

# The Impact of Audience Size on Image Concerns

## Evidence from a Charity Dictator Game

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We care deeply about others' opinion of us (our *image*)

- People have *strategic* reasons to care about their image
- Most also have an *intrinsic* preference for being regarded as “good”

The number of people with a given opinion of us could also matter

- Social media made our life “more observable”
- Audiences have grown larger but more anonymous

- We address the following research questions:
  - ① Can observation from of an anonymous, external, and noninteractive audience motivate prosocial behavior?
  - ② Are audience effects monotonic in the number of onlookers?
- Conflicting answers to ① are influenced by:
  - “Theoretical and Methodological moderators” (Bradley et al. 2018)
  - Active or communicative audiences with stake in the game, repeated games with roles rotation (Regner 2021), ambiguous setting (Attanasi et al. 2023)
  - As a result, these often study different social phenomena

We want to isolate the *psychological, intrinsic* valence of image concerns from the *strategic one* in an unambiguous context

**Image Concerns:** Preferences over others' ex-post beliefs over own private and heterogeneously distributed traits which are imperfectly observed by others (c.f. Battigalli and Dufwenberg 2022)

- Two-players case (Battigalli and Dufwenberg 2022)

$$u_i(z, \alpha_j; \theta_i^{\mathbf{I}}, \theta_i^{\mathbf{R}}) = \pi_i(z) + \theta_i^{\mathbf{I}} [1_i^G(z) - 1_i^B(z)] + \theta_i^{\mathbf{R}} \mathbb{E} [\tilde{\theta}_i^I | z; \alpha_j] \quad (1)$$

- $\theta_i^{\mathbf{I}} \in \mathbb{R}_+$  represents  $i$ 's “Good” trait
- $[1_i^G(z) - 1_i^B(z)]$  qualifies the net “Goodness” of  $i$ 's behaviour
- $\mathbb{E} [\tilde{\theta}_i^I | z; \alpha_j]$  is loosely referred to as  $j$ 's *opinion* about  $i$ 's trait
- $\theta_i^{\mathbf{R}} \in \mathbb{R}_+$  measures how much  $i$  cares about  $j$ 's ex-post estimate of  $\theta_i^{\mathbf{I}}$

# Charity Dictator mini-Game with External Audience

- Two Actions:
  - ① **Donate**  $\rightarrow$  Unambiguously and universally perceived as “**Good**”
  - ② **Don’t donate**  $\rightarrow$  Unambiguously and universally “**Bad**”
- Perfect but **incomplete** information
  - When observed, the choice becomes a signal for the “good” trait
- Additive utility in the opinion of others:

$$u_i(z, \alpha_{-i}; \theta_i^{\mathbf{I}}, \bar{\theta}_i^{\mathbf{R}}) = \pi_i(z) + \theta_i^{\mathbf{I}} [1_i^G(z) - 1_i^B(z)] + \sum_{\{j \neq i\} \in N} \theta_{ij}^{\mathbf{R}} \mathbb{E} [\tilde{\theta}_i^I | z; \alpha_j] \quad (2)$$

- Auxiliary assumption: according to  $i$ , for all  $j \in N \setminus \{i\}$  ,

$$\mathbb{E}[\tilde{\theta}_i^I | z = G; \alpha_j] > \mathbb{E}[\tilde{\theta}_i^I | z = B; \alpha_j]$$

- When observed, even *selfish types may choose to donate* if they are sufficiently worried about others' opinion of them
  - The incentive to donate is *increasing* in the size of the audience
- $\implies$  **Testable predictions**

- H1:** The share of individuals choosing the “Good” action is larger when an Audience is present.
- H2:** The share of individuals choosing the “Good” action is monotonically increasing in the number of Audience members

## Charity Dictator mini-Game with External Audience

- Active players (“Dictators”) can make a **binary choice**
  - donate half of a 10 EUR endowment to a charity (the recipient)
  - or donate only 1 EUR
- **An anonymous passive audience observes** the choice outcome, alongside a random ID for each Dictator
- The monetary payoff of the Audience is fixed at 7 EUR and is the same in all feasible terminal histories.

► See User Interface

# Treatments and Procedures

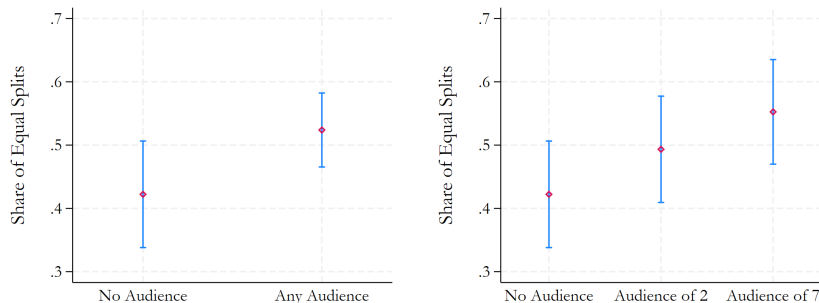
- **Treatments** (Between-subjects design):
  - **Baseline:** No Audience
  - **T1:** Audience of 2
  - **T2:** Audience of 7
- **Procedures**
  - **5 rounds** under a **random stranger** matching protocol with **charitable organizations**.
  - Audience observed all choice outcomes **after every round**
  - One round randomly selected for payments
  - The number of observers choices was common knowledge

► More on Procedures



# Preliminary Results

Figure 1: Average Share of Equal Splits



Note: 95% CIs computed using individual averages across all rounds (one data point per participant).

► More preliminary results

**But:** TE attenuated by a relevant gender unbalance

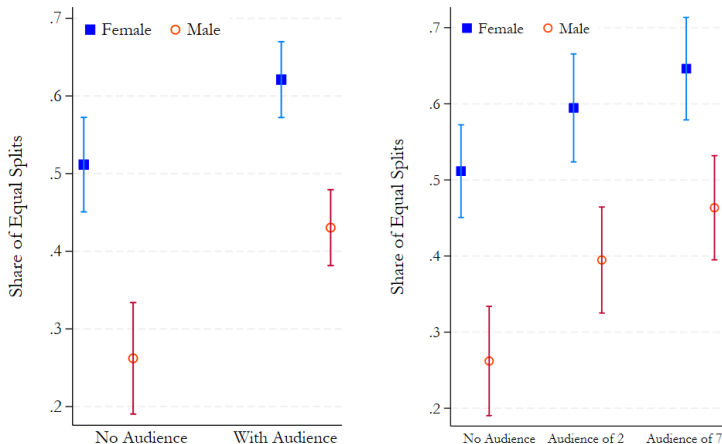
Table 1: Descriptive statistics and Balance by Treatment. Active Players Only.

Variable	(1) No Audience	(2) Audience of 2	(3) Audience of 7	(4) Difference (2-0)	(5) Difference (7-0)	(6) Difference (7-2)
Age	20.741 (1.571)	20.720 (1.849)	20.663 (1.534)	-0.021 (0.274)	-0.078 (0.245)	-0.058 (0.272)
Male	0.358 (0.482)	0.507 (0.503)	0.512 (0.503)	0.149* (0.079)	0.154** (0.078)	0.006 (0.081)
# of Experiments	2.136 (2.042)	2.813 (3.502)	2.987 (3.087)	0.678 (0.455)	0.852** (0.412)	0.174 (0.529)
Econ or Finance	0.481 (0.503)	0.507 (0.503)	0.312 (0.466)	0.025 (0.081)	-0.169** (0.076)	-0.194** (0.078)
Politics	0.062 (0.242)	0.040 (0.197)	0.075 (0.265)	-0.022 (0.036)	0.013 (0.040)	0.035 (0.038)
Law	0.074 (0.264)	0.067 (0.251)	0.100 (0.302)	-0.007 (0.041)	0.026 (0.045)	0.033 (0.045)
Managment	0.358 (0.482)	0.320 (0.470)	0.387 (0.490)	-0.038 (0.076)	0.029 (0.077)	0.068 (0.077)
Stats or CompSci	0.012 (0.111)	0.027 (0.162)	0.075 (0.265)	0.014 (0.022)	0.063* (0.032)	0.048 (0.036)
Other	0.012 (0.111)	0.040 (0.197)	0.050 (0.219)	0.028 (0.025)	0.038 (0.027)	0.010 (0.034)
Year of Study	2.593 (1.367)	2.573 (1.397)	2.513 (1.414)	-0.019 (0.221)	-0.080 (0.219)	-0.061 (0.226)
Observations	81	75	80	156	161	155

Note: Standard deviations/errors in parenthesis. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

# Gender Heterogeneity

Figure 2: Average Share of Equal Splits by Gender



Note: 95% CIs computed using individual averages across all rounds (one data point per participant).

# Regression Analysis

We estimate the following linear random effects model:

$$\text{Choice}_{i,t} = \beta(\text{Audience Size})_{i,t} + \alpha_i + \gamma X_{i,t} + \varepsilon_{i,t}$$

Table 2: GLS Random-Effects Estimator

Dependent Variable	(1)	(2)	(3)	(4)	(5)
	Share of Equal Splits				
Audience Size	0.0174** (0.00818)	0.0211*** (0.00785)	0.0199** (0.00788)	0.0200** (0.00795)	0.0200** (0.00796)
Male		-0.205*** (0.0465)	-0.205*** (0.0463)	-0.203*** (0.0479)	-0.203*** (0.0480)
# of Experiments Control	✗	✗	✓	✓	✓
Year of Study Controls	✗	✗	✗	✓	✓
Field of Study Controls	✗	✗	✗	✓	✓
Round FE	✗	✗	✗	✗	✓
Observations	1,180	1,180	1,180	1,180	1,180
# of individuals	236	236	236	236	236
Avg. Share in Control	0.4222	0.4222	0.4222	0.4222	0.4222

Note: Robust standard errors clustered at the individual level in parentheses. \*\*\*  
p<0.01, \*\* p<0.05, \* p<0.1

By employing a simple design we found:

- Robust evidence that anonymous, external, and noninteractive observers can motivate prosocial behavior
  - 31% increase in “fair” donations when an audience is present
- Suggestive evidence that this effect may be monotonic in the number of observers
  - 5% average increase for each observer  $\Rightarrow$  main novelty!
- Showed promising research topics in the application of Psychological Game Theory in applied work

Thank You!

Questions?

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- Attanasi, Giuseppe, Roberta Dessi, Frederic Moisan, and Donald Robertson**, “Public Goods and Future Audiences,” *Available at SSRN 4471977*, 2023.
- Battigalli, Pierpaolo and Martin Dufwenberg**, “Belief-dependent motivations and psychological game theory,” *Journal of Economic Literature*, 2022, 60 (3), 833–882.
- Bradley, Alex, Claire Lawrence, and Eamonn Ferguson**, “Does observability affect prosociality?,” *Proceedings of the Royal Society B: Biological Sciences*, 2018, 285 (1875), 20180116.
- Regner, Tobias**, “What’s Behind Image? Toward a Better Understanding of Image-Driven Behavior,” *Frontiers in psychology*, 2021, 12, 614575.

## Procedures

- The experiment took place in April and October 2023 at BELSS
- Participants were recruited through BELSS's SONA
- 6 (September) + 4 (October) sessions of 27 participants
- Total of **270 participants**, 236 Active and 34 Audience
- Five, minor and less-known Italian **charitable organizations**
- Random allocation to computer cubicles (and roles)
- Instructions and charities descriptions read aloud and displayed
- Decision-making phase followed by a short survey on basic demographics

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Figure 3: Decision Interface, No Audience

## Fase Decisionale

### Round 1

Hai ricevuto 10€, puoi ripartire questa cifra tra te e un'associazione senza scopo di lucro selezionata casualmente tra la lista presentata in precedenza.

Puoi scegliere tra due possibili allocazioni:

- A. Allocare 9€ a te stesso e 1€ all'associazione benefica;
- B. Allocare 5€ a te stesso e 5€ all'associazione benefica.

Choice

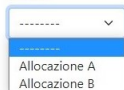




Figure 4: Decision Interface, Audience of 2

## Fase Decisionale

### Round 1

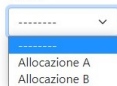
Hai ricevuto 10€, puoi ripartire questa cifra tra te e un'associazione senza scopo di lucro selezionata casualmente tra la lista presentata in precedenza.

Puoi scegliere tra due possibili allocazioni:

- A. Allocare 9€ a te stesso e 1€ all'associazione benefica;
- B. Allocare 5€ a te stesso e 5€ all'associazione benefica.

La tua scelta verrà osservata da 2 spettatori.

Choice

A dropdown menu with a white background and a thin grey border. The top part is a header bar with a blue gradient and a white downward arrow. Below the header, the menu is open, showing two options: 'Allocazione A' and 'Allocazione B', both in black text. The first option is highlighted with a blue background.

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- **Audience Effects**

- +10.1 pp (+24.0%) “fair” splits when an audience wrt baseline
- Two-sided t-test significant at 5% (p-value 0.0478)

- Comparing treatments:

- **Audience of 2 vs Baseline**

- +7.1 pp (+16.8%) “fair” splits
- Two-sided t-test not significant (p-value 0.2364)

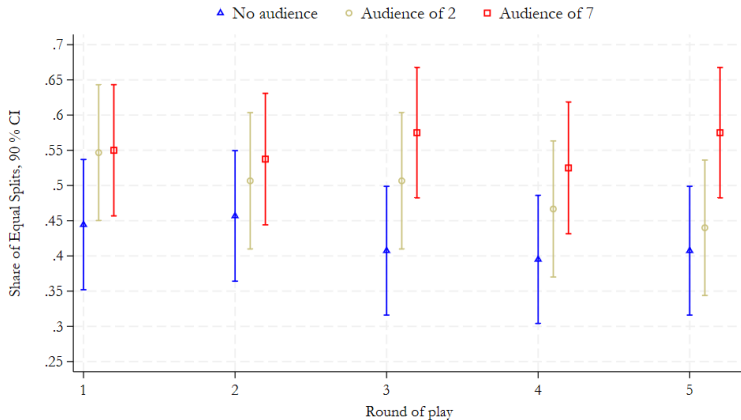
- **Audience of 7 vs Baseline**

- +13.0 pp (+30.9%) “fair” splits
- Two-sided t-test significant at 5% (p-value 0.0295)

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# Preliminary Results (III)

Figure 5: Average Share of Equal Splits by Round of Play

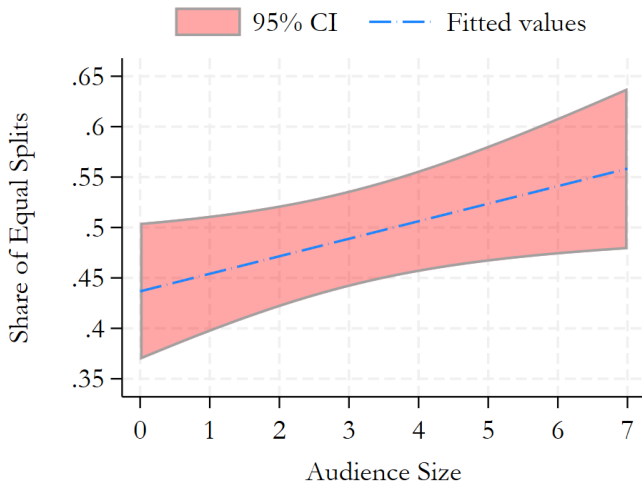


Note: 90% Confidence Intervals, average share of equal splits by treatment assignment.

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# Linear Fit in the Size of the Audience

Figure 6: Linear Fit in the Number of Observers



# Audience Size as a categorical variable

**Table 3:** GLS Random-Effects Estimator, Categorical Treatment

Dependent Variable	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	Share of Equal Splits									
Audience (any size)	0.102** (0.0515)	0.133*** (0.0490)	0.124** (0.0496)	0.131*** (0.0492)	0.131*** (0.0493)					
Audience of 2						0.0711 (0.0595)	0.102* (0.0568)	0.0945* (0.0569)	0.106* (0.0565)	0.106* (0.0566)
Audience of 7						0.130** (0.0591)	0.163*** (0.0565)	0.153*** (0.0569)	0.156*** (0.0572)	0.156*** (0.0573)
Male		-0.210*** (0.0464)	-0.209*** (0.0462)	-0.209*** (0.0479)	-0.209*** (0.0480)		-0.210*** (0.0463)	-0.209*** (0.0461)	-0.209*** (0.0478)	-0.209*** (0.0478)
# of Experiments	✗	✗	✓	✓	✓	✗	✗	✓	✓	✓
Year of Study	✗	✗	✗	✓	✓	✗	✗	✗	✓	✓
Field of Study	✗	✗	✗	✓	✓	✗	✗	✗	✓	✓
Round FE	✗	✗	✗	✗	✓	✗	✗	✗	✗	✓
Observations	1,180	1,180	1,180	1,180	1,180	1,180	1,180	1,180	1,180	1,180
# of individuals	236	236	236	236	236	236	236	236	236	236
Avg. Share in Control	0.4222	0.4222	0.4222	0.4222	0.4222	0.4222	0.4222	0.4222	0.4222	0.4222

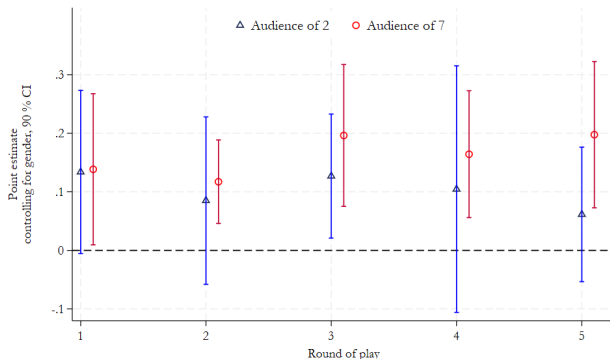
Note: Robust standard errors clustered at the individual level in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

# Treatments comparison by round of play

We estimate the following model for each round of play and plot point estimates on the two audience treatments:

$$\text{Choice}_i = \beta_1(\text{Audience of 2})_i + \beta_2(\text{Audience of 7})_i + \text{Gender}_i + \varepsilon_i$$

Figure 7: Point estimates on the two treatments



Note: 90% Confidence Intervals, SE clustered of the experimental session

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