}}(0, \sigma^2(\Delta^*))$$

# Logistic noise decomposition

Let's decompose

$$X_{log} = X_{norm} + X_{corr}, \quad X_{norm} \sim \mathcal{N}(0, \sigma^2),$$

where  $X_{corr}$  – correction distribution with PDF  $\mathcal{C}_\sigma(x)$

If

$$\Delta^* = \Delta + X_{norm}, \quad X_{norm} \sim \mathcal{N}(0, \sigma^2),$$

Not true!



Then

$$\Delta + X_{log} = \underbrace{\Delta + X_{norm}}_{\Delta^*} + X_{corr} = \Delta^* + X_{corr}$$

# Big picture

1. Evaluate

$$\Delta^*(\theta, \theta') = \frac{N}{b} \sum_{i=1}^b \log \frac{p(x_i|\theta')}{p(x_i|\theta)} + \log \frac{p_0(\theta')q(\theta|\theta')}{p_0(\theta)q(\theta'|\theta)}$$

2. Sample

$$X_{corr} \sim \text{Correction Distribution}(\sigma^2(\Delta^*))$$

3. Accept  $\theta'$  if

$$\Delta^* + X_{corr} > 0$$

4. Otherwise repeat  $\theta$



# We still have some questions

- How to sample from correction distribution?
- What error we have if we assume that

$$\Delta^* = \Delta + X_{norm}, \quad X_{norm} \sim \mathcal{N}(0, \sigma^2(\Delta^*))$$

Instead of

$$\Delta^* = \Delta + X_{norm}, \quad X_{norm} \sim \overline{\mathcal{N}}(0, \sigma^2(\Delta^*))$$

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# PDF of the correction distribution

$$X_{log} = X_{norm} + X_{corr}, \quad X_{norm} \sim \mathcal{N}(0, \sigma^2)$$

$\Phi_\sigma$  – CDF of  $\mathcal{N}(0, \sigma^2)$

$\mathcal{C}_\sigma$  – PDF of corresponding correction distribution

$\cdot * \cdot$  – convolution operation

$$CDF(X_{norm} + X_{corr}) = \Phi_\sigma * \mathcal{C}_\sigma$$

$$\mathbb{P}\{X + Y < t\} = \int_{-\infty}^{+\infty} dx \int_{-\infty}^{t-x} p_x(x) p_y(y) dy = \int_{-\infty}^{+\infty} dx p_x(x) F_y(t - x)$$

# PDF of the correction distribution

$$C_{\sigma}^* = \operatorname{argmin}_{C_{\sigma}} \sup |\Phi_{\sigma} * C_{\sigma} - S|$$

$S$  – CDF of Logistic Distribution

After discretization on the uniform grid  $[-20, 20]$

$$C_{\sigma}^* = \operatorname{argmin}_{C_{\sigma}} \max_{i \in I} \left| \sum_{j \in J} \Phi_{\sigma}(X_i - Y_j) C_{\sigma}(Y_j) - S(X_i) \right|$$

Define

$$M_{ij} = \Phi_{\sigma}(X_i - Y_j), \quad u_j = C_{\sigma}(Y_j), \quad v_i = S(X_i)$$

Then

$$u^* = (M^T M + \lambda I)^{-1} M^T v$$

# PDF of the correction distribution

$S$  – CDF of Logistic Distribution

After discretization on the uniform grid  $[-20, 20]$

$$C_{\sigma}^* = \underset{C_{\sigma}}{\operatorname{argmin}} \left\| \sum_{j \in J} \Phi_{\sigma}(X_i - Y_j) C_{\sigma}(Y_j) - S(X_i) \right\|_2^2 + \lambda \sum_j C_{\sigma}(Y_j)^2$$

Define

$$M_{ij} = \Phi_{\sigma}(X_i - Y_j), \quad u_j = C_{\sigma}(Y_j), \quad v_i = S(X_i)$$

Then

$$u^* = (M^T M + \lambda E)^{-1} M^T v$$

# PDF of the correction distribution

But instead of solving

$$u^* = \operatorname{argmin}_u \max_{i \in I} |Mu - v|, \quad u > 0$$

Let's solve

$$u^* = \operatorname{argmin}_u \|Mu - v\|_2^2 + \lambda \|u\|_2^2$$

$$u^* = (M^T M + \lambda I)^{-1} M^T v$$

And show empirically that error is negligible

# Precomputing correction distribution

Note that PDF  $C_\sigma(x)$  depends on variance  $\sigma^2$  of normal distribution  $\mathcal{N}(0, \sigma^2)$

$$\underbrace{\Delta + \mathcal{N}(0, \sigma^2(\Delta^*))}_{\approx \Delta^*} + X_{corr}(\sigma^2(\Delta^*)) =$$
$$\Delta + \underbrace{\mathcal{N}(0, \sigma^2(\Delta^*)) + \mathcal{N}(0, 1 - \sigma^2(\Delta^*))}_{\mathcal{N}(0, 1)} + X_{corr}(\sigma^2 = 1)$$

We can imitate standard normal noise by adding  $X_{nc} \sim \mathcal{N}(0, 1 - \sigma^2(\Delta^*))$

**Good** We can use precomputed  $X_{corr}(\sigma^2 = 1)$

**Bad** We need to sample minibatches until  $\sigma^2(\Delta^*) < 1$

# We still have some questions

- How to sample from correction distribution?
- What error we have if we assume that

$$\Delta^* = \Delta + X_{norm}, \quad X_{norm} \sim \mathcal{N}(0, \sigma^2(\Delta^*))$$

Instead of

$$\Delta^* = \Delta + X_{norm}, \quad X_{norm} \sim \overline{\mathcal{N}}(0, \sigma^2(\Delta^*))$$



# Bounding acceptance probability error

$$\begin{aligned} X_i &= N \log \frac{p(x_i|\theta')}{p(x_i|\theta)} - \sum_i^N \log \frac{p(x_i|\theta')}{p(x_i|\theta)} \\ &\approx N \log \frac{p(x_i|\theta')}{p(x_i|\theta)} - \frac{N}{b} \sum_i^b \log \frac{p(x_i|\theta')}{p(x_i|\theta)} \end{aligned}$$

Authors bound error of acceptance probability

$$\sup_y |\mathbb{P}\{\Delta^* + X_{nc} + X_{corr} < y\} - S(y - \Delta)| \leq \frac{6.4\mathbb{E}|X|^3 + 2\mathbb{E}|X|}{\sqrt{b}} = \varepsilon$$

$S$  – CDF of Logistic Distribution

# Bounds stationary distribution

$\hat{p}, p$  – stationary distributions of true and approximate transition operators  $\hat{\tau}$  and  $\tau$

If

$$|\hat{\mathbb{P}}\{\text{acceptance}\} - \mathbb{P}\{\text{acceptance}\}| < \varepsilon$$

And true operator has contraction property

$$d(\tau q, p) < \eta d(q, p),$$

Where  $d(q, p)$  – total variation distance

Then

$$d(\hat{p}, p) < \frac{\varepsilon}{1 - \eta}$$

# Algorithm

1. Sample candidate  $\theta' \sim q(\theta'|\theta)$

2. Increase minibatch until

$$\sigma^2(\Delta^*) < 1 \quad \text{and} \quad \varepsilon < \delta$$

3. Accept  $\theta'$  if

$$\Delta^* + X_{nc} + X_{corr} > 0$$

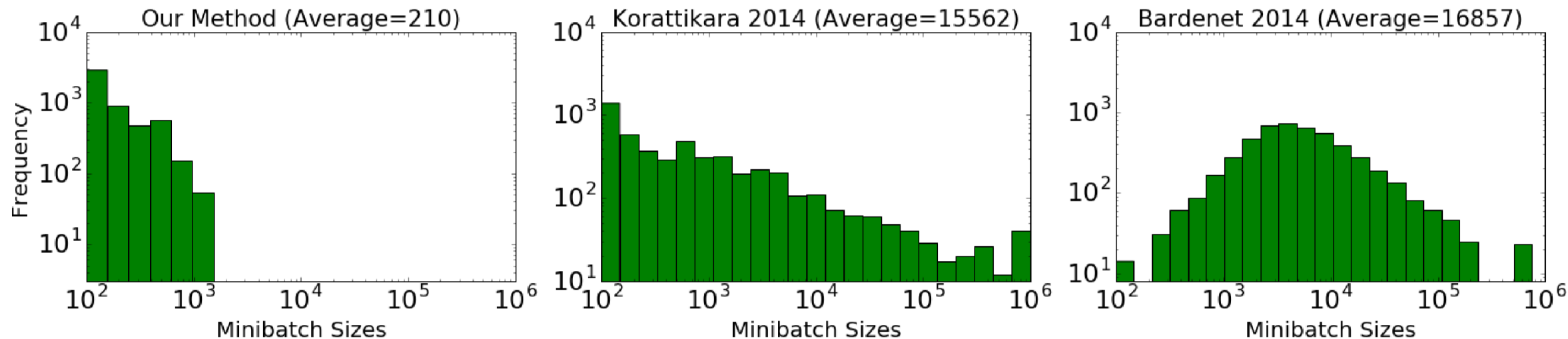
$$X_{nc} \sim \mathcal{N}(0, 1 - \sigma^2(\Delta^*))$$

$$X_{corr} \sim \text{Correction Distribution}(\sigma^2 = 1)$$

4. Otherwise keep old  $\theta$

# Efficiency

Dataset of  $10^6$  points sampled from mixture of Gaussians



5 min break

# Langevin Dynamics

Makes use of the gradient of log-density

$$\Delta\theta_t = \frac{\varepsilon}{2} \nabla \log p(\theta) + \eta_t, \quad \eta_t \sim \mathcal{N}(0, \varepsilon)$$

In Bayesian Inference

$$\Delta\theta_t = \frac{\varepsilon}{2} \left( \nabla \log p(\theta_t) + \sum_{i=1}^N \nabla \log p(x_i | \theta_t) \right) + \eta_t, \quad \eta_t \sim \mathcal{N}(0, \varepsilon)$$

Full dataset!

# Bayesian Learning via Stochastic Gradient Langevin Dynamics

*Max Welling, Yee Whye Teh*

# Stochastic Gradient Langevin Dynamics

Estimate gradient in Langevin Dynamics on minibatch =  $\{x_{t_1}, \dots, x_{t_n}\}$

$$\Delta\theta_t = \frac{\varepsilon_t}{2} \left( \nabla \log p(\theta_t) + \frac{\overset{n}{N}}{\underset{n}{n}} \sum_{i=1}^n \nabla \log p(x_{\textcolor{red}{t}_i} | \theta_t) \right) + \eta_t, \quad \eta_t \sim \mathcal{N}(0, \varepsilon_t)$$



# ~~Proof~~ Intuitive Analysis of SGLD

True gradient

$$g(\theta) = \nabla \log p(\theta) + \sum_{i=1}^N \nabla \log p(x_i | \theta)$$

Deviations

$$h_t(\theta) = \nabla \log p(\theta) + \frac{N}{n} \sum_{i=1}^n \nabla \log p(x_{t_i} | \theta) - g(\theta),$$

$$h_t(\theta) \sim \overline{\mathcal{N}}(0, V_t(\theta))$$

# Intuitive Analysis of SGLD

Given

$$\sum_{t=1}^{\infty} \varepsilon_t = \infty \qquad \sum_{t=1}^{\infty} \varepsilon_t^2 < \infty$$

We can find subsequence  $t_1 < t_2 < \dots$  such that

$$\lim_{s \rightarrow \infty} \sum_{t=t_s+1}^{t_{s+1}} \varepsilon_t = \varepsilon_0,$$

Where  $0 < \varepsilon_0 < 1$  is initial step

After one step

$$\Delta\theta_t = \frac{\varepsilon_t}{2} (g(\theta_t) + h_t(\theta_t)) + \eta_t, \quad \eta_t \sim \mathcal{N}(0, \varepsilon_t)$$

After several steps

$$\Delta\theta = \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) + h_t(\theta_t)) + \mathcal{N}\left(0, \sum_{t=t_s+1}^{t_{s+1}} \varepsilon_t\right) =$$

$$\sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) - g(\theta_{t_s}) + g(\theta_{t_s})) + \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} h_t(\theta_t) + \mathcal{N}\left(0, \sum_{t=t_s+1}^{t_{s+1}} \varepsilon_t\right)$$

For  $s$  big enough

One step of Langevin dynamics



Langevin noise

Has order of  $O(\sqrt{\varepsilon_0})$

$$\Delta\theta = \frac{\varepsilon_0}{2} g(\theta_{t_s}) + \mathcal{N}(0, \varepsilon_0) +$$

$$\sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) - g(\theta_{t_s})) + \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} h_t(\theta_t)$$

Systematic error

Non-zero mean

Has order of  $O(\varepsilon_0^2)$

“Random” error

Zero-mean

Has order of  $O(\varepsilon_0)$

# Bounding systematic error

Firstly we bound  $\|\theta_t - \theta_{t_s}\|_2 \forall t \in [t_s + 1, t_{s+1}]$

$$\begin{aligned} \|\theta_t - \theta_{t_s}\|_2 &\leq \left\| \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) + h_t(\theta_t)) + \mathcal{N}\left(0, \sum_{t=t_s+1}^{t_{s+1}} \varepsilon_t\right) \right\|_2 \\ &\leq \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} \|g(\theta_t)\|_2 + \left\| \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} h_t(\theta_t) \right\|_2 + \|\mathcal{N}(0, \varepsilon_0)\|_2 = O(\varepsilon_0) \end{aligned}$$

Assuming that  $\|g(\theta)\|_2$  and  $\|h_t(\theta_t)\|$  have some upper bounds

# Bounding systematic error

Assuming gradient Lipschitz continuity  $\|g(\theta_t) - g(\theta_{t_s})\|_2 \leq L\|\theta_t - \theta_{t_s}\|_2$

$$\left\| \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) - g(\theta_{t_s})) \right\| \leq O(\varepsilon_0) \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} = O(\varepsilon_0^2)$$

$O(\varepsilon_0^2)$  is negligible compared to  $\frac{\varepsilon_0}{2} g(\theta_{t_s})$

$$\sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) - g(\theta_{t_s})) \ll \frac{\varepsilon_0}{2} g(\theta_{t_s})$$

Systematic error

True gradient

# Analysis of “random” error

$$\sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} h_t(\theta_t) \sim \overline{\mathcal{N}} \left( 0, \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t^2}{4} V_t(\theta_t) \right)$$
$$\sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t^2}{4} V_t(\theta_t) \leq V \left( \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} \right)^2 = V \varepsilon_0^2$$

Variance  $V \varepsilon_0^2$  is negligible compared to  $\varepsilon_0$

$$\sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} h_t(\theta_t) \ll \mathcal{N}(0, \varepsilon_0)$$

“Random” error

Langevin dynamics noise

Several steps of **S**GLD  $\approx$  one step of LD

$$\begin{aligned}\Delta\theta &= \sum_{t=t_s+1}^{t_{s+1}} \frac{\varepsilon_t}{2} (g(\theta_t) + h_t(\theta_t)) + \mathcal{N}\left(0, \sum_{t=t_s+1}^{t_{s+1}} \varepsilon_t\right) \approx \\ &\approx \frac{\varepsilon_0}{2} g(\theta_{t_s}) + \mathcal{N}(0, \varepsilon_0)\end{aligned}$$

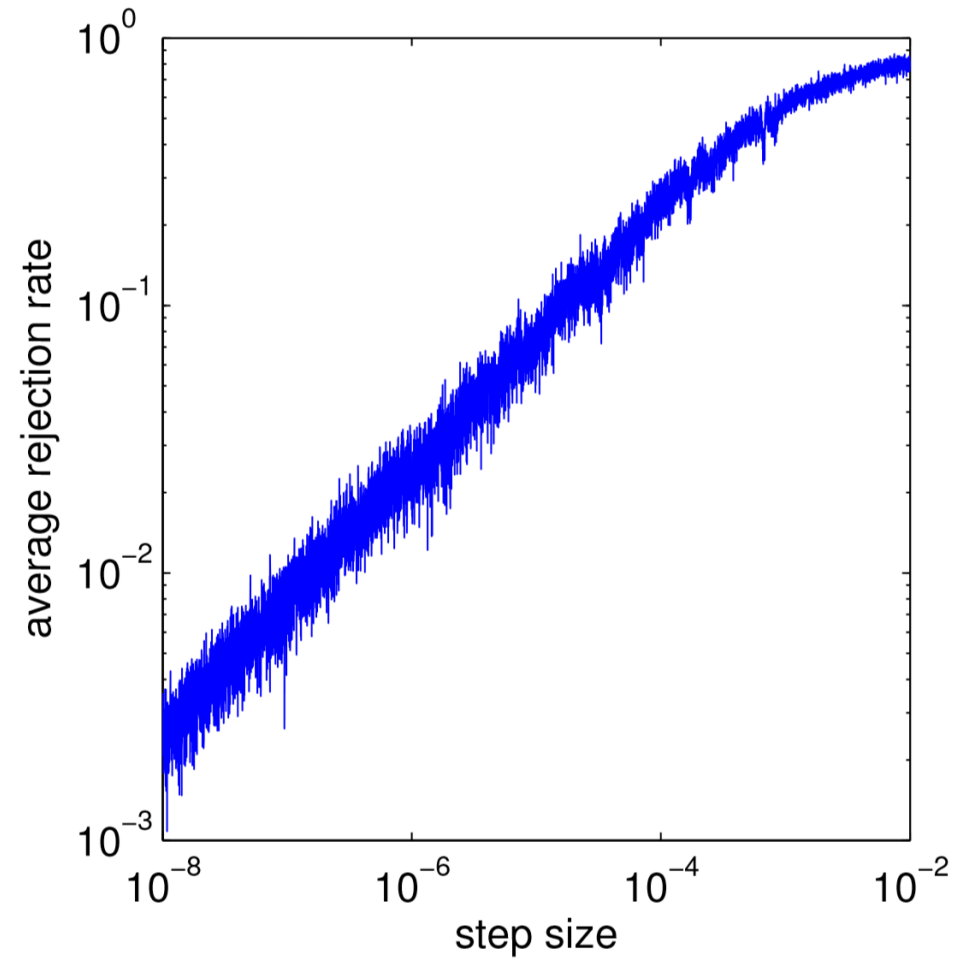
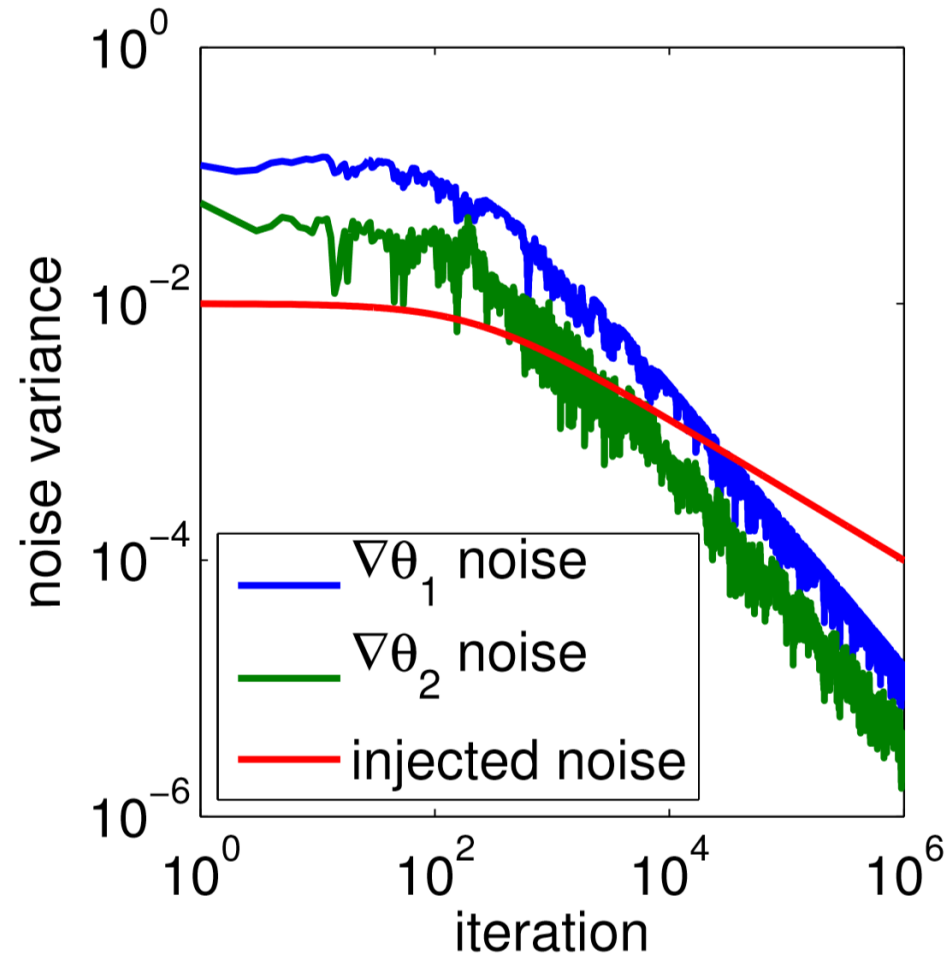
For  $\varepsilon_0$  small enough we can ignore M-H test

OR

Perform minibatch M-H test!



# Empirical analysis on toy problem



# Conclusion

- Stochastic MCMC is a new-generation methods of sampling from posterior conditioned on large dataset
- Makes use of mini-batching and stochastic optimization
- Higher rejection rates but MUCH cheaper iterations

# Acceptance rate

Acceptance probability

