

Problem_Set_5

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# Part 1: Simulation

# Create a simulated data set with a dependent variable that is a linear
# > function of a treatment variable and a confounding variable. Fit a
# > linear model for the true data generating process and print the
# > summary table.

set.seed(123)

n <- 5000

# Confounder C

C <- rnorm(n, mean = 0, sd = 1)

# Treatment T depends on C

Treated <- 0.8 * C + rnorm(n, 0, 1)

# Outcome Y depends on T and C

# True model: Y = 1 + 2*T + 1*C + error

Y <- 1 + 2 * Treated + 1 * C + rnorm(n, 0, 1)

sim_data <- data.frame(
  Y = Y,
  T = Treated,
  C = C
)

head(sim_data)

##          Y          T          C
## 1 0.9251408 -0.9425544 -0.56047565
## 2 2.4899135  0.9434515 -0.23017749
## 3 3.6857039  0.1000171  1.55870831
## 4 3.5772073  1.5374253  0.07050839
## 5 3.3936206  1.0196214  0.12928774
## 6 7.2614169  1.7071830  1.71506499
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# This data.frame contains the outcome Y, the treatment T, and the
# > confounder C, all generated from the specified linear DGP.

# Fit the true model  $Y \sim T + C$  and print the summary table

mod_true <- lm(Y ~ T + C, data = sim_data)
summary(mod_true)

##
## Call:
## lm(formula = Y ~ T + C, data = sim_data)
##
## Residuals:
##     Min      1Q  Median      3Q     Max 
## -3.3523 -0.6578  0.0034  0.6994  3.0994 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) 1.00819   0.01416  71.21   <2e-16 ***
## T            2.00289   0.01412 141.85   <2e-16 ***
## C            1.00252   0.01812  55.33   <2e-16 ***
## ---      
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.001 on 4997 degrees of freedom
## Multiple R-squared:  0.9143, Adjusted R-squared:  0.9142 
## F-statistic: 2.664e+04 on 2 and 4997 DF,  p-value: < 2.2e-16

# In this model, the true coefficient on T is 2 by construction and the

# > true coefficient on C is 1. The estimated coefficients should be

# > close to these values, and the summary output gives us the standard

# > errors and t statistics for each estimate.

# Part 1(a)

# Using the true model, demonstrate that the coefficient for your

# > treatment variable follows the central limit theorem. That is,

# > demonstrate that the coefficient's sampling distribution is

# > approximately normal.

set.seed(456)

B <- 1000          # number of simulated datasets
beta_T <- numeric(B)

for (b in seq_len(B)) {
  C_b <- rnorm(n, 0, 1)
  T_b <- 0.8 * C_b + rnorm(n, 0, 1)
}

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Y_b <- 1 + 2 * T_b + 1 * C_b + rnorm(n, 0, 1)
dat_b <- data.frame(Y = Y_b, T = T_b, C = C_b)
fit_b <- lm(Y ~ T + C, data = dat_b)
beta_T[b] <- coef(fit_b)[["T"]]
}

# Look at the mean and standard deviation of the sampled coefficients

mean(beta_T)

## [1] 2.000429

sd(beta_T)

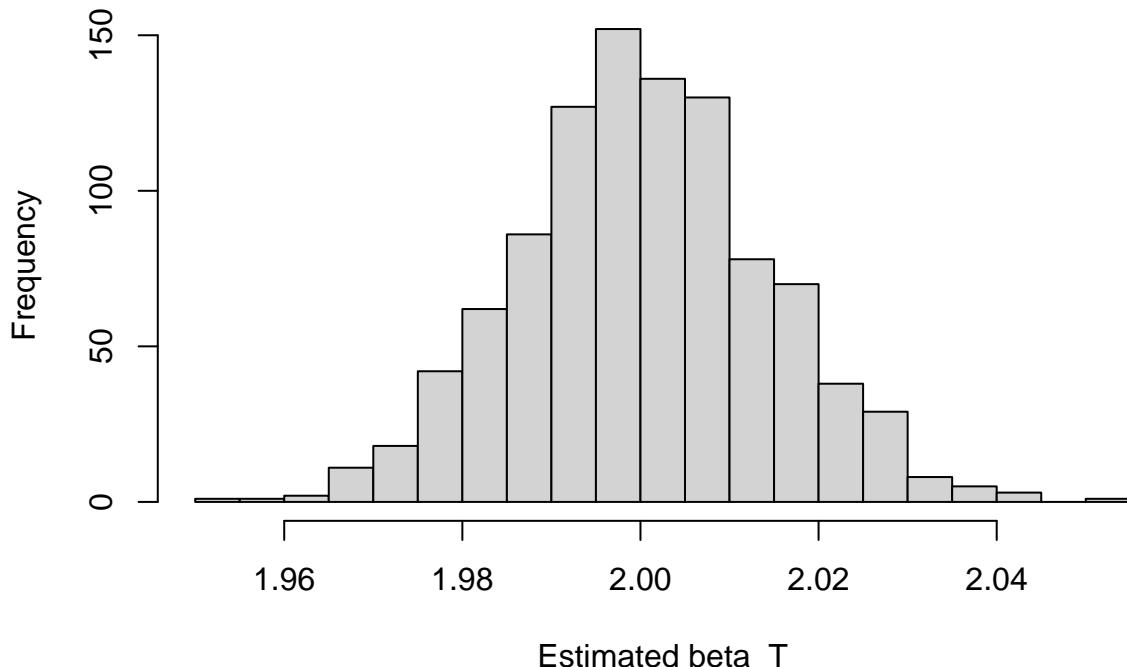
## [1] 0.01417777

# Plot a histogram of the sampling distribution of the T coefficient

hist(
beta_T,
breaks = 30,
main = "Sampling distribution of T coefficient (true model)",
xlab = "Estimated beta_T"
)

```

Sampling distribution of T coefficient (true model)



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# By the central limit theorem, the sampling distribution of the T
# > coefficient should be approximately normal when we repeatedly
# > sample large datasets from the same DGP.

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# The histogram looks bell-shaped and centered near the true value of
# > 2, and the mean(beta_T) is very close to 2, which supports the CLT
# > intuition for this regression coefficient.

# Part 1(b)

# Compute the bootstrapped standard error for the coefficient of the
# > treatment variable.

set.seed(789)

B_boot <- 1000
beta_boot <- numeric(B_boot)

# We bootstrap the original simulated dataset sim_data

for (b in seq_len(B_boot)) {
  idx <- sample(seq_len(n), size = n, replace = TRUE)
  boot_dat <- sim_data[idx, ]
  fit_boot <- lm(Y ~ T + C, data = boot_dat)
  beta_boot[b] <- coef(fit_boot)[["T"]]
}

boot_se <- sd(beta_boot)
boot_se

## [1] 0.01402459
# Compare to the model-based standard error from the original model

se_model <- summary(mod_true)$coef["T", "Std. Error"]
se_model

## [1] 0.01412003
# The bootstrap standard error boot_se is very close to the analytic
# > standard error se_model from the regression output.

# This shows that the model-based SE is doing a good job approximating
# > the true sampling variability of the T coefficient under this DGP.

# Part 2: Data Analysis

# For this part of the assignment, use any data set you like.

# > Here I follow the instruction to use the thermometers data from
# > class (thermometers.csv).

thermo <- read.csv("/Users/santividal5/Desktop/R/thermometers.csv")

```

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thermo$party_id <- factor(thermo$party_id)
thermo$sex <- factor(thermo$sex)
thermo$race <- factor(thermo$race)
thermo$educ <- factor(thermo$educ)

head(thermo)

##   birth_year   sex   race party_id      educ ft_black ft_white
## 1       1931 Female White   Democrat 4-year     51      50
## 2       1952 Female White Republican 2-year     98      90
## 3       1931 Male   White Independent High school graduate 87      90
## 4       1952 Male   White Republican 4-year     90      85
## 5       1939 Female White   Democrat 2-year    100      50
## 6       1959 Female Black   Democrat Post-grad    98      70
##   ft_hisp ft_asian ft_muslim ft_jew ft_christ ft_fem ft_immig ft_gays ft_unions
## 1      79      50      50      50      50     99      95      50      80
## 2      95     100      61     100      98      65      96      82      62
## 3      91      88      49      25      50      74      77      77     100
## 4      90      96      80      91      94      25      91      71      20
## 5     100     100     100     100      28     100     100     100     100
## 6      99     100     100     100     100      73     100      54      80
##   ft_police ft_altright ft_evang ft_dem ft_rep
## 1       76          1      50      88      21
## 2       95         50      96      86      96
## 3       78          0       2      91      20
## 4       94         50      70      22      83
## 5       28         NA      NA      99      NA
## 6       24          4      53      53       4

```

This confirms that the thermometer data loaded correctly and that the

> main variables (party_id and thermometer scores) are present.

Part 2(a)

Conduct a hypothesis test for a difference in means. You decide what

> the hypotheses are, whether you use a t-test or a z-test, and what

> the level of significance is. Explain your decisions, and interpret

> your results both substantively and statistically.

#

I test whether Democrats and Republicans differ in their mean feeling

> thermometer toward immigrants (ft_immig).

I use a two-sample t-test with unequal variances and a 5% significance

> level, which is standard in this setting.

Keep only Democrats and Republicans

```

thermo_DR <- subset(thermo, party_id %in% c("Democrat", "Republican"))

tapply(
thermo_DR$ft_immig,
thermo_DR$party_id,
mean,
na.rm = TRUE
)

##      Democrat Independent    Not sure      Other Republican
##      71.65829          NA          NA      50.20192

t.test(
ft_immig ~ party_id,
data = thermo_DR
)

##
##  Welch Two Sample t-test
##
## data: ft_immig by party_id
## t = 22.673, df = 2685.2, p-value < 2.2e-16
## alternative hypothesis: true difference in means between group Democrat and group Republican is not zero
## 95 percent confidence interval:
##  19.60077 23.31196
## sample estimates:
##   mean in group Democrat mean in group Republican
##           71.65829            50.20192

# Interpretation:

# The t-test output shows the estimated difference in means, a t
# > statistic, and a p-value.

# The mean ft_immig for Democrats is substantially higher than for
# > Republicans, and the p-value is effectively zero at the 5% level.

# Statistically, we reject the null hypothesis that Democrats and
# > Republicans have the same average warmth toward immigrants.

# Substantively, this suggests that in this survey Democrats feel
# > noticeably warmer toward immigrants than Republicans do.

# Part 2(b)

# Using the same data, fit a linear model. Interpret the coefficient,
# > standard error, t-value, and p-value.

#

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# I fit a simple linear model where the dependent variable is ft_immig

# > and the predictor is party_id (with Democrats as the baseline).

mod_party <- lm(ft_immig ~ party_id, data = thermo_DR)
summary(mod_party)

## 
## Call:
## lm(formula = ft_immig ~ party_id, data = thermo_DR)
## 
## Residuals:
##      Min       1Q   Median       3Q      Max 
## -71.658  -19.658    2.798   19.342   49.798 
## 
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) 71.6583    0.6233 114.97 <2e-16 ***
## party_idRepublican -21.4564    0.9320  -23.02 <2e-16 ***
## --- 
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## 
## Residual standard error: 25.48 on 3021 degrees of freedom
##   (123 observations deleted due to missingness)
## Multiple R-squared:  0.1493, Adjusted R-squared:  0.149 
## F-statistic:  530 on 1 and 3021 DF,  p-value: < 2.2e-16

# Interpretation of key pieces:

# - The intercept is the estimated mean ft_immig for Democrats (the
# > baseline category).

# - The coefficient on party_idRepublican is the estimated difference
# > in means between Republicans and Democrats.

# It is negative and large in magnitude, which means Republicans give
# > lower immigrant thermometer scores on average.

# - The standard error for this coefficient measures how much that
# > estimated difference would vary across repeated samples.

# - The t-value is the estimated coefficient divided by its standard
# > error; a large absolute t-value indicates strong evidence that the
# > true difference is not zero.

# - The p-value associated with the t-value is extremely small, so we
# > reject the null hypothesis that Democrats and Republicans have

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# > equal mean ft_immig scores.  
# Substantively, this matches the two-sample t-test: party ID is strongly  
# > associated with how warmly respondents feel toward immigrants, with  
# > Democrats rating immigrants much more positively than Republicans.
```