

a) ~~not~~

$$H = - \int \exp(s) \log \exp(s) dx = - \int \exp(s) s dx$$

$$\Lambda = - \int \exp(s) s dx - \lambda_1 \left(\int \exp(s) dx - 1 \right)$$

$$- \lambda_2 \int x \exp(s) dx$$

$$- \lambda_3 \left[\int x^2 \exp(s) dx - \int x \exp(s) dx - C^2 \right]$$

$$b) \frac{\partial \Lambda}{\partial s} = - \exp(s) s - \exp(s)$$

$$- \lambda_1 \exp(s) - \lambda_2 x \exp(s) - \lambda_3 [x^2 \exp(s) - x \exp(s)] =$$

$$= 0$$

↗
quadratic in x

$$(c) \Delta = - \int p(x) \ln p(x) dx + \lambda_1 \left(\int p(x) dx - 1 \right) + \lambda_2 \int p(x) \cdot x dx \\ + \lambda_3 \left(\int p(x) \cdot x^2 dx - c^2 \right)$$

$$\frac{\partial \Delta}{\partial p(x)} = \int (\ln p(x) + \frac{p(x)}{p(x)} + \lambda_1 + \lambda_2 x + \lambda_3 x^2) dx = 0.$$

$$(\ln p(x) + 1 + \lambda_1 + \lambda_2 x + \lambda_3 x^2) dx = 0.$$

$$p(x) = e^{-1 - \lambda_1 - \lambda_2 x - \lambda_3 x^2}$$

To show $p(x)$ is Gaussian distributed

choose $\lambda_1, \lambda_2, \lambda_3$ as

plugging these values

$$\lambda_1 = -1 - \ln \frac{1}{\sqrt{2\pi} c^{-2}}$$

in $p(x)$ we get $x^2/2c^2$

$$\lambda_2 = 0$$

$$p(x) = \frac{1}{\sqrt{2\pi c^2}} e^{-x^2/2c^2}$$

$$\lambda_3 = \frac{1}{2c^2}$$

then

$$\begin{aligned} \frac{\partial \Delta}{\partial p(x)} &= \ln(p(x)) + 1 + \lambda_1 + \lambda_2 x + \lambda_3 x^2 \\ &= \ln \left(e^{-1 - \lambda_1 - \lambda_2 x - \lambda_3 x^2} \right) + 1 + \lambda_1 + \lambda_2 x + \lambda_3 x^2 \\ &= \ln \left(e^{-1 + 1 + \ln \frac{1}{\sqrt{2\pi} c^{-2}}} + 0 - \frac{x^2}{2c^2} \right) + 1 \left(-1 - \ln \frac{1}{\sqrt{2\pi} c^{-2}} \right. \\ &\quad \left. + 0 + \frac{x^2}{2c^2} \right) \end{aligned}$$

$$= \ln \left(\frac{1}{\sqrt{2\pi} c^{-2}} \right) - \frac{x^2}{2c^2} - \frac{\ln}{\sqrt{2\pi} c^{-2}} + \frac{x^2}{2c^2} = 0$$

\therefore we can show $p(x)$ is gaussian distributed.

If ~~f~~ is random variable with pdf $p(x)$

has zero mean and variance σ^2 , then

[and $y(x) \sim N(0, \sigma^2)$]

$$-\int p(x) \log y(x) dx$$

$$= -\int p(x) \left(\log \frac{1}{\sqrt{2\pi\sigma^2}} - \frac{x^2}{2\sigma^2} \right) dx$$

$$= \frac{1}{2} \log (2\pi\sigma^2) + \frac{E_p[x^2]}{2\sigma^2} = \frac{1}{2} \log (2\pi e\sigma^2)$$

Now it is also true that if X has pdf $p(x)$ with zero mean & variance σ^2 , then

$$\text{H}(X) = \int p(x) \log y(x) dx - \underbrace{\int p(x) \log p(x) dx}_{H(x)}$$

$$= \int p(x) \log \frac{y(x)}{p(x)} dx \geq \log \int p(x) \frac{y(x)}{p(x)} dx$$

Jensen's inequality,
log is convex

$$= \log \int y(x) dx = 0$$

$$\Rightarrow H(X) \leq \int p(x) \log y(x) dx = \frac{1}{2} \log(2\pi e \sigma^2)$$

We just showed $\boxed{\text{ }}$ that if $x^* \sim N(0, \sigma^2)$

then $H(x^*) = \frac{1}{2} \log(2\pi e \sigma^2)$.

We also showed that for a random var.
with zero mean and variance σ^2 it holds that

$$H(x) \leq \frac{1}{2} \log(2\pi e \sigma^2).$$

Therefore $f(x) = H(x^*) - H(x) \geq 0$

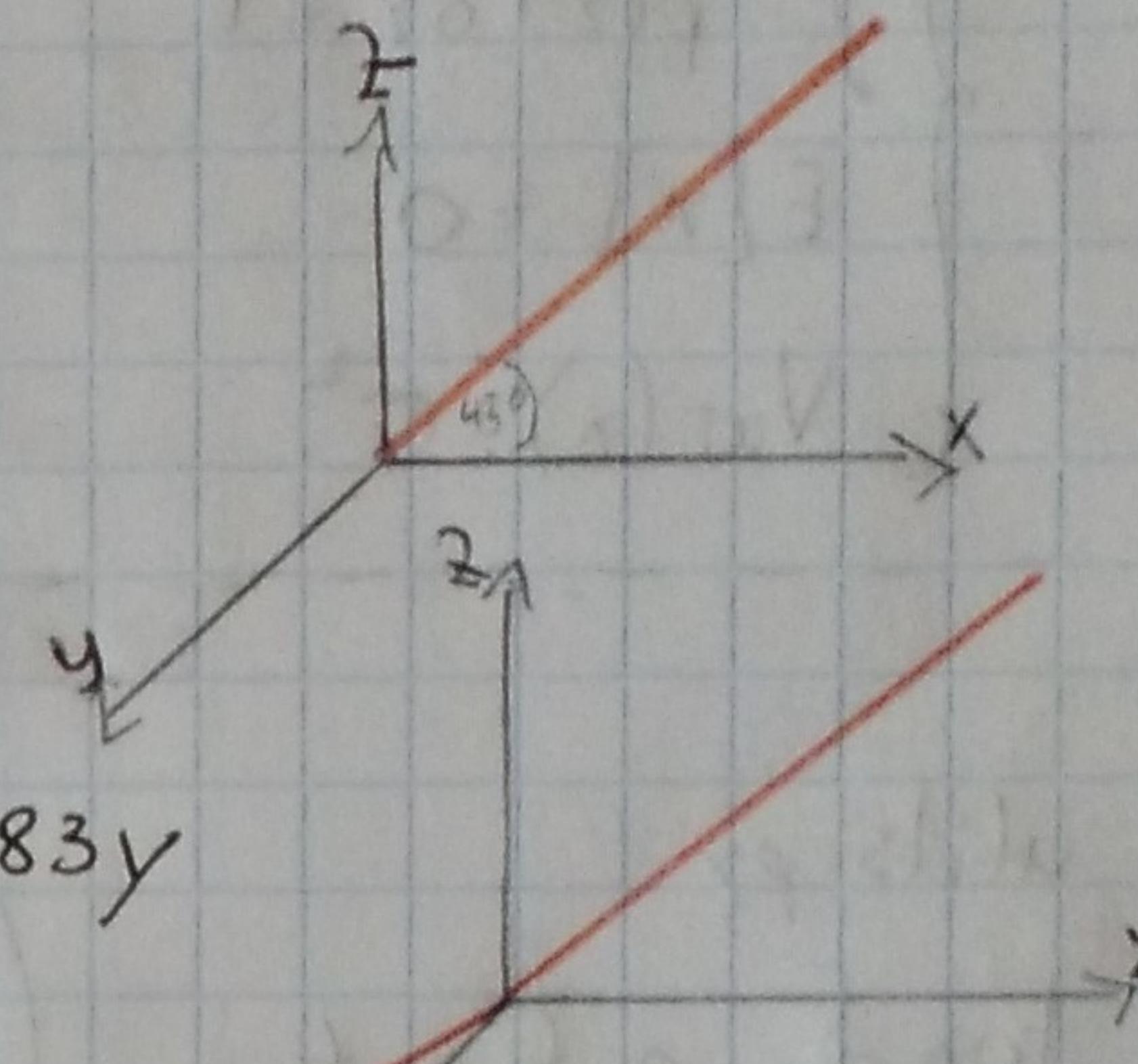
and equal if x Gaussian.

Exercise 2

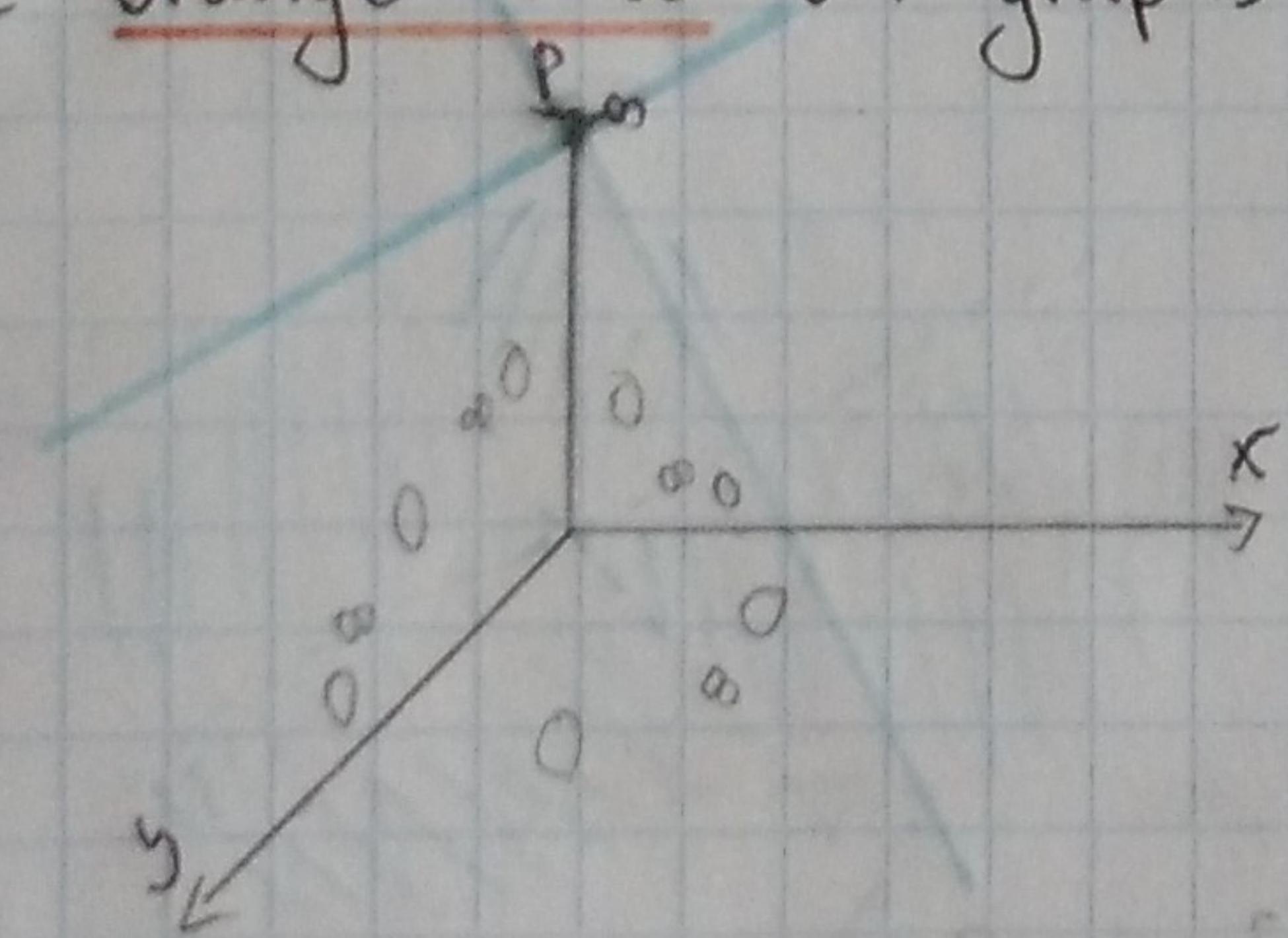
a) $\Theta = 0 \rightarrow \begin{pmatrix} 1 \\ 0 \end{pmatrix} \quad z(0) = x$

$$\Theta = \frac{\pi}{8} \rightarrow \begin{pmatrix} 0,924 \\ 0,383 \end{pmatrix} \quad z\left(\frac{\pi}{8}\right) = 0,924x + 0,383y$$

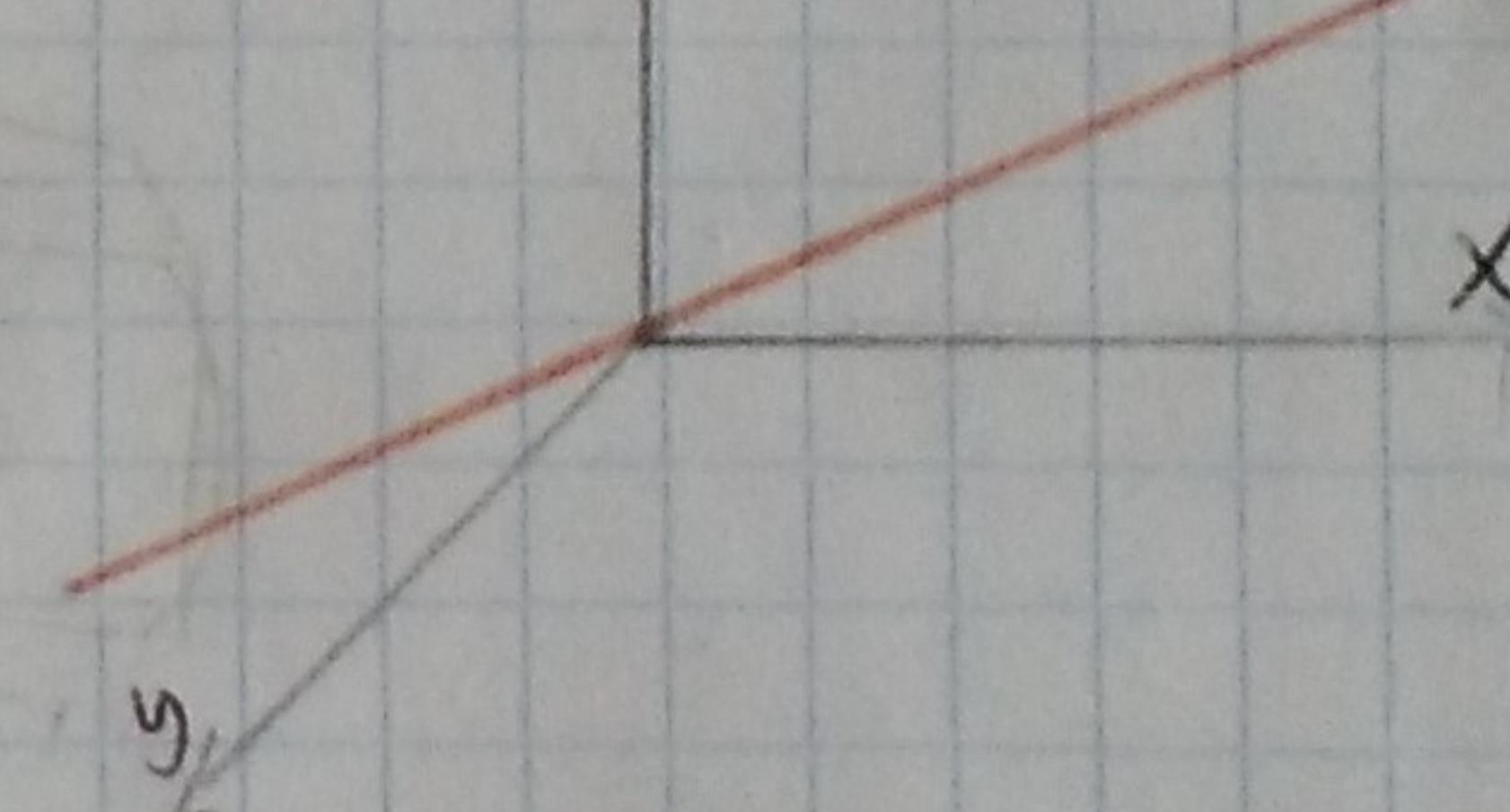
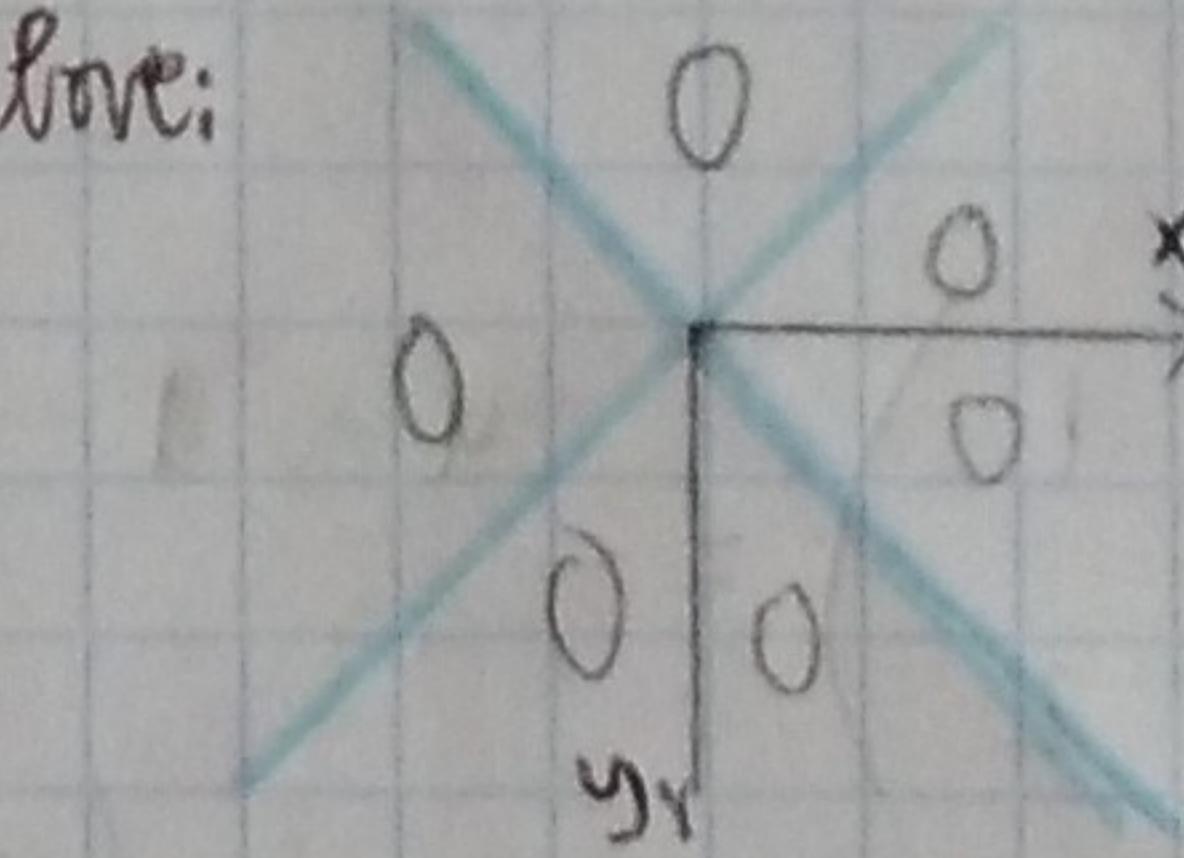
$$\Theta = \frac{\pi}{4} \rightarrow \begin{pmatrix} 0,707 \\ 0,707 \end{pmatrix} \quad z\left(\frac{\pi}{4}\right) = 0,707(x+y)$$



The orange lines on graphs denote planes



From above:



The blue lines denote infinity. On the z axis joint probability $p(x,y)$

$$b) \text{Var}(z(\theta)) = E[z^2(\theta)] - E[z(\theta)]^2$$

$$\begin{aligned} \text{Var}(z(\theta)) &= E[(x\cos\theta + y\sin\theta)^2] - E[x\cos\theta + y\sin\theta]^2 = \\ &= \underbrace{E[x^2\cos^2\theta]}_{(*)} + E[2xy\cos\theta\sin\theta] + E[y^2\sin^2\theta] - \underbrace{(E[x\cos\theta] + E[y\sin\theta])^2}_{(*)} = \end{aligned}$$

$$\begin{aligned} &= 2\cos\theta\sin\theta E[xy] + \sin^2\theta E[y^2] - (E[y\sin\theta])^2 = \\ &= 2\cos\theta\sin\theta E[xy] + \sin^2\theta \text{Var}(y) \end{aligned}$$

Let $\cos^2\theta, \cos\theta$ be const, then $(*)$ is 0 as $x \sim N(0, 1)$

$$\frac{\partial \text{Var}(z(\theta))}{\partial \theta} = -2E[xy]\sin^2\theta + 2E[xy]\cos^2\theta + \sin^2\theta \text{Var}(y) = 0$$

$$\sin^2\theta \text{Var}(y) = 2E[xy](\sin^2\theta - \cos^2\theta)$$

$$\frac{\sin 2\theta}{4\sin^2\theta - 2} = \frac{E[xy]}{\text{Var}(y)}$$

$$\text{Var}[z(\theta)] = \frac{\mathbb{E}[(z(\theta) - \mathbb{E}[z(\theta)])^4]}{(\text{Var}[z(\theta)])^2} - 3$$

$\text{Var}[z(\theta)]$ from part b: $\text{Var}[z(\theta)] = 2\cos(\theta)\sin(\theta)\mathbb{E}[xy] + \sin^2(\theta)\text{Var}(y)$

 $(\text{Var}[z(\theta)])^2 = 4\cos^2(\theta)\sin^2(\theta)\mathbb{E}[xy]^2 + 4\cos(\theta)\sin^3(\theta)\mathbb{E}[xy]\text{Var}(y) + \sin^4(\theta)\text{Var}(y)^2$

$$\mathbb{E}[(z(\theta) - \mathbb{E}[z(\theta)])^4] = \mathbb{E}[(x\cos(\theta) + y\sin(\theta) - \mathbb{E}[x\cos(\theta) + y\sin(\theta)])^4] \quad \text{with } \mathbb{E}[x\cos(\theta)] = 0$$
 $= \mathbb{E}[(x\cos(\theta) + y\sin(\theta) - \mathbb{E}[y])^4] \quad \text{with } \mathbb{E}[y] = y$
 $= \mathbb{E}[(x\cos(\theta))^4]$
 $= \cos^4(\theta)\mathbb{E}[x^4]$

$$\frac{\partial}{\partial \theta} \frac{\cos^4(\theta)\mathbb{E}[x^4]}{4\cos^2(\theta)\sin^2(\theta)\mathbb{E}[xy]^2 + 4\cos(\theta)\sin^3(\theta)\mathbb{E}[xy]\text{Var}(y) + \sin^4(\theta)\text{Var}(y)^2} - 3$$

s.t. $\theta \in [0, 2\pi]$

$$= -4\mathbb{E}[x^4]\sin(\theta)\cos^3(\theta) \cdot (2\cos(\theta)\sin(\theta)\mathbb{E}[xy] + \sin^2(\theta)\text{Var}(y))^2 - \cos^4(\theta)\mathbb{E}[x^4] \cdot (2(2\cos(\theta)\sin(\theta)\mathbb{E}[xy] + \sin^2(\theta)\text{Var}(y)) \\ + 2\sin^2(\theta)\mathbb{E}[xy]) \cdot (-2\sin^2(\theta)\mathbb{E}[xy] + 2\cos^2(\theta)\mathbb{E}[xy] + \sin(2\theta)\text{Var}(y)))$$

$$(2\cos(\theta)\sin(\theta)\mathbb{E}[xy] + \sin^2(\theta)\text{Var}(y))^3$$

$\stackrel{!}{=} 0$