

A Stochastic Variational Inference Approach for Semiparametric Distributional Regression

Final Kolloquium MSc Applied Statistic

Sebastian Lorek

First supervisor: Prof. Dr. Thomas Kneib

Second supervisor: Gianmarco Callegher

2023-11-28

Table of Contents I

Introduction

Theory

Application

Introduction

Statistical Inference

Frequentist inference:

- ML estimation (Newton-Raphson/Fisher-Scoring)
- Hyperparameter estimation
- REML for latent parameters (random-effects)
- SGD, Automatic differentiation in machine/deep learning

Bayesian inference:

- Conjugate models, full conditional conjugate models, non-conjugate models
- Gibbs sampling
- Rejection/Importance sampling
- MCMC (IWLS, HMC, NUTS)
- MAP estimation in combination with the Laplace approximation

Bayesian Inference

- Focal point of interest is the posterior distribution
- General posterior specification for Bayesian regression

$$p(\theta|\mathbf{y}, \mathcal{D}) = \frac{p(\mathbf{y}, \theta | \mathcal{D})}{p(\mathbf{y}|\mathcal{D})} \quad (1)$$

$$= \frac{p(\mathbf{y} | \theta, \mathcal{D})p(\theta)}{p(\mathbf{y}|\mathcal{D})} \quad (2)$$

$$= \frac{p(\mathbf{y} | \theta, \mathcal{D})p(\theta)}{\int p(\mathbf{y} | \theta, \mathcal{D})p(\theta) d\theta} \quad (3)$$

$$\propto p(\mathbf{y} | \theta, \mathcal{D})p(\theta) \quad (4)$$

- Primary challenge is the calculation of the normalizing constant/evidence
- We need approximate methods that bypass the direct calculation of the normalizing constant

- MCMC methods have been **so far** the work horse for Bayesian inference in classical statistics
- Set up Markov chain (adhere to detailed balance and ergodicity) and sample ...
- Enjoys nice properties
 - Assurance of convergence to the true posterior
- But unfortunately does not scale well for modern applications
 - High dimensional parameter spaces (tausands of parameters)
 - Large dataset
 - Many latent and hyper-parameters
- Trade some **accuracy** for **scalability**
- Make use of SGD and automatic differentiation which works well for inference in machine/deep learning (frequentist inference)

Theory

Variational Inference

- VI is a method from machine learning to **approximate** probability densities (Jordan et al. [1999](#))
- Approximate posterior with a variational distribution $q(\theta)$ from a predefined (parametric) variational family \mathcal{Q}

$$q(\theta) \in \mathcal{Q}$$

- Use optimization (SGD) to find a member that is as closely as possible to the true posterior
- What means close in terms of distributions ?

Kullback-Leibler divergence

- Divergence measure (Kullback and Leibler 1951) that quantifies the proximity between two probability distributions

$$\begin{aligned} D_{\text{KL}} &= \int q(\theta) \ln \left(\frac{q(\theta)}{p(\theta|\mathbf{y}, \mathcal{D})} \right) d\theta \\ &= \mathbb{E}_{q(\theta)} \left[\ln \left(\frac{q(\theta)}{p(\theta|\mathbf{y}, \mathcal{D})} \right) \right] \end{aligned}$$

- In short $D_{\text{KL}}(q \parallel p)$
- It holds that $D_{\text{KL}}(q \parallel p) \geq 0$
- Has some nice properties but also drawbacks (not a distance/metric)

Optimization objective

- Use optimization to find a variational distribution that is as close as possible in terms of the divergence measure to the true posterior

$$\begin{aligned} q^*(\theta) &= \arg \min_{q(\theta) \in \mathcal{Q}} D_{\text{KL}}(q(\theta) || p(\theta | \mathbf{y}, \mathcal{D})) \\ &= \arg \min_{q(\theta) \in \mathcal{Q}} \int q(\theta) \ln \left(\frac{q(\theta)}{p(\theta | \mathbf{y}, \mathcal{D})} \right) d\theta \\ &= \arg \min_{q(\theta) \in \mathcal{Q}} \mathbb{E}_{q(\theta)} \left[\ln \left(\frac{q(\theta)}{p(\theta | \mathbf{y}, \mathcal{D})} \right) \right]. \end{aligned}$$

- Flexibility of \mathcal{Q} significantly influences the optimization
 - Complex $\mathcal{Q} \rightarrow$ better approximation, but increased complexity during optimization
 - Simple $\mathcal{Q} \rightarrow$ worse approximation, but simpler optimization
- Objective offers theoretical insights but remains **infeasible** to compute, due to containing the posterior (evidence)

Evidence Lower Bound

- Think about a way to introduce a mathematical **equivalent** objective that does not depend on the evidence
- Lets start with the log-evidence

$$\ln(p(\mathbf{y}|\mathcal{D})) = \int q(\theta|\phi) \ln(p(\mathbf{y}|\mathcal{D})) d\theta \quad (5)$$

$$= \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y}|\mathcal{D})p(\theta|\mathbf{y},\mathcal{D})}{p(\theta|\mathbf{y},\mathcal{D})} \right) d\theta \quad (6)$$

$$= \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y},\theta|\mathcal{D})}{p(\theta|\mathbf{y},\mathcal{D})} \right) d\theta \quad (7)$$

$$= \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y},\theta|\mathcal{D})}{q(\theta|\phi)} \frac{q(\theta|\phi)}{p(\theta|\mathbf{y},\mathcal{D})} \right) d\theta \quad (8)$$

$$= \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y},\theta|\mathcal{D})}{q(\theta|\phi)} \right) d\theta + \int q(\theta|\phi) \ln \left(\frac{q(\theta|\phi)}{p(\theta|\mathbf{y},\mathcal{D})} \right) d\theta \quad (9)$$

$$= \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y},\theta|\mathbf{X})}{q(\theta|\phi)} \right) \right] + D_{\text{KL}}(q(\theta|\phi)||p(\theta|\mathbf{y},\mathcal{D})) \quad (10)$$

- Parameterize q with ϕ

Optimization objective revisited

- Rewrite 10 as

$$\ln(p(\mathbf{y}|\mathbf{X})) - D_{\text{KL}}(q||p) = \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}, \theta|\mathcal{D})}{q(\theta|\phi)} \right) \right] \quad (11)$$

- And take arg max w.r.t. ϕ

$$\arg \max_{\phi} \ln(p(\mathbf{y}|\mathcal{D})) - D_{\text{KL}}(q||p) = \arg \max_{\phi} \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}, \theta|\mathcal{D})}{q(\theta|\phi)} \right) \right]$$

$$\arg \max_{\phi} -D_{\text{KL}}(q||p) = \arg \max_{\phi} \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}, \theta|\mathcal{D})}{q(\theta|\phi)} \right) \right]$$

$$\arg \min_{\phi} D_{\text{KL}}(q||p) = \arg \max_{\phi} \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}, \theta|\mathcal{D})}{q(\theta|\phi)} \right) \right]$$

- New optimization objective

$$\begin{aligned}\hat{\phi} &= \arg \max_{\phi} \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}, \theta | \mathcal{D})}{q(\theta | \phi)} \right) \right] \\ &= \arg \max_{\phi} \text{ELBO}(\phi)\end{aligned}$$

- Take a breath 🥴

■ What does the ELBO ?

$$\text{ELBO}(\phi) = \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}, \theta|\mathcal{D})}{q(\theta|\phi)} \right) \right] \quad (12)$$

$$= \mathbb{E}_{q(\theta|\phi)} \left[\ln \left(\frac{p(\mathbf{y}|\mathcal{D}, \theta)p(\theta)}{q(\theta|\phi)} \right) \right] \quad (13)$$

$$= \mathbb{E}_{q(\theta|\phi)} [\ln(p(\mathbf{y}|\mathcal{D}, \theta))] + \mathbb{E}_{q(\theta|\phi)} [\ln(p(\theta))] - \mathbb{E}_{q(\theta|\phi)} [\ln(q(\theta|\phi))] \quad (14)$$

$$= \mathbb{E}_{q(\theta|\phi)} [\ln(p(\mathbf{y}|\mathcal{D}, \theta))] - D_{\text{KL}}(q(\theta|\phi) || p(\theta)) \quad (15)$$

■ Actually something similar to MAP/ML 🤔

Variational family

- What is the \mathcal{Q} and thus $q(\theta|\phi)$?
- Start simple

$$q(\theta|\phi) = \prod_{j=1}^J q_j(\theta_j | \phi_j)$$

- Known as mean-field variational family
- q_j factors in the variational distribution
 - Some parameteric distribution that respects the parameter space of θ_j
 - All model parameters are independent

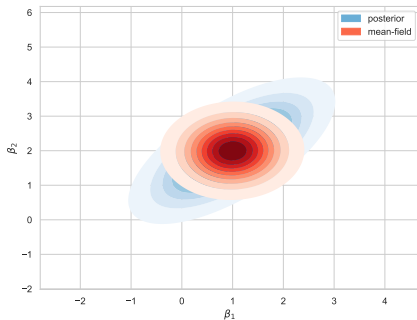


Figure 1: Mean-field approximation to a 2-D multivariate normal distribution, based on Blei et al. ([ibid.](#), p. 9, figure 1).

- Augment the variational distribution to blocks of parameters
- Structured mean-field variational inference (Wainwright and Jordan 2007)

$$q(\boldsymbol{\theta}|\boldsymbol{\phi}) = \prod_{j=1}^J q_j(\boldsymbol{\theta}_j | \boldsymbol{\phi}_j)$$

- Captures interdependencies for blocks of parameters

Coordinate ascent variational inference (side node)

- Traditional way of solving opt. objective is CAVI (Blei et al. [2017](#))
 - However CAVI does not scale well
 - Closely connected to Gibbs sampling
 - Only works for conditional conjugate models
 - If you search for VI CAVI is still all over the place, see f.e. [wikipedia](#)
- Of course we want to be able to also conduct inference in non-conjugate models ✗

Stochastic variational inference

- SGD to optimize the ELBO (Hoffman et al. [2012](#))
- 2 sources of stochasticity
 - We use a subset \mathcal{I} of the data in each iteration (ELBO remains unbiased)
 - We need to evaluate the integral in the ELBO

$$\text{ELBO}(\phi)_{\mathcal{I}} = \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y}_{\mathcal{I}}, \theta|\mathcal{D}_{\mathcal{I}})}{q(\theta|\phi)} \right) d\theta \quad (16)$$

$$\nabla_{\phi} \text{ELBO}(\phi) = \nabla_{\phi} \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y}_{\mathcal{I}}, \theta|\mathcal{D}_{\mathcal{I}})}{q(\theta|\phi)} \right) d\theta \quad (17)$$

- Common method to solve this problem is Monte Carlo integration

$$\nabla_{\phi} \text{ELBO}(\phi)_{\mathcal{I}} = \nabla_{\phi} \int q(\theta|\phi) \ln \left(\frac{p(\mathbf{y}_{\mathcal{I}}, \theta | \mathcal{D}_{\mathcal{I}})}{q(\theta|\phi)} \right) d\theta \quad (18)$$

$$\approx \nabla_{\phi} \frac{1}{S} \sum_{s=1}^S \ln \left(\frac{p(\mathbf{y}_{\mathcal{I}}, \theta^s | \mathcal{D}_{\mathcal{I}})}{q(\theta^s|\phi)} \right), \quad \theta^s \sim q(\theta|\phi). \quad (19)$$

- But this does not work 😞
- If we change ϕ even infinitesimal the samples θ^s are invalid, which we used to calculate $\nabla_{\phi} \text{ELBO}(\phi)_{\mathcal{I}}$ in the first place

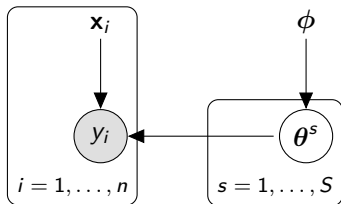


Figure 2: Plate notation of the dependence structure in a Bayesian regression model for VI, when sampling from the variational distribution.

Reparameterization gradient estimator

- Reparameterize $\theta = \mathbf{g}_\phi(\epsilon)$, with a bijective function \mathbf{g}_ϕ such that we can use SGD with Monte Carlo integration (Kingma and Welling 2013; Kucukelbir et al. 2016; Rezende et al. 2014)
- Amounts to using the (multivariate) change of variable theorem for probability density functions

$$\mathbf{g}_\phi : \mathbb{R}^d \rightarrow \mathbb{R}^d; \theta, \epsilon \in \mathbb{R}^d$$

$$\theta = \mathbf{g}_\phi(\epsilon), \epsilon \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$$

$$q(\theta|\phi) = \begin{cases} p_\epsilon(\mathbf{g}_\phi^{-1}(\theta)) \left| \det(\mathbf{J}_{\mathbf{g}_\phi^{-1}}) \right|, & \text{if } \theta \text{ is in the codomain of } \mathbf{g}_\phi \\ 0, & \text{else.} \end{cases}$$

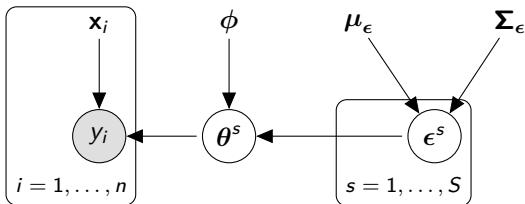


Figure 3: Plate notation of the dependence structure in a Bayesian regression model in VI, using the "reparameterization-trick".

- This “trick” allows us to pull the gradient operator inside of the monte carlo integral and use the chain rule

$$\begin{aligned}\nabla_{\phi} \text{ELBO}(\phi)_{\mathcal{I}} &\approx \frac{1}{S} \sum_{s=1}^S \nabla_{\phi} \ln \left(\frac{p(\mathbf{y}_{\mathcal{I}}, \theta = \mathbf{g}_{\phi}(\epsilon) | \mathcal{D}_{\mathcal{I}})}{q(\theta = \mathbf{g}_{\phi}(\epsilon) | \phi)} \right) \Big|_{\theta=\theta^s} \\ &\approx \frac{1}{S} \sum_{s=1}^S \nabla_{\theta} \ln \left(\frac{p(\mathbf{y}_{\mathcal{I}}, \theta = \mathbf{g}_{\phi}(\epsilon) | \mathcal{D}_{\mathcal{I}})}{q(\theta = \mathbf{g}_{\phi}(\epsilon) | \phi)} \right) \Big|_{\theta=\theta^s} \nabla_{\phi} \mathbf{g}_{\phi}(\epsilon^s), \\ &\text{with } \theta_{\phi}^s = \mathbf{g}_{\phi}(\epsilon^s), \epsilon^s \sim \mathcal{N}(\mathbf{0}, \mathbf{I}), s = 1, \dots, S\end{aligned}$$


- Opens the door for backpropagation and thus automatic diff. 💡

“Black-box” variational inference

- Using SVI with the reparameterization gradient estimator
- Researcher only formulates a probabilistic model and provides a dataset (Kucukelbir et al. [2016](#)), inference algo. is model “agnostic”
- What is \mathbf{g}_ϕ ?
 - Linear and non-linear choices
 - We consider linear choice

$$\theta_j = \mathbf{L}_j \epsilon_j + \mu_j, \epsilon_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$$

$$\theta_j \sim \mathcal{N}(\mu_j, \mathbf{L}_j \mathbf{L}_j^T)$$

- For positive restricted parameters we need to chain another transformation layer via exp transformation
- Chaining variable transformations 
 - Normalizing flows (Rezende et al. [2014](#))
 - Allow for expressive \mathcal{Q} s but are more difficult to optimize

- We use Adam (Kingma and Ba 2014)

$$\hat{\phi}^t = \hat{\phi}^{t-1} + \rho_t \nabla_{\phi} \text{ELBO}(\phi)_{\mathcal{I}^t} \Big|_{\phi=\hat{\phi}^{t-1}}$$

Full Algorithm

Data: $\mathcal{D}_{\text{train}}$; \mathcal{D}_{val}

Require: Learning rate α ; stopping threshold ε ; mini-batch size M ; share train w ; num. var. samples S ; num. epochs: E

Initialize $\hat{\phi}^0$; set $t = 1$

for $e = 1$ **to** E **do**

$n_{\text{train}} = |\mathcal{D}_{\text{train}}| * w$

 create the mini-batches, $\mathcal{B} = \{\dots, \mathcal{I}^k, \dots\}$, $k = 1, \dots, n_{\text{train}} // M (+1)$

for $k = 1$ **to** $n_{\text{train}} // M (+1)$ **do**

 sample noise, $\epsilon_j^s \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$, $s = 1, \dots, S$, $\forall j$

 calculate approx. gradient, $\nabla_{\phi} \text{ELBO}(\phi)_{\mathcal{I}^k} \Big|_{\phi=\hat{\phi}^{t-1}}$

 update variational parameters, $\hat{\phi}^t = \hat{\phi}^{t-1} + \rho_t \nabla_{\phi} \text{ELBO}(\phi)_{\mathcal{I}^k} \Big|_{\phi=\hat{\phi}^{t-1}}$

 calculate approx. ELBO, $\text{ELBO}(\hat{\phi}^t)_{\mathcal{D}_{\text{val}}}$

$t = t + 1$

end

if $t > 200$ **then**

$\Delta \text{ELBO} = |\text{ELBO}(\hat{\phi}^t)_{\mathcal{D}_{\text{val}}} - \text{ELBO}(\hat{\phi}^{t-200})_{\mathcal{D}_{\text{val}}}|$

else

$\Delta \text{ELBO} = \infty$

end

if $\Delta \text{ELBO} < \varepsilon$ **then**

break

end

end

Result: $\hat{\phi}$; $\text{ELBO}(\hat{\phi})_{\mathcal{D}_{\text{val}}}$

Algorithm 1: BBVI algorithm.

Impact Analysis

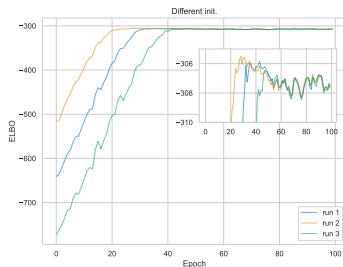


Figure 4: ELBO traces for 3 different SGD runs, using different initializations but the same seed. We use a batch size of 128, 64 samples from the variational distribution and a learning rate of $1e-2$.

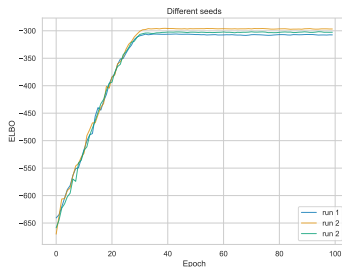


Figure 5: ELBO traces for 3 different SGD runs, using different seeds but the same initialization. Otherwise same configuration as in Figure 4.

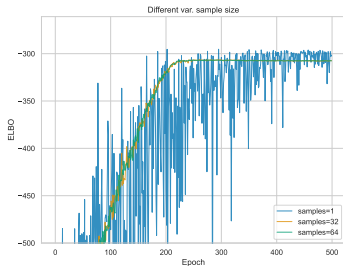


Figure 6: ELBO traces for 3 different SGD runs, using different variational sample sizes. We use batch VI with a learning rate of $1e-2$ and the same seed as in Figure 4.

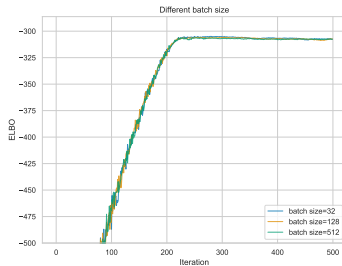


Figure 7: ELBO traces for 3 different SGD runs, using different batch sizes. We use a variational sample size of 64 with a learning rate of $1e-2$ and the same seed as in Figure 4.

Semiparametric distributional regression

- Not only normally distributed responses
- Linear predictors with inverse link function
- Structured additive linear predictors so fixed and smooth effects
 - B-spline basis functions
- Augmented with priors
 - Bayesian P-splines

Application

Implementation

- Developed a small python package `tigerpy`, which consists of two libraries
- A model building library `tigerpy.model`
 - Construct the model
 - Uses the idea of probabilistic graphical models
- An inference library `tigerpy.bbvi`
 - Runs the inference algorithm
- Aligned with concepts found in `liesel` (Riebl et al. [2022](#))

When walking about the countryside of Italy, the people will not hesitate to tell you that JAX has “[una anima di pura programmazione funzionale](#)”. (JAX [docs](#), Bradbury et al. (2018))

- There are great things about JAX 🥰
 - Uses a `numpy` flavored API
 - Closely follows the math (`jax.grad`)
 - Is fast (if you follow JAXs principles)
- There are things that cause headaches 🤔
 - Pure functions
 - Tracing
 - Efficiency considerations in JIT-compiled code

- `tigerpy.model` constructs under the hood a DAG
- Employs the `networkx` package for constructing, traversing and visualizing the DAG

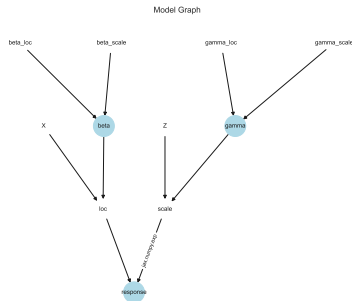









Figure 8: The DAG visualization for location-scale regression from the method `.visualize_graph()`.

Simulation Studies

Open Problems

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