# Confidence Intervals in the Normal Model and the Binomial Model

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### Confidence Interval - Normal Distribution

- Example: In connection with the behaviour of children in road traffic, the reaction time of 10-year-old children is investigated. Experience has shown that reaction times are normally distributed, but the paramaters  $\mu$  and  $\sigma$  are unknown.
- Sample of n=51 observations:

```
## [1] 0.53 1.29 0.65 0.56 0.15 1.47 0.26 0.46 1.12 0.63 1.80 0.59 0.89 0.87 0.39 
## [16] 0.40 1.25 1.47 1.19 1.13 0.30 1.16 0.25 0.23 0.88 1.00 1.25 0.95 1.28 0.69 
## [31] 0.83 0.36 1.25 1.24 1.03 0.65 0.70 1.26 0.85 0.12 0.80 0.92 0.74 0.45 1.52 
## [46] 1.21 0.82 1.46 0.62 0.80 0.73
```

## Confidence Interval for $\mu$ , $1 - \alpha = 0.95$

We assume that  $\sigma = 0.4$ .

 $\left[\bar{X}_{(n)} - u_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{X}_{(n)} + u_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right]$ 

```
alpha <- 0.05

q <- qnorm(1-alpha/2,0,1)

mv <- mean(values)

lb1 <- mv-q*sigma/(n^0.5)

ub1 <- mv+q*sigma/(n^0.5)
```

- From n=51,  $u_{1-\alpha/2}=1.959964$ ,  $\bar{x}_{(n)}=0.8519608$  we get as the confidence interval for  $\mu$ : [0.7421808, 0.9617407]
- If we increase the confidence level to 99% we get a wider confidence interval

```
alpha <- 0.01
q <- qnorm(1-alpha/2,0,1)
mv <- mean(values)
1b2 <- mv-q*sigma/(n^0.5)
ub2 <- mv+q*sigma/(n^0.5)</pre>
```

- 99% confidence interval for  $\mu$ : [0.7076855, 0.9962361]
- The package TeachingDemos includes a function z.test which can be used to evaluate the confidence interval directly.

```
library(TeachingDemos)
```

## Warning: package 'TeachingDemos' was built under R version 4.0.5

```
z.test(x = values, stdev = sigma, alternative = "two.sided", conf.level = 0.95)$conf.int

## [1] 0.7421808 0.9617407

## attr(,"conf.level")

## [1] 0.95

z.test(x = values, stdev = sigma, alternative = "two.sided", conf.level = 0.99)$conf.int

## [1] 0.7076855 0.9962361

## attr(,"conf.level")

## [1] 0.99
```

#### $\sigma$ is unknown

•

$$\left[\bar{X}_{(n)} + t_{n-1,1-\alpha/2} \frac{S_{(n)}}{\sqrt{n}}, \bar{X}_{(n)} + t_{n-1,1-\alpha/2} \frac{S_{(n)}}{\sqrt{n}}\right]$$

```
qt <- qt(1-alpha/2,n-1)
mv <- mean(values)
s <- sd(values)
lb3 <- mv-qt*s/(n^0.5)
ub3 <- mv+qt*s/(n^0.5)</pre>
```

- From n = 51,  $t_{n-1,1-\alpha/2} = 2.6777933$ ,  $\bar{x}_{(n)} = 0.8519608$ ,  $s_{(n)} = 0.4034006$  we get as the confidence interval for  $\mu$ : [0.7006992, 1.0032223]
- Mention that the length of the interval is bigger as in the case if  $\sigma$  is known.
- With the function t.test the confidence interval can be evaluated directly.

```
t.test(x = values, alternative = "two.sided", conf.level = 1-alpha)$conf.int
## [1] 0.7006992 1.0032223
## attr(,"conf.level")
## [1] 0.99
```

## Confidence interval for $\sigma^2$ , $1 - \alpha = 0.95$

•  $\mu$  is known

$$\left[\frac{Q_n}{\chi_{n,1-\alpha/2}^2}, \frac{Q_n}{\chi_{n,\alpha/2}^2}\right]$$

```
mu <- 0.8
Qn <- sum((values - mu)^2)
qchi1 <-qchisq(alpha/2,n)
qchi2 <-qchisq(1-alpha/2,n)
lb3 <- Qn/qchi2
ub3 <- Qn/qchi1</pre>
```

- From n = 51,  $\chi^2_{n,1-\alpha/2} = 80.746659$ ,  $\chi^2_{n,\alpha/2} = 28.734712$ ,  $s_{(n)} = 0.4034006$  we get as the confidence interval for  $\sigma^2$ : [0.1024724, 0.2879549]
- Taking the square root of the bounds we get as a confidence interval for the standard deviation  $\sigma$ : [0.320113, 0.5366143]

•  $\mu$  is unknown

$$\left[\frac{(n-1)S_{(n)}^2}{\chi_{n-1,1-\alpha/2}^2}, \frac{(n-1)S_{(n)}^2}{\chi_{n-1,\alpha/2}^2}\right]$$

```
s <- sd(values)
qchi3 <-qchisq(alpha/2,n-1)
qchi4 <-qchisq(1-alpha/2,n-1)
lb4 <- (n-1)*s^2/qchi4
ub4 <- (n-1)*s^2/qchi3</pre>
```

- From  $n=51, \chi^2_{n-1,1-\alpha/2}=27.9907489, \chi^2_{n-1,\alpha/2}=79.4899785, s_{(n)}=0.4034006$  we get as the confidence interval for  $\sigma^2$ : [0.1023601, 0.290689]
- Taking the square root of the bounds we get as a confidence interval for the standard deviation  $\sigma$ : [0.3199377, 0.5391559]
- Applying the function sigma.test of the package TeachingDemos the confidence interval can be evaluated directly.

```
library(TeachingDemos)
sigma.test(x = values, alternative = "two.sided", conf.level = 1-alpha)$conf.int

## [1] 0.1023601 0.2906890
## attr(,"conf.level")
## [1] 0.99
```

### One sided confidence intervals

 $1-\alpha=0.95$  upper bound for  $\mu$  if  $\sigma$  is unknown

 $\bar{X}_{(n)} + t_{n-1,1-\alpha} \frac{S_{(n)}}{\sqrt{n}}$ 

- From n = 51,  $t_{n-1,1-\alpha} = 2.4032719$ ,  $\bar{x}_{(n)} = 0.8519608$ ,  $s_{(n)} = 0.4034006$  we get as the upper confidence bound for  $\mu$ : **0.9877153**
- Applying the t.test function we get directly

```
t.test(x = values, alternative = "less", conf.level = 1-alpha)$conf.int
```

```
## [1] -Inf 0.9877153
## attr(,"conf.level")
## [1] 0.99
```

 $1 - \alpha = 0.95$  lower bound for  $\sigma$  if  $\mu$  is unknown

 $\sqrt{\frac{(n-1)S_{(n)}^2}{\chi_{n-1,1-\alpha}^2}}$ 

- From n = 51,  $\chi^2_{n-1,1-\alpha} = 76.1538912$ ,  $s_{(n)} = 0.4034006$  we get as the lower confidence bound for  $\sigma$ : 0.3268704
- Applying the function sigma.test of the package TeachingDemos we get

```
library(TeachingDemos)
(sigma.test(x = values, alternative = "greater", conf.level = 1-alpha)$conf.int)^0.5
```

```
## [1] 0.3268704
                        Inf
## attr(,"conf.level")
## [1] 0.99
```

What is the minimum sample size so that the length of the  $1-\alpha=0.95$  confidence interval for  $\mu$  is 0.1 or less? für Normalverteilunger

 $\sigma = 0.4$  is known

• length of the confidence interval

$$l = 2u_{1-\alpha/2} \frac{\sigma}{\sqrt{n}} \le 0.1 \Rightarrow n \ge \left(\frac{2u_{1-\alpha/2}\sigma}{0.1}\right)^2$$

• From  $u_{1-\alpha/2}=1.959964$  we get as the lower bound for the sample size:  $n\geq 424.6333825$ 

 $\sigma$  is unknown

• length of the confidence interval

$$l = 2t_{n-1,1-\alpha/2} \frac{S_{(n)}}{\sqrt{n}} \le 0.1$$

• Mention that the length l depends on the sample values, i.e. it is a random variable. Since  $S_{(n)}$  is an unbiased estimator for  $\sigma$  and  $t_{n-1,1-\alpha/2}$  is decreasing with respect to n we get in the average for  $n \geq 51$ 

$$E(l) = 2t_{n-1,1-\alpha/2} \frac{\sigma}{\sqrt{n}} \le 2t_{51-1,1-\alpha/2} \frac{\sigma}{\sqrt{n}} \le 0.1 \Rightarrow n \ge \left(\frac{2t_{50,1-\alpha/2}\sigma}{0.1}\right)^2$$

• From  $t_{50,1-\alpha/2}=2.0085591$  we get in the average as the lower bound for the sample size:  $n\geq$ 458.9169153

## Confidence Interval for an unknown proportion

• Example: In a survey, 75 of the 500 respondents voted for Party XY.

### Two sided confidence interval

• approximative confidence interval for the unknown proportion p

$$\left[ \hat{p} - u_{1-\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}, \hat{p} + u_{1-\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \right]$$

- $1-\alpha=0.95$ : From n=500,  $\hat{p}=\frac{75}{500}=0.15, u_{1-\alpha/2}=1.959964$  we get as the confidence for the unknown proportion [0.1187019, 0.1812981]
- Checking the rule of thumb:  $n\hat{p} \ge 10$  and  $n(1-\hat{p}) \ge 10$  we see that the approximative calculation is ok.
- Applying the function binom.test() we can evaluate the exact confidence interval.

```
binom.test(x=75, n=500, alternative = "two.sided", conf.level = 1-alpha)$conf.int
```

## [1] 0.1198486 0.1843596

## attr(,"conf.level")

## [1] 0.95

### One sided confidence interval

• approximative upper confidence bound for the unknown proportion p

$$\hat{p} + u_{1-\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

- $1 \alpha = 0.95$ : From n=500,  $\hat{p} = \frac{75}{500} = 0.15$ ,  $u_{1-\alpha} = 1.6448536$  we get as the upper confidence bound for the unknown proportion **0.1812981**]
- Applying the function binom.test() we can evaluate the exact confidence interval.

binom.test(x=75, n=500, alternative = "less", conf.level = 1-alpha)\$conf.int

## [1] 0.0000000 0.1788032 ## attr(,"conf.level") ## [1] 0.95

What is the minimum sample size so that the length of the  $1-\alpha=0.95$  confidence unknown proportion interval for p is 0.05 or less?

• length of the confidence interval

$$l = 2u_{1-\alpha/2}\sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \le 0.05$$

• Mention that the length l depends on the sample values, i.e. it is a random variable. Since we have that for every  $p \in [0,1] : p(1-p) \le 0.25$  we get

$$l \le 2u_{1-\alpha/2}\sqrt{\frac{0.25}{n}} \le 0.05 \Rightarrow n \ge \left(\frac{u_{1-\alpha/2}}{0.05}\right)^2$$

- From  $u_{1-\alpha/2}=1.959964$  we get that if  $n\geq 1536.5835283$  the length of the confidence interval will be less than 0.05.
- We have a better estimation if we know an upper bound for p. If for example  $p \leq 0.2$  we get

$$l \le \left(\frac{2 * u_{1-\alpha/2}\sqrt{0.2 \cdot 0.8}}{0.05}\right)^2.$$

• Thus we have  $n \ge 983.4134581$  if  $p \le 0.2$ .