



# Long working hours and depressive symptoms: moderating effects of gender, socioeconomic status, and job resources

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Received: 23 February 2018 / Accepted: 13 January 2019 / Published online: 12 March 2019  
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## Abstract

**Purpose** Systematic reviews and meta-analyses have found inconsistent associations between working hours and depressive symptoms. The purpose of this study was to investigate the possible moderators of this association, using data from a large-scale cross-sectional survey.

**Methods** A total of 16,136 Japanese employees (men 83.5%; women 16.5%) responded to a self-administered questionnaire inquiring about overtime working hours during the previous month and depressive symptoms (Center for Epidemiologic Studies Depression scale), as well as moderating factors including gender, age, marital status, socioeconomic status, commuting time, sleeping hours per day, job control and worksite social support (Job Content Questionnaire), neuroticism (Eysenck's Personality Questionnaire Revised), and social desirability (Social Desirability Scale) (response rate, 85%). We conducted sequential regression analyses to investigate the main effects and interaction effects of all moderating variables.

**Results** The association between overtime working hours and depressive symptoms was significantly moderated by gender (interaction effect:  $\beta = 0.03$ ), age ( $\beta = -0.02$ ), manager ( $\beta = 0.03$ ), sleeping hours ( $\beta = -0.02$ ), job control ( $\beta = -0.03$ ), and neuroticism ( $\beta = 0.02$ ). Among workers engaged in 80+ hours of overtime, higher depressive symptoms were reported by women, younger employees, non-managers, employees with low job control, low worksite social support, and high neuroticism. A significant main effect of long overtime working hours on depressive symptoms was also observed even after controlling for all independent variables ( $\beta = 0.02$ ).

**Conclusions** Long overtime working hours is associated with depressive symptoms. We also found significant heterogeneity in the association according to employee characteristics, which may explain the inconsistent findings in previous literature.

**Keywords** Depression · Japan · Overwork Mental Health · Workplace

The members of JSTRESS Group are listed in “Acknowledgements”.

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## Introduction

Long working hours exceeding regular working hours (40 h/week) remain a persistent problem in Japan, despite recent attention in the mass media to issues such as “karoshi” (death from over-work). Although average total working hours have been declining, daily working hours are actually on the rise among certain groups such as full-time male workers (Kuroda 2010).

Epidemiologic research indicates that long working hours are a cause of various physical health problems. For example, two systematic reviews (Bannai and Tamakoshi 2014; Kivimäki et al. 2015) showed that long working hours ( $\geq 41$  h/week or  $\geq 55$  h/week) are associated with high blood pressure, cardiovascular diseases, and all-cause mortality. In contrast, several systematic reviews or meta-analyses

have reported inconsistent findings for the association between working hours and mental health problems such as depression (Bannai and Tamakoshi 2014; Fujino et al. 2006; Sparks et al. 1997; Van der Hulst 2003; Virtanen et al. 2018; Watanabe et al. 2016). Specifically, although Bannai and Tamakoshi (2014), Sparks et al. (1997), and Watanabe et al. (2016) found a positive association between working hours and depression, Fujino et al. (2006) and Van der Hulst (2003) found inconsistent results. However, the most recent systematic review and meta-analysis (Virtanen et al. 2018) found an association between long working hours and the onset of depressive symptoms. Additionally, they found a stronger association in Asia than in European countries and no association in North America and Australia. However, this study did not reveal the reason for this regional difference nor what moderating factors affect this association.

Inconsistencies in the relation between long work hours and depression may stem from the potential heterogeneity of the association according to employee characteristics. For example, one possible modifier of the association between long working hours and mental illness is gender. Several studies reported that the association between overtime work and depressive disorder for women was greater than in men (Artazcoz and Gutiérrez 2012; Shields 1999). Among workers reporting the same working hours, it is likely that female workers—especially married women—are also engaged in more domestic labor compared to male workers, as captured by the concept of the “double shift” (Milner et al. 2015; Spurgeon et al. 1997). In Japan, gender differences in the division of household labor are even more marked than in many western societies. For example, working married women perform domestic chores for 5 h, while the average working married men report less than 1 h even if they have children under 6 years old (Ministry of Internal Affairs and Communications 2016).

Since there are only 24 h in a day, long work hours will be inversely correlated with sleeping time. Insufficient sleep could potentiate the adverse effects of long working hours by increasing negative affect (Kahn-Greene et al. 2007). Indeed, Nakata (2011) reported that combination of long work and short sleep posed the strongest risk for depressive symptoms among full-time employees. A similar argument applies to commuting time. Hence, sleep or commute time is a potential moderator of long working hours and depression.

Another possible reason for inconsistent results is the buffering role of job resources. According to the job demands–control model by Karasek (1979), the combination of high job demands and low job control generates the highest levels of worker stress. This situation which is called job strain was reported in a systematic review as a strong and consistent predictor for onset of clinical depression (Madsen et al. 2017). Thus, job control is critical when considering the impact of overwork on depressive symptoms (Hino et al.

2015). Additionally, as suggested by the iso-strain model (Johnson and Hall 1988), worksite social support (supervisor support and coworker support) may also buffer the association between long working hours and depressive symptoms. Thus, the effect of longer working hours on mental health is likely to be moderated by the availability of these job resources (Bannai and Tamakoshi 2014).

Socioeconomic status (SES), which is defined by educational status, income, or occupational status, is also robustly lined with health outcomes such as mortality or depression and anxiety (Glymour et al. 2014). Although there is no clear gradient in the association between SES and long work hours, a census-based longitudinal study revealed that the adverse effect of long working hours on mortality from cardiovascular disease was more marked in individuals engaged in low SES occupations (O’reilly and Rosato 2013). Thus, the adverse impact of long working hours on depressive symptoms might be observed among low SES workers, possibly because of low job control.

Finally, the association between long working hours and depressive symptoms might be moderated by the worker’s personality characteristics. Neuroticism is strongly associated with depressive and anxiety symptoms, while introversion is moderately associated with depressive symptoms (Jylhä and Isometsä 2006). On the other hand, social desirability has been negatively correlated with depression (Cole 1988).

In summary, the purpose of this study was to investigate a comprehensive range of possible moderators, i.e., gender, age, marital status, socioeconomic status (SES), commuting time, sleeping hours per a day, job control, worksite social support, neuroticism, and social desirability, of the association between overtime working hours and depressive symptoms, using data from the large-scale cross-sectional survey (Kawakami et al. 2004). We hypothesized that women, younger employees, lower SES employees, as well as workers with longer commuting time, shorter sleeping hours, fewer workplace resources (e.g., job control or social support), higher neuroticism, and lower social desirability are more likely to report depressive symptoms than men, older employees, higher SES employees, as well as workers with shorter commuting time, longer sleeping hours, larger workplace resources, lower neuroticism, and higher social desirability in situations involving severe overtime working hours.

## Methods

### Participants

The present study was conducted using the baseline survey of the Japan Work Stress and Health Cohort Study

(JSTRESS Study) which was conducted from April 1996 until May 1998 and recruitment strategies were as previously described (Kawakami et al. 2004). A questionnaire was distributed to 29,471 workers at nine manufacturing companies and factories, and a total of 22,770 questionnaires were returned (average response rate, 85.2%). After excluding participants with at least one missing response for variables relevant to the study, data from 16,136 respondents (13,467 men and 2669 women) were analyzed in this study. The study procedure was reviewed and approved by the Human Research Ethics Committee of Gifu University Graduate School of Medicine, before the surveys were conducted.

## Measures

### Overtime working hours per month

All study participants were full-time workers and their scheduled working hours were 40 h/week. Overtime working hours exceeding scheduled working hours were measured by asking participants average overtime working hours per month. First, we used overtime working hours as a continuous variable, but we also classified them into three groups (< 45 h, 45–79 h, and  $\geq 80$  h/month) to compare depressive symptoms scores among sub categories. This classification is generally used in the Japanese society and research, and equal to < 51.25 h, 51.25–59.75 h,  $\geq 60$  h/week (e.g., Hino et al. 2015).

### Depressive symptoms

Depressive symptoms were measured by Japanese version of the Center for Epidemiologic Studies Depression (CES-D) scale (Radloff 1977; Shima et al. 1985). This scale consists of 20 items asking participants to rate how often they experienced symptoms associated with depression over the past week. The response options were from 0 (“rarely” or “never”) to 3 (“most” or “almost all of the time”). The Japanese version of the CES-D has a good reliability and validity (Shima et al. 1985). In this study, Cronbach’s alpha coefficient for internal consistency was 0.82.

### SES and demographic characteristics

Gender, age, marital status, and SES variables (education and occupation) were assessed in the questionnaire. Marital status was classified into currently married and unmarried, which consists of never married, divorced, or widowed. Education was dichotomized into under high school graduates and higher education. Occupation was classified into manager, white-collar, and blue-collar.

### Commuting and sleeping time

One-way commuting time and average sleeping hours per a day were assessed in the questionnaire. One-way commuting time was classified into two groups with median time of the participants: short ( $\leq 40$  min) and long ( $> 40$  min). Average sleeping hours per a day during previous 1 year were grouped into short ( $\leq 6$  h) and long ( $\geq 7$  h) because the median daily sleeping hours in this sample are 7 h.

### Job control and worksite social support

Job control and worksite social support were measured using the Japanese version of the Job Content Questionnaire (JCQ) (Karasek 1985; Kawakami et al. 1995). Job control consists of skill discretion (six items) and decision authority (three items) (total Cronbach alpha for nine items = 0.79). Social support consists of supervisor support (four items) and coworker support (four items) (total Cronbach alpha for eight items = 0.84). Response options ranged from 1 (“strongly disagree”) to 4 (“strongly agree”).

### Neuroticism

Neuroticism was measured using scales from the short version of Eysenck’s Personality Questionnaire Revised (SEPQ-R) (Eysenck et al. 1985). This scale consists of 12 items with a yes/no response yielding a total score 0–12: a high score reflects the degree of nervousness or anxiety experienced by the respondent for neuroticism or the degree of participation and interaction with others in social situations for extroversion. The Japanese version of SEPQ-R was proven to be reliable and valid (Hosokawa and Ohyama 1993). Respondents were classified into two groups: low score (0.0–5.0) and high score (5.1–12.0). In this study, Cronbach’s alpha coefficient for internal consistency was 0.77.

### Social desirability

Social desirability was measured using the Japanese version of the Social Desirability Scale (Crowne and Marlowe 1960; Kitamura and Suzuki 1986). This scale consists of ten items with a yes/no response. A high score reflects the high degree of social desirability. In this study, Cronbach’s alpha coefficient for internal consistency was 0.53.

### Statistical analyses

We conducted sequential regression analyses to investigate the association of long overtime working hours on depressive symptoms. All continuous variables were grand-mean centered to avoid multicollinearity when entering interaction effects into the model. In Step 1, we first entered long

overtime working hours, demographic characteristics, SES, commuting time, and sleeping hours (Model 1). In Step 2, job resources and personality traits were additionally entered (Model 2). Finally, interaction terms (gender, age, marital status, education, manager, blue-collar, commuting time, sleeping hours, job control, worksite social support, neuroticism, social desirability  $\times$  working hours) were additionally entered into the model (Model 3). Subsequently, two-way ANCOVA and post hoc test (Bonferroni) were also conducted on the significant interaction effects to see the score differences of depressive symptoms. The level of significance was 0.05 (two tailed). The statistical analyses were conducted using IBM SPSS 23.0J for Windows.

## Results

### Characteristics of participants

Table 1 shows the demographic and occupational characteristics, overtime working hours, job resources, personality traits, and depressive symptoms scores of respondents in this study. Most of them were men (83.5%), married (77.6%), high school graduates (52.9%), and blue-collar workers (47.2%). The average overtime working time was 27.9 h. The prevalence of depressive symptoms when using cut-off point of 16+ was 24.0% in this study.

### Main effects of overtime working hours and characteristics on depression

A significant main effect of long overtime working hours on depression was observed after adjusting for gender, age, marital status, education, occupational status, commuting time, and sleeping hours in Model 1 (Table 2). In Model 2, the main effect of long overtime working hours on depression remained significant, while main effects were also found for all other variables including job resources and personality traits.

### Interaction effects of long overtime working hours and depressive symptoms

Significant interaction effects of overtime hours and age, woman, manager, sleeping hours, job control, and neuroticism were observed for depressive symptoms (Table 3). On the other hand, we did not find significant interaction effects of overtime hours and unmarried ( $p=0.417$ ), no college ( $p=0.960$ ), blue-collar ( $p=0.257$ ), long commuting time ( $p=0.877$ ), worksite social support ( $p=0.633$ ), or social desirability ( $p=0.808$ ).

Table 4 and Figs. 1, 2, 3, 4, 5 and 6 show score differences of CES-D among the groups which showed significant

interaction effects. We found that women, younger workers (30–49 years), workers in non-managerial white-collar occupations, and employees reporting low job control were more likely to report depressive symptoms when experiencing  $\geq 80$  h overtime working hours. Employees with high neuroticism were also more likely to experience depressive symptoms under conditions of long working hours. On the other hand, short-sleepers were more likely to have depressive symptoms regardless of how many hours they had worked overtime.

## Discussion

In this large-scale cross-sectional study, we found a significant main effect of long work hours on depressive symptoms. However, we also found significant heterogeneity in the relation according to gender, age, occupational status, sleeping hours, job control, and neuroticism. Women, younger employees, non-managers, employees with low control or high neuroticism are more likely to have depressive symptoms when they suffer from heavy overwork, compared to men, older employees, managers, employees with high control or low neuroticism, respectively. In other words, if an employee is a man,  $\geq 50$  years old, manager, having high job control, and having low neuroticism, he/she is less likely to have depressive symptoms even though working  $\geq 80$  h overtime per a month. The results of the current study revealed that the effect of severe overtime working time is heterogeneous, which helps to understand why some existing literature had inconsistent results on the association between working hours and depression. The result differences highly depend on the characteristics of participants.

In the current study, long overtime working hours themselves were one of the predictors of depressive symptoms. This is consistent with the systematic reviews or meta-analyses, which reported a positive association between long working hours and depression (Bannai and Tamakoshi 2014; Sparks et al. 1997; Virtanen et al. 2018; Watanabe et al. 2016). Additionally, the severity of depressive symptoms was obvious in working overtime 80+ h in the current study. This is also in line with the studies which showed that risk of psychiatric morbidity predominated in longest working hour group such as  $\geq 60$  h/week (Kato et al. 2014), 49–100 h/week (Kleppa et al. 2008), or  $\geq 280$  h/month (Nagashima et al. 2007). Working overtime  $\geq 80$  h means employees work at least 60 h/week or 12 h/day. Sleeping, bathing, or eating time are part of employees' free time so that if they work longer their free time will be reduced, resulting in less time to relax or recover efficiently.

Both men and women who work overtime  $\geq 80$  h had significantly higher depressive symptoms compared with those working less hours; however, women had much higher

**Table 1** Demographic and occupational characteristics, over time working hours per month, job resources, personality traits, and depressive symptoms among Japanese employees ( $N = 16,136$ )

Variables	Average (SD)	<i>n</i>	%
Demographic characteristics			
Gender			
Men		13,467	83.5
Women		2669	16.5
Age (years)	40.1 (9.2)		
Marital status			
Married		12,528	77.6
Unmarried		3608	22.4
SES			
Educational status			
Junior high school/high school graduate		10,395	64.4
College/university/graduate school graduate		5741	35.6
Occupational status			
Manager		2208	13.7
White-collar		6346	39.3
Profession		2250	13.9
Engineer		2238	13.9
Clerical job		1858	11.5
Blue-collar		7582	47.0
Service worker		178	1.1
Craft worker		1848	11.5
Machinery operator		3340	20.7
Physical worker		1406	8.7
Others		810	5.1
Over time working hours/month	27.9 (30.9)		
Less than 45 h		12,950	80.3
45–79 h		2326	14.4
80 h and over		860	5.4
One-way commuting time	38.2 (7.9)		
Less than 40 min		5723	35.5
40 min and more		10,413	64.5
Average sleeping hours/day			
6 h and less		6578	40.8
7 h and more		9558	59.2
Variables (scale)	Range	Average (SD)	Cronbach's $\alpha$
Job control (JCQ)	24–96	66.2 (11.1)	0.79
Worksite social support (JCQ)	12–60	45.2 (7.17)	0.84
Neuroticism (EPQ-R)	0–12	5.4 (3.3)	0.77
Social desirability (SDS)	0–10	4.5 (2.0)	0.53
Depressive symptoms (CES-D)	0–60	12.3 (6.7)	0.82

*SD* standard deviation, *JCQ* Job Content Questionnaire, *EPQ-R* Eysenck Personality Questionnaire-Revised, *SDS* Social Desirability Scale, *CES-D* Center for Epidemiologic Scale for Depression

depressive scores than men when they worked  $\geq 80$  h over-time, as expected. This finding is consistent with previous studies (Artazcoz and Gutiérrez 2012; Kleppa et al. 2008; Shields 1999). Although the current study did not ask participants about their household labor, it seems reasonable to assume that women had more responsibility for domestic chores, child care, and elderly care compared to men

(Milner et al. 2015; Spurgeon et al. 1997). Shimazu et al. (2013) reported that the association between family-to-work or work-to-family conflict and psychological distress among wives is greater than among husbands, which also supports our results. Moreover, long working hours are hazardous for pregnant women, because especially if associated with heavy lifting and prolonged standing, it may also increase



**Table 2** Sequential regression analyses with depressive symptoms ( $N=16,136$ )

	Model 1				Model 2			
	<i>b</i>	SE	$\beta$	<i>p</i>	<i>b</i>	SE	$\beta$	<i>p</i>
<b>Step 1</b>								
Overtime working hours	0.01	0.00	0.03	0.000	0.00	0.00	0.02	0.026
Age	−0.03	0.01	−0.04	0.000	−0.01	0.01	−0.02	0.022
Gender (woman = 1, man = 0)	0.35	0.15	0.02	0.017	−0.56	0.13	−0.03	0.000
Marital status (unmarried = 1, married = 0)	1.96	0.14	0.12	0.000	1.26	0.12	0.08	0.000
Education (no college = 1, college = 0)	0.92	0.14	0.07	0.000	0.82	0.12	0.06	0.000
Manager (manager = 1, non-manager = 0)	−1.60	0.17	−0.08	0.000	−0.85	0.15	−0.04	0.000
Blue-collar (blue-collar = 1, others = 0)	1.13	0.13	0.08	0.000	0.63	0.11	0.05	0.000
Commuting time (> 40 min = 1, ≤ 40 min = 0)	0.51	0.11	0.04	0.000	0.30	0.09	0.02	0.001
Sleeping hours (≤ 6 h = 1, ≥ 7 h = 0)	0.77	0.11	0.06	0.000	0.45	0.09	0.03	0.000
<b>Step 2</b>								
Job control					−0.06	0.01	−0.09	0.000
Worksite social support					−0.16	0.01	−0.17	0.000
Neuroticism					0.90	0.01	0.44	0.000
Social desirability					−0.09	0.02	−0.03	0.000
Adjusted $R^2$				0.060**				0.326**
$\Delta R^2$								0.266
<i>F</i> change				115.380				1594.173

*b* partial regression coefficient, *SE* standard error,  $\beta$  standard partial regression coefficient,  $R^2$  coefficient of determination

\*\* $p < 0.01$

**Table 3** Sequential regression analyses with depressive symptoms: interaction effects ( $N=16,136$ )

	Model 3			
	<i>b</i>	SE	$\beta$	<i>p</i>
<b>Step 3</b>				
Overtime hours × age	0.00	0.00	−0.02	0.034
Overtime hours × woman	0.02	0.01	0.03	0.014
Overtime hours × unmarried	0.00	0.00	0.01	n.s.
Overtime hours × no college	0.00	0.00	0.00	n.s.
Overtime hours × manager	0.01	0.01	0.02	0.049
Overtime hours × blue-collar	−0.01	0.00	−0.01	n.s.
Overtime hours × long commuting time	0.00	0.00	0.00	n.s.
Overtime hours × short sleeping hours	0.01	0.00	0.03	0.004
Overtime hours × job control	−0.00	0.00	−0.03	0.000
Overtime hours × worksite social support	0.00	0.00	−0.00	n.s.
Overtime hours × neuroticism	0.00	0.00	0.02	0.003
Overtime hours × social desirability	0.00	0.00	0.00	n.s.
Adjusted $R^2$				0.328**
$\Delta R^2$				0.003
<i>F</i> change				5.265

*b* partial regression coefficient, *SE* standard error,  $\beta$  standard partial regression coefficient,  $R^2$  coefficient of determination

\*\* $p < 0.01$

the risk of preterm delivery and low birthweight (Palmer et al. 2013). Thus, the results of the current study indicate long working hours impose a heavy burden on women and reducing overtime working hours may contribute to gender equality and occupational safety.

Occupation and age were significant moderators of the relationship between long working hours and depressive symptoms. In the present study, managers and employees aged  $\geq 50$  years were less likely to report depressive symptoms even when they worked  $\geq 80$  h overtime, while blue-collar employees had higher depressive symptoms regardless of how many hours they worked overtime. This finding is in line with previous studies (Glymour et al. 2014; O'reilly and Rosato 2013; Rawshani et al. 2016) and can be explained by job control, since low SES workers have less flexibility in the choice of overtime work or autonomy. In other words, workers in higher positions have more discretion over the content and hours of work. In this sample, the average age of managers was older (45.6 vs. 39.3 years old) and their score for job control was much higher than among non-managers (74.6 vs. 64.9). The traditional Japanese companies are notable for promoting workers based on seniority (i.e., age, or duration of tenure in the company)—as opposed to a system whereby talented employees can leap-frog over more senior colleagues. As O'Reilly and Rosato (2013) reported, men, managers, and own-account workers tend to work longer hours. Since workers with high job control tend to be more

**Table 4** Two-way interaction between overtime and other covariates on depression with post hoc tests results: adjusted means of CES-D score

	Overtime hours per month			<i>F</i>	<i>p</i>
	< 45	45–79	≥ 80		
Overtime hours × gender				4.82	< 0.01
Man	12.2 <sup>a</sup>	12.1 <sup>b</sup>	13.6 <sup>ab</sup> **		
Woman	12.3 <sup>c</sup>	9.9 <sup>d</sup>	21.2 <sup>cd</sup> ]		
Overtime hours × age				1.02	0.411
29 and less	12.6 ] **	13.3 ] *	14.3 ]		
30–39	12.6 <sup>**</sup> ] **	12.3 <sup>b</sup> ] **	14.2 <sup>ab</sup> ] *		
40–49	11.9 <sup>c</sup> ] **	12.0 <sup>c</sup> ]	13.4 <sup>cd</sup> ]		
50 and over	11.7 ]	11.4 ]	11.9 ]		
Overtime hours × occupation				1.83	0.118
Manager	10.3 <sup>a</sup> ] **	10.6 <sup>b</sup> ] **	11.9 <sup>ab</sup> ] **		
White-collar	11.8 <sup>c</sup> ] **	12.1 <sup>d</sup> ] **	13.7 <sup>cd</sup> ]		
Blue-collar	13.0 ]	12.5 ]	13.6 ]		
Overtime hours × sleeping hours				16.3	< 0.01
	**			5	
Short sleeping hours (≤ 6)	12.5 <sup>a</sup> ]	12.8 <sup>b</sup> ] *	14.0 <sup>ab</sup> ]		
Long sleeping hours (≥ 7)	12.0 <sup>c</sup> ]	11.5 <sup>d</sup> ]	13.3 <sup>cd</sup> ]		
Overtime hours × job control				4.42	0.012
Low job control (0.0–66.0)	13.0 <sup>a</sup> ] **	13.2 <sup>b</sup> ] **	15.5 <sup>ab</sup> ] **		
High job control (66.1–96.0)	11.4 <sup>c</sup> ]	11.3 <sup>d</sup> ]	12.4 <sup>cd</sup> ]		
Overtime hours × neuroticism				9.23	< 0.01
Low neuroticism (0.0–5.0)	9.8 <sup>a</sup> ] **	9.7 <sup>b</sup> ] **	10.0 <sup>c</sup> ] **		
High neuroticism (5.1–12.0)	14.9 <sup>a</sup> ]	14.8 <sup>b</sup> ]	16.9 <sup>ab</sup> ]		

Adjusted for gender, age, education, marital status, occupation, commuting time, and sleeping time

<sup>abcd</sup>*p* < 0.05; \**p* < 0.05, \*\**p* < 0.01 from post hoc tests (Bonferroni)

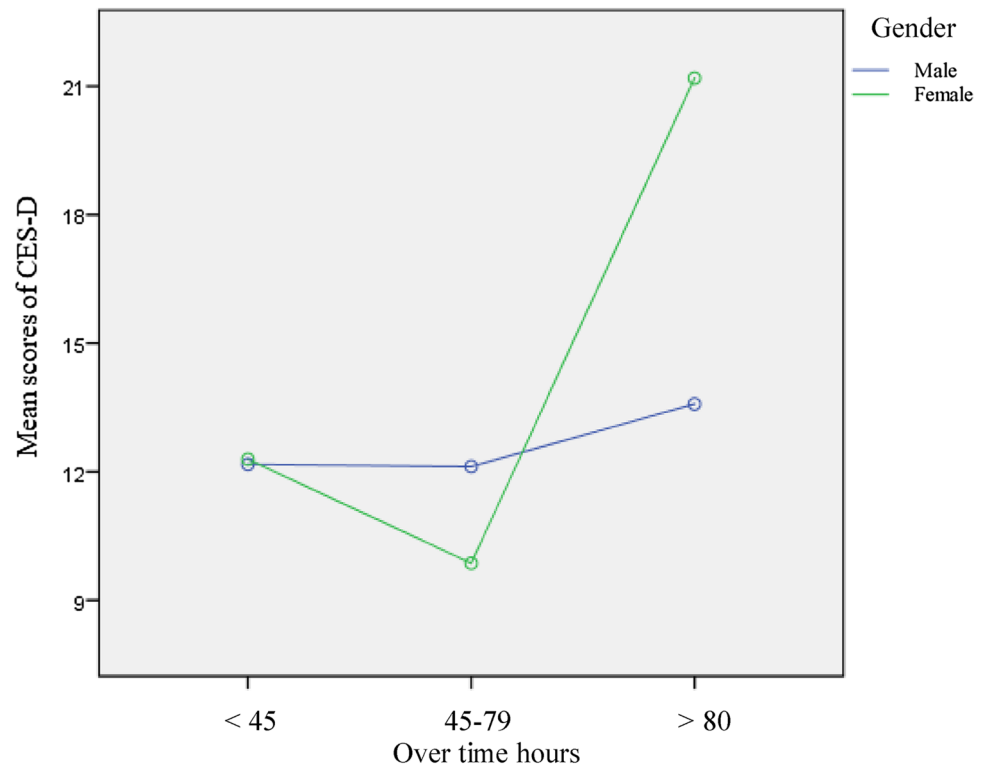
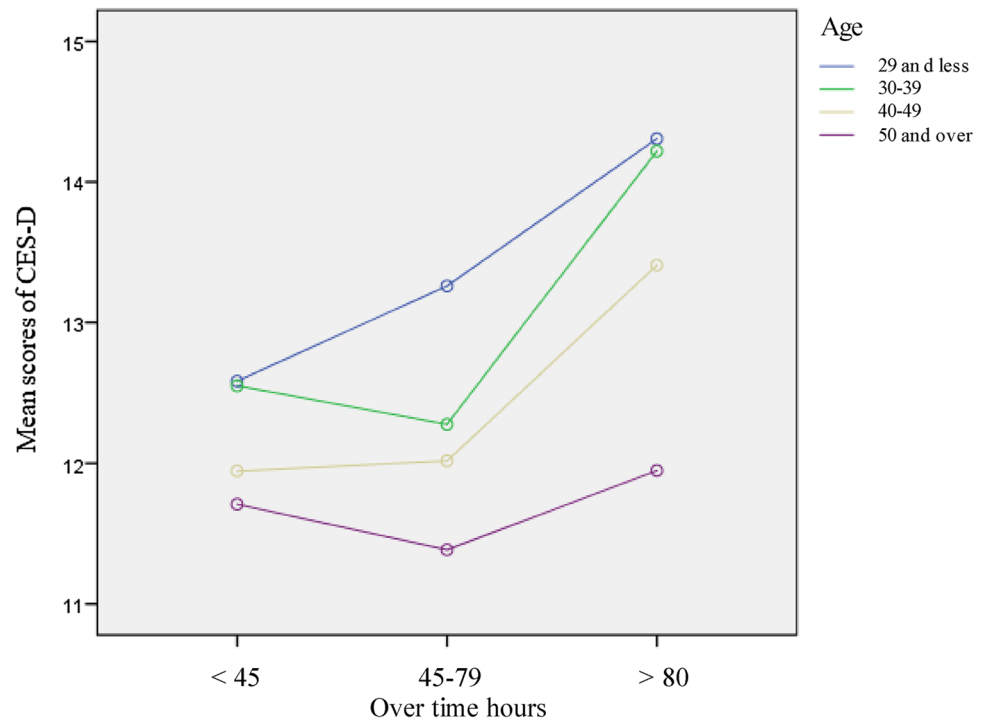
highly engaged in their work (Bakker et al. 2007), many of them contribute longer hours on a voluntary basis and thus they may experience less psychological stress. According to a model put forward by Shimazu and Schaufeli (2009), workaholism (working excessively and working compulsively) is positively associated with distress, while work engagement (vigor, dedication, and absorption) is protectively associated with psychological distress.

Employees with low job control and low worksite social support were more likely to suffer from depressive symptoms, and the score difference was more marked among those working 80+ overtime working hours. This is congruent with job demands–control support model itself and those reporting buffering effects of job resources on health outcomes (Bakker and Demerouti 2007; Hino et al. 2015). As we discussed in the previous paper (Hino et al. 2015), this finding may be explained by the concept of worktime control (WTC), which refers to the self-determination of worktime aspects including control over whether, when, and how many hours one should work overtime (Nijp et al. 2012). The results of the current study also suggest that the

amount of social support from supervisors or coworkers may influence depression risk regardless of how many hours they work overtime. Although we did not find a significant interaction effect of overtime working hours and worksite social support on depressive symptoms, our findings indicate that reducing long working hours and improving job control and worksite social support reduce depressive symptoms.

High neuroticism was a risk factor of depressive symptoms and its effect was more pronounced when employees work ≥ 80 h overtime. This pattern is consistent with the finding that neurotic traits are strongly related to depression (Jylhä and Isometsä 2006). We did not find an interaction effect of overtime working hours and social desirability on depressive symptoms but did a main effect, which is in line with the study that reported a negative association between social desirability and depression (Cole 1988).

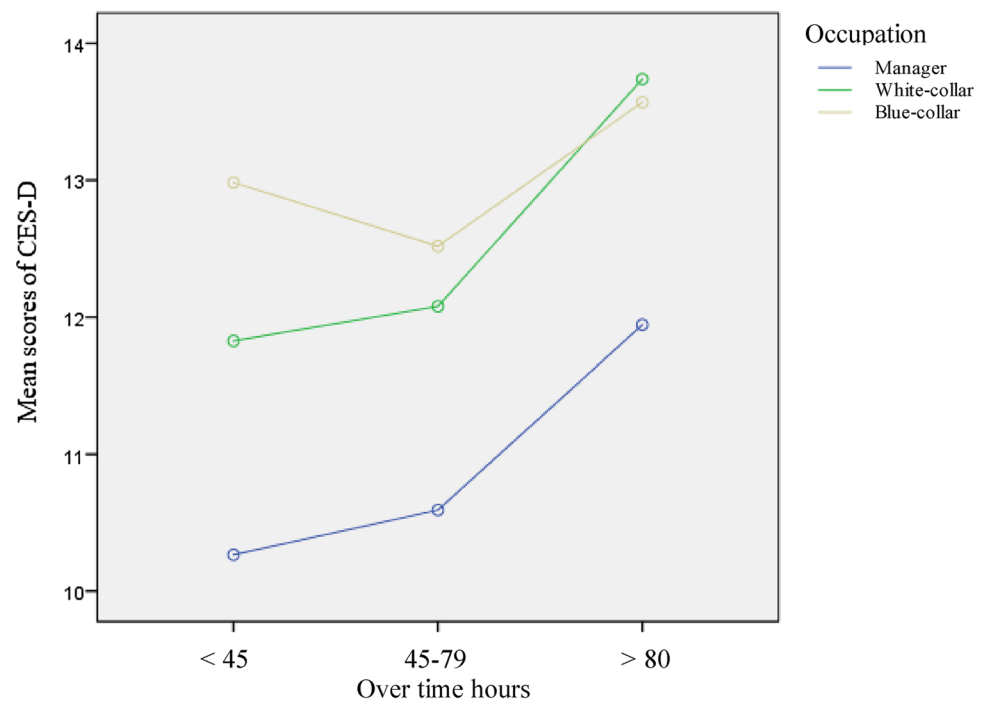
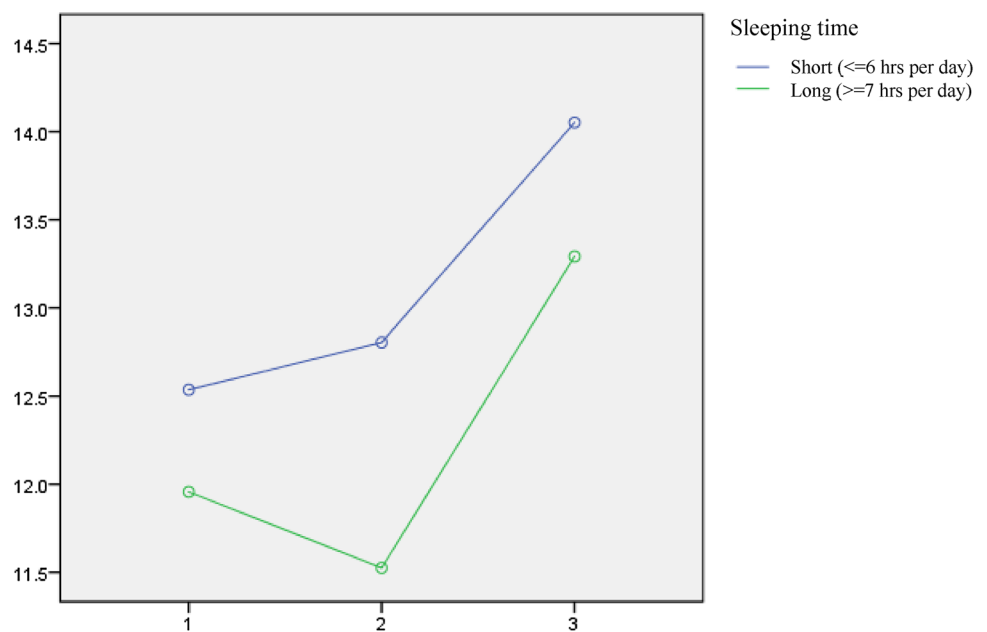
The interaction of overtime working and sleeping time was significant, while a significant interaction effect of overtime working hours and commute time was not found on depressive symptoms. As many studies reported (Kahn-Greene et al. 2007), the results of the current study suggest

**Fig. 1** Overtime hours × gender**Fig. 2** Overtime hours × age

that sleep deprivation would associate with depressive symptoms because short-sleepers had higher depressive symptoms than long-sleepers even among employees who worked overtime for < 45 h. Our study results are also in line with Nakata's (2011) study that reported combination

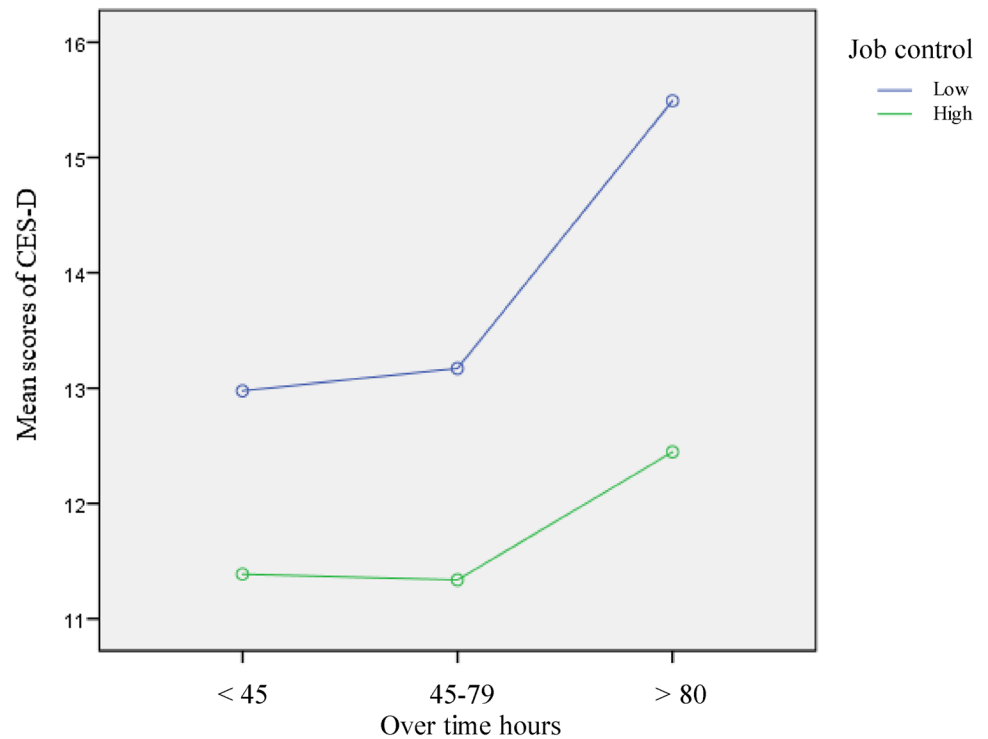
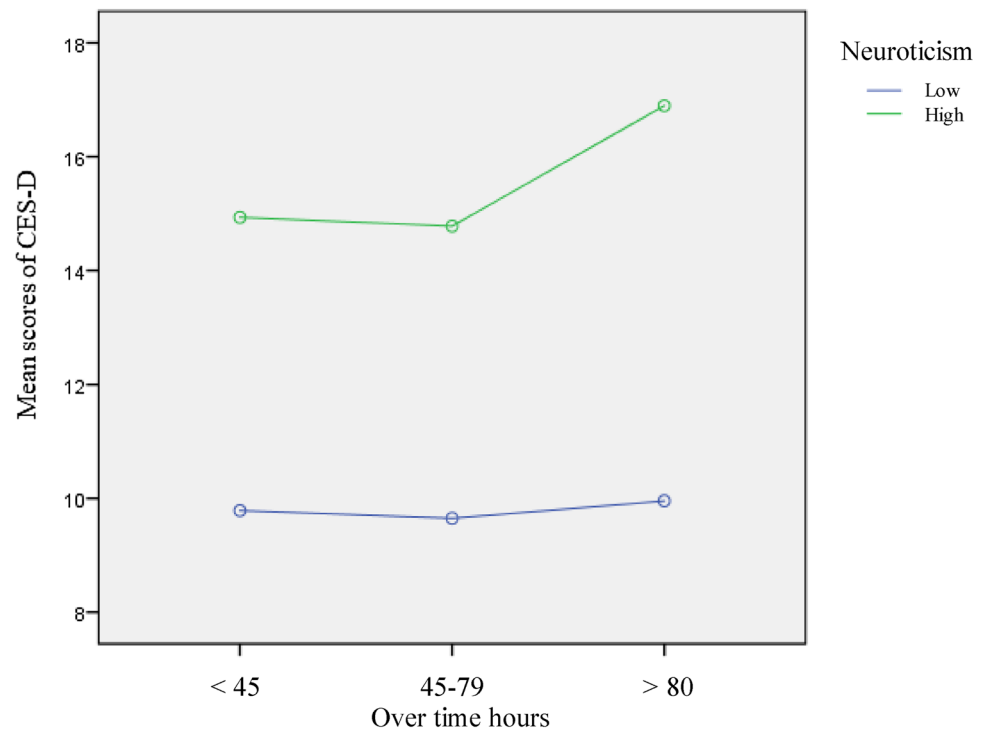
of long work and short sleep had a strongest risk for depressive symptoms among full-time employees. This might be because sleep may moderate the association between long work hours and depressive symptoms, and also short sleep is often a consequence of depression.



**Fig. 3** Overtime hours × occupation**Fig. 4** Overtime hours × sleeping time

Some possible limitations of our study should be considered. First, a causal relationship cannot be determined because this is a cross-sectional study. A prospective study is required to elucidate the causal relationship between overtime working hours and depressive symptoms and its moderators. Second, overtime working hours were measured by self-report, which may cause recall bias. Information of overtime working hours should obtain from company data or from employees on a daily basis to avoid this

bias. Finally, our sample was only from male-dominated manufacturing companies in Japan; therefore, there may be some degree of some health worker effect among female employees and generalization of the findings should be done with caution.

**Fig. 5** Overtime hours  $\times$  job control**Fig. 6** Overtime hours  $\times$  neuroticism

## Conclusion

Long overtime working hours were a predictor of depressive symptoms, independent of employee characteristics, commute or sleeping time, job resources, and personality

traits. However, severity of depressive symptoms was greater among women, younger employees, non-managers, employees with low job control, employees with low worksite social support, and employees with high neuroticism when they suffer from long overtime working hours.

**Acknowledgements** The following are members of JSTRESS (the Japan Work Stress and Health Cohort Study) group: Takashi Haratani (National Institute of Occupational Safety and Health, Japan), Fumio Kobayashi (Aichi Medical University, Japan), Masao Ishizaki (Kanazawa Medical University, Japan), Takeshi Hayashi (Hitachi Health Care Center, Japan), Osamu Fujita (Kariya Toyota General Hospital, Japan), Yoshiharu Aizawa (Kitasato University, Japan), Shogo Miyazaki (Kanda Christian Clinic, Japan), Hisanori Hiro (University of Occupational and Environmental Health, Japan), Takeshi Masumoto (Kimitsu Health Service Center, Japan), Shuji Hashimoto (Fujita Health University School of Medicine, Japan), and Shunichi Araki (Saitama Occupational Health Promotion Center, Japan).

**Funding** This study was supported by a special research grant for the prevention of work-related diseases in 1995–1999 from the Ministry of Health, Labour and Welfare, Japan.

## Compliance with ethical standards

**Conflict of interest** The authors declare that they have no competing interests.

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