

- 17 Show that the distribution

$$F(z_1, z_2) = \exp \left[- \left\{ z_1^{-1} + z_2^{-1} + (z_1 z_2)^{-1} \right\} \right], \quad z_1, z_2 > 0,$$

has unit Fréchet margins. Is it max-stable? If not, what is the limiting distribution of appropriately rescaled componentwise maxima from F ?

Setting $z_1 = \infty$ shows that the margin for z_2 is unit Fréchet, and the result for z_1 follows by symmetry. As $t \rightarrow \infty$,

$$F(tz_1, tz_2)^t = \exp \left[- \left\{ z_1^{-1} + z_2^{-1} + (tz_1 z_2)^{-1} \right\} \right] \rightarrow \exp \left[- \left\{ z_1^{-1} + z_2^{-1} \right\} \right],$$

so it is not max-stable, and the limiting distribution for rescaled maxima is independence.

- 18 Derive (5.28) by noting that if the intensity function $\dot{\mu}$ exists, then

$$V(0) = V(z) + \int_0^{z_1} \cdots \int_0^{z_D} \dot{\mu}(x_1, \dots, x_D) dx_1 \cdots dx_D.$$

If this argument makes you queasy, find a better one.

As $V(0) = \mu(\mathcal{E})$ is constant, differentiation by z_{C+1}, \dots, z_D yields

$$0 = \frac{\partial^{D-C}}{\partial z_{C+1} \cdots \partial z_D} V(z_1, \dots, z_D) + \int_0^{z_1} \cdots \int_0^{z_C} \dot{\mu}(x_1, \dots, x_C, z_{C+1}, \dots, z_D) dx_1 \cdots dx_C,$$

which gives the required result on setting $z_1 = u_1, \dots, z_C = u_C$. Queasiness arises because the left-hand side and the integral on the right-hand side both equal infinity. Alternatively one can integrate the equality

$$\dot{\mu}(z_1, \dots, z_D) = - \frac{\partial^D V(z_1, \dots, z_D)}{\partial z_1 \cdots \partial z_D}$$

over the set $\{(z_1, \dots, z_D) : 0 \leq z_d \leq u_d, d = 1, \dots, C\}$.

- 19 (a) Check the argument below (5.29), noting that the constraints $\sum_d w_d^* = \sum_d w_d = 1$ imply that

$$w_d^* = \frac{m_d w_d}{m_D + \sum_{c=1}^{D-1} (m_c - m_D) w_c}, \quad d = 1, \dots, D-1.$$

and the Jacobian is

$$\frac{\partial w^*}{\partial w} = \frac{\partial(w_1^*, \dots, w_{D-1}^*)}{\partial(w_1, \dots, w_{D-1})}.$$

You may require the matrix determinant lemma $|A + ab^T| = |A|(1 + b^T A^{-1} a)$, where A is an invertible $p \times p$ matrix and a and b are $p \times 1$ vectors.

(b) Show that $\max_d m_d \geq m^T w \geq \min_d m_d$ for all $w \in \mathcal{S}_{D-1}$, and give an algorithm for simulation from $\dot{\nu}$ by noting that

$$\dot{\nu}(w) \propto \Pr(U \leq \min_d m_d / m^T w \mid W = w) \times f(w),$$

where $f(w)$ is the density of the variable W defined in (5.29) and $U \sim U(0, 1)$ is independent of W . What is the efficiency of your algorithm?

(Coles and Tawn, 1991, Theorem 2)

(a) If we write $w_d^* = m_d w_d / m^T w$, where $m^T w = \sum_{c=1}^{D-1} (m_c - m_D) w_c$, then

$$\begin{aligned} \partial w_d^* / \partial w_d &= m_d / (m^T w) - w_d (m_d - m_D) / (m^T w)^2, \\ \partial w_d^* / \partial w_c &= m_d w_d (m_c - m_D) / (m^T w)^2, \quad c \neq d. \end{aligned}$$

Hence the Jacobian is the determinant of a matrix of the form $A + ab^T$, where

$$\begin{aligned} A &= \text{diag}(m_1, \dots, m_{D-1}) / (m^T w), \\ a &= (m_1 w_1, \dots, m_{D-1} w_{D-1})^T, \\ b &= -(m_1 - m_D, \dots, m_{D-1} - m_D)^T / (m^T w)^2. \end{aligned}$$

Now $|A| = (m^T w)^{-(D-1)} \prod_{d=1}^{D-1} m_d$ and a little algebra gives $1 + b^T A^{-1} a = m_D / (m^T w)$, so the matrix determinant lemma gives $|\partial w^* / \partial w| = (m^T w)^{-D} \prod_d m_d$. The probability density function for W is obtained on rewriting w^* in terms of w , and the rest of the argument is straightforward.

(b) As $m^T w$ is a linear function of $w \in \mathcal{S}_{D-1}$, the inequalities follow at once. Let $m_{\min} = \min_d m_d$. A realisation of W is obtained by generating W^* from ν^* and then applying the transformation (5.29), but this has density f rather than $\dot{\nu}$. Now $m_{\min} / m^T w \leq 1$, so conditional on $W = w$, the event $U \leq m_{\min} / m^T w$ has probability $m_{\min} / m^T w$, and thus the marginal density of those W for which this event occurs is proportional to $f(w) / (m^T w)$ and must therefore be $\dot{\nu}$. The efficiency of the algorithm is the probability that a generated W is accepted, and this equals

$$\int \frac{m_{\min}}{m^T w} f(w) dw = D m_{\min} \int \dot{\nu}(w) dw = D m_{\min},$$

which is greatest when all the m_d are equal to $1/D$, as might be expected; then $m^T w = 1/D$ and hence $\dot{\nu} = \dot{\nu}^*$. The more the m_d vary, the less efficient the algorithm becomes.

- 20 Suppose that $\varepsilon_1, \varepsilon_2, \varepsilon_3$ are independent unit Fréchet variables, and let

$$Z_1 = \max\{\alpha \varepsilon_1, (1 - \alpha) \varepsilon_2\}, \quad Z_2 = \max\{\beta \varepsilon_1, (1 - \beta) \varepsilon_3\}, \quad 0 \leq \alpha, \beta \leq 1.$$

- (a) Show that $\Pr(Z_1 \leq z_1, Z_2 \leq z_2)$ equals

$$\exp\{-(1 - \alpha)/z_1 - (1 - \beta)/z_2 - \max(\alpha/z_1, \beta/z_2)\}. \quad z_1, z_2 > 0,$$

and verify that the marginal distributions of Z_1 and Z_2 are unit Fréchet.

- (b) Deduce that this joint distribution is bivariate extreme-value with discrete angular measure

$$\nu(\{0\}) = (1 - \beta)/2, \quad \nu(\{1\}) = (1 - \alpha)/2, \quad \nu(\{\alpha/(\alpha + \beta)\}) = (\alpha + \beta)/2,$$

and check that this has mean $1/2$ for all choices of α and β .

- (c) Compute the extremal coefficient for this distribution, and discuss how it depends on the parameter values.

- (d) Sketch the angular measure ν and investigate special cases corresponding to independence and complete dependence. Sketch typical datasets of values (z_1, z_2) corresponding to different possible choices of the parameters.

- 21 The multivariate asymmetric logistic extreme-value model has exponent function

$$V(z_1, \dots, z_D) = \sum_{d=1}^D \frac{1 - \psi_d}{z_d} + \sum_{d=1}^D \left\{ \left(\frac{\psi_d}{z_d} \right)^{1/\alpha} \right\}^\alpha, \quad z_1, \dots, z_D > 0,$$

where $\psi_1, \dots, \psi_D, \alpha \in (0, 1]$. Investigate how the placement of point masses and densities on corners and edges of \mathcal{S}_{D-1} depends on the parameters, in the case $D = 3$.