

1.

$$\hat{\sigma}_m^2 = \frac{1}{m} \sum_{i=1}^m (x^{(i)} - \hat{\mu}_m)^2$$

$$\text{Bias}(\hat{\sigma}_m^2) = \frac{1}{m} E(\hat{\sigma}_m^2) - \cancel{(\sigma_m^2)} \sigma_m^2$$

$$= E \left[\frac{1}{m} \sum_{i=1}^m (x^{(i)} - \hat{\mu}_m)^2 \right] - \sigma_m^2 \quad \text{--- ①}$$

$$\text{Using } E \left[\sum_{i=1}^m (x^{(i)} - \mu_m)^2 \right] = (m-1) \cdot \sigma^2 \quad -$$

we have from ①

$$\cancel{E} \frac{1}{m} E \left[\sum_{i=1}^m (x^{(i)} - \mu_m)^2 \right] - \hat{\sigma}_m^2$$

$$= \frac{1}{m} \times (m-1) \times \sigma_m^2 - \hat{\sigma}_m^2$$

$$= -\frac{\hat{\sigma}_m^2}{m}$$

Since the difference between expectation and true value is not zero, the estimator $\hat{\sigma}_m^2$ is biased.

$$\tilde{\sigma}_m^2 = \frac{1}{m-1} \sum_{i=1}^m (x^{(i)} - \hat{\mu}_m)^2$$

Using the bias definition and hint as above -

$$\text{Bias}(\tilde{\sigma}_m^2) = \frac{m-1 \times \tilde{\sigma}_m^2}{m-1} - \sigma_m^2$$

$$= 0$$

$\tilde{\sigma}_m^2$ is unbiased as the difference between expectation & true value is zero.

$$2. \text{MSE}(\hat{\theta}_m) = E[(\hat{\theta}_m - \theta)^2]$$

$$= E[\hat{\theta}_m^2 + \theta^2 - 2\hat{\theta}_m\theta] = E[\hat{\theta}_m^2] + \theta^2 - 2E[\hat{\theta}_m]\theta \quad \text{--- ①}$$

Adding and subtracting $E^2[\hat{\theta}_m]$, we have from ①

$$E[(\hat{\theta}_m)^2] + \theta^2 - 2E[\hat{\theta}_m]\theta + E^2[\hat{\theta}_m] - E^2[\hat{\theta}_m]$$

$$= \underbrace{E^2[\hat{\theta}_m] + \theta^2 - 2E[\hat{\theta}_m]\theta}_{\downarrow \text{Bias}(\hat{\theta}_m)^2} + \underbrace{E[(\hat{\theta}_m)^2] - E^2[\hat{\theta}_m]}_{\downarrow \text{Var}(\hat{\theta}_m)}$$

Hence,

$$\text{MSE}(\hat{\theta}_m) = \text{Bias}(\hat{\theta}_m)^2 + \text{Var}(\hat{\theta}_m).$$

3.

3. $P(y|x, \theta) \rightarrow$ Applying the Bayes Rule and neglecting the normalizing term \rightarrow

$$\rightarrow P(y|x) P(\theta|x) \quad \hookrightarrow \text{not dependent}$$

Hence, $P(y|x, \theta) P(\theta)$

$$\theta_{\text{MAP}} = \underset{\theta}{\operatorname{argmax}} \left[\sum_{i=1}^n \log \{ p(y^{(i)} | x^{(i)}, \theta) \} + \log P(\theta) \right]$$