## Panel Data Methods

BS1802 Statistics and Econometrics

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# Roadmap

- Regression analysis with cross-sectional data
  - The multiple regression analysis
    - Basics: estimation, inference, analysis with dummy variables
    - More technically involved: asymptotics, heteroskedasticity, specification and data issues
- Advanced topics
  - Limited dependent variable models
  - Panel data analysis
  - Regression analysis with time series data

# Outline (Wooldridge, Ch. 13.3, 13.5, 14.1)

- Two period panel data
- First-differenced estimation
- Fixed effects estimation

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### What is Panel Data?

- A set of panel data
  - has both a cross-sectional and a time series dimension
  - is collected by following the same individuals over a number of time periods
- E.g., a panel data set for wage, educ, exper, ...
  - Randomly select a sample of people from the population and collect data for 2016
  - 2 The same people are re-interviewed to collect data for 2017, 2018, ...
- It is possible to use a panel just like cross sections, but can do more than that
- Panel data allows us to address issues related to unobserved factors, which are difficult to handle with cross sectional data

#### Two Period Panel Data

- Example 9.4. City Crime Rates
  - Data: crime rates (*crmrte*) and unemployment rates (*unem*) from a sample of 46 cities in 1982 (t=1) and 1987 (t=2).
  - Question: Did *unem* influence *crmrte*?
  - Regressing crmrte on unem using the sample from 1987, we have

$$\widehat{\textit{crmrte}}_{87} = \underset{(20.76)}{128.38} - \underset{(3.42)}{4.16} \, \textit{unem}_{87},$$

$$n = 46, R^2 = .033$$

 The result is likely biased because many relevant factors (e.g., city, police, ...) are not controlled for

#### Two Period Panel Data

- An alternative way to look at the data
  - If the omitted variables are fixed over time, then we can decompose the error into two parts: factors that vary over time and those do not
- Consider the previous example in the panel setting

$$crmrte_{it} = \beta_0 + \delta_0 d2_t + \beta_1 unem_{it} + a_i + u_{it}, \quad t = 1, 2$$

#### where

- *i* is the city
- t is the time period
- ullet  $d2_t$  is the dummy variable indicating the second time period
- A time-constant component is added to the error  $v_{it} = a_i + u_{it}$

#### Fixed Effects Model

• In general, the fixed-effects model can be written as

$$y_{it} = \beta_0 + \delta_0 d2_t + \beta_1 x_{it1} + \dots + \beta_k x_{itk} + a_i + u_{it}, \quad t = 1, 2$$

#### where

- $a_i$  is the fixed effect (invariant to t) that represents factors specific to individual i (allowed to be correlated with  $\mathbf{x}_{it}$ )
- $u_{it}$  is called the idiosyncratic error that represents unobserved factors varying both overtime and across sections (typically assumed to be uncorrelated with  $\mathbf{x}_{it}$ )

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#### First-Differenced Estimation

Write the model separately

$$y_{i1} = \beta_0 + \delta_0 \cdot 0 + \beta_1 x_{i11} + \dots + \beta_k x_{i1k} + a_i + u_{i1}, \quad (t = 1)$$
  
$$y_{i2} = \beta_0 + \delta_0 \cdot 1 + \beta_1 x_{i21} + \dots + \beta_k x_{i2k} + a_i + u_{i2}, \quad (t = 2)$$

Subtracting the first equation from the second one gives

$$\Delta y_i = \delta_0 + \beta_1 \Delta x_{i1} + \dots + \beta_k \Delta x_{ik} + \Delta u_i,$$

(first-differenced equation) which is a cross-section model and is free of  $a_i$ 

- $\bullet$  When  $u_{it}$  is uncorrelated with regressors in both periods
  - There is no correlation between  $\Delta x_i$ 's and  $\Delta u_i$
  - OLS will be unbiased

#### Panel Data Estimation in R

- The command to perform panel data estimation in R is plm(formula, data, effect, model, index, ...)
  - effect
    - fixed effects for cross-sectional units ("individual")
    - time effects ("time")
    - both ("twoways")
  - model
    - first-differences ("fd")
    - fixed effects ("within")
    - random effects ("random")

# First-Differenced Estimation: An Example

- Example 9.7. City Crime Rates.
  - First-differenced estimation

$$\widehat{\Delta \textit{crmrte}} = \underset{(4.70)}{\widehat{15.40}} + \underset{(.88)}{2.22} \Delta \textit{unem},$$

$$n = 46, R^2 = .127$$

- There is a positive and significant relationship between unem<sub>it</sub> and crmrte<sub>it</sub>
- One percentage point rise in unemployment rate increases 2.22 crimes per 1,000 people
- The crimes per 1,000 people increased by 15.4 in 1987, in comparison to 1982

# First-Differenced Estimation: A Shortcoming

Consider the log wage model with two-period panel

$$\log(wage_{it}) = \beta_0 + \delta_0 d2_t + \beta_1 educ_{it} + a_i + u_{it}, \quad t = 1, 2,$$

where ai represents unobserved factors, say abilityi

• The first-differenced equation is

$$\Delta \log(wage_i) = \delta_0 + \beta_1 \Delta educ_i + \Delta u_i$$

- However, for most adult workers,  $\Delta educ_i$  is zero. The overall variation in  $\Delta educ_i$  is small, and thus OLS estimator will have a large standard error
- Using the first-differenced estimation is a good idea for "returns to eduction". But, frequently, it does not work well because of the lack of variation in  $\Delta educ_i$

### Panel Data with More than Two Periods

- For the panel data with T periods
  - Subtract period 1 from period 2
  - **2** Subtract period (T-1) from period T
  - $oldsymbol{3}$  We have (T-1) observations per individual
  - **4** Estimate by OLS, assuming the  $\Delta u_{it}$  are uncorrelated over time
- ullet The key assumption about the idiosyncratic error  $u_{it}$  is

$$Cov(x_{itj}, u_{is}) = 0$$

- When using more than two time periods, we must assume that  $\Delta u_{it}$  is uncorrelated over time for the usual standard errors and test statistics to be valid
  - To deal with serial correlation in  $\Delta u_{it}$ , GLS method may be used

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#### Fixed Effects Estimation

- When there is an unobserved fixed effect, an alternative to first differences is fixed effects estimation
- Consider a model with a single explanatory variable

$$y_{it} = \beta_0 + \beta_1 x_{it} + a_i + u_{it}$$

• The average over time for individual *i* is

$$\bar{y}_i = \beta_0 + \beta_1 \bar{x}_i + a_i + \bar{u}_i$$

• The average of  $a_i$  will be  $a_i$ . So if we subtract the average from  $y_{it}$ , we have

$$y_{it} - \bar{y}_i = \beta_1(x_{it} - \bar{x}_i) + (u_{it} - \bar{u}_i)$$

 Each individual has been "de-meaned" for all variables, which eliminates the fixed effects.

### First Difference vs Fixed Effects?

- When T=2, first difference and fixed effects estimators will be exactly the same
- For T > 2
  - Both are unbiased (with T fixed as  $N \to \infty$ )
  - The relative efficiency of the estimators is determined by the serial correlation in  $u_{it}$ 
    - When u<sub>it</sub> are serial uncorrelated, fixed effects is typically more efficient
    - Serial correlation tests in R: pwartest()

# Fixed Effects Estimation: An Example

• Example 14.1. Effect of Job Training on Firm Scrap Rates

$$\begin{split} \log(\textit{scrap}_{\textit{it}}) &= \beta_0 + \delta_0 \textit{d}88 + \delta_1 \textit{d}89 \\ &+ \beta_1 \textit{grant}_{\textit{it}} + \beta_2 \textit{grant}_{\textit{i,t-1}} + \textit{a_i} + \textit{v_{\textit{it}}} \end{split}$$

- Data description
  - 54 firms reported scrap rates in each of the three years, 1987, 1988, 1989
  - No firms received grants prior to 1988
  - In 1988, 19 firms received grants; in 1989, 10 different firms received grants
  - A lagged value of the grant indicator  $(grant_{i,t-1})$  is included to allow for the possibility that the additional job training in 1988 made workers more productive in 1989

# Fixed Effects Estimation: An Example

Dependent Variable: log(scrap)	
Independent Variables	Coefficient (Standard Error)
d88	080 (.109)
d89	247 (.133)
grant	252 (.151)
$grant_{-1}$	422 (.210)
Observations Degrees of freedom <i>R</i> -squared	162 104 .201