

Exercise 2: Numerical Stability of the Posterior.

We consider the logistic differential equation

$$\dot{x}(t) = r x(t) (k - x(t)), \quad x(0) = x_0,$$

describing the growth of a population in a habitat with capacity k . The initial state $x_0 = 0.1$ and the reproduction rate $r = 0.25$ are known. However, we are uncertain about the capacity k .

We propose a uniform prior for $k \sim \pi_0 = \mathcal{U}[10, 20]$ and aim to update that prior given noisy observations $y = (y_1, y_2, y_3)^\top \in \mathbb{R}^3$ of

$$y_i = x(i) + \eta_i, \quad \eta_i \sim \mathcal{N}(0, 1)$$

at the 'times' $i = 1, 2, 3$ with stochastically independent η_i . The given data is $y = (3, 14, 17)^\top$.

The logistic equation can be solved analytically

$$x(t) = \frac{k}{1 + \exp(-r k t) \left(\frac{k}{x_0} - 1 \right)}, \quad t \geq 0,$$

but to explore the stability of the posterior distribution with respect to approximations of the forward operator, we will solve it numerically using a time stepping scheme.

Euler's method: Consider

$$\dot{x}(t) = f(t, x(t)) \quad x(0) = x_0.$$

Then, choose a stepsize $h > 0$ and generate approximations $x_i \approx x(t_i)$ for $t_i = i h$, $i \in \mathbb{N}_0$, recursively by

$$x_{i+1} = x_i + h f(t_i, x(t_i)).$$

Under suitable assumptions, the Euler method has a linear convergence rate, i.e.,

$$\sup_{i=1, \dots, T/h} |x_i x(t_i)| \propto h.$$

Thus, using the analytic solution we get the true posterior $\pi_{k|y}$ and using the Euler method we get approximate posteriors $\pi_{k|y}^h$ on the uncertain carrying capacity k .

- (a) Using Bayes' rule, explicitly compute and plot the (normalized) posterior (densities) $\pi_{k|y}$ and $\pi_{k|y}^h$ for $h \in \{0.5, 0.25, 0.1, 0.05, 0.025, 0.01, 0.005, 0.0025, 0.001\}$.

Comments: You may use the Matlab files `Euler.m` and `Exercise_01_template.m` provided in the git repository. To estimate the normalizing constant (i.e., the reciprocal of the evidence), you may use numerical quadrature on $[10, 20]$, e.g., the trapezoidal rule

$$\int_a^b f(k) dk \approx \frac{\Delta k}{2} f(a) + (\Delta k) \sum_{j=1}^{N-1} f(a + j \Delta k) + \frac{\Delta k}{2} f(b), \quad N = \frac{b-a}{\Delta k}$$

for, e.g., $\Delta k = 0.01$.

- (b) Compute the corresponding Hellinger distances $d_{\text{Hell}}(\pi_{k|y}, \pi_{k|y}^h)$ and plot them in a log-log-plot versus the stepsize h in order to verify

$$d_{\text{Hell}}(\pi_{k|y}, \pi_{k|y}^h) \propto h$$

in accordance to the theoretical result from the lectures.

Exercise 3: MCMC Performance.

We consider again the logistic differential equation

$$\dot{x}(t) = r x(t) (k(t) - x(t)), \quad x(0) = x_0,$$

describing the growth of a population in a habitat, but now with time-dependent capacity $k(t)$, modelling the influence of the seasons on resources in the habitat. The initial state is now $x_0 = 1$ and the reproduction rate is again $r = 0.25$.

We want to infer the capacity k again. To this end we observe the data $y_i = x(i) + \eta_i$ with stochastically independent $\eta_i \sim \mathcal{N}(0, 0.25)$ at 'times' $i = 1, \dots, 12$.

For your experiments you may use the Matlab template `Exercise_03_template.m` provided in the git repository. The logistic equation is solved numerically using Euler's method on $[0, 12]$ with stepsize $h = 0.01$ using the special function `EulerLogEq.m` also provided. This function just needs to know $k(t_i)$ at times $t_i = i h$, $i = 0, \dots, 1200$, i.e., it requires a vector $\mathbf{k} \in \mathbb{R}^{1201}$ as input.

- (a) In order to generate (synthetic data) and for comparison we take as ground truth

$$k_{\text{true}}(t) = 12 \exp\left(-\frac{1}{4} \cos\left(\frac{\pi}{6} t\right)\right).$$

In our simulations we add noise to the true observations and set the seed `rng(17)` for the pseudo random number generator in Matlab when drawing η_1, \dots, η_{12} .

Use Euler's method to compute the ground truth $x(t)$ in $t \in [0, 12]$ and then plot it together with the perturbed data y_1, \dots, y_{12} .

- (b) We use a *lognormal* prior for k given by the series expansion

$$\log(K(t)) = \log(10) + \sum_{j=0}^{\infty} U_j \phi_j(t), \quad U_j \sim \mathcal{N}(0, \lambda_j)$$

where $\phi_0(t) \equiv 1/12$, $\lambda_0 = \frac{1}{\pi^2}$, and $\phi_j(t) = \frac{1}{\sqrt{6}} \sin(j \frac{\pi}{12} t)$, $\lambda_j = \frac{1}{\pi^2 j^2}$, for $j \geq 1$.

In our simulations, we truncate the expansion after $N = 1000$ terms. Further, we need to know the ϕ_j only at the grid points $t_i = i h$, $i = 0, \dots, 1200$. Thus, our Bayesian inference problem is to infer the coefficients $U = (U_1, \dots, U_N)^\top$ where the prior distribution is

$$U \sim \pi_0 = \mathcal{N}(0, \Lambda), \quad \Lambda = \text{diag}(\lambda_1, \dots, \lambda_N)$$

Plot 100 prior realizations for k and compare them visually to the ground truth.

- (c) Run an MCMC simulation using the pCN or random walk proposals discussed in lectures and use it to compute the posterior mean of the following quantity of interest

$$f(k) = \int_0^{12} k(t) dt.$$

Tune the stepsize to an average acceptance rate of roughly 0.25 before running the full Markov chain simulation with 20,000 iterations. **Comments:** You can again apply the trapezoidal rule to approximate $f(k)$. To estimate the average acceptance rate, you can simply average the values of the acceptance probability $\alpha(u_i, v_i)$ over the iterations.

- (d) Store the states of the Markov chain in a matrix and use them to plot (the last) 100 realizations of the posterior for k . Compare them visually to the ground truth. Also plot the pointwise mean and the (centered) 95% credibility intervals.
- (e*) Compute an asymptotic confidence interval for $\mathbb{E}_{\pi_{k|y}}[f]$ using `autocorr` to estimate the autocorrelation.