Prosperies of the OLS estimators on any sample of data.

emple of data.

$$\hat{J}_{i} = \hat{\beta}_{0,ous} + \hat{\beta}_{1,ous}$$
;  $\hat{J}_{i} = \hat{J}_{i} + \hat{U}_{i}$ 

· due to the OLS estimators based on the green sample.
· not the same as the error term in the original model.

All Ols estimators: B, B, ûi 2 = 2 = (y: - Bo - Bai) = 0  $\sum_{i=1}^{N} \hat{U}_i = 0$  $\leq_{i=1}^{N} \hat{u}_i x_i = 0$ 8 = 2 = (yi-Bo-B, xi) xi = 0 The point ( \( \tau , \text{y} \) is always on the OLS line. 岁=第十月元

In essense,

Goodness of fit

Define:  

$$SST = \sum_{i=1}^{N} (f_i - \bar{f})^2 \rightarrow Sum g squares - both$$

SST to total random in data Interpreteation: SSE - explained variation in date SSR -> residual variation in data

$$\sum_{i=1}^{N} (y_i - y_i)^2 = \sum_{i=1}^{N} (\hat{y}_i - y_i)^2 + \sum_{i=1}^{N} \hat{u}_i^2$$

Start w/ LHS!

$$SST$$
 $\sum_{i=1}^{N} (y_i - y_i)^2 = \sum_{i=1}^{N} (y_i - y_i + y_i - y_i)^2$ 

$$= 2 \left[ \left( \frac{1}{3} - \frac{1}{3} \right) + \left( \frac{1}{3} - \frac{1}{3} \right) \right]^{2}$$

$$= \frac{1}{3} \left[ \left( \frac{1}{3} - \frac{1}{3} \right) + \left( \frac{1}{3} - \frac{1}{3} \right) \right]^{2}$$

$$= \sum_{i=1}^{N} \left[ \hat{u}_{i}^{2} + (\hat{y}_{i} - \bar{y})^{2} + 2 \hat{u}_{i} (\hat{y}_{i} - \bar{y}) \right]$$

$$= \sum_{i=1}^{N} \hat{u}_{i}^{2} + \sum_{i=1}^{N} (\hat{y}_{i} - \bar{y}) + 2 \sum_{i=1}^{N} \hat{u}_{i} (\hat{y}_{i} - \bar{y})$$

$$= \sum_{i=1}^{N} \hat{u}_{i}^{2} + \sum_{i=1}^{N} \hat{u}_{i} (\hat{y}_{i} - \bar{y})$$

$$= \sum_{i=1}^{N} \hat{u}_{i}^{2} + \sum_{i=1}^{N} \hat{u}_{i} (\hat{y}_{i} - \bar{y})$$

$$= \sum_{i=1}^{N} \hat{u}_{i}^{2} + \sum_{i=1}^{N} \hat{u}_{i} (\hat{y}_{i} - \bar{y})$$

(How and

$$\sum_{i=1}^{N} \hat{U}_{i} \left( \hat{y}_{i} - \hat{y}_{i} \right) = \sum_{i=1}^{N} \hat{U}_{i} \left( \hat{y}_{i} + \hat{y}_{i} \right) = 0$$

$$= \sum_{i=1}^{N} \hat{U}_{i} \left( \hat{y}_{i} + \hat{y}_{i} \right) = 0$$

$$= \sum_{i=1}^{N} \hat{U}_{i} + \hat{y}_{i} = 0$$

$$= \sum_{i=1}^{N} \hat{U}_{i} + \hat{y}_{i}$$

Hence:

Signification index splained rendual

Total variation index splained rendual

Goodness of fit o SSF x100
SST
variation in yi explained
by xi via SLRM.

 $R^{2} = SSE = | - SSR |$  SST  $R^{e} \in (0,1)$  Sample variation explained by the SLRM.

Note: Low R<sup>2</sup> values are not un common in social a ciences

In cross-sectional reprossion, love  $\ell^2$  does not mean that the regression is useless.

HUIX Somethy en cetero panho experiment

The function form of the simple linear repression model and interpretation of & -> yi = Bo + B, xi, + Ui Di = B values Interpretation (8)  $\mathcal{X}$ Model Sy = BIAX Hi Cevel-level Ji W % Dy = BIDX log-level b)(Ji) Dy = B, (), sx) log(yi) log(xi) level-log % Ay = B, (% Dx) log-log  $\Rightarrow \beta_1 = \frac{\partial \log y_i}{\partial x_i} = \frac{1}{2} \frac{\partial x_i}{\partial x_i}$ = Bo+B, xi + Uî los (yi) B= di dyi = % change in y di drii upon 1 unit di upon 1 unit change =) [Pigi] = dyi/dx; = /(yi)

log-log model

log(yi) = Bot B, log(xi) + Wi tenis a SLAM

linear in parameteral

B = Dlog(yi) = y, dyi = 1, chargeing

Dlog(xi) = 1/2 dxi % chargeing

Plasticity

Ly writ.

Charge in X.

dyi = Byi
di

