

# CS6111: Foundations of Cryptography

## Assignment 3

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CS16B039

### Instructions

- Deadline: Wednesday, Oct 16.
- We encourage submissions by Latex. Paper is also accepted.

### References

- Introduction to Cryptography - Delfs and Knebl
- A Graduate Course in Applied Cryptography - Boneh and Shoup ([link](#))
- Introduction to Modern Cryptography - Katz and Lindell

## 1 Number Theory

Let  $G$  be some group and let  $g \in G$  be an element of prime order  $n$ . That is, the set of elements generated by  $g$  is a cyclic group of prime order, denoted as follows:  $G_n = \langle g \rangle = \{g^i \mid i \geq 0\}$ . We have that  $\langle g \rangle^* = \{g^x : x \in \mathbb{Z}_n^*\}$  is the set of all generators of  $G_n$ . Clearly,  $|\langle g \rangle^*| = \phi(n)$ , which is known to be the number of generators of any cyclic group of order  $n$ .

1. (8 points) Show that for any  $h \in \langle g \rangle$ , the following conditions are equivalent:

- (1)  $h \in \langle g \rangle^*$
- (2)  $\text{ord}(h) = n$
- (3)  $\langle h \rangle = \langle g \rangle$
- (4)  $\langle h \rangle^* = \langle g \rangle^*$

To show equivalency, you must prove the following:

- (a) (2 points) (1)  $\implies$  (2)
- (b) (2 points) (2)  $\implies$  (3)
- (c) (2 points) (3)  $\implies$  (4)
- (d) (2 points) (4)  $\implies$  (1)

**Solution:** Across all parts, we have the following common information:

$$g^n = e$$

$$\langle g \rangle^* = \{g^x : x \in \{1, 2, \dots, n-1\}\}$$

This is because, actual definition has  $x \in \mathbb{Z}_n^*$  and  $n$  is prime. Therefore, every number less than  $n$  is co-prime to  $n$  and thus

$$\mathbb{Z}_n^* = \{1, 2, 3, \dots, n-1\}$$

(a) We need to prove that  $h \in \langle g \rangle^* \implies \text{ord}(h) = n$ .

If  $h \in \langle g \rangle^*$ , we have  $x$  in  $\{1, 2, 3, \dots, n-1\}$  such that

$$h = g^x$$

Since  $h \in \langle g \rangle$ , it belongs to the finite group as  $g$  itself has a finite order  $n$ . This implies  $h$  has a finite order as otherwise  $\{h^1, h^2, h^3, \dots\}$  will become an infinite group. This is a contradiction because each of these terms can be also expressed as powers of  $g$  as  $h = g^x$ . This is a contradiction as  $g$  is of finite order and so it cannot have infinite distinct number of powers.

Therefore,

$$\exists a; h^a = e$$

We have shown that  $h$  has a finite order. Now we will show that this order must be equal to  $n$ .

**If  $a < n$ :**

In this case, we have

$$e = h^a = g^{ax}$$

As  $a, x < n$ , we have:  $g^{ax} = e$ . As inverse exists in  $\langle g \rangle$ , consider  $g^{ax-kn}$  such that  $ax-kn < n$ . In this case, we have a  $y = ax - kn < n$  such that

$$g^y = e$$

This is a contradiction as we know that  $g$ 's order is  $n$ . Therefore  $a \geq n$ .

**If  $a > n$ :** This cannot be the case because  $h^n = g^{nx} = e$ . Therefore,  $a$  has to be less than  $n$ .

Hence we have proved that  $a = \text{ord}(h) = n$ .

(b) We need to prove that  $\text{ord}(h) = n \implies \langle h \rangle = \langle g \rangle$

We have  $\text{ord}(h)$  is  $n$ . Firstly, note that as  $h \in \langle g \rangle$ , we have  $\exists a$  such that  $h = g^a$ .

Now, returning to the question, we need to basically prove that:

$$\langle h \rangle = \{h, h^2, h^3, \dots\} = \{g, g^2, g^3, \dots\} = \langle g \rangle$$

As  $\text{ord}(h) = n = \text{ord}(g)$ , this is same as showing:

$$\{h, h^2, h^3, \dots, h^n\} = \{g, g^2, g^3, \dots, g^n\}$$

As  $n$  is prime, all the elements in  $\{h, h^2, \dots, h^n\}$  are different

This is because if  $h^i = h^j$ , then  $h^{i-j} = 1$  but  $i - j < n$ . Also, observe that any power of  $h$  is a power of  $g$  as  $h = g^a$ . As  $g$ 's power is  $n$ , there are only ' $n$ ' distinct powers of  $g$  possible (which are simply  $\{g, g^2, \dots, g^n\}$ ). Thus from the above observations,

$$\langle h \rangle = \{h, h^2, \dots, h^n\} = \{g, g^2, \dots, g^n\} = \langle g \rangle$$

$$\text{Hence, } \langle h \rangle = \langle g \rangle$$

(c) We need to prove that  $\langle h \rangle = \langle g \rangle \implies \langle h \rangle^* = \langle g \rangle^*$

Since  $n$  is prime, all the numbers in  $\{1, 2, \dots, n-1\}$  are co-prime to it. We need to show that if

$$\{h, h^2, h^3, \dots\} = \{g, g^2, g^3, \dots\}$$

then

$$\{h, h^2, \dots, h^{n-1}\} = \{g, g^2, \dots, g^{n-1}\}$$

From the common information, we are given that  $h \in \langle g \rangle$ . Therefore, for some  $a \in \mathbb{N}$ ,  $h = g^a$ .

From this, we can arrive at the proof by considering three facts:

- 1) All of  $\{h, h^2, \dots, h^{n-1}\}$  are all different because  $n$  is prime.
- 2) Each power of  $h$  is also a power of  $g$ .
- 3) There are at most  $n$  distinct powers of  $g$ .

The first one is true because  $\langle h \rangle = \langle g \rangle \implies h \neq 1$  and  $g$ 's order is  $n$ . Other two claims are also easy to prove. From the above three claims, following similar proof as above, we get

$$\{h, h^2, \dots, h^{n-1}\} = \{g, g^2, \dots, g^{n-1}\}$$

(d) We need to prove that  $\langle h \rangle^* = \langle g \rangle^* \implies h \in \langle g \rangle^*$

$\langle h \rangle^*$  has all generators of  $\langle h \rangle$ . Hence, obviously  $h \in \langle h \rangle^*$ . As  $\langle h \rangle^* = \langle g \rangle^*$ , then we directly have  $h \in \langle g \rangle^*$  as  $h \in \langle h \rangle^*$ .

2. (2 points) Show that  $a^{\log_h b} = b^{\log_h a}$  for any  $a, b \in \langle g \rangle$  and  $h \in \langle g \rangle^*$ .

**Solution:** As  $h$  generates  $\langle g \rangle$ , we have  $i, j$  such that  $a = h^i$  and  $b = h^j$ .

$$a^{\log_h b} = a^j = h^{ij}$$

$$b^{\log_h a} = b^i = h^{ij}$$

Hence  $a^{\log_h b} = b^{\log_h a}$ .

## 2 Probability Theory

Let  $X, Y$  be two random variables and  $V$  denote the set of possible values for  $X$  and  $Y$ .  $\Delta(X; Y)$  represents the statistical distance between  $X$  and  $Y$ .

1. (2 points) Prove the following proposition:  $\Delta(X; Y) = 1 - \sum_{v \in V} \min(\Pr[X = v], \Pr[Y = v])$ .

**Solution:** "The" statistical distance refers to Total Variation Distance. In simple terms, it is defined as the largest difference between the probabilities assigned by two distributions to an event, i.e., a subset in the probability space.

**Claim:** Largest differentiating subset is the set of  $a$ 's such that  $\Pr(Y=a) > \Pr(X=a)$ . i.e.,

$$S_{max} = \{a | \Pr(Y = a) > \Pr(X = a)\}$$

**Proof:** To begin with, if any set maximizing this probability difference has at least one element such that  $\Pr(X=a) > \Pr(Y=a)$ , then simply removing such elements would increase the probability difference. Therefore, any maximizing set should not have such elements.

Moreover, if a set does not contain at least one element that satisfies  $\Pr(Y=a) > \Pr(X=a)$ , then simply adding such an element would increase the probability. Therefore, all such elements must be present in such a maximizing set.

Hence we have

$$S_{max} = \{v | \Pr(Y = v) > \Pr(X = v)\}$$

(Note: it can be symmetrically opposite if we consider  $\Pr(Y=v) < \Pr(X=v)$ .)

Therefore, we know that:

$$\Delta(X; Y) = \sum_{v \in S_{max}} (\Pr(X = v) - \Pr(Y = v))$$

Considering the complement event of  $S_{max}$  which elicits the same difference in probabilities, we have:

$$\Delta(X; Y) = \sum_{v \in S_{max}} (\Pr(X = v) - \Pr(Y = v)) = \sum_{v \notin S_{max}} (\Pr(Y = v) - \Pr(X = v))$$

From the two equalities above, we have:

$$\Delta(X; Y) = \sum_{v \in V} \frac{1}{2} |\Pr(X = v) - \Pr(Y = v)|$$

$$\Delta(X; Y) = \sum_{v \in V} \frac{1}{2} \{ \max(\Pr(X = v), \Pr(Y = v)) - \min(\Pr(X = v), \Pr(Y = v)) \}$$

We know that  $\sum_{v \in V} \{ \max(\Pr(X = v), \Pr(Y = v)) + \min(\Pr(X = v), \Pr(Y = v)) \}$  is 2 as the probabilities over all elements in probability space add to two. By substituting this and cancelling half, we have:

$$\Delta(X; Y) = 1 - \sum_{v \in V} \min(\Pr[X = v], \Pr[Y = v])$$

2. (2 points) Prove the triangle inequality:  $\Delta(X; Z) \leq \Delta(X; Y) + \Delta(Y; Z)$

**Solution:** To prove this, we need to use a slightly different formula from that derived in Q2.1.

We use,

$$\Delta(X; Z) = \sum_{v \in V} \frac{1}{2} \{ \max(\Pr(X = v), \Pr(Z = v)) - \min(\Pr(X = v), \Pr(Z = v)) \}$$

Now, consider what happens when a new distribution '**Y**' comes into play. At a given  $v \in V$  assume that  $\Pr(Z = v) \geq \Pr(X = v)$ , we arrive at three cases:

**Case1:**  $\Pr(Y = v) \geq \Pr(Z = v) \geq \Pr(X = v)$ .

In this case,

$$(\Delta(X; Y) + \Delta(Y; Z))|_v \geq \Delta(X; Z)|_v$$

This is because,

$$\max(\Pr(X = v), \Pr(Y = v)) - \min(\Pr(X = v), \Pr(Y = v)) = \Pr(Y = v) - \Pr(X = v)$$

and its already  $\geq$

$$\max(\Pr(X = v), \Pr(Z = v)) - \min(\Pr(X = v), \Pr(Z = v)) = \Pr(Z = v) - \Pr(X = v)$$

**Case2:**  $\Pr(Z = v) \geq \Pr(Y = v) \geq \Pr(X = v)$ .

In this case,

$$(\Delta(X; Y) + \Delta(Y; Z))|_v = \Delta(X; Z)|_v$$

This is because,

$$\max(\Pr(X = v), \Pr(Y = v)) - \min(\Pr(X = v), \Pr(Y = v)) = \Pr(Y = v) - \Pr(X = v)$$

$$\text{and } \max(\Pr(Z = v), \Pr(Y = v)) - \min(\Pr(Z = v), \Pr(Y = v)) = \Pr(Z = v) - \Pr(Y = v)$$

and their sum is exactly same as

$$\max(\Pr(X = v), \Pr(Z = v)) - \min(\Pr(X = v), \Pr(Z = v)) = \Pr(Z = v) - \Pr(X = v)$$

**Case3:**  $\Pr(Z = v) \geq \Pr(X = v) \geq \Pr(Y = v)$

In this case,

$$(\Delta(X; Y) + \Delta(Y; Z))|_v \geq \Delta(X; Z)|_v$$

This is because,

$$\max(\Pr(Z = v), \Pr(Y = v)) - \min(\Pr(Z = v), \Pr(Y = v)) = \Pr(Z = v) - \Pr(Y = v)$$

and its already  $\geq$

$$\max(\Pr(X = v), \Pr(Z = v)) - \min(\Pr(X = v), \Pr(Z = v)) = \Pr(Z = v) - \Pr(X = v)$$

We have shown that for all  $v$  such that  $\Pr(Z = v) \geq \Pr(X = v)$ ,  $(\Delta(X; Y) + \Delta(Y; Z))|_v \geq \Delta(X; Z)|_v$ . By symmetry, we can see that this holds also if  $\Pr(Z = v) \leq \Pr(X = v)$ .

Hence taking  $\sum_{v \in V}$ , we have:

$$\sum_{v \in V} (\Delta(X; Y) + \Delta(Y; Z))|_v \geq \Delta(X; Z)|_v$$

Such addition is justified because we are simply adding the (max-min of distributions) at each element in the system of events. Therefore, we have:

$$\Delta(X; Y) + \Delta(Y; Z) \geq \Delta(X; Z)$$

3. (2 points) Prove that  $\Delta(f(X); f(Y)) \leq \Delta(X; Y)$  for any function  $f$  defined on  $V$ .

**Solution:** Suppose the function  $f$  maps the space of events in  $V$  to the space of events in  $W$ , i.e.,  $f: V \rightarrow W$ . Now, consider  $W$ , the co-domain and let  $W'$  be the range of  $f$ . Hence  $W' \subseteq W$

The contribution to variational distance will come only from the elements in  $W'$ . This is because both  $f(X)$  and  $f(Y)$  will have zero probability in  $W \setminus W'$ .

If  $f$  is one-one, then there is an isomorphism from  $V$  to  $W'$  (bijection) hence, there will be no difference in  $\Delta(f(X); f(Y))$  and  $\Delta(X; Y)$ . To precisely prove this, we can consider  $Pr(f(X) = a) = Pr(X = f^{-1}(a))$  as inverse can be properly defined.

$$\text{Hence, } f \text{ is one-one} \implies \Delta(f(X); f(Y)) = \Delta(X; Y)$$

In the case that  $f$  is not a one-one function, there is at least one element in  $W'$  that is mapped to two elements in the domain.

**Claim:** If  $f$  is not one-one, then  $\Delta(f(X); f(Y)) \leq \Delta(X; Y)$ .

**Proof:** From Q2.1, we have

$$\Delta(X; Y) = 1 - \sum_{v \in V} \min(\Pr[X = v], \Pr[Y = v])$$

For the element  $y$  that is mapped has two pre-images, say  $a, b$ , we have:

$$\Pr(f(X) = y) = \Pr(X = a) + \Pr(X = b) \text{ and } \Pr(f(Y) = y) = \Pr(Y = a) + \Pr(Y = b).$$

Therefore, we have:

$$\min(\Pr[f(X) = y], \Pr[f(Y) = y]) > \min(\Pr[X = a], \Pr[Y = a])$$

Summing over all  $y \in W'$ , we have:

$$\sum_{y \in W'} [\min(\Pr[f(X) = y], \Pr[f(Y) = y])] > \sum_{a \in V} \min(\Pr[X = a], \Pr[Y = a])$$

Note that ' $a$ ' cannot obviously have two images as  $f$  is a function and hence the above summations cover all elements in  $V$  in the LHS.

$$\Delta(f(X); f(Y)) \leq \Delta(X; Y)$$

We have shown that whether  $f$  is one-one or not,  $\Delta(f(X); f(Y)) \leq \Delta(X; Y)$ . Therefore, we can state absolutely that  $\Delta(f(X); f(Y)) \leq \Delta(X; Y)$ .

4. (4 points) For  $n \geq 1$ , let  $X \in_R \mathbb{Z}_n$  and  $Y \in_R \mathbb{Z}_n^*$ .

- (a) (2 points) Determine  $\Delta(X; Y)$ .

**Solution:**  $\mathbb{Z}_n = \{0, 1, 2, \dots, n-1\}$  and  $\mathbb{Z}_n^* = \{1, \dots, n-1\}$  ( $\mathbb{Z}_n^*$  considers only co-primes).

Now, using result from Q2.1, we have:

$$\begin{aligned}\Delta(X; Y) &= 1 - \sum_{n \in \mathbb{N}} \min(\Pr(X = n), \Pr(Y = n)) \\ \implies \Delta(X; Y) &= 1 - \sum_{n \in \{0, 1, 2, \dots, n-1\}} \min(\Pr(X = n), \Pr(Y = n)) \\ \implies \Delta(X; Y) &= 1 - \sum_{n \in \mathbb{Z}_n^*} \min(\Pr(X = n), \Pr(Y = n)) - \sum_{n \in \mathbb{Z}_n \setminus \mathbb{Z}_n^*} \min(\Pr(X = n), \Pr(Y = n))\end{aligned}$$

(Dividing them into disjoint sets)

As  $\Pr(Y=n) = 0$  for  $n \in \mathbb{Z}_n \setminus \mathbb{Z}_n^*$ , we have the right most term as zero.

$$\implies \Delta(X; Y) = 1 - \sum_{n \in \mathbb{Z}_n^*} \min(\Pr(X = n), \Pr(Y = n))$$

Now, we observe that within  $\mathbb{Z}_n^*$ ,  $\Pr(X=n)$  is lesser than  $\Pr(Y=n)$ , i.e.,

$$\forall n \in \mathbb{Z}_n^*, \Pr(X = n) < \Pr(Y = n)$$

Substituting this result and considering that  $|\mathbb{Z}_n^*| = \varphi(n)$ , we get:

$$\Delta(X; Y) = 1 - \frac{\varphi(n)}{n}$$

- (b) (2 points) Show that  $\Delta(X + Y; XY) = 0$ , where addition and multiplication are done modulo  $n$ .

**Solution:** To show this, we need to rely on an interesting observation about  $\mathbb{Z}_n^*$ .

$$\forall x \in \mathbb{Z}_n^*, \forall y \in \mathbb{N} \quad n|xy \implies n|y$$

**Claim1:** For an  $x \in \mathbb{Z}_n^*$ ,  $\{x, 2x, 3x, 4x, \dots, nx\}$  must all represent different numbers.

**Proof:** This is true because if any of  $ax = bx$ , then  $(a-b)x = 0$ . As  $n|0$ , and  $n \nmid x$ ,  $n$  must divide  $a-b$ . However, as  $a-b < n$  and  $a, b$  are distinct,  $n$  cannot divide  $a-b$ . Hence a contradiction.

From here, the proof of  $\Delta(X + Y; XY) = 0$  is straightforward. Consider any element 'a' in  $\mathbb{Z}_n$ . In the modulo 'n' world, we see that it is displaced by a unique amount w.r.t any element in  $\mathbb{Z}_n^*$ , i.e.,

$$\forall a \in \mathbb{Z}_n \quad \forall n \in \mathbb{Z}_n^* \quad \exists! d \in \mathbb{Z}_n \text{ such that } a = n + d$$

Now, from claim1, we have that all of  $\{x, 2x, 3x, \dots, nx\}$  must be different in modulo-n world. This means that this set should cover all of  $\{0, 1, 2, \dots, n-1\}$  in the modulo world as the cardinalities of  $\{0, 1, 2, \dots, n-1\}$  and  $\{x, 2x, \dots, nx\}$  are equal. From the above two arguments, we have:

$$\forall a \in \mathbb{Z}_n \quad \forall n \in \mathbb{Z}_n^* \quad \exists! d \in \mathbb{Z}_n \text{ such that } a = n * d$$

Note that we don't use the term 'nx', instead we use '0\*x' if we need to arrive at 0 in  $\mathbb{Z}_n$ .

**Claim2:**  $X + Y$  and  $XY$  define two probability distributions over  $0, 1, \dots, n - 1$  and they are exactly equal.

**Proof:** We have seen that there are unique elements  $d_1, d_2$  in  $\mathbb{Z}_n^*$  for any element  $n$  in  $\mathbb{Z}_n^*$  to reach any element  $a$  in  $\mathbb{Z}_n$  through modulo addition and multiplication respectively.

Therefore,

$$Pr(X + Y = a) = \sum_{n \in \mathbb{Z}_n^*} Pr(X = n) * Pr(Y = d_1) = \sum_{n \in \mathbb{Z}_n^*} Pr(X = n) * \frac{1}{n} = \sum_{n \in \mathbb{Z}_n^*} \frac{1}{\varphi(n)} * \frac{1}{n} = \frac{1}{n}$$

Also,

$$Pr(XY = a) = \sum_{n \in \mathbb{Z}_n^*} Pr(X = n) * Pr(Y = d_2) = \sum_{n \in \mathbb{Z}_n^*} Pr(X = n) * \frac{1}{n} = \sum_{n \in \mathbb{Z}_n^*} \frac{1}{\varphi(n)} * \frac{1}{n} = \frac{1}{n}$$

Here, we have used that the variables  $X, Y$  are randomly chosen from their respective domains. Therefore, as our choice of 'a' has been arbitrary, we can conclude that  $X + Y$  and  $XY$  define the same probability distribution on all  $a \in \{0, 1, \dots, n - 1\}$ . Hence, from claim-2,  $\Delta(X + Y; XY) = 0$

5. (2 points) For  $n \geq d \geq 1$ , let random variable  $X$  take on values in  $\{0, \dots, d - 1\}$ , and let  $U \in_R \{0, \dots, n - 1\}$ . Show that  $\Delta(U; X + U) \leq (d - 1)/n$ , and that this bound is tight.

**Solution:** We can see that the random variable  $X + Y$  ranges from 0 to  $n + d - 2$ . As  $U$  is in  $\{0, 1, \dots, n - 1\}$ , only these values contribute to the Total variational distance  $\Delta(U; X + U)$ .

Hence, we can determine  $\Delta(U; X + U)$  by considering the distribution of  $X + U$  only in this range of  $\{0, 1, \dots, n - 1\}$ .

Let us determine this probability distribution:

$$Pr(X + U = a) = \begin{cases} \frac{1}{n} Pr(X \leq a) & \text{if } 0 \leq a \leq d - 1 \\ \frac{1}{n} & \text{if } d \leq a \leq n - 1 \end{cases} \quad \text{This is because, if } a \geq d, \text{ there is always a 'b' such}$$

that  $a + b = n$ , hence the probability is simply  $Pr(U = b) = \frac{1}{n}$ . However, if the case that  $a \leq d - 1$ , there is a possibility that  $X > a$  and hence there would be no  $U$  such that  $X + U = a$ . Therefore, we have the term  $Pr(X \leq a)$ .

Now, from Q2.1, we have:

$$\Delta(X; Y) = 1 - \sum_{v \in V} \min(Pr[X = v], Pr[Y = v])$$

On substituting  $Y$  by  $X + U$  and considering the appropriate range, we have:

$$\Delta(X; X + U) = 1 - \sum_{a \in \{0, 1, \dots, n - 1\}} \min(Pr[U = a], Pr[X + U = a])$$



Considering RHS, we can split it as:

$$\Delta(X; X+U) = 1 - \sum_{a \in \{0,1,\dots,d-1\}} \min(\Pr[U = a], \Pr[X+U = a]) - \sum_{a \in \{d,d+1,\dots,n-1\}} \min(\Pr[U = a], \Pr[X+U = a])$$

Considering the distribution of  $X + U$  as defined above, we have the last term as :

$$\Delta(X; X + U) = 1 - \sum_{a \in \{0,1,\dots,d-1\}} \min(\Pr[U = a], \Pr[X + U = a]) - \frac{1}{n} * (n - d)$$

Now, expanding the middle term, we have:

$$\begin{aligned} \Delta(X; X + U) &= 1 - \sum_{a \in \{0,1,\dots,d-1\}} \frac{1}{n} (\Pr(X \leq a)) - \frac{1}{n} * (n - d) \\ \Rightarrow \Delta(X; X + U) &= \frac{d}{n} - \frac{1}{n} * \sum_{a \in \{0,1,\dots,d-1\}} (\Pr(X \leq a)) \end{aligned}$$

Now, for a fixed 'd', we know that the right term in RHS is greater than 1. This is because, when  $a = d - 1$ , the probability  $\Pr(X \leq a) = 1$  for any distribution of 'X' and the other probabilities in the summation are at least zero. Hence, we have, **over all distributions of X**:

$$\Delta(X; X + U) \leq \frac{(d - 1)}{n}$$

The argument above gives us the idea on why the bound must be tight. It is tight when  $\Pr(X \leq a) = 1$  when  $a = d - 1$  and is zero when  $a < d - 1$ .

We can **construct a distribution where  $\Pr(X = d - 1) = 1$  and is 0 otherwise**. For this, from the formula in Q2.1, we have:

$$\begin{aligned} \Delta(X; X + U) &= 1 - 0 * (d - 1) - \frac{1}{n} * ((n - 1) - (d - 1) + 1) \\ \Delta(X; X + U) &= \frac{(d - 1)}{n} \end{aligned}$$

Hence the bound is tight.

6. (4 points) For  $n \geq 1$ , consider distributions  $X, Y, Z$  given by

$$\begin{aligned} X &= \{u : u \in_R \{0, \dots, n - 1\}\}, \\ Y &= \{2u : u \in_R \{0, \dots, n - 1\}\}, \\ Z &= \{2u + 1 : u \in_R \{0, \dots, n - 1\}\}. \end{aligned}$$

Clearly,  $\Delta(Y; Z) = 1$ .

(a) (2 points) Show that  $\Delta(X; Y) = \Delta(X; Z) = 1/2$  for even  $n$

**Solution:** We have:

$$X \in_R \{0, 1, 2, \dots, n-1\},$$

$$Y \in_R \{0, 2, 4, \dots, 2n-2\},$$

$$Z \in_R \{1, 3, 5, \dots, 2n-1\}.$$

Considering the distributions in range of  $X$  will be enough when we consider the formula in Q2.1. When  $n$  is even, say  $n = 2k$ , then number of values of  $Y$  in  $\{0, 1, \dots, 2k-1\}$  are  $k$  exactly. Using Q2.1, we have:

$$\Delta(X; Y) = k * (0) + k * \left(\frac{1}{n}\right)$$

$$\text{Hence, } \Delta(X; Y) = k * \left(\frac{1}{n}\right) = \frac{1}{2}$$

Similarly, then number of values of  $Z$  in  $\{0, 1, \dots, 2k-1\}$  are  $k$  exactly. Using Q2.1, we have:

$$\Delta(X; Z) = k * (0) + k * \left(\frac{1}{n}\right)$$

$$\text{Hence, } \Delta(X; Z) = k * \left(\frac{1}{n}\right) = \frac{1}{2}$$

(b) (2 points) Determine  $\Delta(X; Y)$  and  $\Delta(X; Z)$  for odd  $n$ .

**Solution:** Similar to the above question, When  $n$  is odd, say  $n = 2k+1$ , then we have exactly  $k+1$  even numbers in  $0, 1, \dots, 2k$  and  $k$  odd numbers in  $0, 1, \dots, 2k$ . Therefore, the corresponding statistical distances are:

$$\Delta(X; Y) = k * (0) + (k+1) * \left(\frac{1}{n}\right) = \frac{n+1}{2n}$$

$$\Delta(X; Z) = (k+1) * (0) + k * \left(\frac{1}{n}\right) = \frac{n-1}{2n}$$

7. (5 points) For  $n$  prime, let  $h$  and  $M_0$  be arbitrary, fixed elements of  $G_n = \langle g \rangle, h \neq 1$ . Consider distributions  $X, Y, Z$  given by

$$X = \{(A, B) : A \in_R \langle g \rangle, B \in_R \langle g \rangle\},$$

$$Y = \{(g^u, h^u M) : u \in_R \mathbb{Z}_n, M \in_R \langle g \rangle\},$$

$$Z = \{(g^u, h^u M_0) : u \in_R \mathbb{Z}_n\}.$$

**Solution:** Across, all parts of the question, as  $h$  and  $M_0$  are fixed, let us consider them as  $h = g^i$  and  $M_0 = g^j$ . Also, as  $\langle g \rangle = G_n$ , we have:  $\langle g \rangle = \{g, g^2, g^3, \dots, g^n\}$ .

(a) (2 points) Show that  $\Delta(X; Y) = 0$

**Solution:** The random variable  $X$  is an ordered pair of two powers taken uniformly randomly from  $\langle g \rangle$ . The left coordinate of  $Y$  is also uniformly randomly chosen power of  $g$ . Hence the distributions are identical with respect to the left coordinate. We claim that the RHS's distributions are also identical.

**Claim:**  $h^u M$  is equivalent to the uniform distribution as defined by B.

**Proof:** From our pre-assumptions, we have:  $h = g^i$ . Let us say  $M = g^k$ , then  $h^u M = g^{iu+k}$  where  $u, k \in_R \mathbb{Z}_n$ . ( $\mathbb{Z}_n$  because when  $k = n$ , we can consider  $k$  as 0 instead and the distributions wouldn't change.)

Analogous to 'Affine cipher' we will show that the randomness offered by 'k' will be enough to make the distribution of ' $iu + k$ ' uniformly random.

**Proof:** As order of  $g$  is 'n' all the powers can be considered modulo  $n$ . Therefore,  $Pr(iu + k = a) = \sum_{y \in \mathbb{Z}_n} Pr(i = y) * Pr(k = (a - yu) \bmod n)$ . As  $(a - yu) \bmod n$  lies in  $\mathbb{Z}_n$  for all  $y$ , this probability must be simply  $\frac{1}{n}$ . Therefore, we have:

$$Pr(iu + k = a) = \sum_{y \in \mathbb{Z}_n} Pr(i = y) * \frac{1}{n}$$

$$Pr(iu + k = a) = \frac{1}{n} * \sum_{y \in \mathbb{Z}_n} Pr(i = y)$$

As this summation is 1, we have:

$$Pr(iu + k = a) = \frac{1}{n}$$

Hence, the RHS is equivalent to uniform distribution. Therefore, as left and right coordinates of  $X, Y$  match in their respective ranges, the joint probabilities of their coordinates also match over the cross product of ranges. Hence the distributions are **exactly the same** and thus:

$$\Delta(X; Y) = 0$$

We did not even use  $h \neq 1$  because that will not affect the summation of probabilities which will be 1 anyway.

(b) (2 points) Show that  $\Delta(Y; Z) = 1 - 1/n$

**Solution:** Let us look at the Probability distributions:

$$Pr(X = (g^a, g^b)) = \left(\frac{1}{n} * \frac{1}{n}\right)$$

$$Pr(Z = (g^a, g^b)) = \text{Let us calculate}$$

To calculate the probability distribution of  $Z$ , let us first look at the ranges of  $a, b, u, i, j$ :

As  $h \neq 1$ ,  $i \in \{1, 2, \dots, n-1\}$ .

$a, b, j \in \{1, 2, \dots, n\}$

$u \in \{0, 1, \dots, n-1\}$

Now, for  $g^u = g^a \implies u = a$  as we have considered their ranges only in the modulo world. Also,

$$h^u \mathbb{M}_0 = g^b \implies iu + j = b.$$

For  $iu+j=b$  to be true,

$$u = i^{-1} * (b - j)$$

As  $n$  is prime and  $i^{-1}$  is defined uniquely. Hence we have the condition that

$$a = i^{-1} * (b - j)$$

This happens only in  $n$  pairs of  $(g^a, g^b)$ . And for each of this pair, the probability that  $Pr(Z = (g^a, g^b))$  is only  $\frac{1}{n}$  as 'u' is chosen randomly. Therefore, from formula in Q2.1,

$$\Delta(Y; Z) = 1 - (n^2 - n) * 0 - n * \left(\frac{1}{n^2}\right)$$

$$\Delta(Y; Z) = 1 - \frac{1}{n}$$

(c) (1 point) Show that  $\Delta(X; Z) = 1 - 1/n$

**Solution:** We have shown that distributions of  $X, Y$  are identical at all points. Since statistical distance strictly depends only on the probability value at each point in the domain, we can infer that

$$\Delta(X; Z) = 1 - 1/n$$

## Omitted

8. (0 points) Prove the following proposition:

1.  $0 \leq \Delta(X; Y) \leq 1$ , “nonnegativity” and “boundedness”.
2.  $\Delta(X; Y) = 0$  if and only if  $\forall_{v \in V} \Pr[X = v] = \Pr[Y = v]$ , “identical distributions”.
3.  $\Delta(X; Y) = \Delta(Y; X)$ , “symmetry”.
4.  $\Delta(X; Z) \leq \Delta(X; Y) + \Delta(Y; Z)$ . “triangle inequality”.

9. (0 points) Prove the following proposition:

1.  $\Delta(X; Y) = \sum_{v \in V^+} (\Pr[X = v] - \Pr[Y = v])$ , with  $V^+ = \{v \in V : \Pr[X = v] > \Pr[Y = v]\}$ .
2.  $\Delta(X; Y) = \sum_{v \in V} (\Pr[X = v] \dot{-} \Pr[Y = v])$ , with  $x \dot{-} y = \max(x - y, 0)$  (“x minus y”).
3.  $\Delta(X; Y) = 1 - \sum_{v \in V} \min(\Pr[X = v], \Pr[Y = v])$ .
4.  $\Delta(X; Y) = \max_{W \subseteq V} |\Pr[X \in W] - \Pr[Y \in W]|$ .

10. (0 points) For  $n, d \geq 1$ , consider distributions  $X$  and  $Y$  given by

$$X = \{u : u \in_R \{0, \dots, n-1\}\},$$

$$Y = \{u + d : u \in_R \{0, \dots, n-1\}\}.$$

Determine  $\Delta(X; Y)$ , assuming  $d \leq n$ . Also, what is  $\Delta(X; Y)$  if  $d > n$ ?

### 3 One Way Functions/ Permutations

1. (2 points) Let  $g_1 : \{0, 1\}^n \rightarrow \{0, 1\}^n$  and  $g_2 : \{0, 1\}^n \rightarrow \{0, 1\}^n$  be two length preserving one-way functions. Define  $f(x) = g_1(x) || g_2(x)$ . Show that  $f$  is not necessarily a one way function.

**Solution:**

It will be enough if we show one counterexample wherein  $f(x)$  will not be a OWF.

**Construction:**

Consider 'g' to be an arbitrary OWF from  $\{0, 1\}^{\frac{n}{2}}$  to  $\{0, 1\}^{\frac{n}{2}}$ . Now, select  $g_1(x_1 \circ x_2) = x_1 \circ g(x_2)$  and  $g_2(x_1 \circ x_2) = g(x_1) \circ x_2$ .

**Claim:**  $g_1$  and  $g_2$  are both OWFs.

**Proof:** We will first show that  $g_1$  is a OWF. By symmetry,  $g_2$  will also be a OWF.

Assume that  $g_1$  is not one-way. Then we have a PPT Adversary A for  $g_1$  which can invert with non-negligible probability. From here, we have:

$$g(A(g_1(x_1 \circ x_2))_{[\frac{n}{2}, \dots, n]}) = g(x_2)$$

whenever A succeeds. Therefore, by taking the right half of the A's inversion, we can always succeed in inverting 'g' when A succeeds in inverting  $g_1$ .

Hence, a polynomial blackbox reduction exists between the adversary A of  $g_1$  to the shown adversary of  $g$  (say B). Also we have shown that whenever A inverts  $g_1$ , B inverts  $g$ . From our assumption that A inverts with a probability :  $\text{non-negl}(\frac{n}{2})$ , we can tell that B must invert with the same non-negl probability. But,  $\text{non-negl}(\frac{n}{2})$  is also  $\text{non-negl}(n)$ . This is a contradiction that  $g$  is a OWF.

Hence,  $g_1$  is a OWF. And, by symmetry, we have  $g_2$  is also a OWF.

From here, it is easy to see why  $f(x) = g_1(x) || g_2(x)$  is not a one-way function. We can rewrite the equation, by plugging in definitions as:  $f(x_1 \circ x_2) = x_1 \circ g(x_2) || g(x_1) \circ x_2$ . For a given  $y$ ,  $f^{-1}(y)$  can simply be found by considering first and last quarters of the  $y$  and invert with a probability 1, i.e.,

$$Pr(f(y_{[0,1,\dots,\frac{n}{2}]} || y_{[\frac{3n}{2}, \dots, 2n-1]})) = y = 1$$

Hence, there exists a perfect inversion for 'f' by just looking at the output in linear time. Hence **f is not a one-way function, although  $g_1$  and  $g_2$  are one-way.**

2. (4 points) Let  $f_1$  and  $f_2$  be one way functions, where  $\exists |x_1| = |x_2| \implies |f(x_1)| = |f(x_2)|$  (same sized output for same sized inputs). Let  $f(x) = f_1(x_1) \oplus f_2(x_2)$  where  $x = x_1 || x_2$  and  $|x_1| = |x_2|$  (assume even inputs).
- (a) Give an example where  $f_1$ ,  $f_2$  and  $f$  are one way functions.

**Solution:** To begin with, we need to assume that OWFs exist. Let  $g$  be one such OWF. It is easy to see that

$$f_1(x) = 0^n || g(x)$$

and

$$f_2(x) = g(x) || 0^n$$

are OWFs.

**Claim:**  $f(x) = f_1(x_1) \oplus f_2(x_2)$  must be a OWF.

**Proof:** We have

$$f(x_1 \circ x_2) = g(x_2) || g(x_1)$$

To prove that this  $f$  is a OWF, we can follow the proof strategies as done in assignment-2. To begin with, assume that 'f' is not a OWF. Then there exists a PPT non-negligible inverter say A. Let us use A to construct B (a PPT, non-negl inverter for g).

**Reduction from B to A:**

For a given  $y$  where  $|y| = n$ , we need to find a  $x \in g^{-1}(y)$  to successfully invert 'g'. Suppose the adversary 'B' has the access to the blackbox encrypter of 'g' (considering Kirchhoff's security). Then he can simply consider an arbitrary string 'z' such that  $|z| = n$  and encrypt it using 'g's blackbox. Now, B can append  $y$  and  $g(z)$  to get a '2n' string. He can simply feed this string to 'A' which will give him say 'r'.

$$\text{i.e., Say } A(g(z) || y) = r$$

Now, we have, if  $r = r_1 \circ r_2$ ,  $f(r) = g(r_2) || g(r_1)$ , in the scenarios where 'A' successfully inverts 'f'. Also, as  $f(r)$  is originally  $y || g(z)$ , we have  $g(r_2) = y$ . This means that whenever A succeeds, B can succeed by considering the second half on input to invert g. As A is non-negl adversary in '2n', B is also a non-negl adversary in '2n' and this is non-negl(n). Hence a contradiction that g is OW. **Therefore, f must be a OWF.**

To put simply, B's algorithm is:

B(y):

Consider a random n-bit string z

Find  $g(z)$

Give A the input  $g(z) \circ y$ , say  $r = A(g(z) \circ y)$

If  $r = r_1 \circ r_2$ , return  $r_2$

PS: For odd size inputs to  $f_1$  or  $f_2$ , split them unevenly, i.e., use  $f_1(x_1 \circ x_2) = 0^{n+1} || g(x_1)$  and  $f_2(x_1 \circ x_2) = g(x_1) || 0^{n+1}$  instead and ignore  $f[0]$  and  $f[-1]$ .

- (b) Give an example where  $f_1, f_2$  are one way but not  $f$ .

**Solution:** Following the hint given, let us set  $x$  to be equal to  $x_1 || x_2 || x_3 || x_4$  and the output  $f(x)$  to be  $y_1 || y_2$ . Now, let us consider the one-way functions  $f_1$  and  $f_2$  identical to those in question-1. i.e.,

$$f_1(x_1 \circ x_2) = x_1 \circ g(x_2)$$

and

$$f_2(x_3 \circ x_4) = g(x_3) \circ x_4$$

where 'g' is a original one-way function from  $\{0, 1\}^{\frac{n}{4}} \rightarrow \{0, 1\}^{\frac{n}{4}}$ .

From here, we can see that  $y_1 || y_2 = f(x_1 || x_2 || x_3 || x_4) = x_1 \oplus g(x_3) || x_4 \oplus g(x_2)$ . This means that for any given  $y_1 || y_2$ , we have the adversary A with access to g's encryption (considering Kirckhoff's security):

**Algorithm for A( $y_1 || y_2$ ):**

Consider random  $\frac{n}{2}$  bit strings a,b

Calculate  $y_a = g(a)$  and  $y_b = g(b)$

Consider  $c = y_1 \oplus y_b$  and  $d = y_2 \oplus y_a$

Return  $(c || a || b || d)$

Obviously, when we find  $f(c || a || b || d)$ , we see that it is equal to  $c \oplus g(b) || g(a) \oplus d$ , i.e.,

$$f(c || a || b || d) = c \oplus g(b) || g(a) \oplus d$$

$$f(c || a || b || d) = c \oplus y_b || y_a \oplus d$$

From our choices of c and d, this is same as:

$$f(c || a || b || d) = y_1 || y_2$$

Therefore, the given adversary A always succeeds in inverting 'f' with probability of 1. Hence **f is not a one-way function although  $f_1$  and  $f_2$  are.**

3. (8 points) Let  $g(x)$  be a length preserving one-way function. Let  $x = x_1 || x_2$  where  $x_1 = x_2$  (assume even inputs). Which of following are one-way functions? Prove your answers.

- (a)  $f_a(x) = g(\bar{x})$ , where  $\bar{x}$  is the bitwise compliment of  $x$

**Solution:** If there is a PPT, non-negligible inverter for  $f_a$ , say A, then we will try to construct a non-negligible PPT inverter B for g.

B(y):

$z := A(y)$

$x := \bar{z}$

**Claim:** This algorithm of B successfully inverts g whenever A inverts  $f_a$ .

**Proof:** Say  $A(g(\bar{x})) = y$ . If A inverts  $f_a$ , i.e.,  $g(\bar{x})$  successfully

$$g(\bar{y}) = g(\bar{x})$$

Therefore, given an input encryption of  $\bar{x}$ , it found a  $\bar{y}$  such that:

$$g(\bar{y}) = g(\bar{x})$$

Now, replace  $a = \bar{x}$  and  $b = \bar{y}$  in the above proposition without loss of generality.

Then we find get the new proposition:  $A(g(a)) = \bar{b}$ . If A inverts  $f_a$ , i.e.,  $g(a)$  successfully

$$g(b) = g(a)$$

$$\implies g(B(g(a))) = g(a)$$

Hence both A,B invert  $f_a, g$  respectively the same probability. As A inverts with non-negl, so does B and this creates a contradiction that  $g$  is one-way. **Therefore,  $f_a$  is a one-way function.**

(b)  $f_b(x) = g(x_1 \oplus x_2)$

**Solution:**  $f_b(x_1 || x_2) = g(x_1 \oplus x_2)$ . Suppose  $f_b$  is not a OWF. A PPT non-negl adversary for  $f_b$  say A, can take in  $g(x_1 \oplus x_2)$  and give us  $x' = x'_1 || x'_2$  such that  $g(x'_1 \oplus x'_2) = g(x_1 \oplus x_2)$ . i.e.,

$$A(g(x_1 \oplus x_2)) = x' \text{ such that } g(x_1 \oplus x_2) = g(x'_1 \oplus x'_2)$$

with non-negl probability. Let us construct B, an adversary for  $g$  as:

Algorithm for B( $g(x)$ ):

Find  $x' := A(g(x))$

Consider  $x' = x'_1 || x'_2$

Return  $x'_1 \oplus x'_2$

B will succeed whenever A does because of the definition of A's success that  $A(g(x_1 \oplus x_2)) = x'$  such that  $g(x_1 \oplus x_2) = g(x'_1 \oplus x'_2)$ . Therefore, as A is non-negl( $2n$ ), PPT( $2n$ ) it is also non-negl( $n$ ) and also PPT( $n$ ). Hence, a contradiction as B will now succeed with non-negl probability over input space. Therefore,  $f_b$  is a OWF.

$$\text{A part } f_c(x) = \begin{cases} 0^{|x|} & \text{if exactly one bit of } x_1 \text{ is 1} \\ 0^{|x_1|} \cdot g(x_2) & \text{otherwise} \end{cases}$$

**Solution:** If  $x = x_1 \circ x_2$ , and  $|x_1| = n = |x_2|$  the number of  $x_1$ 's that have exactly one bit as '1' are ' $n$ '. We have already proven that  $f(x) = 0^n || g(x_2)$  is a OWF.

Now we will show that setting  $f(x)$  to  $0^{2n}$  for  $\frac{n}{2^n}$  fraction of the input does not affect the one-wayness of  $f$ . Assume that an arbitrary adversary A of  $f$  could not invert any of these  $x$ 's successfully. Now the



adversary of  $f_c(x)$  say  $A_c$  can invert this fraction successfully with  $\Pr=1$ . This means that

$$\Pr[x \leftarrow \{0,1\}^{2n}, A_c(1^{2n}, f_c(x)) \in f_c^{-1}(f_c(x))] \leq \Pr[x \leftarrow \{0,1\}^{2n}, A(1^{2n}, f(x)) \in f^{-1}(f(x))] + \frac{n}{2^n} * 1$$

As  $A$  can be any adversary, let us take it equal to  $A_c$  except over this fraction of input and wrong-inverting over this  $\frac{n}{2^n}$  fraction of the input. In this case,  $f$  is a OWF, the RHS term is negligible. Hence the LHS cannot be non-negligible.

Hence  $f_c(x)$  is a OWF.

$$(c) f_d(x) = \begin{cases} 0^{|x|} & \text{if at least one bit of } x_1 \text{ is 1} \\ 0^{|x_1|} \cdot g(x_2) & \text{otherwise} \end{cases}$$

**Solution:** Just invert everything to  $0^{2n}$ , we will be correct with a probability of  $1 - \frac{1}{2^n}$  as only when  $x_1 = 0^n$ , we might be wrong. Therefore,  $f_d$  cannot be a OWF.

4. (2 points) We know that  $f$  may be one way but  $f(f(x))$  may not be one way. What about  $f(x)||f(f(x))$ ?

**Solution:**  $g(x) = f(x)||f(f(x))$  must be a OWF. We can give a reduction where if we have a non negligible PPT inverter black box for  $g$ , we can invert  $f$  also.

#### Reduction using Black Box Inverter of $g$ :

Suppose  $g$  is not a OWF, it has an PPT non-negl inverting adversary. Let  $B$  be one such inverter of ' $g$ '. Let us construct an adversary  $A$  for ' $f$ '. On receiving an string ' $y$ ',  $A$  can use  $f$ 's blackbox (considering Kirchhoff's security) to encrypt that string again. It can then append to get  $y||f(y)$ . Now, if this string is fed to  $B$ , we have, say :

$$B(y||f(y)) = x$$

Then, if  $B$  successfully inverter ' $g$ ', this means that:

$$y||f(y) = f(x)||f(f(x))$$

Hence, this means that  $f(x)$  must be  $y$ . This means that if ' $A$ ' simply returns whatever  $B$  returns, it has to follow:  $f(B(y||f(y))) = y$  whenever  $B$  is successful.

Hence,  $A$  succeeds whenever  $B$  does. This is a contradiction as now  $f$  can be inverted with non-negl probability. **Hence  $g$  must be a OWF.**

5. (3 points) Given a strong one way function  $f$ , construct a weak one way function  $g$  that is NOT a strong one way function.

**Solution:** Let us begin by defining two classes of One-Way Functions.

**Strong One-Way Functions:** A strong one-way function can be computed by a PPT and  $\forall$  PPT  $A$ ,  $\exists \text{negl } \epsilon$  and  $\forall n \in \mathbb{N}$  such that  $\Pr[x \leftarrow \{0,1\}^n, A(1^n, f(x)) \in f^{-1}(f(x))] \leq \epsilon(n)$

**Weak One-Way Functions:** A weak one-way function can be computed by a PPT and  $\exists$  polynomial  $q(x)$   $\forall$  PPT  $A$ ,  $\forall n \in \mathbb{N}$  such that  $\Pr[x \leftarrow \{0,1\}^n, A(1^n, f(x)) \in f^{-1}(f(x))] \leq 1 - \frac{1}{q(x)}$

We need to construct a weak one-way function from a strong one.

**Idea:** We construct a function  $g$  that acts like identity function on  $\frac{3}{4}$ <sup>th</sup> of the input and acts like 'f' on the rest of  $\frac{1}{4}$ <sup>th</sup>. This quartile is chosen tactically, as we will see. Now we show that any adversary that gets 'considerably' close to probability '1' of inversion, can invert 'f' with a non-negligible probability, thereby forming a contradiction.

**Construction:** Map  $\frac{3}{4}$ <sup>th</sup> of the input ( $I$ ) to itself, i.e.,  $g(x) = x$  and map  $\frac{1}{4}$ <sup>th</sup> of the input ( $I'$ ) to  $f$ , i.e.,  $g(x) = f(x)$ . **This subset of domain, the quartile  $I'$ , is chosen to be non-polynomial.** Such quartile must exist from pigeon hole principle as OWFs (i.e  $f$ ) cannot have a polynomial range.

**Claim:**  $G$  is a Weak One-Way Function.

**Proof:** We prove that it is a OWF and that it is not a strong OWF.

**It is weak and it is OW:**

We show that there cannot be an adversary that gets within inverse polynomial range to probability of 1. From the definition of weak one-way functions, we have  $g$  is a weak OWF if:

$$\exists \text{ polynomial } q(x), \forall \text{PPTA}, \forall n \in \mathbb{N} \text{ such that } \Pr[x \leftarrow \{0,1\}^n, A(1^n, f(x)) \in f^{-1}(f(x))] \leq 1 - \frac{1}{q(x)}$$

Suppose, this definition is not satisfied, i.e., we have an inverter  $A$  that can invert with a probability that is  $> 1 - \frac{1}{q(x)}$  for each polynomial  $q(X)$  for at least one  $n \in \mathbb{N}$ . Then, it has to invert the total range leaving out the  $g(x) = x$  edges with a probability  $> \frac{1}{4} - \frac{1}{q(x)}$  for all polynomials  $q(x)$ . All these inversions find a preimage into  $I'$  as we did not consider  $g(x)=x$  images.

**Arriving at the contradiction:** Now, considering the original function 'f', the same inverter can invert successfully with a probability  $> \frac{1}{4} - \frac{1}{q(x)}$ . This is because the edges of  $I'$  are not modified but only edges from  $I \setminus I'$  are put back. Clearly  $> \forall q \frac{1}{4} - \frac{1}{q(x)}$  implies a non-negligible fraction. Therefore, this **contradicts the assumption that  $f$  is a OWF.** Hence, no inverter can invert the constructed  $f(I')$  with a probability  $> \frac{1}{4} - \frac{1}{q(x)}$  and therefore, cannot invert 'g' with probability  $> 1 - \frac{1}{q(x)}$ . **Therefore,  $g$  is weak and  $g$  is OW.**

**It is not strong:**

This is the case because a trivial inverter that returns  $A^{-1}(x) = x$  would give us a probability of at least  $\frac{3}{4}$  which is not negligible. **Hence,  $g$  is not strong.**  
**Thus, we have constructed  $g$ , a weak OWF that is not strong.**