

★ The Identification proc. require certain assumption, known as identification assumption.

{ Consistency
 Exchangeability

Def. The observable outcome Y is defined as

$$Y = \begin{cases} Y(1) & \text{if } A=1 \\ Y(0) & \text{if } A=0 \end{cases}$$

In general, $Y = Y(a)$ if $A = a$

(Remark. paired definition

$$Y = Y(1)A + Y(0)(1-A)$$

★ If two groups are exchangeable then any difference in outcomes between them can be attributed to the treatment, not to pre-existing difference.

$$P(Y(a)=1 | A=1) = P(Y(a)=1 | A=0) \quad \forall a=0,1$$

or equivalently

$$\begin{cases} Y(1) \perp\!\!\!\perp A \\ Y(0) \perp\!\!\!\perp A \end{cases}$$

In general, $Y(a) \perp\!\!\!\perp A \quad \forall a$ (counterfactual outcome and actual treatment are indep)

Subject :

No. : 11

Date :

★ Causal Inference as a fundamental missing data challenge.

Econ Program \longrightarrow Income
 $A = \begin{cases} 1 & \text{enter} \\ 0 & \text{not enter} \end{cases}$ Y

A	$Y(1)$	$Y(0)$
1	100	?
1	120	?
1	220	?
1	200	?
0	?	170
0	?	90
0	?	190
0	?	80

Representative?

★ What does representative look like?

\Rightarrow obs data and missing data take some values

A	$Y(1)$	$Y(0)$
1	100	90 120
1	120	80 110
1	220	170 200
1	200	190 200
0	200 170	170
0	120 90	90
0	220 170	190
0	100 70	80

• For both potential outcomes, the distⁿ of outcomes is the same regardless of the value of the treatment variable.

$$Y(a) \perp\!\!\!\perp A \quad \forall a=0,1$$

• Lack of exchangeability

★ $Y(a) \perp\!\!\!\perp A \not\Rightarrow Y \perp\!\!\!\perp A$

$$E[Y(a) | A=1] = E[Y(a) | A=0] \\ \forall a=0,1 \text{ (mean exchangeability)}$$

$$\begin{cases} Y(1) \perp\!\!\!\perp A \\ Y(0) \perp\!\!\!\perp A \end{cases} \text{ (exchangeability)}$$

$$(Y(1), Y(0)) \perp\!\!\!\perp A \text{ (full exchangeability)}$$

$$\text{Let } \phi(a) = E[Y(a)] \quad a=0,1$$

$$\text{Causal Effect} = \phi(1) - \phi(0)$$

$$\stackrel{\text{Identify}}{=} E(Y|A=1) - E(Y|A=0)$$

$$\begin{aligned} \phi(a) &= E[Y(a)] \\ &= E[Y(a) | A=a] \quad \left(\begin{array}{l} \text{Exchangeability} \\ Y(a) \perp\!\!\!\perp A \quad \forall a \end{array} \right) \\ &= E(Y | A=a) \quad \left(\begin{array}{l} \text{Consistency} \\ Y=Y(a) \text{ if } A=a \end{array} \right) \end{aligned}$$

Latin \rightarrow confundere which mean mixing
 } Confounding

- Effect or association between exposure and outcome is distorted by the presence of another variable.

- Confounding is also form a bias.

- An informal definition

A confounder is any variable that can be used to adjust for confounding.

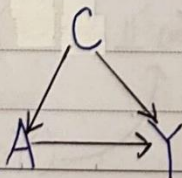
- In traditional approach.

- 1) It is associated w/ treatment.

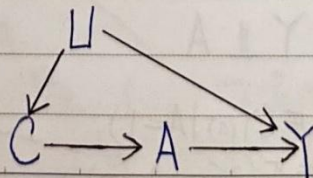
- 2) " outcome conditional on treatment.

- 3) It does not lie on a causal pathway between treatment and outcome.

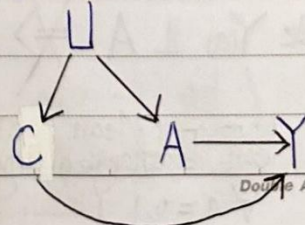
(a)



(b)



(c)



Case (a)

C has causal effect on A

C has direct causal effect on Y

C does not lie on causal pathway between A and Y

Case (b)

C has causal effect on A

C shares cause \perp w/ Y

C does not lie on the causal pathway between A and Y

Case (c)

C shares cause \perp w/ A

C has causal effect on Y

C does not lie on the causal pathway between A and Y

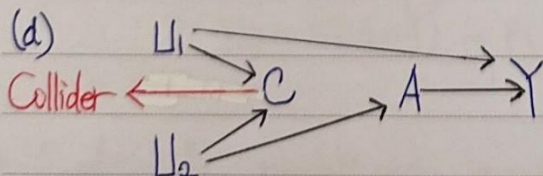
Thm.

Consider a set of nodes $C = \{v_1, \dots, v_k\}$ s.t. $C \cap A = \emptyset$

Assume that following conditions hold:

i) No element of C is a descendent of A

ii) C blocks all back door paths from A to Y

then $Y(a) \perp\!\!\!\perp A \mid C \quad \forall a$  \Rightarrow No confounding

There are no common causes of A and Y.

The backdoor path between A and Y through C is blocked because C is a collider.

The C in (d) meets the rules for traditional approach

C shares common cause U_2 ✓ A
 " " U_1 " Y

C does not lie on the causal pathway between A and Y

★ Adjust for C results in a biased estimator of the causal effect due to selection bias.

How to correct it? ↙ Randomized Control Trials

(1) Experimental Study { RCT
 Matching (必需要知道有兴趣的干扰因子)
 Restriction / 分层分析

(2) Observational Study { Stratification { 模型法
 G-method { Standardization, g-formula
 IPW estimator and MSM
 G-estimation and structural nested Model (Semi-parametric)

{ Randomization

• Randomized experiments, like other real study general data w/ missing value of counterfactual outcome.

• Randomization ensures that those missing values occurred by chance.

★ A ideal randomized experiment


i) No loss to follow-up

ii) Full adherence to the assigned treatment

iii) A single version of treatment

iv) Double blind assignment (双盲实验)

design 1.



critical

$C=1$

75% → get new ♡

25% → Not

non-critical

$C=0$

50% → get new ♥

50% → Not

design 1 \Rightarrow single unconditional randomization probly that is common to all individuals.
marginally randomized experiments

conditionally

- Several randomization probabilities that depend on the value of the variable e

* A marginally randomized experiment is expected to result in exchangeability of treated and untreated

$$Y(a) \perp\!\!\!\perp A \quad \forall a=0,1$$

or

$$P(Y_{(a)} = 1 | A = 1) = P(Y_{(a)} = 1 | A = 0) \quad \forall a = 0, 1$$

- ★ Conditionally randomized experiments is simply the combination of separate marginally randomized experiments. It guarantees

$$\rightarrow Y(a) \perp A \mid C \quad \forall a=0,1$$

Conditional exchangeability

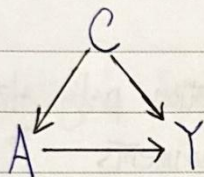
Subject :

Standardization

$$\begin{aligned}
 E(X) &= \int x f_X(x) dx \\
 &= \int x \left(\int_y f_{X|Y}(x,y) dy \right) dx \\
 &= \int_y \left(\int_x x f_{X|Y}(x,y) dx \right) f_Y(y) dy \\
 &= \int_y E(X|Y=y) f_Y(y) dy
 \end{aligned}$$

$$E[Y(a)] = \begin{cases} \sum_c E[Y(a)|C=c] P(C=c) & \text{if } C \text{ discrete} \\ \int_c E[Y(a)|C=c] f_C(c) dc & \text{if } C \text{ conti} \end{cases}$$

DAGs



$$\begin{aligned}
 \phi(a) &= E[Y(a)] = \sum_c E[Y(a)|C=c] P(C=c) && \text{conditional exchangeability by } Y(a) \perp\!\!\!\perp A | C \\
 &= \sum_c E[Y(a)|A=a, C=c] P(C=c) \\
 &\stackrel{\text{g-formula}}{=} \sum_c E(Y|A=a, C=c) P(C=c) && \text{by consistency}
 \end{aligned}$$

the causal effect

$$\begin{aligned}
 &E[Y(1)] - E[Y(0)] \\
 &= \sum_c \{ E(Y|A=1, C=c) - E(Y|A=0, C=c) \} P(C=c)
 \end{aligned}$$

★ Inverse probability weighting (IPW)

An individual's IP weight depends on + value of $\begin{cases} \text{Covariate } C \\ \text{treatment } A \end{cases}$

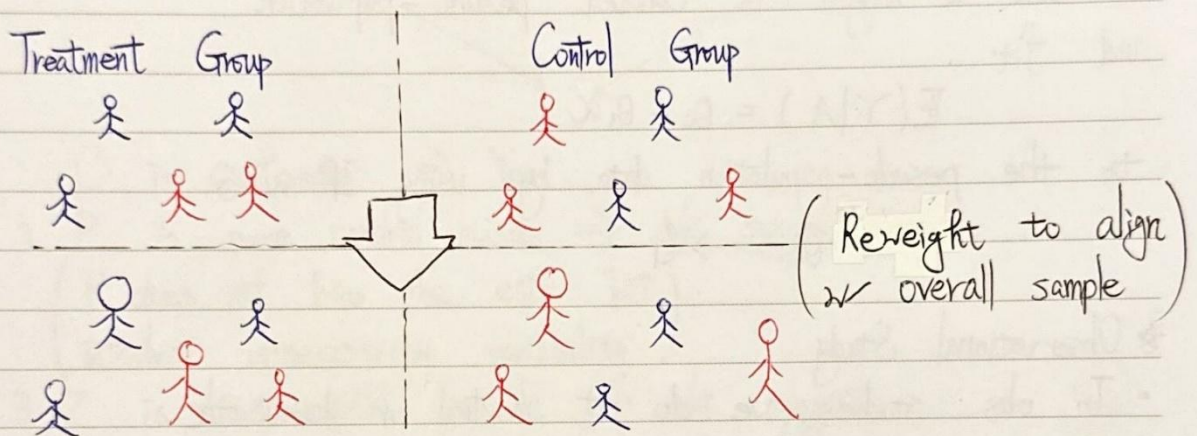
$$\begin{aligned}
 &\sum_c E(Y|A=a, C=c) P(C=c) \\
 &= \sum_c \left\{ \sum_y y \frac{P(Y=y, A=a, C=c)}{P(A=a|C=c)} \right\}
 \end{aligned}$$

Subject :

$$= \int_C \left\{ \sum_y \sum_{a^*} y I\{a^*=a\} \frac{P(Y=y, A=a^*, C=c)}{P(A=a^* | C=c)} \right\}$$

$$= E \left[\frac{I\{A=a\}}{P(A|C)} Y \right]$$

{ A treated individual w/ $C=c$ receives weight $\frac{1}{P(A=1|C=c)}$
 { An untreated " $C=c^*$ " " $\frac{1}{P(A=1|C=c^*)}$



★ Goal of IP weight is to create a pseudo-population the IP weight

$$w^A = \frac{1}{P(A|C)}$$

are referred to as nonstabilized weights.

and

$$SW^A = \frac{P(A)}{P(A|C)}$$

narrower CI

are referred to as stabilized weights.

2 Marginal Structural Model (MSM)

The model attempts to solve

$$E[Y(a)] = \mu(a; \xi)$$

(1) Linear model

$$E[Y|a] = \beta_0 + \beta_1 a$$

(2) Log-linear model

$$\log\{E[Y|a]\} = \beta_0 + \beta_1 a \Leftrightarrow E[Y|a] = e^{\beta_0 + \beta_1 a}$$

(3) Logistic model

$$\text{logit}\{E[Y|a]\} = \beta_0 + \beta_1 a \Leftrightarrow E[Y|a] = \frac{e^{\beta_0 + \beta_1 a}}{1 + e^{\beta_0 + \beta_1 a}}$$

We use IP weight to conduct pseudo-population and fit

$$E(Y|A) = \theta_0 + \theta_1 a$$

to the pseudo-population data by using IP WLS

$$\hat{\theta}_1 \rightarrow \beta_1$$

3 Observational Study

- In obs. studies, we do not control or know the assignment mechanism.
- Presence of measured and unmeasured confounders
 - unbalanced between group.
- Some structural (often untestable) assumption must be made

★ Informally, an obs. study can be conceptualized as a conditionally randomized experiments if

1. Well-defined treatment \Leftrightarrow Consistency

The treatments being compared are clearly defined and match the ver. in data.

2. No unmeasured confounding \Leftrightarrow Exchangeability

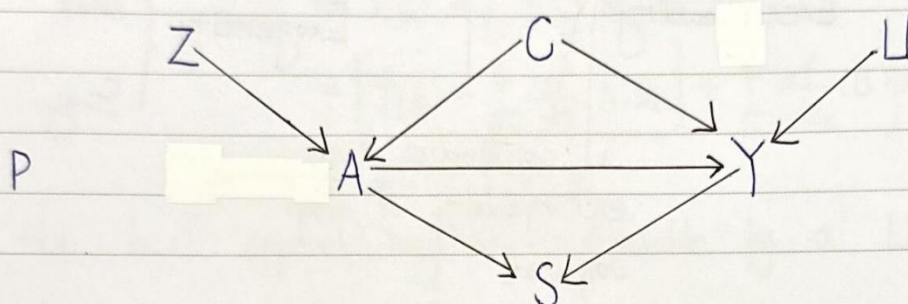
Treatment assign depends only on observed covariates C

3. Positivity

Every treatment has non-zero probly for all value of C

- These conditions are often heroic which explain why causal inferences from obs. study are viewed w/ suspicion.

★ What covariates should be adjust for obs. study?



1. C is confounder \Rightarrow Yes
 2. P is pure random noise \Rightarrow Not suggest
(P does not bias the estimate but introduce unnecessary variability)
 3. Z is instrumental variable \Rightarrow Not suggest
(Z does not bias the estimate although it increases variability. However, w/ unmeasured confounding, include it amplifies bias)
 4. U affect Y only \Rightarrow Yes
(Since U is predictive to Y, including them often improve precision)
 5. S is collider \Rightarrow No
- \Rightarrow We should adjust for at least C to remove bias and more ideally, further adjust for U to reduce variance.

2 Causality in Econometric:

Choice vs chance

Endogenous Decision Exogenous Variation