

Lecture 18
Econometrics - I

02.04.2021

JOINT

Hypothesis testing w/ > 1 linear restriction
[for a set of linear restrictions]

Testing a set of linear restrictions jointly

$$y_i = \beta_1 + \beta_2 x_{2i} + \beta_3 x_{3i} + \dots + \beta_K x_{Ki} + u_i$$

$i=1, 2, \dots, N$

Test

$$H_0: \gamma_{11} \beta_1 + \gamma_{12} \beta_2 + \dots + \gamma_{1K} \beta_K = \gamma_1$$

$$\text{and } \gamma_{21} \beta_1 + \gamma_{22} \beta_2 + \dots + \gamma_{2K} \beta_K = \gamma_2$$

\vdots

$$\text{and } \gamma_{Q1} \beta_1 + \gamma_{Q2} \beta_2 + \dots + \gamma_{QK} \beta_K = \gamma_Q$$

Q : # of
non-redundant
linear
restrictions.

$$\begin{matrix} R & \cdot & \beta \\ Q \times K & & K \times 1 \end{matrix} = \begin{matrix} \gamma \\ Q \times 1 \end{matrix}$$

Assume: $Q \leq K$
 R is full rank.

$$R = \begin{bmatrix} r_{11} & r_{12} & \dots & r_{1k} \\ r_{21} & r_{22} & \dots & r_{2k} \\ \vdots & \vdots & \ddots & \vdots \\ r_{Q1} & r_{Q2} & \dots & r_{Qk} \end{bmatrix}; \underline{\beta} = \begin{bmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_k \end{bmatrix}$$

$Q \times k$ $k \times 1$

$$\underline{y} = \begin{bmatrix} r_1 \\ r_2 \\ \vdots \\ r_Q \end{bmatrix}$$

$Q \times 1$

Simplest example:

$H_0: \beta_1 = 0 \text{ and } \beta_2 = 0 \text{ and } \beta_3 = 0 \text{ and } \dots \text{ and } \beta_k = 0$

$$R = \begin{bmatrix} 1 & 0 & 0 & \dots & 0 \\ 0 & 1 & 0 & \dots & 0 \\ 0 & 0 & 1 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & 1 \end{bmatrix}$$

$Q \times k$ $k \times k$

$Q = k$

example

Now, let's work out how to test > 1 linear restriction.

$$H_0: \underset{Q \times K}{R} \underset{K \times 1}{\beta} = \underset{Q \times 1}{\gamma} \quad Q \leq K.$$

$$\underset{OLS}{\hat{\beta}} \sim \mathcal{N} \left(\overset{E(\hat{\beta}_{OLS})}{\underset{||}{\beta}}, \overset{V(\hat{\beta}_{OLS})}{\underset{||}{\sigma^2 (X'X)^{-1}}} \right)$$

$$R \underset{OLS}{\hat{\beta}} \sim \mathcal{N} \left(R \underset{||}{\beta}, R [\sigma^2 (X'X)^{-1}] R' \right)$$

$$R \underset{OLS}{\hat{\beta}} - \underset{||}{\gamma} \sim \mathcal{N} \left(R \underset{||}{\beta} - \underset{||}{\gamma}, \sigma^2 R (X'X)^{-1} R' \right)$$

Under $H_0: R \underset{||}{\beta} = \underset{||}{\gamma}$ or $R \underset{||}{\beta} - \underset{||}{\gamma} = \underset{Q \times 1}{\underline{0}}$

$$\underbrace{R \underset{OLS}{\hat{\beta}} - \underset{||}{\gamma}}_{Q \times 1} \sim \mathcal{N} \left(\underbrace{\underline{0}}_{Q \times 1}, \underbrace{\sigma^2 R (X'X)^{-1} R'}_{Q \times Q} \right)$$

Under H_0 , we have

$$\underbrace{R \hat{\beta}_{OLS} - \underline{y}}_{Q \times 1} \sim \underline{NT} \left(\underbrace{\underline{0}}_{Q \times 1}, \underbrace{\sigma^2 R (X'X)^{-1} R'}_{Q \times Q} \right)$$

Aside 1: $\underline{y}_{m \times 1} \sim \underline{NT} (0, \Sigma_{m \times m})$

then $\underbrace{c}_{1 \times 1} = \underbrace{\omega' \Sigma^{-1} \omega}_{1 \times m \quad m \times m \quad m \times 1} \sim \chi_m^2$
 matrix form of quadratic sum of standard normal r.v.

Aside 2: $\underline{y} \quad c_1 \sim \chi_{m_1}^2$
 $c_2 \sim \chi_{m_2}^2$; c_1, c_2 are independent

then $\frac{c_1/m_1}{c_2/m_2} \sim F_{m_1, m_2}$

Aside 3: $\underbrace{\frac{\hat{u}' \hat{u}}{\sigma^2}}_{\substack{\text{residual} \\ \text{vector for MLRM}}} \sim \chi_{N-K}^2$ } prev. lecture had scalar form:
 $\sum_{i=1}^N \frac{\hat{u}_i^2}{\sigma^2} \sim \chi_{N-K}^2$

Aside 1 \Rightarrow

$$\left(R \hat{\beta}_{\text{OLS}} - \underline{y} \right)' \left[\sigma^2 R (X'X)^{-1} R' \right]^{-1} \left(R \hat{\beta}_{\text{OLS}} - \underline{y} \right) \sim \chi^2_Q$$

Aside 3 $\Rightarrow \quad \underline{\hat{u}}' (\sigma^2)^{-1} \underline{\hat{u}} \sim \chi^2_{N-K}.$

Overall,

$$\frac{\left[\left(R \hat{\beta}_{\text{OLS}} - \underline{y} \right)' \left(\sigma^2 R (X'X)^{-1} R' \right)^{-1} \left(R \hat{\beta}_{\text{OLS}} - \underline{y} \right) \right] / Q}{\left[\underline{\hat{u}}' (\sigma^2)^{-1} \underline{\hat{u}} \right] / (N-K)} \sim F_{Q, N-K}$$

$$\underbrace{\left[\underline{\hat{u}}' (\sigma^2)^{-1} \underline{\hat{u}} \right] / (N-K)}_{D^2} = \frac{\hat{\underline{u}}' \hat{\underline{u}}}{N-K} = \sigma^2$$

So,

$$\frac{\left(R \hat{\beta}_{\text{OLS}} - \underline{y} \right)' \left[\sigma^2 R (X'X)^{-1} R' \right]^{-1} \left(R \hat{\beta}_{\text{OLS}} - \underline{y} \right)}{Q} \sim F_{Q, N-K}.$$

⇒ Conducting hypothesis tests by estimating a restricted regression model.

Consider a full model:

① *Full model* $y = \beta_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_4 + u$

Test: $H_0 : \beta_2 + 2\beta_4 = 4$ *estimate:* $\hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3, \hat{\beta}_4$

Recover a restricted model by substituting the restriction in the original model (1)

② *Restricted model* $y = \beta_1 + (4 - 2\beta_4) x_2 + \beta_3 x_3 + \beta_4 x_4 + u$

$$\underbrace{y - 4x_2}_{=\tilde{y}} = \beta_1 + \beta_3 x_3 + \beta_4 \underbrace{(x_4 - 2x_2)}_{=z} + u$$

$$\tilde{y} = \beta_1 + \beta_3 x_3 + \beta_4 z + u$$

estimate: $\tilde{\beta}_1, \tilde{\beta}_3, \tilde{\beta}_4$

Test statistic: CHOW TEST

$$\textcircled{A} \frac{(SSE_{\text{Restricted}} - SSE_{\text{Full}}) / Q}{(SSE_{\text{Full}}) / (N - K)} \sim F_{Q, N-K}$$

$$\textcircled{B} \frac{(R^2_{\text{Full}} - R^2_{\text{Restricted}}) / Q}{(1 - R^2_{\text{Full}}) / (N - K)} \sim F_{Q, N-K}$$

For our example: $Q = 1$

Final step.
Inference :

if $F\text{-stat} > F_{0, N-k, \alpha}^*$ then reject H_0 .

if $F\text{-stat} \leq F_{0, N-k, \alpha}^*$ then fail to
reject H_0 .

Next lecture: Maximum Likelihood
Estimation