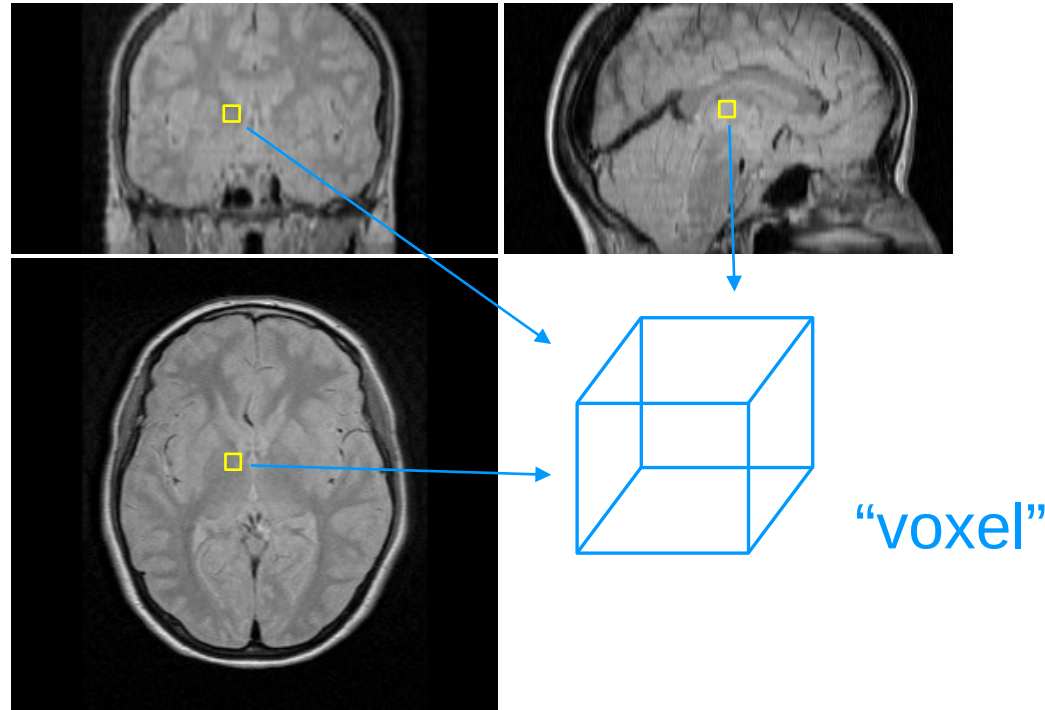


Model-based Segmentation: Part I

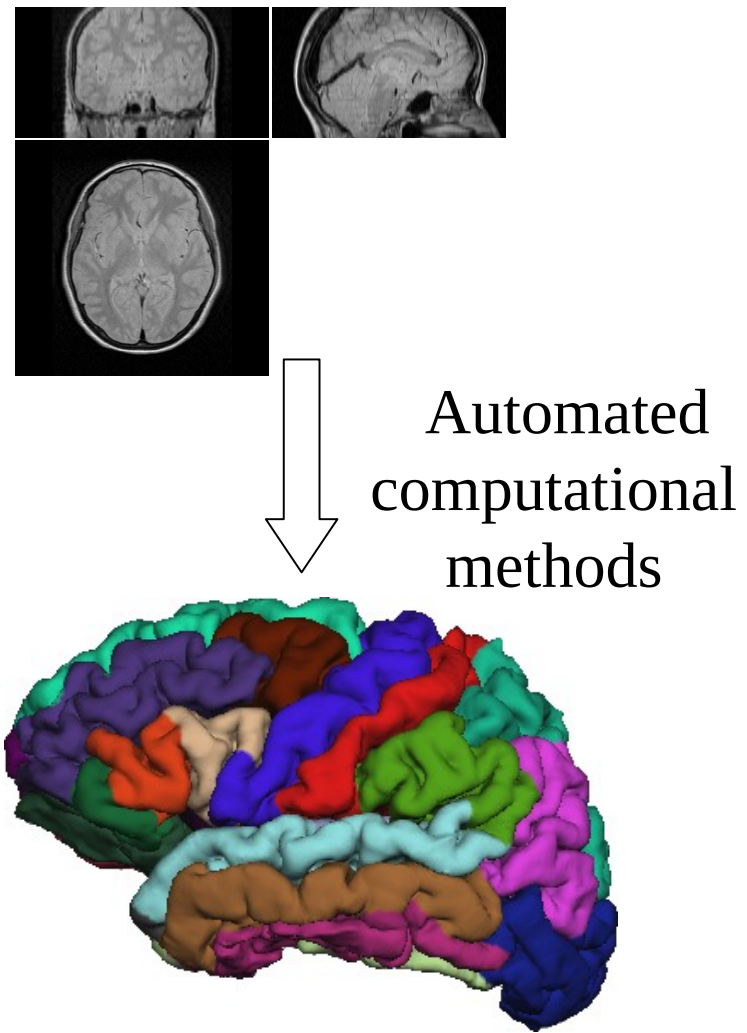
Voxel-based segmentation



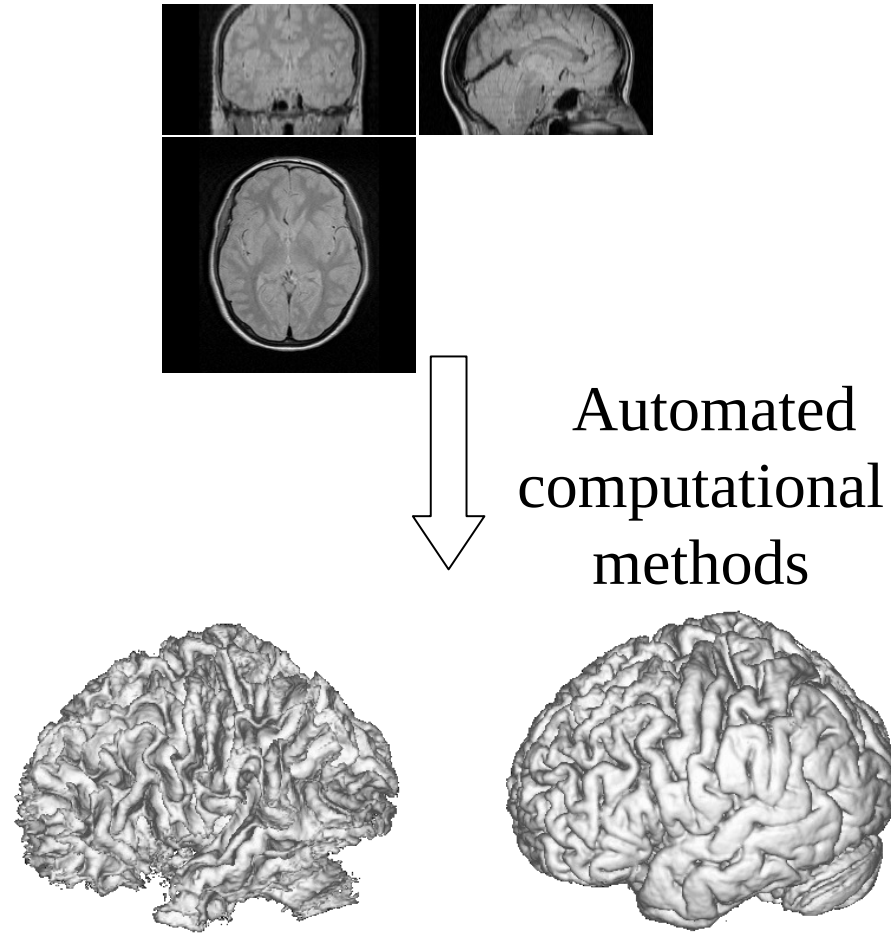
Determine to which anatomical structure each voxel in the image belongs:

- Think “LEGO bricks”
- Outer surfaces can easily be extracted if needed

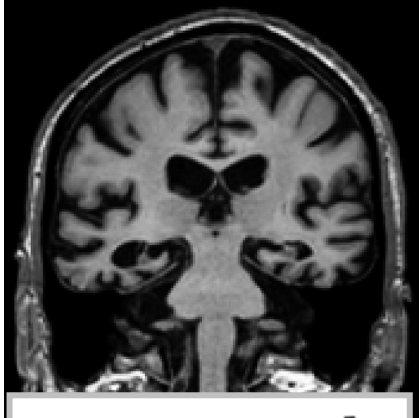
Voxel-based segmentation



This and next lecture



The problem to be solved



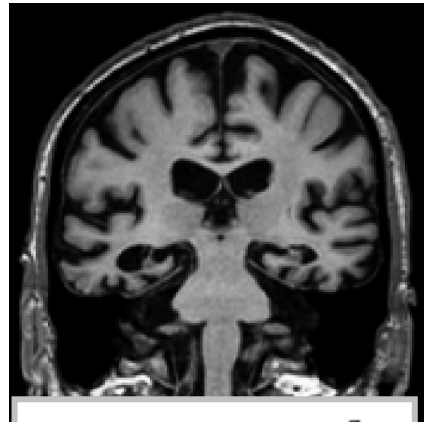
MRI image \mathbf{d}

N voxels

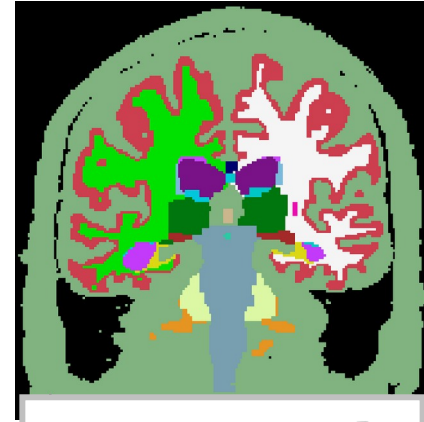
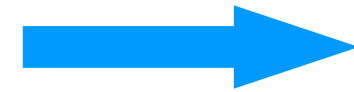
$$\mathbf{d} = (d_1, \dots, d_N)^T$$

d_n : intensity in voxel n

The problem to be solved



MRI image \mathbf{d}



Label image \mathbf{l}

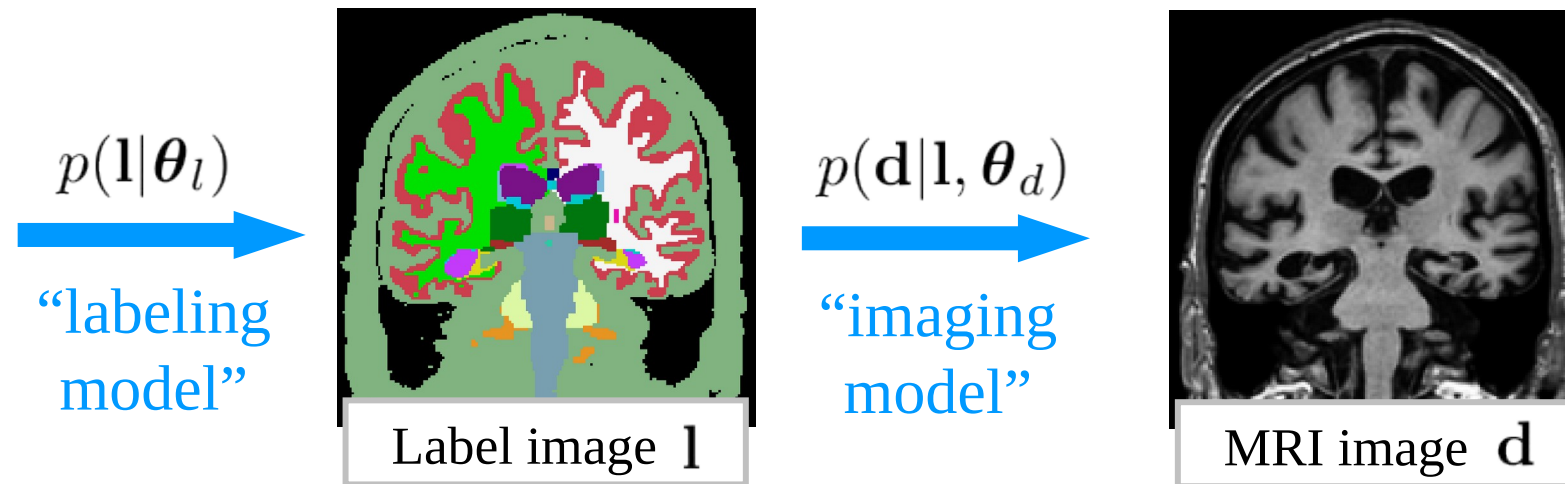
$$\mathbf{l} = (l_1, \dots, l_N)^T$$

$$l_n \in \{1, \dots, K\}$$

K : number of classes

One solution: generative modeling

- Formulate a statistical model of how a medical image is formed



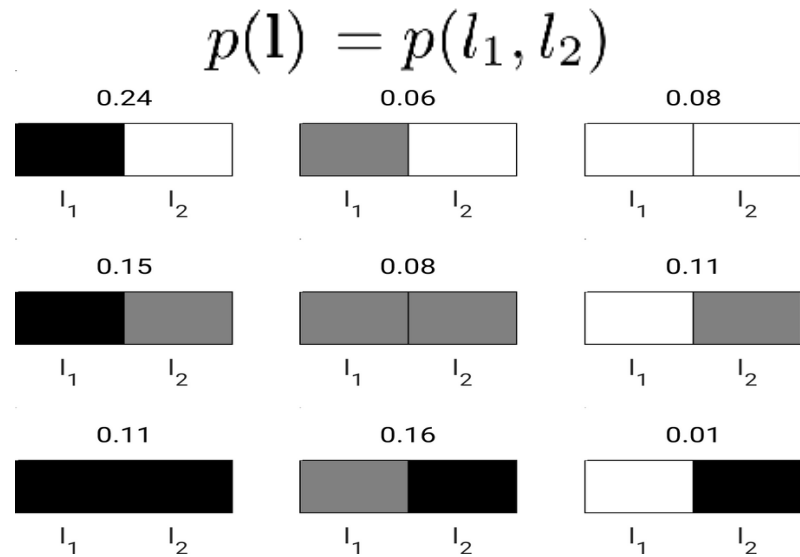
- The model depends on some parameters $\boldsymbol{\theta} = (\boldsymbol{\theta}_l^T, \boldsymbol{\theta}_d^T)^T$
- Appropriate values $\hat{\boldsymbol{\theta}}$ are assumed to be known for now...

Toy example

$N = 2$ voxels

$K = 3$ classes

$$\mathbf{l} = \begin{pmatrix} l_1 \\ l_2 \end{pmatrix}$$



$p(\mathbf{l}) = p(l_1, l_2)$

| | | | | |
|-------|---|-------|------|------|
| | | | | |
| | 2 | 0.24 | 0.06 | 0.08 |
| l_2 | 1 | 0.15 | 0.08 | 0.11 |
| | 0 | 0.11 | 0.16 | 0.01 |
| | | 0 | 1 | 2 |
| | | l_1 | | |

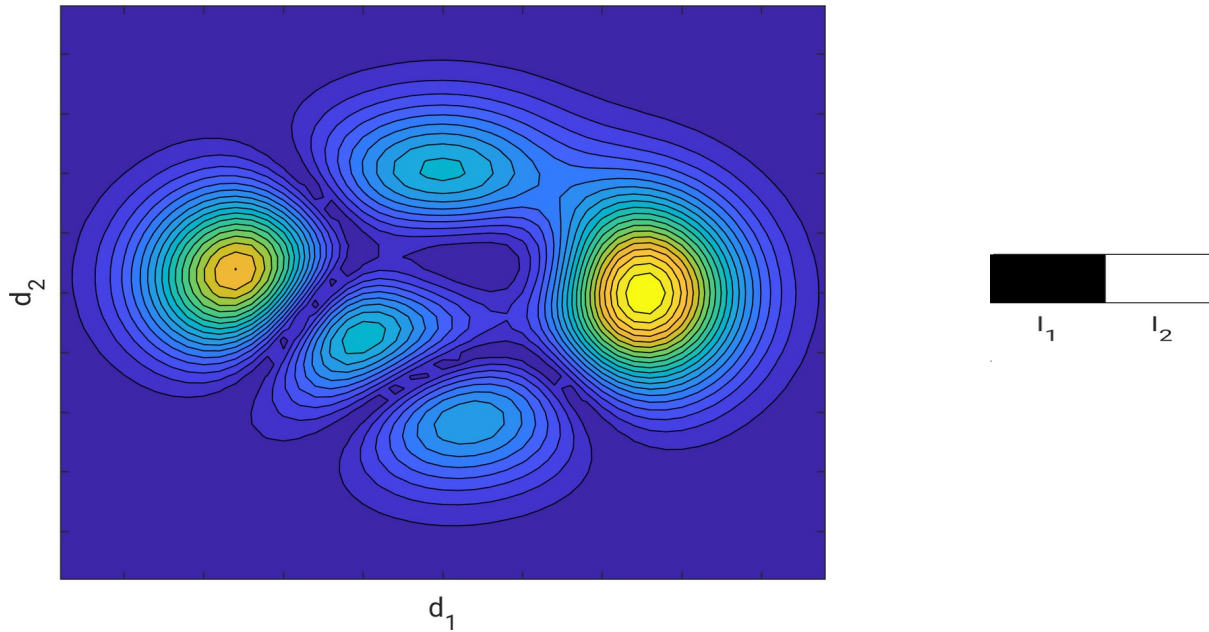
Toy example

$N = 2$ voxels

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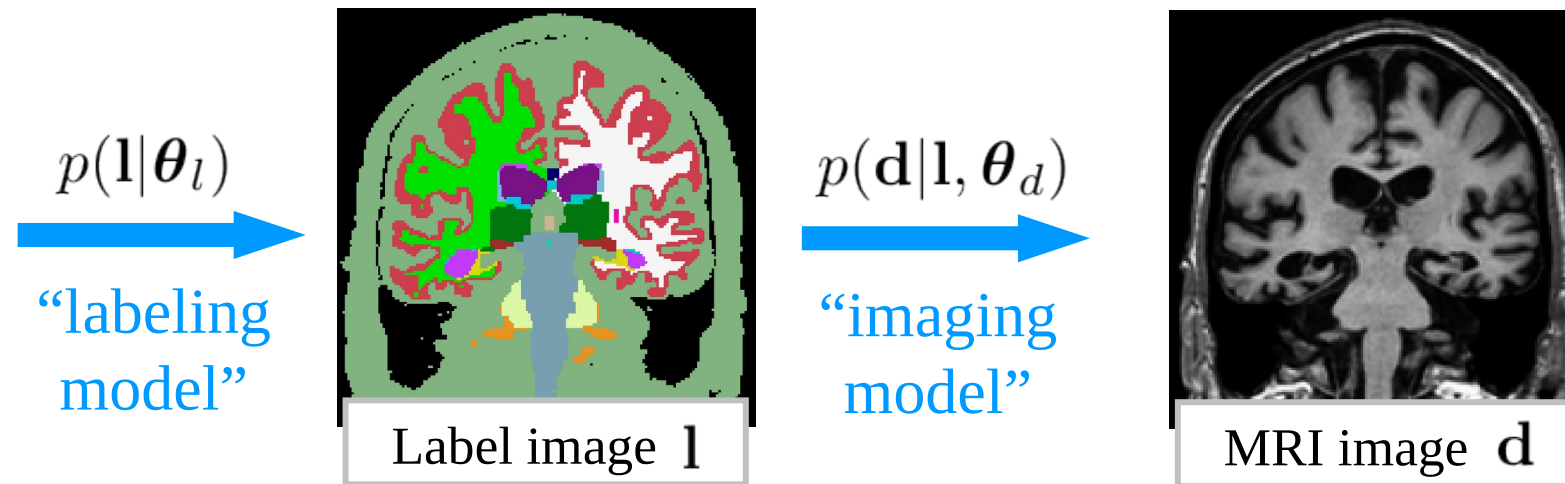
$$\mathbf{d} = \begin{pmatrix} d_1 \\ d_2 \end{pmatrix}$$

$$p(\mathbf{d}|\mathbf{l}) = p(d_1, d_2|l_1, l_2)$$



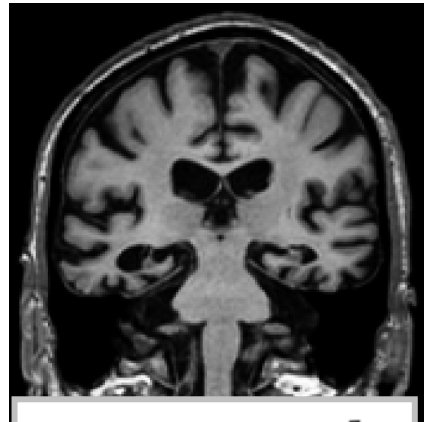
One solution: generative modeling

- Formulate a statistical model of how a medical image is formed



- The model depends on some parameters $\boldsymbol{\theta} = (\boldsymbol{\theta}_l^T, \boldsymbol{\theta}_d^T)^T$
- Appropriate values $\hat{\boldsymbol{\theta}}$ are assumed to be known for now...

Segmentation = inverse problem

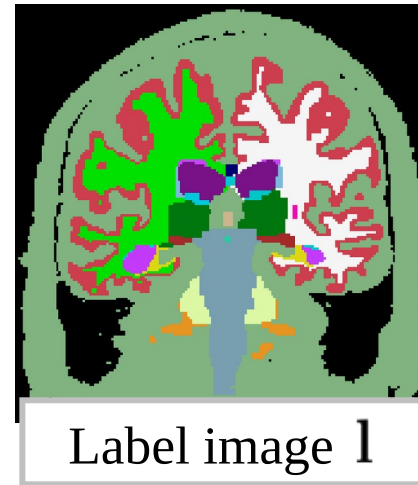
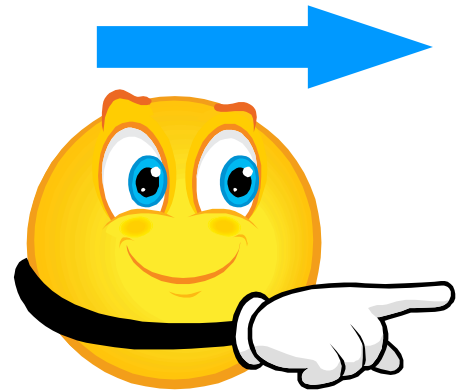
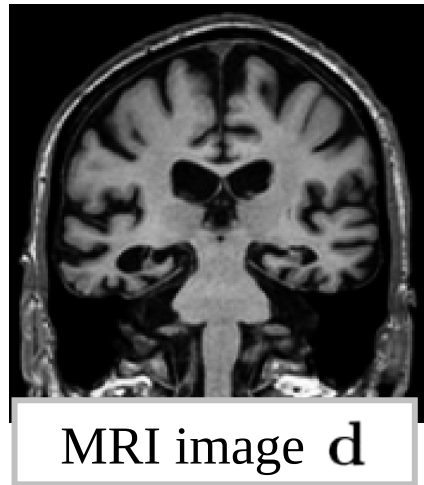


MRI image d



Label image l

Segmentation = inverse problem

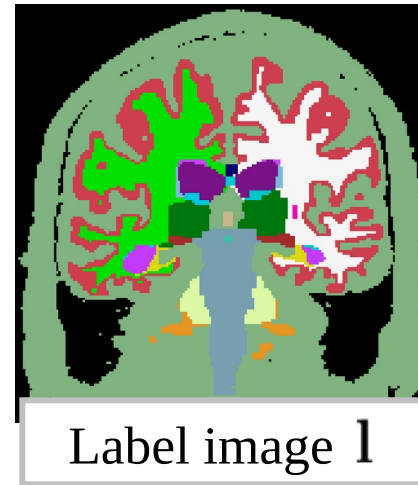
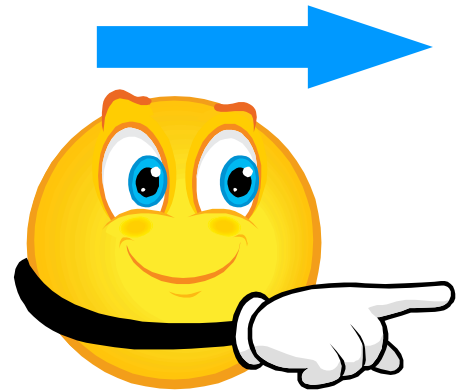
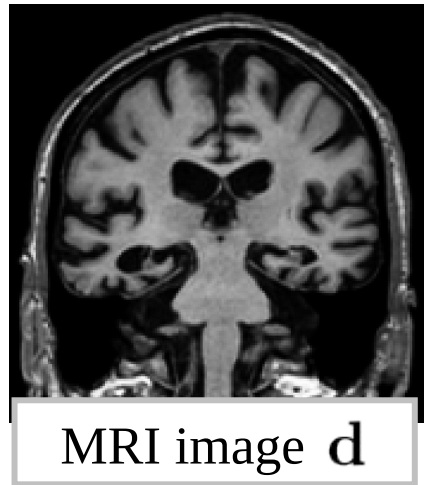


$$\hat{\mathbf{l}} = \arg \max_{\mathbf{l}} p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$$

The posterior distribution $p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$ is given by Bayes rule:

$$p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) = \frac{p(\mathbf{d}|\mathbf{l}, \hat{\boldsymbol{\theta}}_d)p(\mathbf{l}|\hat{\boldsymbol{\theta}}_l)}{p(\mathbf{d}|\hat{\boldsymbol{\theta}})}$$

Segmentation = inverse problem

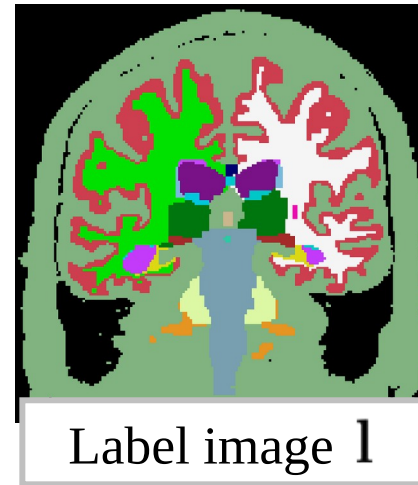
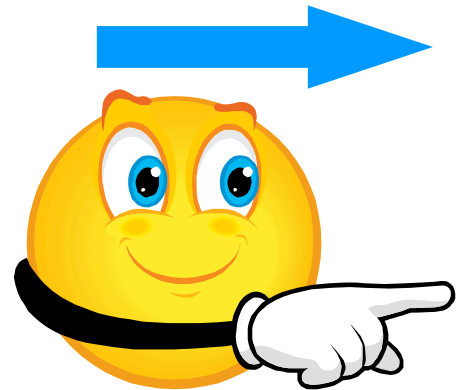
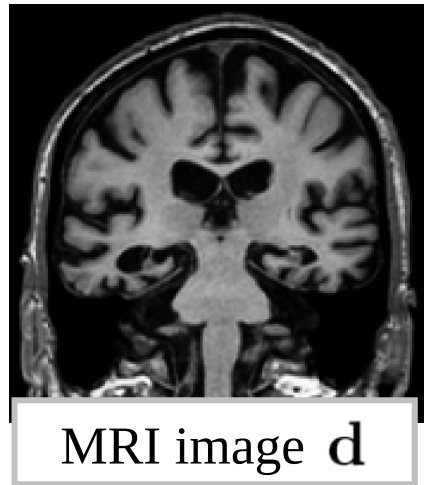


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$$p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) = \frac{p(\mathbf{d}|\mathbf{l}, \hat{\boldsymbol{\theta}}_d) \underbrace{p(\mathbf{l}|\hat{\boldsymbol{\theta}}_l)}_{\text{labeling model}}}{p(\mathbf{d}|\hat{\boldsymbol{\theta}})}$$

Segmentation = inverse problem



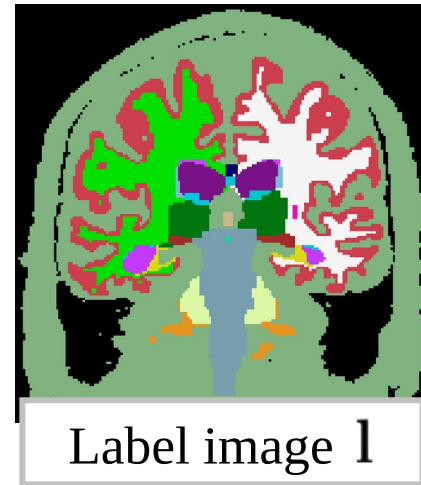
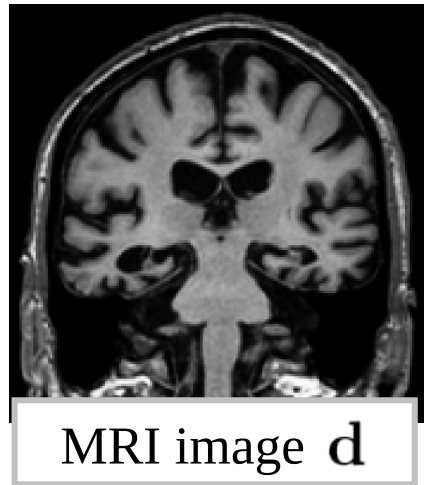
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imaging model

$$p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) = \frac{p(\mathbf{d}|\mathbf{l}, \hat{\boldsymbol{\theta}}_d) p(\mathbf{l}|\hat{\boldsymbol{\theta}}_l)}{p(\mathbf{d}|\hat{\boldsymbol{\theta}})}$$

Segmentation = inverse problem

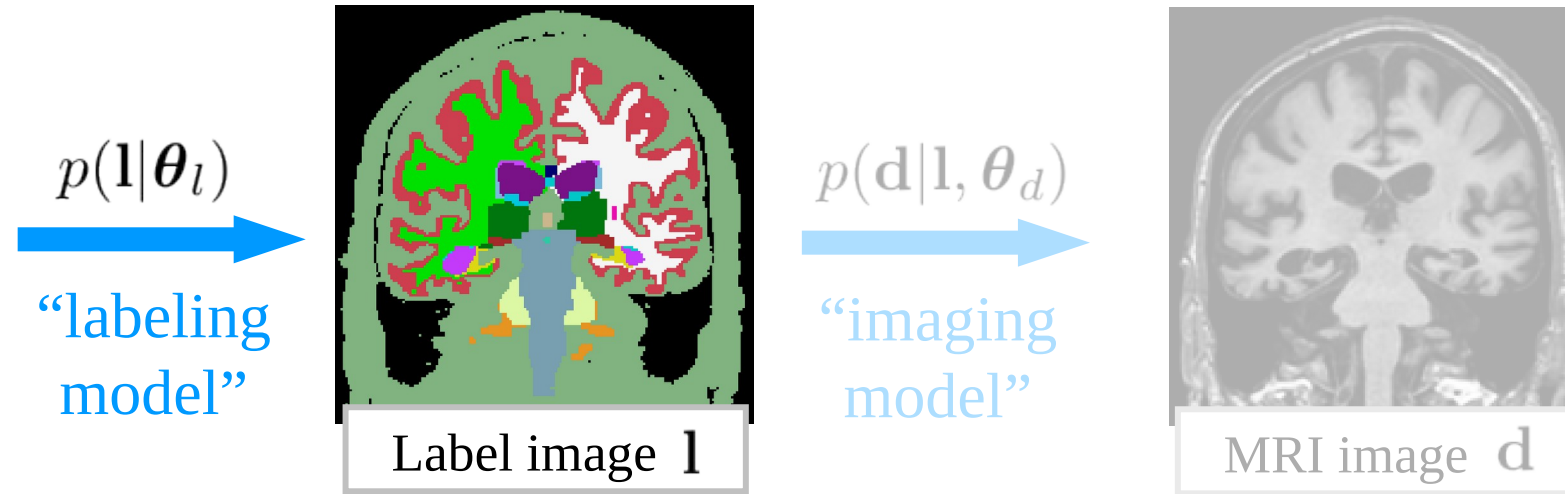


$$\hat{\mathbf{l}} = \arg \max_{\mathbf{l}} p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$$

The posterior distribution $p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$ is given by Bayes rule:

$$p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) = \frac{p(\mathbf{d}|\mathbf{l}, \hat{\boldsymbol{\theta}}_d)p(\mathbf{l}|\hat{\boldsymbol{\theta}}_l)}{p(\mathbf{d}|\hat{\boldsymbol{\theta}})} = \sum_{\mathbf{l}} p(\mathbf{d}|\mathbf{l}, \hat{\boldsymbol{\theta}}_d)p(\mathbf{l}|\hat{\boldsymbol{\theta}}_l) \quad (\text{but not needed})$$

Gaussian mixture model



- Assign a label to each voxel independently
- Probability of assigning label k is π_k

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \prod_n \pi_{l_n} , \quad \boldsymbol{\theta}_l = (\pi_1, \dots, \pi_K)^T$$

Toy example

$N = 2$ voxels

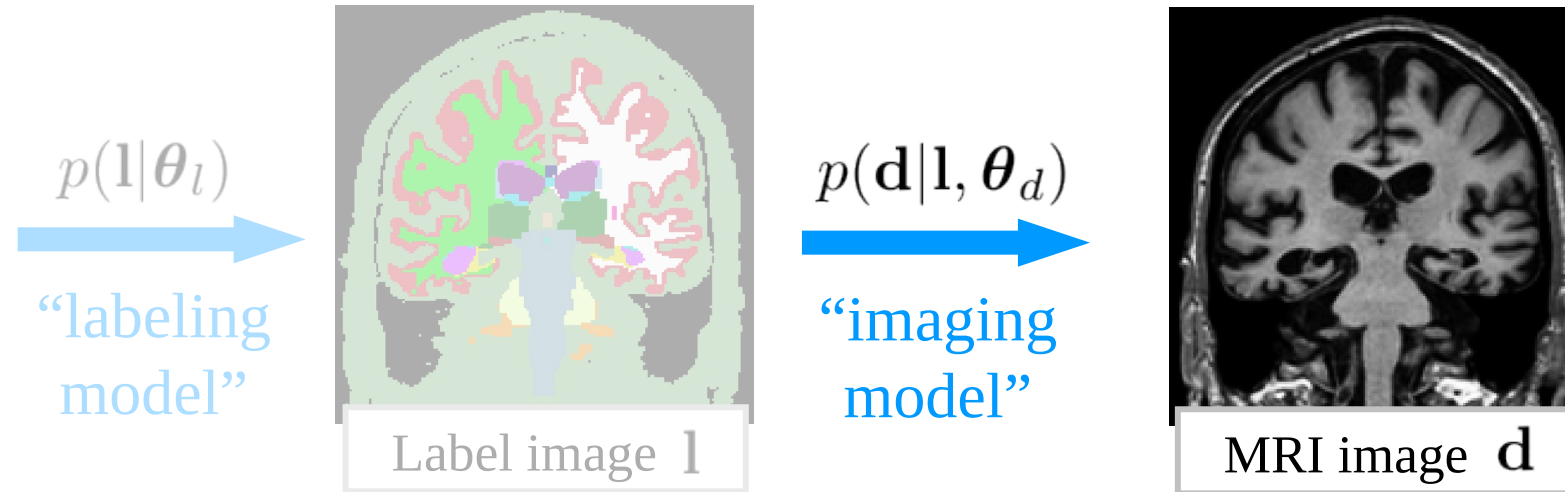
$K = 3$ classes

$$\mathbf{l} = \begin{pmatrix} l_1 \\ l_2 \end{pmatrix}$$

$$p(\mathbf{l}) = p(l_1, l_2) = p(l_2 | l_1) p(l_1)$$

| | | | | | | | | | | | | |
|-------|---|-------|------|------|-------|---|-------|---|-----|-----|-----|---|
| | | = | | | | | | | | | | |
| l_2 | 2 | 0.1 | 0.15 | 0.25 | l_2 | 2 | 0.5 | ▪ | 0.2 | 0.3 | 0.5 | |
| | 1 | 0.06 | 0.09 | 0.15 | | 1 | 0.3 | | | | | |
| | 0 | 0.04 | 0.06 | 0.1 | | 0 | 0.2 | | | | | |
| | | 0 | 1 | 2 | | | | | | 0 | 1 | 2 |
| | | l_1 | | | | | l_1 | | | | | |

Gaussian mixture model



- Drawn the intensity in each voxel with label k from a Gaussian distribution with mean μ_k and variance σ_k^2

$$p(\mathbf{d}|\mathbf{l}, \boldsymbol{\theta}_d) = \prod_n \mathcal{N}(d_n | \mu_{l_n}, \sigma_{l_n}^2), \quad \boldsymbol{\theta}_d = (\mu_1, \dots, \mu_K, \sigma_1^2, \dots, \sigma_K^2)^T$$

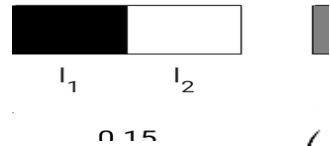
$$\mathcal{N}(d|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[-\frac{(d - \mu)^2}{2\sigma^2} \right]$$

Toy example

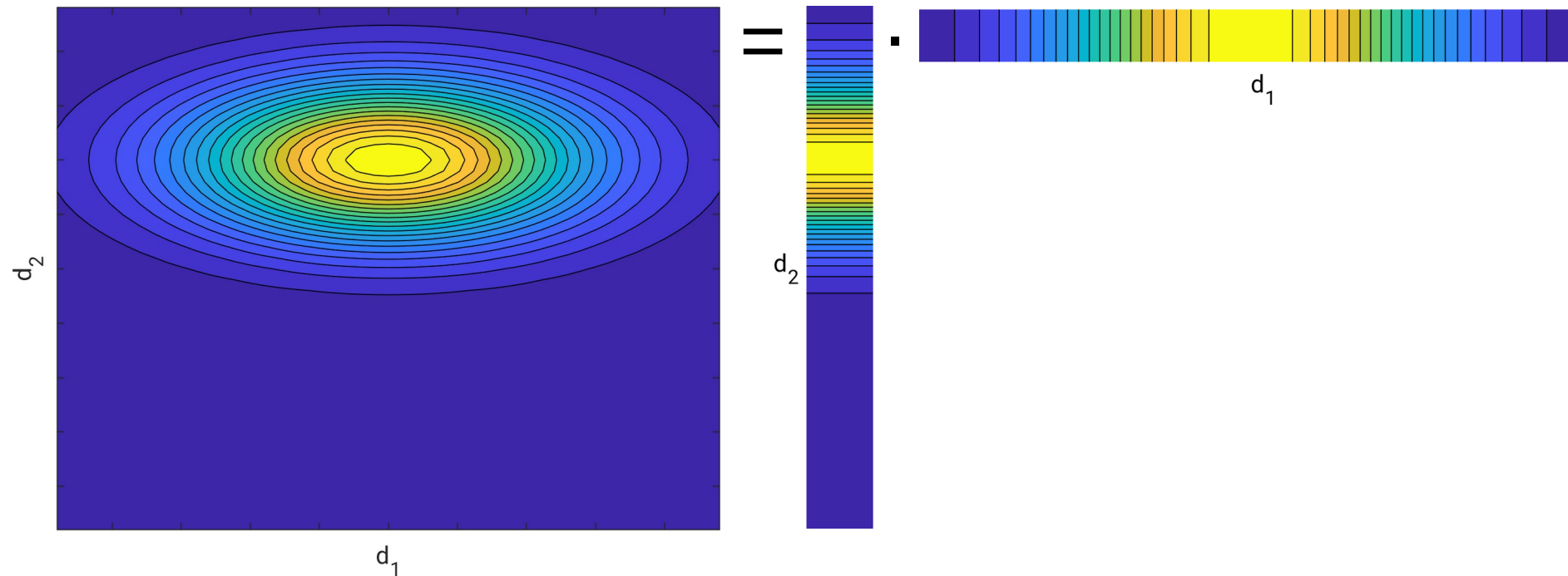
$N = 2$ voxels

$K = 3$ classes

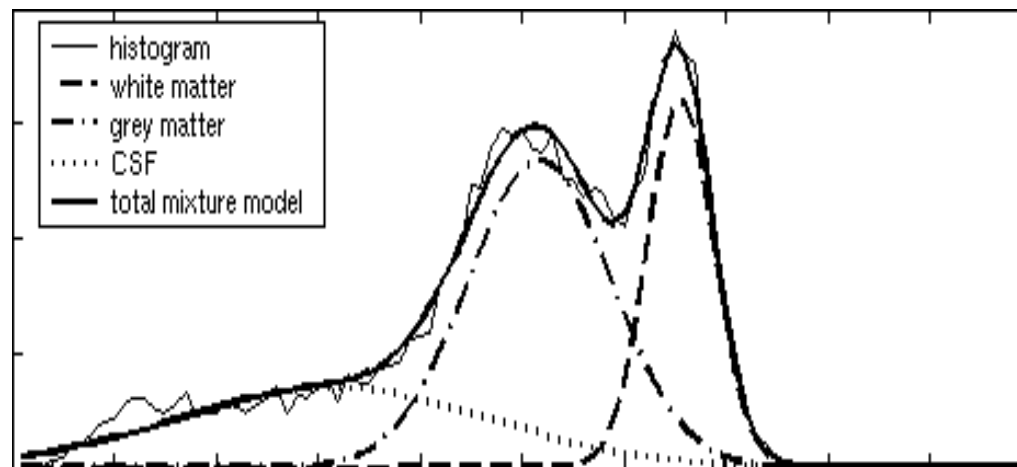
$$\mathbf{d} = \begin{pmatrix} d_1 \\ d_2 \end{pmatrix}$$



$$p(\mathbf{d}|\mathbf{l}) = p(d_1, d_2|l_1, l_2) = p(d_2|l_1, l_2, d_1)p(d_1|l_1, l_2)$$



Gaussian mixture model



$K = 3$ labels

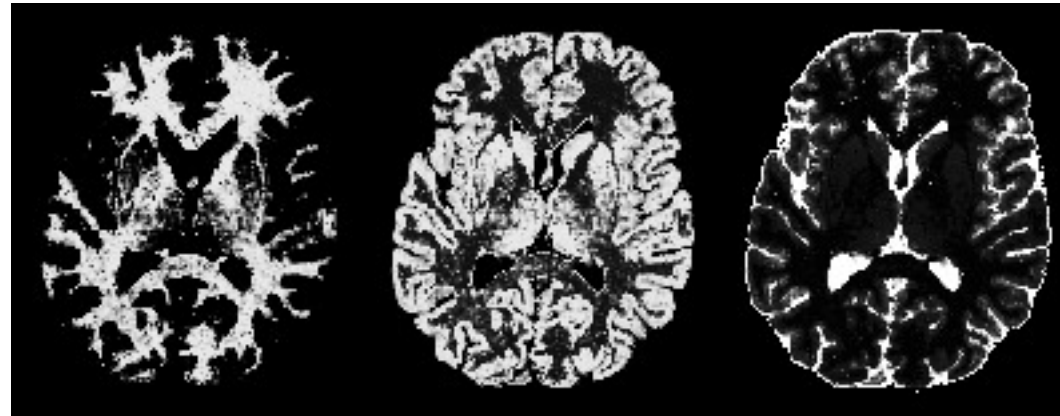
$$p(\mathbf{d}|\boldsymbol{\theta}) = \prod_n \left(\sum_k \mathcal{N}(d_n|\mu_k, \sigma_k^2) \pi_k \right)$$

$$\boldsymbol{\theta} = (\mu_1, \dots, \mu_K, \sigma_1^2, \dots, \sigma_K^2, \pi_1, \dots, \pi_K)^T$$

Posterior probability distribution



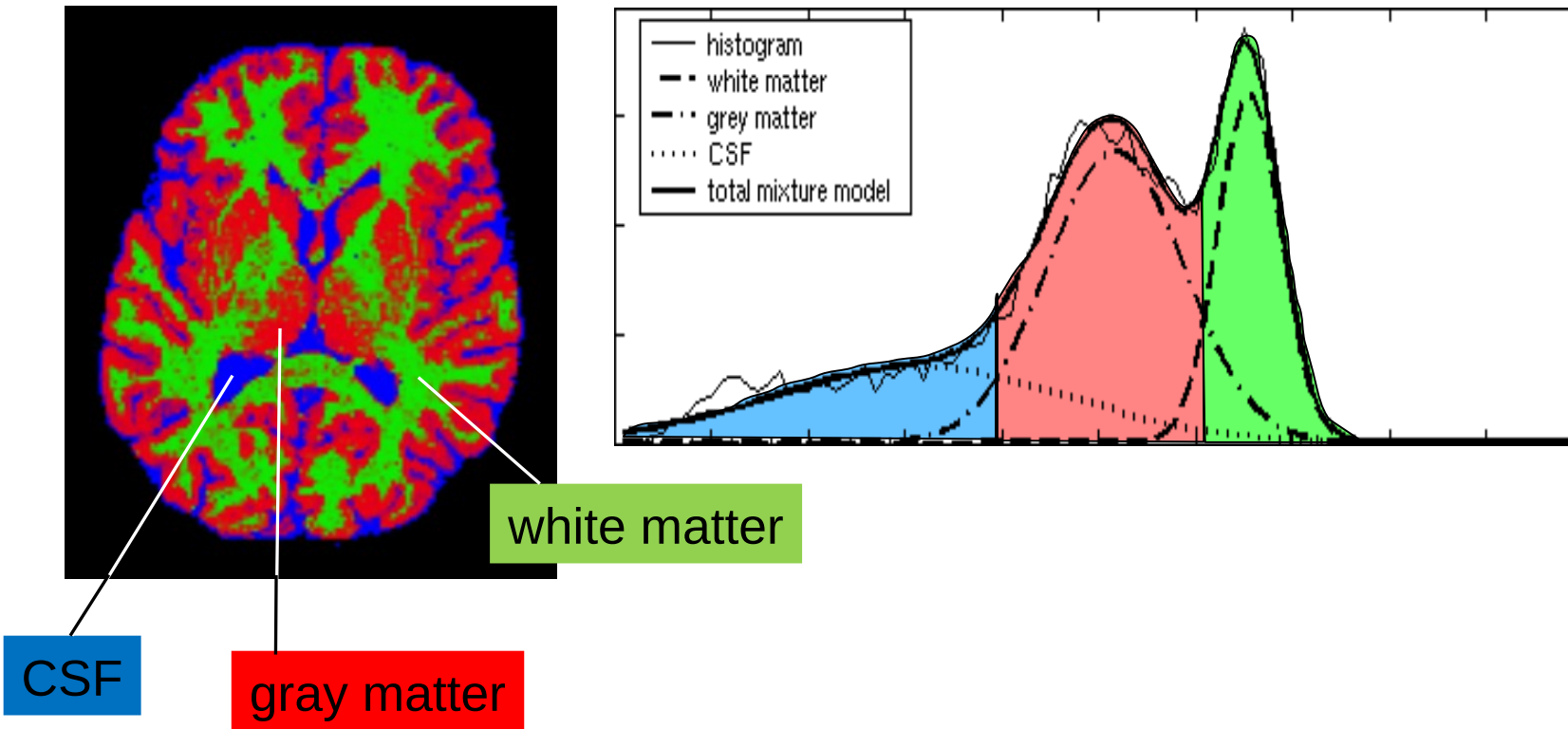
$$\begin{aligned} p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) &= \frac{p(\mathbf{d}|\mathbf{l}, \hat{\boldsymbol{\theta}}_d)p(\mathbf{l}|\hat{\boldsymbol{\theta}}_l)}{p(\mathbf{d}|\hat{\boldsymbol{\theta}})} \\ &= \frac{\prod_n \mathcal{N}(d_n|\hat{\mu}_{l_n}, \hat{\sigma}_{l_n}^2) \prod_n \hat{\pi}_{l_n}}{\prod_n \sum_k \mathcal{N}(d_n|\hat{\mu}_k, \hat{\sigma}_k^2) \hat{\pi}_k} \\ &= \prod_n p(l_n|d_n, \hat{\boldsymbol{\theta}}) \end{aligned}$$



$$p(l_n|d_n, \hat{\boldsymbol{\theta}}) = \frac{\mathcal{N}(d_n|\hat{\mu}_{l_n}, \hat{\sigma}_{l_n}^2) \hat{\pi}_{l_n}}{\sum_k \mathcal{N}(d_n|\hat{\mu}_k, \hat{\sigma}_k^2) \hat{\pi}_k}$$

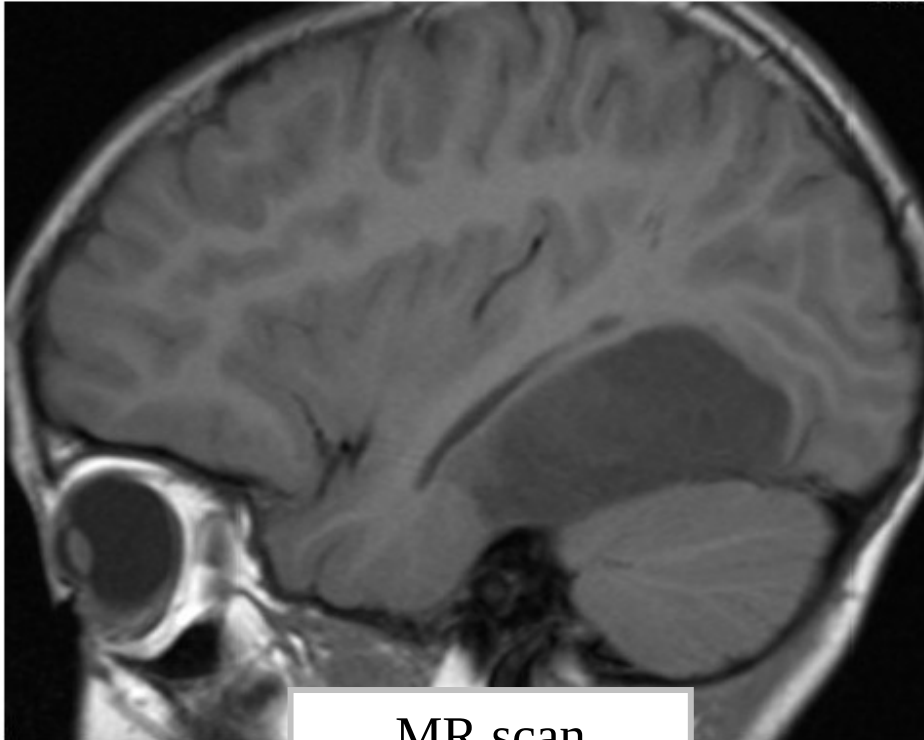
Maximum a posteriori segmentation

$$\hat{\mathbf{l}} = \arg \max_{\mathbf{l}} p(\mathbf{l} | \mathbf{d}, \hat{\boldsymbol{\theta}}) = \arg \max_{l_1, \dots, l_I} p(l_n | d_n, \hat{\boldsymbol{\theta}})$$

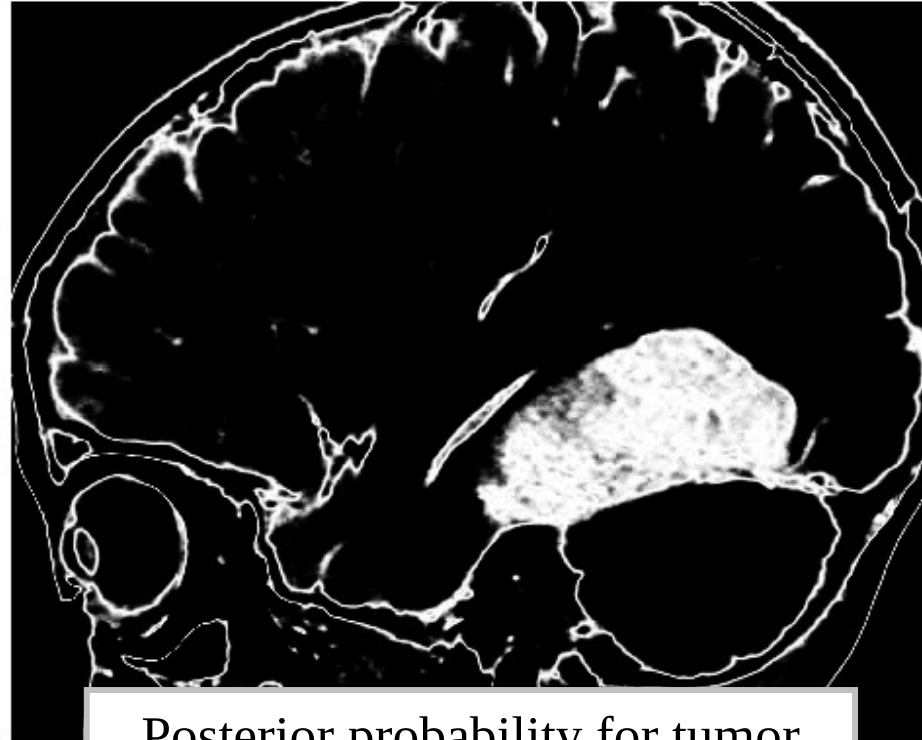


Problem solved?

Two-component Gaussian mixture model:
tumor vs. “other”

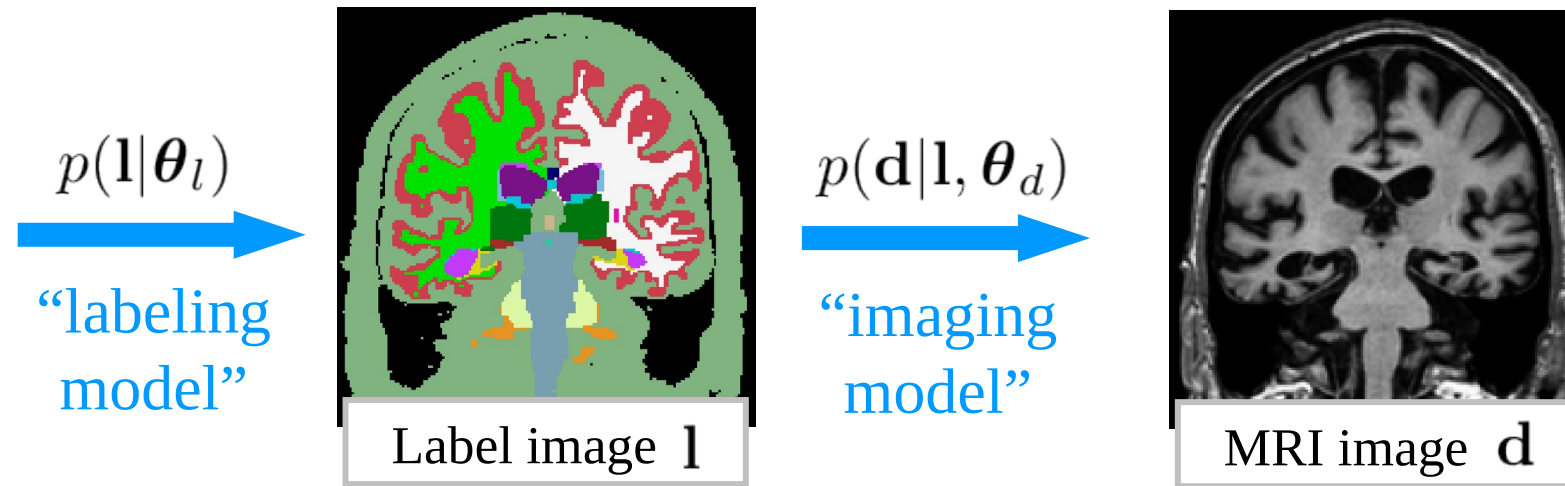


MR scan

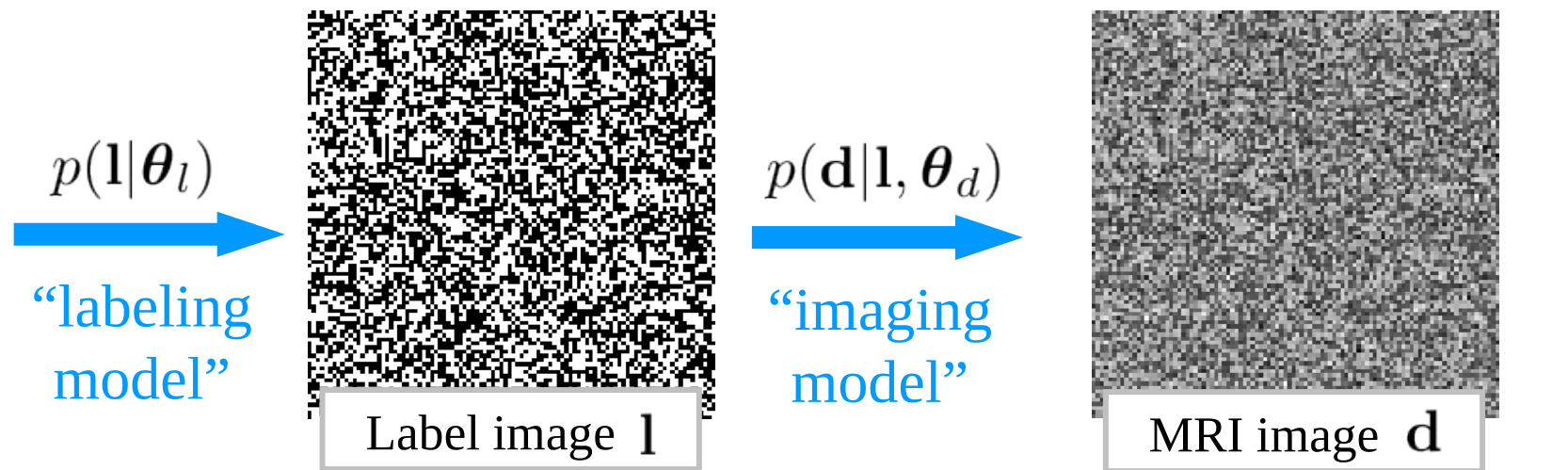


Posterior probability for tumor

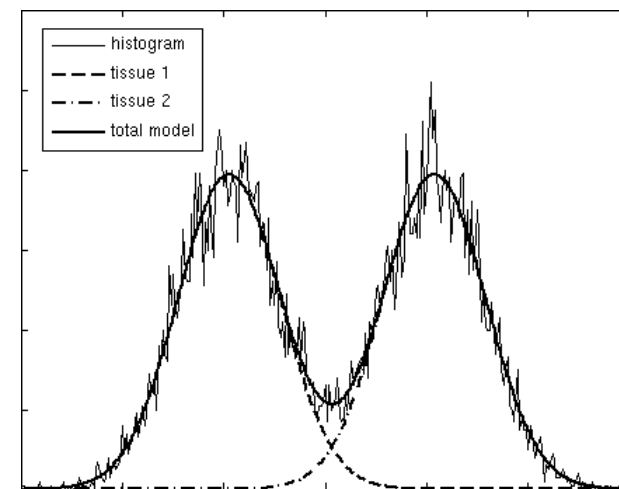
Gaussian mixture model



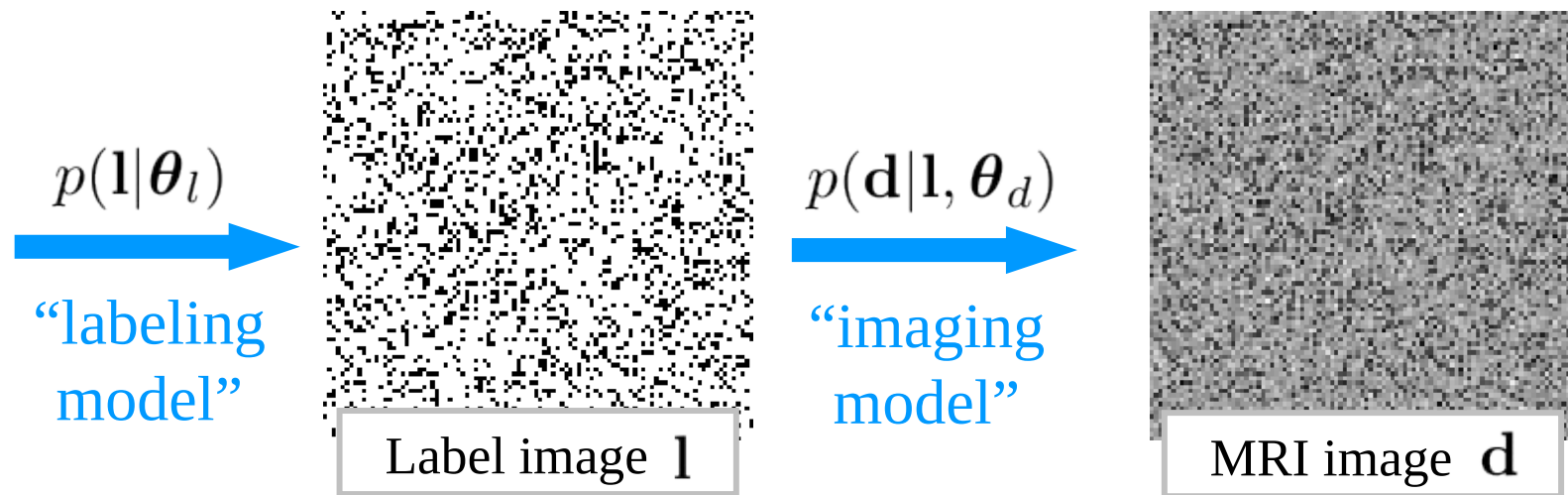
Gaussian mixture model



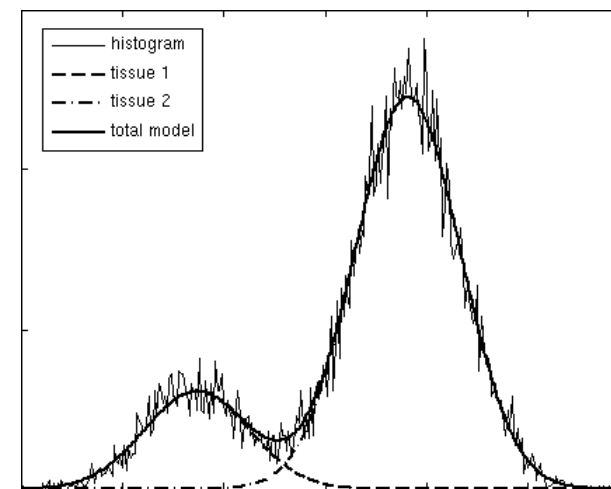
$$\begin{aligned}\mu_1 &= 70, \mu_2 = 90 \\ \sigma_1 &= 5, \sigma_2 = 5 \\ \pi_1 &= 0.5, \pi_2 = 0.5\end{aligned}$$



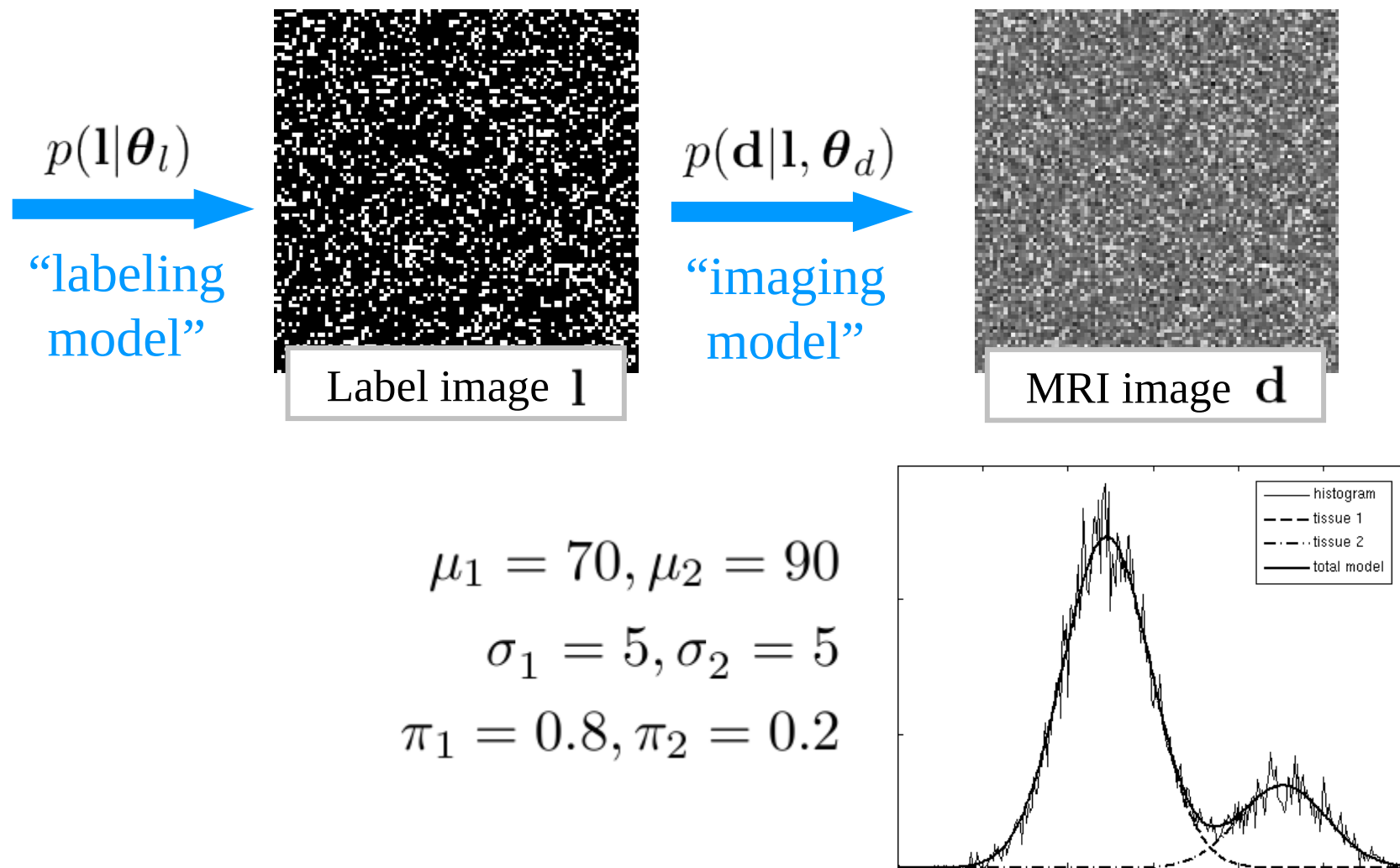
Gaussian mixture model



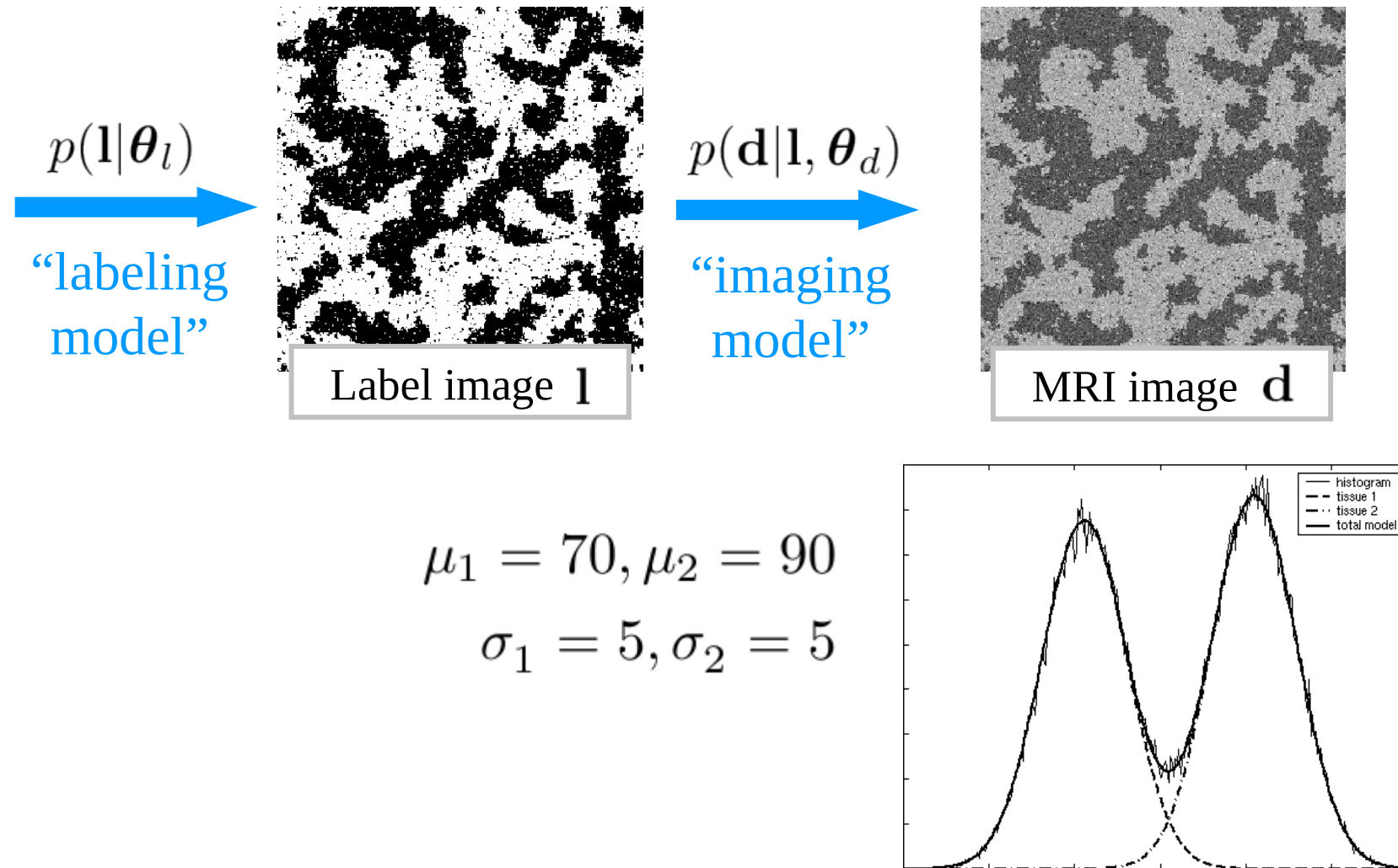
$$\begin{aligned}\mu_1 &= 70, \mu_2 = 90 \\ \sigma_1 &= 5, \sigma_2 = 5 \\ \pi_1 &= 0.2, \pi_2 = 0.8\end{aligned}$$



Gaussian mixture model



Markov random field model



Markov random field model

- Prior that prefers voxels with the same label to be spatially clustered

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \frac{1}{Z(\boldsymbol{\theta}_l)} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$$

$$U(\mathbf{l}|\boldsymbol{\theta}_l) = \beta \sum_{(i,j)} \delta(l_i \neq l_j)$$

- $Z(\boldsymbol{\theta}_l) = \sum_{\mathbf{l}} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$ is a normalizing constant

Markov random field model

- Prior that prefers voxels with the same label to be spatially clustered

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \frac{1}{Z(\boldsymbol{\theta}_l)} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$$

$$U(\mathbf{l}|\boldsymbol{\theta}_l) = \beta \sum_{(i,j)} \delta(l_i \neq l_j)$$

sum over all neighboring voxels

- $Z(\boldsymbol{\theta}_l) = \sum_{\mathbf{l}} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$ is a normalizing constant

Markov random field model

- Prior that prefers voxels with the same label to be spatially clustered

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \frac{1}{Z(\boldsymbol{\theta}_l)} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$$

$$U(\mathbf{l}|\boldsymbol{\theta}_l) = \beta \sum_{(i,j)} \delta(l_i \neq l_j)$$

zero if labels are the same,
one otherwise

- $Z(\boldsymbol{\theta}_l) = \sum_{\mathbf{l}} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$ is a normalizing constant

Markov random field model

- Prior that prefers voxels with the same label to be spatially clustered

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \frac{1}{Z(\boldsymbol{\theta}_l)} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$$

$$U(\mathbf{l}|\boldsymbol{\theta}_l) = \beta \sum_{(i,j)} \delta(l_i \neq l_j)$$

Parameter controlling

strength of penalization

- $Z(\boldsymbol{\theta}_l) = \sum_{\mathbf{l}} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$ is a normalizing constant

Markov random field model

- Prior that prefers voxels with the same label to be spatially clustered

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \frac{1}{Z(\boldsymbol{\theta}_l)} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$$

$$U(\mathbf{l}|\boldsymbol{\theta}_l) = \beta \sum_{(i,j)} \delta(l_i \neq l_j)$$

- $Z(\boldsymbol{\theta}_l) = \sum_{\mathbf{l}} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$ is a normalizing constant

Not needed in practice

Markov random field model

- Slightly more general:

$$p(\mathbf{l}|\boldsymbol{\theta}_l) = \frac{1}{Z(\boldsymbol{\theta}_l)} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))$$

$$U(\mathbf{l}|\boldsymbol{\theta}_l) = \beta \sum_{(i,j)} \delta(l_i \neq l_j) - \sum_i \log(\pi_{l_i})$$

- $\boldsymbol{\theta}_l = (\beta, \pi_1, \dots, \pi_K)^T$ are the model parameters
- Reduces to Gaussian mixture model prior $p(\mathbf{l}|\boldsymbol{\theta}_l) = \prod_n \pi_{l_n}$ for $\beta = 0$!

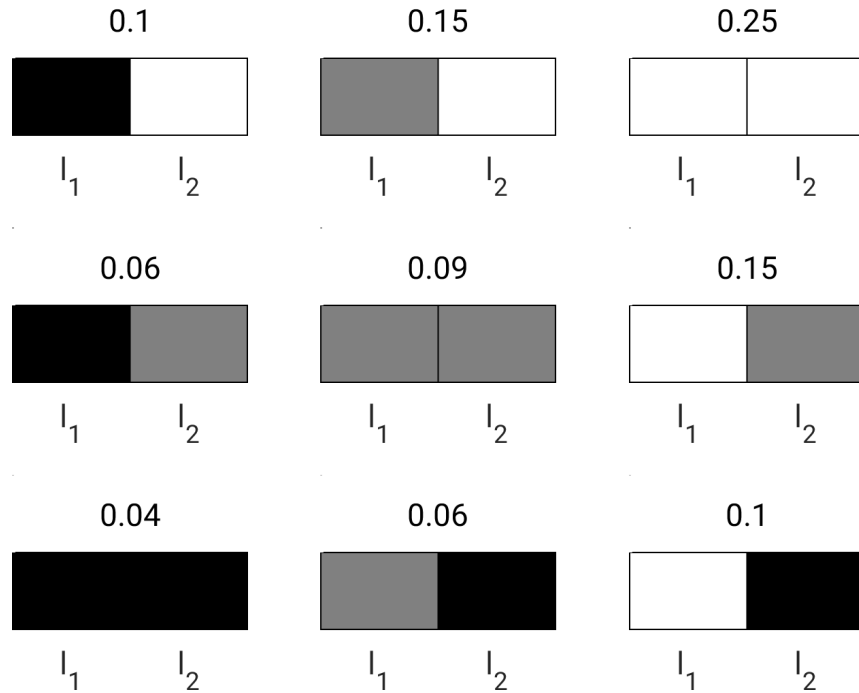
Toy example

$N = 2$ voxels

$K = 3$ classes

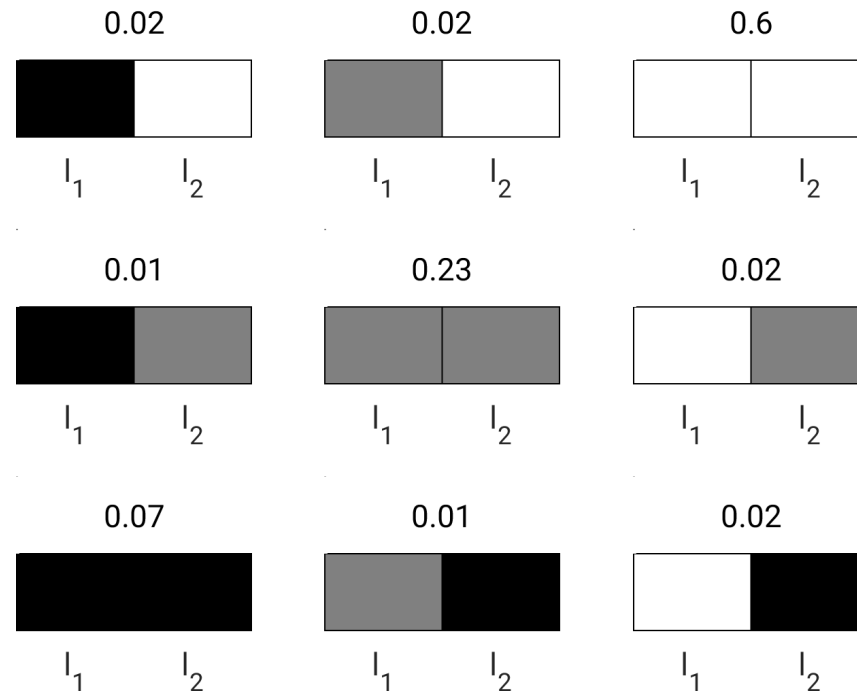
$$\mathbf{l} = \begin{pmatrix} l_1 \\ l_2 \end{pmatrix}$$

$\beta = 0$



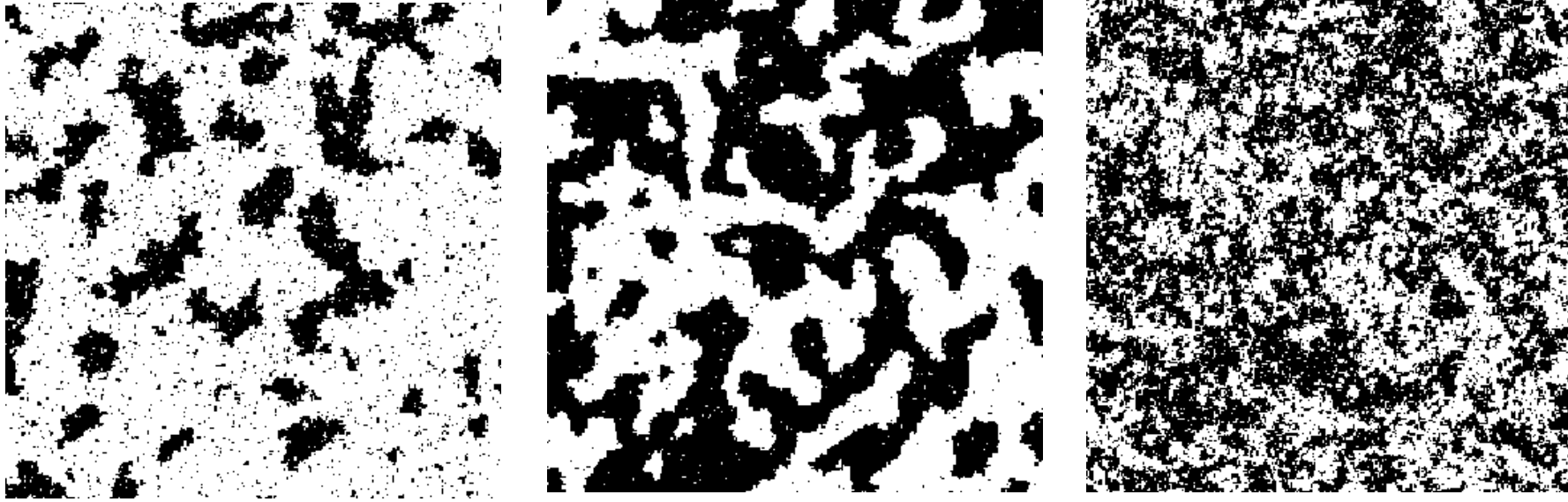
$$p(\mathbf{l}) = p(l_2 | \cancel{l_1}) p(l_1)$$

$\beta = 3.0$



$$p(\mathbf{l}) = p(l_1, l_2)$$

Samples



Different values for model parameters $\theta_l = (\beta, \pi_1, \dots, \pi_K)^T$

Why exactly this model?

- Long-range statistical dependencies between voxels
- Local computations (efficient!):

$$\begin{aligned} p(l_i | \mathbf{l}_{\setminus i}) &= \frac{p(\mathbf{l})}{p(\mathbf{l}_{\setminus i})} \\ &= \frac{p(\mathbf{l})}{\sum_{l_i} p(\mathbf{l})} \\ &= \frac{\exp(-U(\mathbf{l} | \boldsymbol{\theta}_l))}{\sum_{l_i} \exp(-U(\mathbf{l} | \boldsymbol{\theta}_l))} \\ &= \frac{\pi_{l_i} \cdot \exp(-\beta \sum_{j \in \mathfrak{N}_i} \delta(l_i \neq l_j))}{\sum_k \pi_k \cdot \exp(-\beta \sum_{j \in \mathfrak{N}_i} \delta(l_j \neq k))} \end{aligned}$$

Why exactly this model?

- Long-range statistical dependencies between voxels
- Local computations (efficient!):

$$\begin{aligned} p(l_i \mathbf{l}_{\setminus i}) &= \frac{p(\mathbf{l})}{p(\mathbf{l}_{\setminus i})} \\ \text{All labels} &= \frac{p(\mathbf{l})}{\sum_{l_i} p(\mathbf{l})} \\ \text{except the one} &= \frac{\exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))}{\sum_{l_i} \exp(-U(\mathbf{l}|\boldsymbol{\theta}_l))} \\ \text{of voxel } i &= \frac{\pi_{l_i} \cdot \exp(-\beta \sum_{j \in \mathfrak{N}_i} \delta(l_i \neq l_j))}{\sum_k \pi_k \cdot \exp(-\beta \sum_{j \in \mathfrak{N}_i} \delta(l_j \neq k))} \end{aligned}$$

Why exactly this model?

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$$\begin{aligned} p(l_i | \mathbf{l}_{\setminus i}) &= \frac{p(\mathbf{l})}{p(\mathbf{l}_{\setminus i})} \\ &= \frac{p(\mathbf{l})}{\sum_{l_i} p(\mathbf{l})} \\ &= \frac{\exp(-U(\mathbf{l} | \boldsymbol{\theta}_l))}{\sum_{l_i} \exp(-U(\mathbf{l} | \boldsymbol{\theta}_l))} \quad \begin{array}{l} \text{neighbors} \\ \text{of voxel } i \end{array} \\ &= \frac{\pi_{l_i} \cdot \exp(-\beta \sum_{j \in \mathfrak{N}_i} \delta(l_i \neq l_j))}{\sum_k \pi_k \cdot \exp(-\beta \sum_{j \in \mathfrak{N}_i} \delta(l_j \neq k))} \end{aligned}$$

Mean field approximation

- In the Gaussian mixture model, the posterior was of the form

$$p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) = \prod_n p(l_n|d_n, \hat{\boldsymbol{\theta}})$$

- With the Markov random field model, the posterior no longer “factorizes” that way
- For a 2-label model in a standard 256x256x128 MR scan, there are over $10^{1000000}$ unique label images with each its own posterior probability!
- Solution: approximate $p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$

Mean field approximation

- Approximate $p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$ with something of the form

$$q(\mathbf{l}) = \prod_n q_n(l_n)$$

- Find the voxel-wise distributions $q_n(k)$ that minimize the difference between $q(\mathbf{l})$ and $p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})$
- Quantify the difference between the two distributions using the “Kullback-Leibler divergence”

$$KL \left(q(\mathbf{l}) || p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}}) \right) = - \sum_{\mathbf{l}} q(\mathbf{l}) \log \frac{p(\mathbf{l}|\mathbf{d}, \hat{\boldsymbol{\theta}})}{q(\mathbf{l})}$$

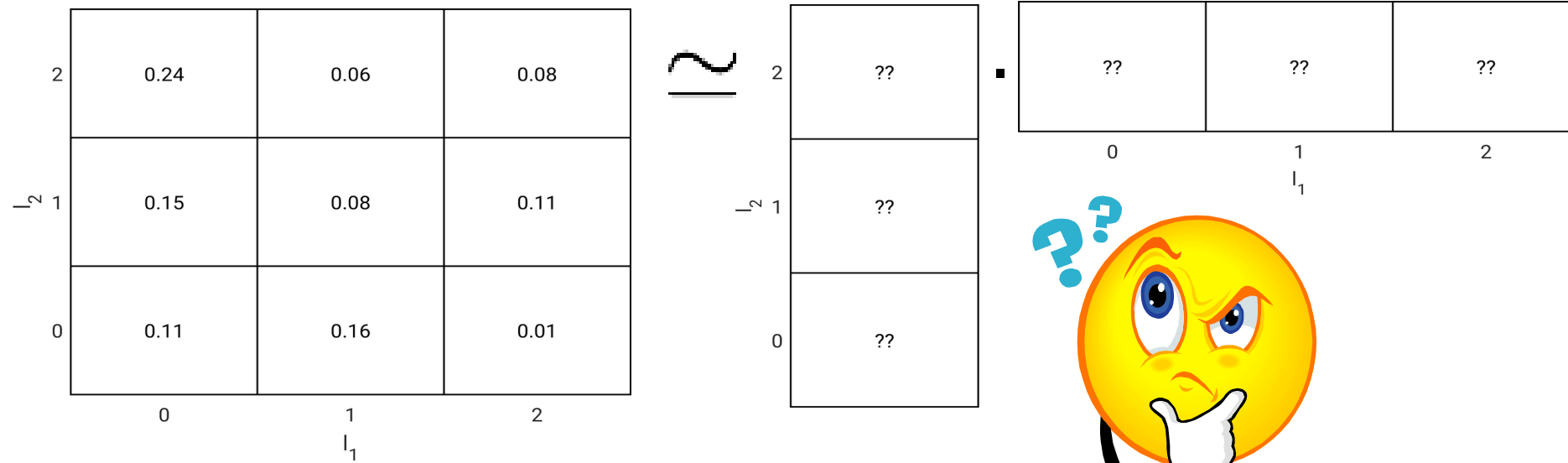
Toy example

$N = 2$ voxels

$K = 3$ classes

$$\mathbf{l} = \begin{pmatrix} l_1 \\ l_2 \end{pmatrix} \quad \mathbf{d} = \begin{pmatrix} d_1 \\ d_2 \end{pmatrix}$$

$$p(\mathbf{l}|\mathbf{d}) = p(l_1, l_2 | d_1, d_2) \simeq q(l_1)q(l_2)$$



Mean field approximation

- Solution for one voxel i :

$$q_i(l_i) = \frac{\mathcal{N}(d_i | \hat{\mu}_{l_i}, \hat{\sigma}_{l_i}^2) \gamma_i(l_i)}{\sum_k \mathcal{N}(d_i | \hat{\mu}_k, \hat{\sigma}_k^2) \gamma_i(k)}$$

$$\text{where } \gamma_i(k) = \frac{\hat{\pi}_k \cdot \exp \left(-\beta \sum_{j \in \mathfrak{N}_i} (1 - q_j(k)) \right)}{\sum_{k'} \hat{\pi}_{k'} \cdot \exp \left(-\beta \sum_{j \in \mathfrak{N}_i} (1 - q_j(k')) \right)}$$

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Influenced by the result
in neighboring voxels:
spatial context!!!!

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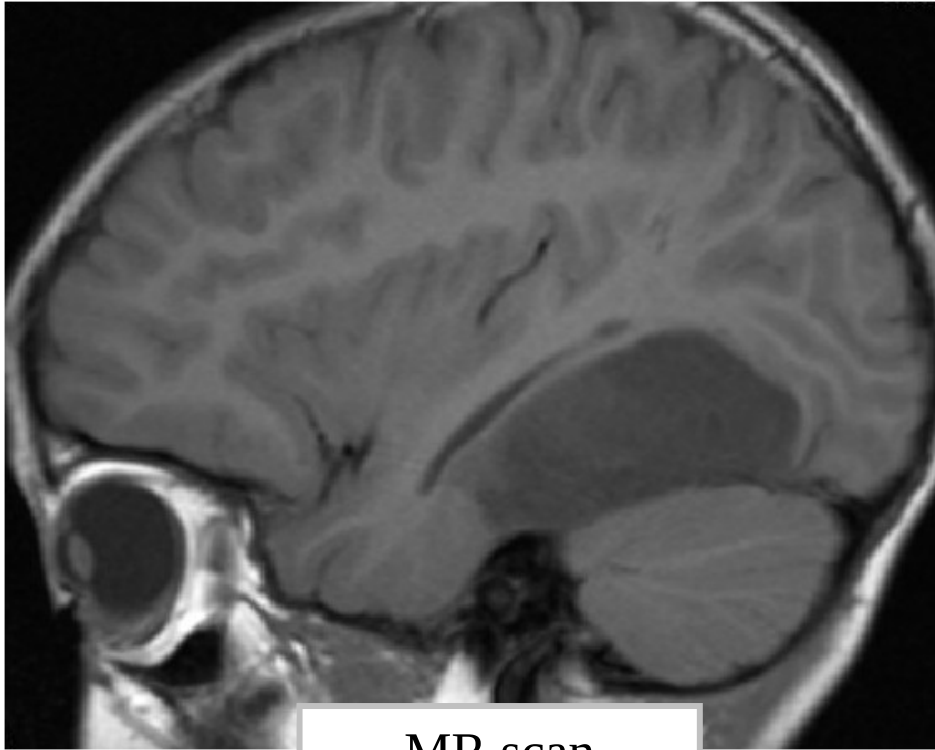
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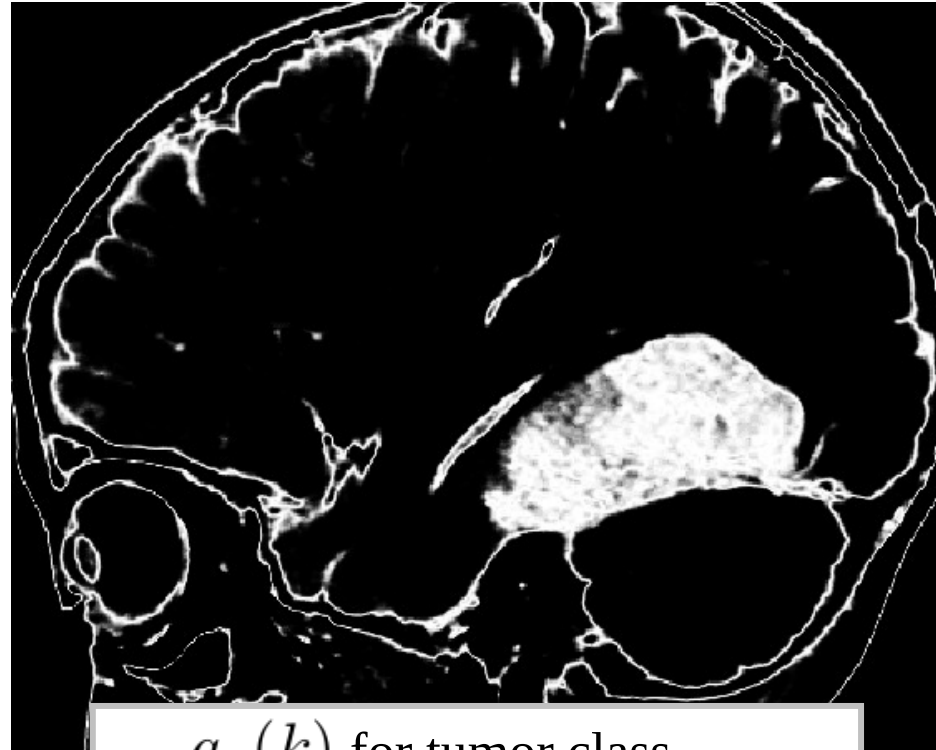
- Need to iterate across all voxels

Example

Two-component Gaussian mixture model:
tumor vs. “other”



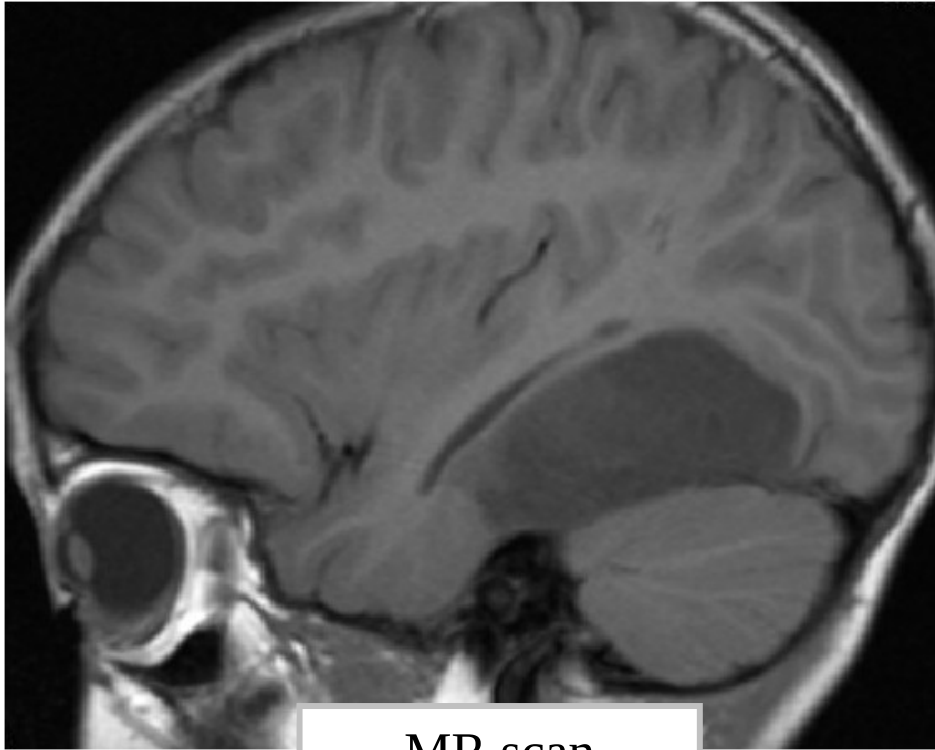
MR scan



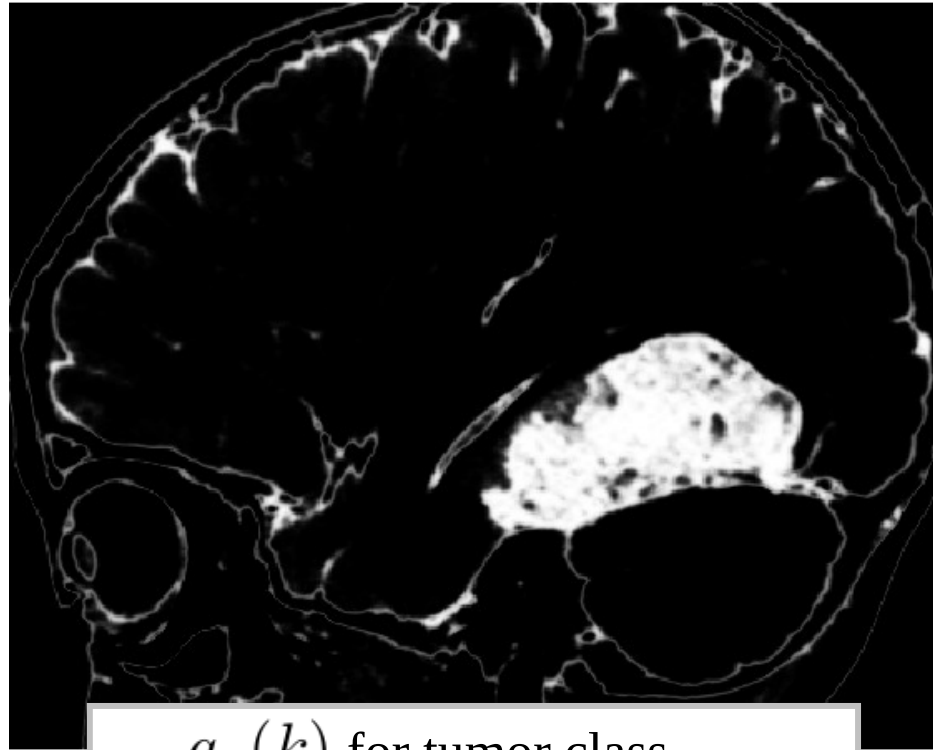
$q_n(k)$ for tumor class
 $\beta = 0$

Example

Two-component Gaussian mixture model:
tumor vs. “other”



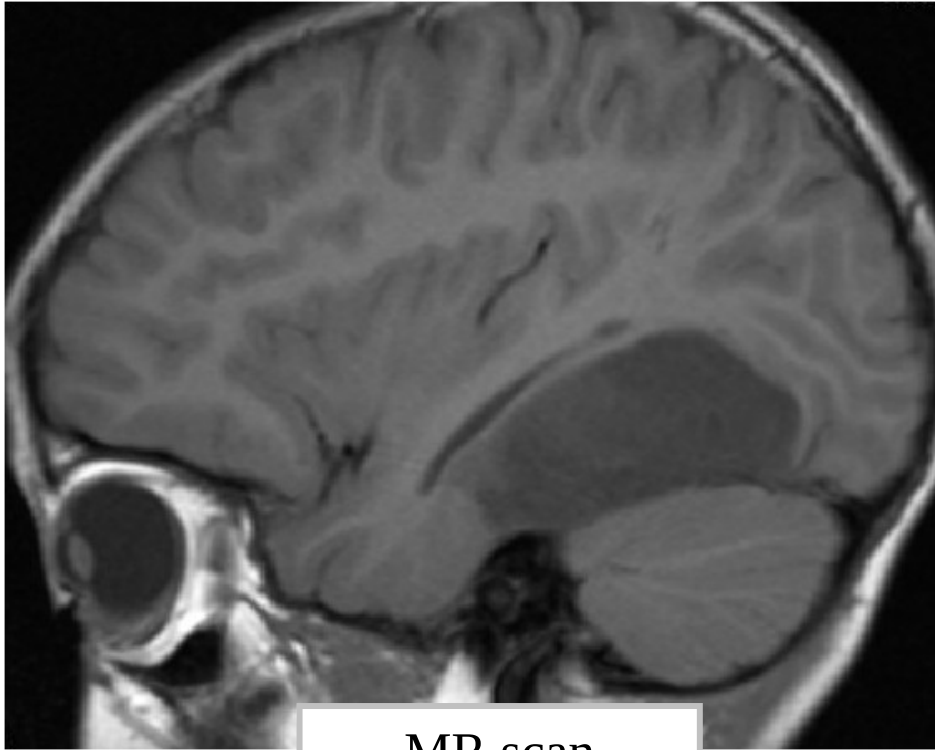
MR scan



$q_n(k)$ for tumor class
 $\beta = 0.25$

Example

Two-component Gaussian mixture model:
tumor vs. “other”



MR scan



$q_n(k)$ for tumor class
 $\beta = 0.55$