Who are ASM Journals? A Gender-based Analysis

Running title: A gender-based analysis of ASM journals

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Abstract

2 Importance

3 Introduction

Evidence has accumulated over the decades that academic research has a representation problem. While at least 50% of biology Ph.D. graduates are women, the number of women in postdoctoral positions and tenure-track positions are less than 40 and 30%, respectively @article{sheltzer elife 2014}. Studies examining other metrics such as race and ethnicity find that less than 10% of all science and engineering doctorates were awarded to underrepresented minorities, while less than 25% of science and engineering doctorates in early career academia identify as non-white (NSF ADVANCE, 2014). Predictabily, the disparities increase alongside academic rank @article{potvin diversity 2018}. There have been many proposed reasons for 11 these disparities (particularly against women) that include biases in training and hiring, the impact of children on career trajectories, a lack of support for primary caregivers, a lack of recognition, and less productivity as measured by research publications. Add citations These issues do not act independant of each other, instead they are cumulative over time for both individuals and the 15 community. Accordingly, addressing these issues necessitates multi-level approaches from all insititutions and members of the scientific community.

Recently, scientific societies and publishers have begun examining internal submissions data to evaluate representation of, and bias against, women in their peer review processes. The American Geological Union found that while the acceptance rate of women-authored publications was greater than that for publications authored by men, women submitted fewer manuscripts than men and were used as reviewers only 20% of the time (Lerback, 2017), a factor influenced by the gender of the editor (Fox, 2016). Despite the disproportional representation of lead women authors, several studies have concluded that there is no significant bias aginst papers authored by women (C&W, 2011; Fox, 2016; Handly, 2015; Edwards, 2018). Conversely, two recent studies—one of the peer review process at eLife, a broad scope biology journal, and the other

of outcomes at six ecology and evolution journals—found that women-authored papers are less likely to have positive reviews and outcomes (Murray, 2018; Fox and Paine, 2019).

However, representation and attitudes differ by scientific field and no studies to-date seem to have investigated academic publishing in the field of microbiology. The American Society for Microbiology (ASM) is one of the largest life science societies, with an average membership of 41,000 since 1990. In its mission statement, the ASM notes that it is "an inclusive organization, engaging with and responding to the needs of its diverse constituencies" and pledges to "address all members' needs through development and assessment of programs and services." One of these services is the publication of microbiology research through a suite of 13 journals. The goal of this research study is to describe the representation of authors, reviewers, and editors at ASM journals by gender and associated peer review outcomes.

38 Results

The term gatekeepers collectively refers to those that facilitate the peer review process, such as editors-in-chief (EIC), editors, and reviewers. Between January 2012 and August 2018, ASM 40 published 15 different journals: Antimicrobial Agenst and Chemotherapy (AAC), Applied and 41 Environmental Microbiology (AEM), Clinical and Vaccine Immunology (CVI), Clinical Microbiology Reviews (CMR), Eukaryotic Cell (EC), Infection and Immunity (IAI), Journal of Bacteriology 43 (JB), Journal of Clinical Microbiology (JCM), Journal of Virology (JV), mBio, Microbiology and 44 Molecular Biology Reviews (MMBR), Genome Announcements (GA, now Microbiology Resource Annoucements), Molecular and Cellular Biology (MCB), mSphere and mSystems. This study 46 only examines original research manuscripts, which eliminates three journals from the remaining 47 analyses (CMR, GA, and MMBR). Given the relatively small number of editors at ASM journals, their presenting genders where identified by hand while the genders of reviewers and authors 49 were predicted from their first names. Assigning gender by first name resulted in 3 possible 50 outcomes: men, women, and unknown (when gender could not be assigned with confidence, 51 see Methods for validation). + describe # of manuscripts and authors + descriptive study of a population - no stats

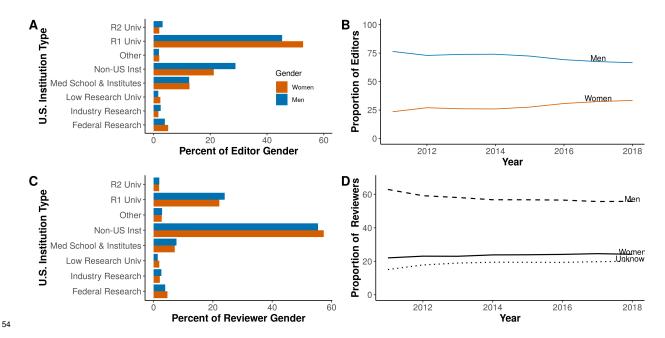


Figure 1. Gendered representation among gatekeepers. Proportion of editors from (A) insitution types and (B) over time from 2012 to 2018. Editors and senior editors are pooled together. Proportion of reviewers from (C) institution types and (D) over time from 2012 to 2018. Each individual was counted once per calendar year.

Men dominate as gatekeepers and senior authors. Each journal is led by an editor-in-chief who manages journal scope and quality standards. Two journals, EC and CVI were retired during the period under study. In total, there were 17 EICs, 17.65% of which were women. In 2013, the leadership of CVI transferred from a man EIC, to a woman. The *Journal of Virology* (JVI) has had the same woman as EIC since 2012. The EICs manage a board of editors with field expertise to help manage the peer review process as needed. The total number of editors over the duration of our study (senior editors and editors pooled) was 1016 and 28.74% were women.

Over 40% of both men and women editors are from US-based R1 institutions, with non-US and uncategorized US institutions supplying the next largest proportions of editors (Fig. 1A). Since the start of our study, there has been a slow trend toward gender parity of editors (Fig. 1B). The trends for each journal studied vary considerably, though most have slow trends toward parity (Fig. S1). CVI and *mSphere* are the only ASM journals to have accomplished equivalent representation of both genders, with CVI having a greater proportion of women editors than men before it was

retired. EC is the only journal with an increasing parity gap.

Our dataset contained 30704 reviewers, 24.61% of which were women. As with editors, the greatest proportion of reviewers (about 50% of both men and women) come from non-US institutions, while R1 institutions supply the next largest cohort of reviewers (Fig. 1C). Over the time period studied, the proportions of each gender have held steady among reviewers at ASM journals (Fig. 1D) and is representative of both reviewer proportions at each journal, and the potential reviewers at all journals combined (Fig. S2).

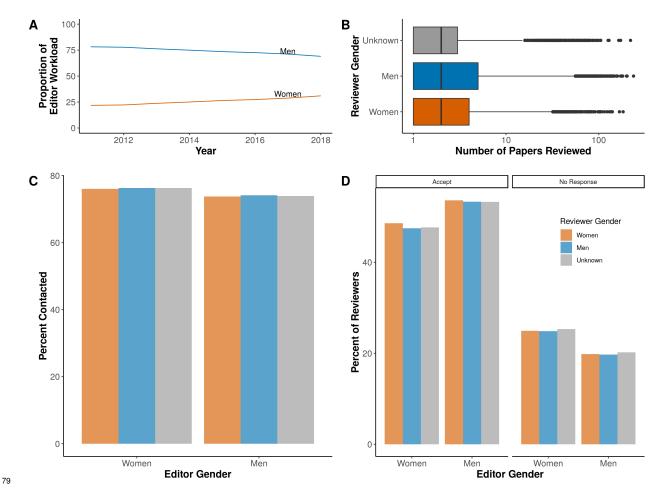


Figure 2. Reviewer representation, workload, and response to requests to review. (A) Proportion of manuscript workloads by men and women editors. Editorial rejections excluded. (B) Boxplot comparison of total papers reviewed by each individual according to gender. (C) Percent of each reviewer gender contacted to review, according to the editor's gender. (D) The percent of reviewers by gender that either accepted the opportunity to review or did not respond to a request

85 to review, split according to the editor's gender.

Editorial workloads are not proportionate Across all journals, men handle a slightly greater proportion of manuscripts (blue) and women a slightly smaller proportion (orange), relative to their respective editorial representations (Fig. 1A). This trend continues accross most journals with varying degrees of difference between workload and representation (Fig. S1). There are exceptions. At *mBio* and *mSphere*, workload and proportions are identical. However, at CVI and JVI, the workload for women editors is much higher than their representation, while the workload of men is considerably less than their representation would suggest. In the years preceding its retirement, the representation of women at CVI increased, which acted to decrease the gap with their workload. However, representations and relative workloads for men and women editors at JVI have held steady over time.

The median number of papers reviewed by individuals in each gender group is equivalent, with a trend to more men reviewing a greater quantity of manuscripts (Fig. 2B). 3275, 6413, and 3176% of men, women, and unknown reviewers have reviewed only one manuscript. Editors of both genders contact reviewers from all three gender groups at equivalent proportions, though women editors contact more reviewers than men (Fig. 2C). Reviewers of all genders, accept fewer, and ignore more, requests to review from women editors than men (Fig. 2D).

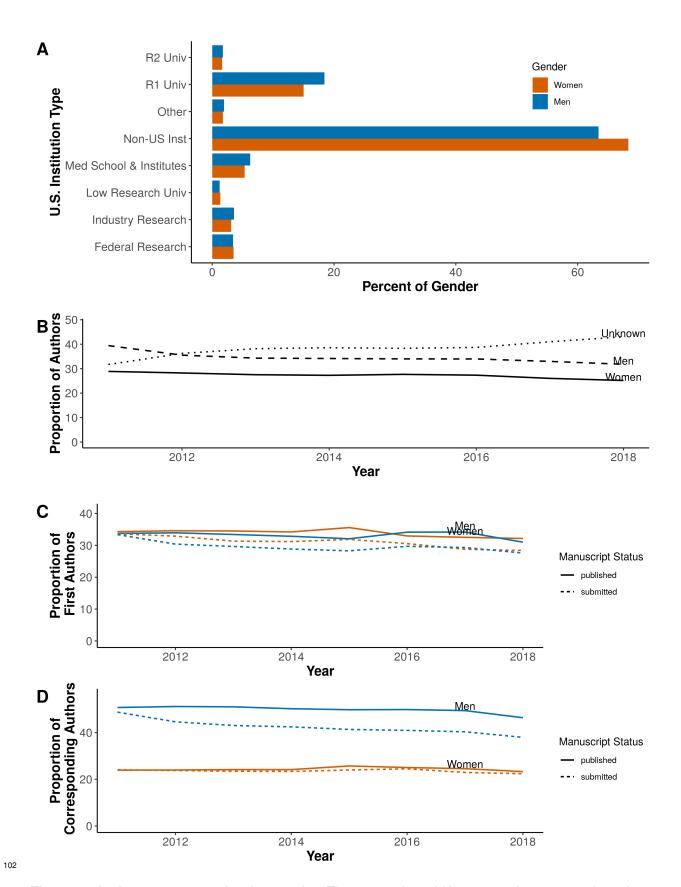


Figure 3. Author representation by gender. The proportion of (A) men and women authors from

US institutions, (B) men, women and unknown authors from 2012 - 2018. Unique manuscripts submitted from 2012 to 2018. The proportion of (C) first authors and (D) corresponding authors from 2012 - 2018. Solid lines indicate individuals, dashed indicate proportion of manuscripts submitted. Men indicated by blue and women by orange.

Women are underrepresented as authors The institution orgin of submitting men and women authors is similar to that of editors and reviewers, where over 60% were from non-US institutions, followed by 20% from R1 institutions (Fig. 3A). The proportions of men and women authors at ASM have decreased over time at equivalent rates, with a ratio of men to women authors of 4:3 since 2012 (or, 57% men) (Fig. 3B). This decrease corresponds with an increase in the proportion of unknown authors. Globally, microbiology researchers are 60% men and 40% women (Elsevier report). At ASM in September 2018, 38.37% of members who reported their gender, were women.

The proportion of papers submitted with men (blue dashed) and women (orange dashed) first 115 authors have remained constant with an average of 29.64 and 31.08 percent, respectively (Fig. 116 3C). Their respective proportions of published manuscripts are nearly identical at 33.16% for men 117 and 33.85% for women. The proportion of submitted papers with men corresponding authors has 118 remained steady at an average of 42.45% and the proportion with women corresponding authors 119 at 23.56%. However, their respective proportions of published manuscripts, 49.85% for men and 120 24.35% for women, are dissimilar (Fig. 3D). The published manuscript proportion where men are 121 corresponding authors has a 7.4 gap in percentage points, while the gap for women corresponding 122 authors is only 0.79. This trend is similar for middle and last authors (Fig. S3AB). 123

To better visualize the 7.52 decrease in the proportion of women who are first authors to those who are corresponding authors, we asked the proportions at which women have been retained through the peer review system at ASM journals.

There were 84482 men and 76215 women who were junior authors at ASM journals during the period of time under study. Of those junior authors, 13.59 of the men were also senior authors, 16.72 considered as reviewers, 11.11 actually reviewed, and 0.66 were editors at ASM journals.

At 8.25, 8.91, and 5.39, 0.25, just half as many women progressed through each role at ASM journals.

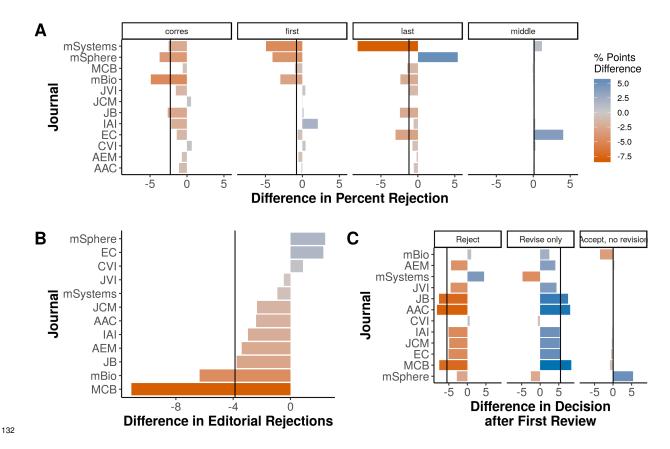


Figure 4. Difference in rejection rates by corresponding author gender. The percent of manuscripts rejected by author gender and type (e.g., corresponding, first, last, middle) at (A) all journals combined or at (B) each journal, which shows the difference in percent rejection rates. (C) The difference in percent editorial rejection rates at each journal, vertical line indicates the difference for all journals combined. (D) The difference in percentage points between each decision type following the first peer review, vertical lines indicate the difference value for all journals combined. The difference in rejection rates was determined by subtracting the rejection rate of women-authored papers from men-authored papers within each category. The shade (ranging from orange to blue) indicates the outperforming gender. No bar indicates no difference in percentange points.

Papers submitted by women have more negative outcomes than those submitted by men.

To better understand the percentage point difference in gendered performance (Fig. 3D), we next compared the rejection rates at each author stage. Middle authors were rejected at similar rates for men and women, a 0.11 percentage point difference across all journals combined. However,

senior woman-authored manuscripts are rejected more frequently than those authored by men with percentage point differences of -2.26 and -1.2 for corresponding and last authors, respectively (Fig. 4A). Breaking it down by individual journals, there are several instances where the overall trend is repeated or even amplified (e.g., AAC, IAI, JB, *mBio*, MCB). The greatest effect was observed when comparing the gender of corresponding authors, so we used this sub-population to further examine the difference in acceptance/rejection rates (Fig. 4A/S4).

We next compared the rejection rates for men and women corresponding authors at two different 153 bottlenecks, before and after the first peer review. Papers authored by women are editorially 154 rejected as much as 12 percentage points more often than those authored by men (Fig. 4B). The 155 percentage point difference at all ASM journals combined is -3.87 (vertical line), and two journals, 156 MCB and mBio, have more extreme percentage point differences. Papers authored by men and 157 women are equally likely to be accepted after the first round of review (Fig. 4C, right panel). 158 However, women-authored papers were rejected more often (left panel) while men-authored 159 papers were more often given revision decisions (center panel). Three journals, JB, AAC, and 160 MCB, have percentage point differences more extreme than for all ASM journals combined in both 161 rejection (-5.5) and revision (5.46) decisions (vertical lines). 162

In addition to manuscript decisions, other disparate outcomes may occur during the peer review process. To determine whether women-authored papers spent more time between being submitted and ready for publication, we compared the number of revisions, days spent in the ASM peer review system, and the number of days from submission to being ready for publication to those authored by men. Papers authored by women take slightly longer (from submission to ready for publication) than men at some journals (*mSphere*, *mBio*, *mSystems*, CVI, JB, JCM, AEM) despite spending similar amounts of time in the ASM journal peer review system, and having equivalent median number of revisions prior to acceptance (Fig. S5).

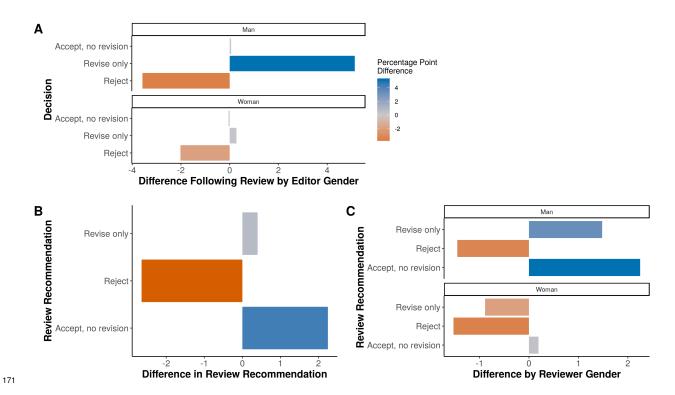


Figure 5. Difference in decisions or recomendations according to the gatekeeper gender.

(A) Effect of editor gender on the difference in percentage points for decisions following review at all journals combined. (B) Difference in percentage points for review recommendations and (C) how that is affected by reviewer gender.

To understand how gatekeeper (editor/reviewer) genders influence decisions (e.g., Fig. 4C), we grouped editor decisions and reviewer suggestions according to the gatekeeper gender. Both men and women editors rejected proportionally more women-authored papers, with men editors making revise decisions on papers authored by men more often than those authored by women (Fig. 5A). Reviewers are more likely to suggest rejections for women as compared to men, though no difference in revise suggestions were observed (Fig. 5B). Both men and women reviewers recommended rejection more often for women-authored manuscripts though only men recommended acceptance more often for men-authored manuscripts (Fig. 5C). Women reviewers suggested revision on women-authored papers more often than men-authored manuscripts.

To evaluate whether or not manuscript decisions are random when gender is taken into account, we used a logistic regression model to predict whether or not a manuscript was reviewed (e.g., editorially rejected or not). Our variables included the genders of the senior editor, editor, and

corresponding author, as well the proportion of authors that were women. The median AUC of this model was 0.59, indicating that the decisions are not completely random. However, because the value was below 0.6 the included variables are not sufficent to create a reliable model. This suggests that other factors influence gender-based decisions.

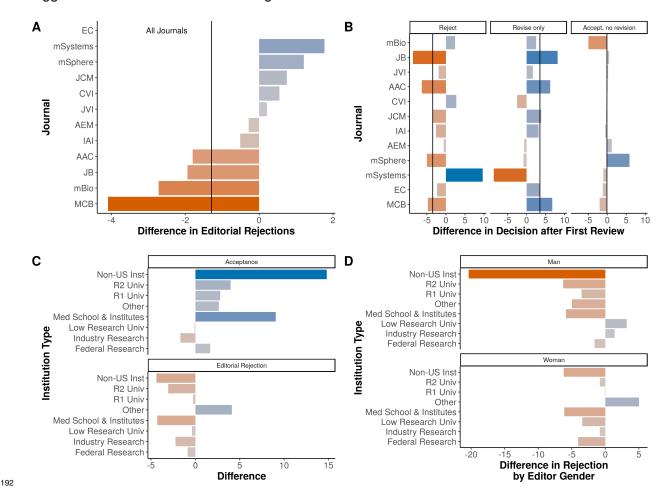


Figure 6. Impact of origin and U.S. institution type on manuscript decisions by gender. Difference in percentage points for corresponding authors in the U.S. (A) editorial rejections, (B) following first review, and by U.S. institution types (C) acceptance and editorial rejections (D) acceptance decisions according to editor gender. Vertical line indicates value for all ASM journals combined. NA indicates non-categorized US institutions or non-US institutions.

Country and institute of origin contribute to overperformance by men. The issue of non-random, gender-based manuscript decisions could be attributed to gender bias by journal gatekeepers, however, there are other types of bias that may contribute to, or obsure, overt

gender bias. For instance, a recent evaluation of peer-review outcomes at *eLife* found evidence of geographic homophiliy, that is, reviewers exhibited preference for research submitted by authors from their own country or region (Murray, 2019). Other studies have documented prestige bias, where men are overrepresented in more prestigous (i.e., more respected and competent) programs (Weeden, 2017). It is therefore possible, that what seems to be gender bias could be geographic or prestige bias interacting with the increased proportion of women submitting from outside the US or at lower prestige institutions (e.g., low research institutions) (Fig. 3A).

To try to separate how these factors affect manuscript decisions among corresponding authors, we next looked at the outcome of papers submitted only by corresponding authors at US institutions. When only considering US-based authors, the difference in percentage points for editorial rejections drops from -3.87 to -1.29, though trends across journals are consistent (Fig. 6A). The difference in percentage points for decisions after review mirror those of Figure 5, at the journal level (Fig. 6B). There are also changes in the values for all journals combined. The overperformance of women in rejection decisions decreases from -5.5 to -3.47, and from 5.46 to 3.51, for the overperformance of men in revise only decisions. The rate of accept decisions changed from -3.87 to -0.08 after restricting the analysis to US-based authors. These results suggest that the country of origin (e.g, US versus not) accounts for some gender bias, particulary for editorial rejections, but not all of it.

To address prestige bias, we next split the US-based corresponding authors according to their institution and re-evaluted the difference in percentage points for men and women. Editorial rejections occured most often for women from industry, R2, and medical schools or institutes while manuscripts submitted by men from medical schools or institutes, R1, and R2 instutitions were accepted more often than those submitted by women (Fig. 6C). The occurence of rejections after review occured more often for manuscripts submitted by women than those submitted by men, regardless of editor gender (Fig. 6D). There are a couple of exceptions where women authors from low research and industry research institutions recieved more positive decisions by men editors. The institution types from which manuscripts submitted by women had the greatest difference in percentage points from those submitted by men are medical schools or institutes and R2 institutions, with values of -20.37 and -6.28, respectively.

To understand if these factors affect manuscript decisions in a non-random manner, we used logistic regression model that took into account both origin (US vs non), institution (US institution 23 type) and the genders of both gatekeepers and authors. This model predicted whether 232 or not a manuscript was published and had a median auc of 0.61 indicating non-random 233 interaction between these factors. Those factors with the greatest postive impacts on liklihood of 234 publication were X.inst.gender.Industry.Research.female., X.inst.gender.Industry.Research.male., 235 X.inst.gender.Institute.or.Medical.School.male., US.inst.yes. The factors with the greatest negative 236 impacts on publication were US.inst.no, X.inst.gender.R2..Institution.female., X.inst.gender.R2..Institution.male., 237 inst.gender.NA.female. These results confirm that the country of origin and prestige bias impact 238 decisions in a non-random manner, but gender-based factors are still at play. 239

Discussion

- summary of results
- discuss representation
- cite lit

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- interpret results
 - decreased response to F editors may increased editorial burden for men?
 - does fewer collaborations contributed to decreased retainment
- remember to hammer trends & diversity of journals
 - why does it matter?

The under representation of women as corresponding authors in publication at ASM journals
has negative consequences for their careers and microbiology. Buckley et al, suggest that
being selected as a reviewer increases visibility of a researcher, which has a direct & significant
impact on salary. Therefore, the underrepresentation of women as reviewers hampers their
career progression and even their desire to progress since reviewing also signals adoption of the

researcher into the scientific community (Buckley et al, 2014). This is supported by Lerback and Hanson who noted that "It [reviewing] provides positive feedback that a scholar is respected and participating in their field and fosters self-confidence, all of which lead to increased retention of 256 women." (Lerback & Hanson, 2017) Retention of women in science is important to the progress of microbiology as a field since less diversity in researchers limits the diversity of perspectives, 258 approaches, and thus stunts the search for knowledge. In addition to boosting productivity and 259 knowledge, more diverse and equitable organizations are more inclusive and supportive for all 260 members (Potvin, 2018). It is thus a moral and scientific imperity for scientific societies and journals, such as ASM, to improve its own diversity, equity and inclusion efforts. The remainder of 262 this manuscript will focus on actions that can be taken at multiple levels of the peer review system 263 to support these efforts. 264

Discuss acceptance/rejection rates

- submission vs decision question of representation
- W asked to do more work by reviewers? (increased rej + increased time + same vers) 267
- cites 268

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 why it matters, compare 269

Addressing bias (gender, geographic, prestige or otherwise) during peer review process is a more difficult challenge, since it is partially the result of accumulated disadvangates and microaggressions (the actions resulting from implict biases). Implict bias training for gatekeepers is a start, as might be double-blinded peer review, a common practice in social sciences. To support efforts of making peer review more transparent, the review process could be unblinded following the editor's final decision on a manuscript. However, these solutions are only bandaids on a deeply infected wound, since both focus on the superficial issue of individuals instead of the underlying structure of the system that has selected for the bias at hand.

Few papers have found disparities between rejection rates of men and women and to our 278 knowledge, this is the first paper to collectively examine this issue with either submissions data 279 from 10+ journals or on the field of microbiology. Critics might argue that the effect size is 280

too small to really matter or that there are too many unaccounted factors to draw conclusions. 281 We acknowledge that these are limitations of our study along with a limited journal dataset, an 282 absence of reviewer comments for sentiment analysis, and that many ASM journals have a narrow 283 focus while the broad scope journals are relatively new. All of these factors prevent us from 284 generalizing our results across microbiology as a field. However, the consistency of the trends 285 to benefit men corresponding authors over women, across all journals included and literature 286 to-date confirms that this study is highly relevant for the ASM as a society and offers opportunities 287 to address both gendered representation in microbiology and systemic barriers to peer review at 288 our journals. + descriptive papers don't have confounding factors!

Data and Methods

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Data All manuscripts handled by ASM journals (e.g., mBio, Journal of Virology) that received an editorial decision between January 1st, 2012 and August 31st, 2018 were supplied as XML 292 files by ASM's publishing platform, eJP. Data were extracted from the XML documents provided using R statistical software (version 3.4.4) and the XML package (R citation). Data manipulation 294 was handled using the tidyverse, lubridate, and xml2 packages for R. Variables of interest 295 included: the manuscript number assigned to each submission, manuscript type (e.g., full length 296 research, erratum, editorial), category (e.g., microbial ecology), related (previously submitted) manuscripts, versions submitted, dates (e.g., submission, decision), author data (e.g., first, 298 last, and corresponding authorship, total number of authors), reviewer data (e.g., reviewer 299 score, recommendation, editor decision), and person data (names, institutions, country) of the 300 editors, authors, and reviewers. For this analysis, only original, research-based manuscripts 301 were included, e.g., long- and short-form research articles, New-Data Letters, Observations, 302 Opinion/Hypothesis articles, and Fast-Track Communications.

Data were visualized using the ggplot, scales, RColorBrewer, and ggalluvial packages for R. 304

Defining manuscript outcomes Many papers are immediately rejected by editors/EICs instead 305 of being sent to peer review, often due to issues of scope or percieved quality. These were defined as editorial rejections and identified as manuscripts rejected without record of review. Alternately, editors could send papers out for review by two or three experts in the field. The reviewers make suggestions to the editor who decides whether the manuscript in question should be accepted, rejected, or sent back for revision. At ASM journals, manuscripts with suggested revisions that are expected to take more than 30 days are rejected, but generally encouraged to resubmit. If resubmitted, the authors are asked to note the previous (related) manuscript and the resubmission is assigned a new manuscript number. Multiple related manuscripts were tracked together by generating a unique grouped manuscript number based on the recorded related manuscript numbers. This grouped manuscript number served multiple purposes including: tracking a single manuscript through multiple rejections or transfers between ASM journals and to avoid duplicate counts of the same authors for the same manuscript.

Institution classification To identify the communities represented, we used Carnegie classifications to group US-based institutions into R1, R2, low (not R1 or R2), and medical research. Medical schools and institutions (e.g., Mayo clinic) were grouped together. Two other categories were added to represent industry and federal research. The other category represents uncategorized US institutions.

Bias analysis and presentation To identify potential biases, we calculated the difference in percentage points between a given outcome for men and women, e.g., the percentage point difference in acceptance rates is the acceptance rate for men minus the acceptance rate for women. A positive value indicates that men receive the outcome more often than women, whereas a negative value indicates that women outperform men in the given metric. To correct for the disparity in the participation of women relative to men at ASM journals, all percentage point comparisions are made relative to the gender and population in question.

Logistic regression models For the L2-regularized logistic regression models, we established modeling pipelines for a binary prediction task. First, we randomly split the data into training and test sets so that the training set consisted of 80% of the full dataset while the test set was composed of the remaining 20% of the data. To maintain the distribution of the 2 model outcomes that was found with the full dataset, we performed stratified splits. The training data

was used to build the models and the test set was used for evaluating predictive performance. To build the models, we performed an internal five-fold cross-validation where we tuned the **cost** hyperparameter which determines the regularization strength where smaller values specify stronger regularization. This internal cross-validation was repeated 100 times. Then, we trained the full training dataset with the selected hyperparameter values and applied the model to the held-out data to evaluate the testing predictive performance of each model. The data-split, hyperparameter selection, training and testing steps were repeated 100 times to get a reliable and robust reading of model performance. Models were trained using the machine learning wrapper caret package (v.6.0.81) in R (v.3.5.0).

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Gender prediction and assignment The gender assignment API genderize.io was used to predict an individual's gender based on their given names, and country where possible. The genderize io platform uses data gathered from social media to predict gender based on given names with the option to include an associated language or country to enhance the odds of successful prediction. Since all manuscripts are submitted in English, precluding language association for names with special characters, names were standardized to ASCII coding (e.g., "José" to "Jose"). We next matched each individuals country against the list of X country names accepted by genderize.io. Using the GenderGuesser package for R, all unique given names associated with an accepted country were submitted to the genderize.io API and any names returned without a predictive assignment of either male or female were resubmitted without an associated country. All predictive assignments of either male or female are returned with a probability match of 0.50 or greater. The predicted genders of all given names (with and without an associated country) whose probabilities were greater or equal to our arbitrary success cut off of 0.65 were used to assign predicted gender to the individuals in our dataset. Predicted genders were assigned to individuals in the following order: first names and country, first names, middle names and country, middle names (Fig. S7). The presenting gender (man/woman) of editors and senior editors in our dataset was hand validated using Google where possible.

We recognize that biological sex (male/female) is not always equivalent to the gender that an individual presents as (man/woman), which is also distinct from the gender(s) that an individual may self-identify as. For the purposes of this manuscript, we choose to focus on the presenting

gender (man/woman/unknown) based on their first names and/or appearance (for editors). In the interest of transparency, we include those individuals whose names don't allow a high degree of confidence for gender assignment in the "unknown" category of our analysis.

Validation of gender prediction We first validated the algorithm using a set of 3265 names 367 whose gender had been hand-coded based on appearance and were generously provided to us by ____ (preprint cite). The names were supplied to the genderize algorithm both with and 369 without the accompanying country data. The data returned include the name, predicted gender 370 (male, female, na), the probability of correct gender assignment (ranging from 0.5 to 1.0), and 371 the number of instances the name and gender were associated together (1 or greater). The 372 genderize algorithm returned gender predictions for 2899 when first names were given and 2167 373 when country data was also supplied (732 names were associated with countries unsupported by 374 genderize).

Sensitivity and specificity, are measurements of the algorithm's tendency to return correct answers instead of false positives (e.g., a man incorrectly gendered as a woman) or false negatives (e.g., a woman incorrectly gendered as a man). The closer these values are to 1, the smaller the chance that the algorithm will return the correlating false response. Accuracy is a composite measure of the algorithm's ability to differentiate the genders correctly. These measurements were calculated from the datasets (with and without country data supplied) at three different probability threshold cutoffs: the default genderize (0.5), a probability threshold of 0.85 (0.85), and a modified probability of 0.85, which factors in the number of instances returned (pmod0.85)(citations).

At the 0.5 threshold, the dataset returned a sensitivity of 0.8943 and specificity of 0.9339 for an accuracy of 0.911, compared to a marginally higher accuracy of 0.9146 for the dataset where country data were included (Table S1). Generally speaking, the accuracy increases as the threshold increases along with slight trade offs between sensitivity and specificity. For the purposes of our analysis, we opted to use the pmod0.85 threshold moving forward (Table S1, in bold).

To understand the extent of geographic bias in our gender assignment against regions and languages with genderless naming conventions, or that lack social media for incorporation

into the genderize algorithm, we compared the number of names predicted without associated country data to when country data was also supplied. In our test dataset, the top five countries associated with names were United States, Germany, United Kingdom, France, and China and the countries with the highest proportion of un-predicted genders when country data were supplied are Cambodia, Iceland, Indonesia, Ireland, and Mexico, where the maximum number of names supplied ranged from 1 to 15. To determine the impact of each country towards the overall percentage of names whose genders were not predicted (27.14%), we found the difference between the percent of names unpredicted for each country and the overall percentage, multiplied by the proportion of observations from that country to the total observations and finally divided by the overall percentage of unpredicted names (Fig. S8). The top five countries with the greatest impact on unpredicted names, and thus the countries receiving the most negative bias from genderize were Canada, China, Ireland, Belgium, and Sweden (Fig. S9). These data suggest that there is likely some bias against countries with gender-neutral naming conventions (China), and indicates the stringency with which the algorithm applies gender to names that are accompanied by country data. For instance, strongly gendered names such as Peter and Pedro were not assigned gender when associated with Canada.

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We next applied the genderize algorithm at the pmod0.85 threshold to our journals dataset and tested its validity on a small portion. All first names collected from our dataset were submitted to genderize both with and without country data. Only those predictions whose pmod were equivalent or greater than 0.85 were carried to the next step. The predicted genders were assigned to individuals in the following order: first names and country, first names, middle names and country, middle names. Given the relatively small number of editors and senior editors in our dataset, the presenting gender (man/woman) of editors and senior editors in our dataset was hand-validated using Google where possible. Of the 1072 editor names, 938 were predicted by genderize for an accuracy of 0.9989339, thus increasing our confidence in the gender predictions where made.

In our full dataset, the five countries with the most individuals were United States, China, Japan,
France, and Germany and the countries with the highest proportion of un-predicted genders were
Burundi, Chad, Kingman Reef, Korea (North), Democratic People's Republic of, and Maldives,
where the maximum number of names supplied ranged from 1 to 4. Proportionally, fewer names

- in our full dataset were assigned gender than in our validation dataset (40.01% unpredicted versus 27.14% unpredicted, respectively). Since adjusting the workflow to predict the gender of names both with and without country data, the countries receiving the most negative bias from genderize were China, Japan, Korea, Republic of, India, Taiwan, Province of China (Fig. S10). These data indicate what we previously predicted, that the genderize algorithm has bias against countries with gender-neutral naming conventions.
- Code availability The code for all analysis steps, logistic regression pipeline, and an Rmarkdown
 version of this manuscript, is available at https://github.com/SchlossLab/Hagan_Gender_mBio_
 2019/
- Acknowledgements We would like to thank Nicole Broderick and Arturo Casadevall for providing
 their dataset for genderize validation and acknowledge Arianna Miles-Jay and Joshua MA Stough
 for their comments.
- A.K.H. was responsible for project and experiment design, data analysis, and drafting the manuscript. B.T. completed the logistical regression models. P.S. provided funding, resources and conceptual advice. All authors contributed to the final manuscript. P.S. is Chair of ASM Journals and A.K.H. was ASM staff prior to publication of the analysis. B.T. reports no conflict of interest.
- Funding for this work was provided by the American Society for Microbiology. The sponsors had no role in study design or interpretation.

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