

# LEARNING BAYESIAN NETWORKS

# Outline

- 1 Introduction
- 2 Learning Parameters
- 3 Learning Structures
- 4 Summary

# Outline

**1 Introduction**

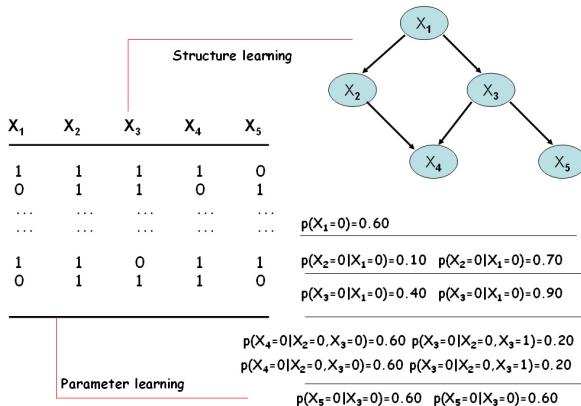
2 Learning Parameters

3 Learning Structures

4 Summary

# From data to Bayesian networks

## Learning structure and parameters



# Discovering associations

## The task of learning Bayesian networks from data

- Given a **data set of cases**  $D = \{\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(N)}\}$  drawn at random from a joint probability distribution  $p_0(x_1, \dots, x_n)$  over  $X_1, \dots, X_n$ , and possibly some domain expert background knowledge
- The **task** consists of identifying (**learning**) a DAG (directed acyclic graph) **structure**  $S$  and a set of corresponding **parameters**  $\Theta$

# Discovering associations

## The task of learning Bayesian networks from data

- When discovering associations **all the variables have the same treatment**
- There is **not a target variable**, as in supervised classification
- There is **not a hidden variable**, as in clustering

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# Maximum likelihood estimation

## Parameter space

- Consider a variable  $X$  with  $r$  possible values:  $\{1, 2, \dots, r\}$
- We have  $N$  observations (cases) of  $X$ :  $D = \{x_1, \dots, x_N\}$ , that is a sample of size  $N$  extracted from  $X$ 
  - Example:  $X$  variable measuring the result obtained after rolling a dice five times.  $D = \{1, 6, 4, 3, 1\}$ ,  $r = 6$ , and  $N = 5$
- We are interested in estimating:  $P(X = k)$
- The parametric space is
$$\Theta = \{\theta = (\theta_1, \dots, \theta_r) | \theta_i \in [0, 1], \sum_{i=1}^r \theta_i = 1\}$$
- $P(X = k | \theta_1, \dots, \theta_r) = \theta_k$



# Maximum likelihood estimation

## Likelihood function

- $L(D : \theta) = P(D|\theta) = P(X = x_1, \dots, X = x_N|\theta)$
- The likelihood function measures **how probable** is to obtain the dataset of cases for a concrete value of the parameter  $\theta$
- Assuming that the **cases are independent**:

$$P(D|\theta) = \prod_{i=1}^N P(X = x_i|\theta) = \prod_{k=1}^r \theta_k^{N_k}$$

$N_k$  = number of cases in the dataset for which  $X = k$

# Likelihood function

## Example

	$X$
1	0
2	0
3	0
4	0
5	0
6	1
7	1
8	1
9	1
10	1

$$\theta = P(X = 1) = \frac{1}{4}$$

$$L(D : \frac{1}{4}) = P(D | \frac{1}{4})$$

$$= P(X = 0, \dots, X = 1 | \frac{1}{4}) = \frac{3}{4}^5 \frac{1}{4}$$

$$\theta = P(X = 1) = \frac{1}{2}$$

$$L(D : \frac{1}{2}) = P(D | \frac{1}{2})$$

$$= P(X = 0, \dots, X = 1 | \frac{1}{2}) = \frac{1}{2}^5 \frac{1}{2}$$

$$= \frac{1}{2}^{10} > \frac{3}{4}^5 \frac{1}{4}$$

# Maximum likelihood estimation

## Categorical distribution: relative frequencies

- $\theta^* = (\theta_1^*, \theta_2^*, \dots, \theta_r^*) = \arg \max_{(\theta_1, \theta_2, \dots, \theta_r)} P(D|\theta)$
- In a **categorical distribution**, the maximum likelihood estimator for  $P(X = k)$  is:

$$\theta_k^* = \frac{N_k}{N}$$

i.e., the relative frequency

- In the previous example, the maximum likelihood estimator of  $\theta = P(X = 1)$  is  $\theta^* = \frac{5}{10}$

# Bayesian estimation

## Prior, posterior distributions

- $\theta = (\theta_1, \theta_2, \dots, \theta_r)$  is assumed to be a random variable
- $f(\theta|S) \sim \text{Dir}(a_1, \dots, a_r)$  **PRIOR** distribution
- $\Rightarrow f(\theta|D, S) \propto p(D|S, \theta)f(\theta|S) \sim \text{Dir}(a_1 + N_1, \dots, a_r + N_r)$   
**POSTERIOR** distribution
- The Bayesian estimation is the posterior mean:

$$\theta_k^* = \frac{N_k + a_k}{N + \sum_{i=1}^r a_i}$$

- $\text{Dir}(\theta_1, \dots, \theta_r; a_1, \dots, a_r) = \frac{\Gamma(\sum_{i=1}^r a_i)}{\prod_{i=1}^r \Gamma(a_i)} \theta_1^{a_1-1} \dots \theta_r^{a_r-1}$

# Bayesian estimation

## Many rules for estimation

- **Lindstone rule**, with  $a_k = \lambda, \forall k$ :

$$\theta_k^* = \frac{N_k + \lambda}{N + r\lambda}$$

- **Laplace rule** with  $\lambda = 1$ :

$$\theta_k^* = \frac{N_k + 1}{N + r}$$

- **Jeffreys-Perks rule** with  $\lambda = 0.5$ :

$$\theta_k^* = \frac{N_k + 0.5}{N + \frac{r}{2}}$$

- **Schurmann-Grassberger rule** with  $\lambda = \frac{1}{r}$ :  $\theta_k^* = \frac{N_k + \frac{1}{r}}{N + 1}$

# Estimation of parameters

## Parameters $\theta_{ijk}$

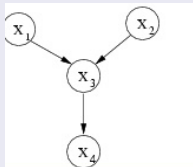
- Bayesian network structure  $S = (\mathbf{X}, A)$  with  $\mathbf{X} = (X_1, \dots, X_n)$  and  $A$  denoting the set of arcs
- Variable  $X_i$  has  $r_i$  possible values:  $x_i^1, \dots, x_i^{r_i}$
- Local probability distribution  $P(x_i \mid \mathbf{pa}_i^{j,S}, \theta_i)$ :

$$P(x_i^k \mid \mathbf{pa}_i^{j,S}, \theta_i) = \theta_{x_i^k \mid \mathbf{pa}_i^j} \equiv \theta_{ijk}$$

- The parameter  $\theta_{ijk}$  represents the conditional probability of variable  $X_i$  being in its  $k$ -th value, knowing that the set of its parent variables is in its  $j$ -th value
- $\mathbf{pa}_i^{1,S}, \dots, \mathbf{pa}_i^{q_i,S}$  denotes the values of  $\mathbf{Pa}_i^S$ , the set of parents of the variable  $X_i$  in the structure  $S$ 
  - The term  $q_i$  denotes the number of possible different instances of the parent variables of  $X_i$ . Thus,  $q_i = \prod_{X_g \in \mathbf{Pa}_i} r_g$
- The local parameters for variable  $X_i$  are given by  $\theta_i = ((\theta_{ijk})_{k=1}^{r_i})_{j=1}^{q_i}$
- Global parameters:  $\theta = (\theta_1, \dots, \theta_n)$

# Maximum likelihood estimation of parameters

## Parameters $\theta_{ijk}$ example



Local probabilities

$$\theta_1 = (\theta_{1-1}, \theta_{1-2})$$

$$\theta_2 = (\theta_{2-1}, \theta_{2-2}, \theta_{2-3})$$

$$\theta_3 = (\theta_{311}, \theta_{321}, \theta_{331},$$

$$\theta_{341}, \theta_{351}, \theta_{361},$$

$$\theta_{312}, \theta_{322}, \theta_{332},$$

$$\theta_{342}, \theta_{352}, \theta_{362})$$

$$\theta_4 = (\theta_{411}, \theta_{421}, \theta_{412}, \theta_{422})$$

$$P(x_1^1), P(x_1^2)$$

$$P(x_2^1), P(x_2^2), P(x_2^3)$$

$$P(x_3^1|x_1^1, x_2^1), P(x_3^1|x_1^1, x_2^2), P(x_3^1|x_1^1, x_2^3),$$

$$P(x_3^2|x_1^2, x_2^1), P(x_3^2|x_1^2, x_2^2), P(x_3^2|x_1^2, x_2^3),$$

$$P(x_3^3|x_1^3, x_2^1), P(x_3^3|x_1^3, x_2^2), P(x_3^3|x_1^3, x_2^3),$$

$$P(x_4^1|x_3^1), P(x_4^1|x_3^2), P(x_4^1|x_3^3),$$

$$P(x_4^2|x_3^1), P(x_4^2|x_3^2), P(x_4^2|x_3^3)$$

Factorisation of the joint mass probability

$$P(x_1, x_2, x_3, x_4) = P(x_1)P(x_2)P(x_3|x_1, x_2)P(x_4|x_3)$$

**Figure:** Structure, local probabilities and resulting factorization for a Bayesian network with four variables ( $X_1$ ,  $X_3$  and  $X_4$  with two possible values, and  $X_2$  with three possible values)

variable	possible values	parent variables	possible values of the parents
$X_i$	$r_i$	$Pa_i$	$q_i$
$X_1$	2	$\emptyset$	0
$X_2$	3	$\emptyset$	0
$X_3$	2	$\{X_1, X_2\}$	6
$X_4$	2	$\{X_3\}$	2

**Table:** Variables ( $X_i$ ), number of possible values of variables ( $r_i$ ), set of variable parents of a variable ( $Pa_i$ ), number of possible instantiations of the parent variables ( $q_i$ )

# Maximum likelihood estimation of parameters

## Global independence of the parameters

- Assuming **global independence of the parameters**:

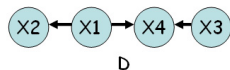
$$L(D : \theta) = \prod_{i=1}^n L(D_i : \theta_i)$$

- It is possible to **estimate the parameter for each variable  $X_i$  independently** of the other variables



# Maximum likelihood estimation of parameters

Global independence:  $L(D : \theta) = \prod_{i=1}^n L(D_i : \theta_i)$



X1	X2	X4	X3
1	0	0	1
0	1	0	1
0	1	1	0
0	1	0	0
1	0	0	1

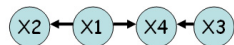


X1	X2	X4	X3
1	0	0	1
0	1	0	1
0	1	1	0
0	1	0	0
1	0	0	1

**Figure:** Dataset  $D_2$  for estimating the parameters of variable  $X_2$

# Maximum likelihood estimation of parameters

Global independence:  $L(D : \theta) = \prod_{i=1}^n L(D_i : \theta_i)$



D

X1	X2	X4	X3
1	0	0	1
0	1	0	1
0	1	1	0
0	1	0	0
1	0	0	1



D1

X1	X2	X4	X3
1	0	0	1
0	1	0	1
0	1	1	0
0	1	0	0
1	0	0	1

**Figure:** Dataset  $D_1$  for estimating the parameters of variable  $X_1$

# Maximum likelihood estimation of parameters

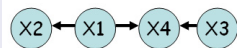
## Local independence of the parameters

- Assuming **local independence of the parameters**:

$$L(D : \theta) = \prod_{i=1}^n \prod_{j=1}^{q_i} L(D_{ij} : \theta_{ij})$$

# Maximum likelihood estimation of parameters

Local independence:  $L(D : \theta) = \prod_{i=1}^n \prod_{j=1}^{q_i} L(D_{ij} : \theta_{ij})$



D

X1	X2	X4	X3
1	0	0	1
0	1	0	1
0	1	1	0
0	1	0	0
1	0	0	1



$D_{21}=D_2|X_1=0$

X1	X2	X4	X3
1	0	0	1
0	1	0	1
0	1	1	0
0	1	0	0
1	0	0	1

**Figure:** Dataset  $D_{21}$  for estimating the parameters of variable  $X_2$  when  $X_1 = 0$

# Maximum likelihood estimation of parameters

$$L(D : \theta) = \prod_{i=1}^n \prod_{j=1}^{q_i} \prod_{k=1}^{r_i} \theta_{ijk}^{N_{ijk}}$$

- $P(X_i = x_i^k \mid \mathbf{pa}_i^j) = \theta_{ijk}$  with  $i = 1, \dots, n; j = 1, \dots, q_i$  and  $k = 1, \dots, r_i$
- $N_{ij}$  number of cases in  $D$  where the configuration  $\mathbf{pa}_i^j$  has been observed
- $N_{ijk}$  number of cases in  $D$  where simultaneously  $X_i = x_i^k$  and  $\mathbf{Pa}_i = \mathbf{pa}_i^j$  has been observed ( $N_{ij} = \sum_{k=1}^{r_i} N_{ijk}$ )

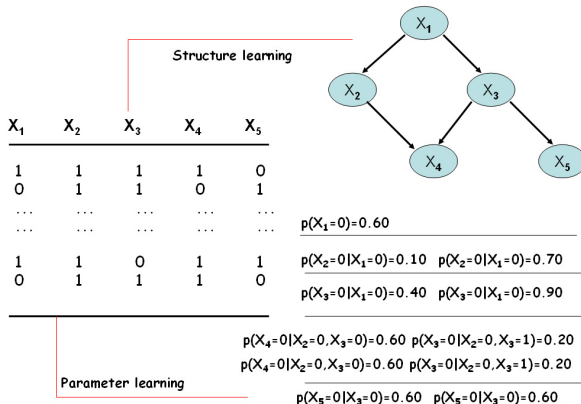
$$L(D : \theta) = \prod_{i=1}^n \prod_{j=1}^{q_i} \prod_{k=1}^{r_i} \theta_{ijk}^{N_{ijk}}$$

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# Introduction

## Learning structure (DAG) and parameters (conditional tables)



# Introduction

## Three types of methods

- Based on **detecting conditional independencies** (Constraint-based methods)
  - ① Study the (in)dependence relationships between the variables by means of **statistical tests**
  - ② Try to **find the structure(s)** that represents the most (or all) of these relationships
- Based on **score + search**
  - Try to find the structure that best **"fit" the data**
  - They need:
    - ① A **score** (metric or evaluation function) to measure the fitness of each candidate structure
    - ② A **search method** (heuristic) to explore in an intelligent manner the space of possible solutions
    - ③ **Several types of spaces** can be considered
- **Hybrid** methods
  - Based on a search technique guided by a score and the detection of conditional independencies



# Testing conditional independencies

## PC algorithm (Spirtes et al. 1993)

- General idea is based on generating a **skeleton** derived **through statistical tests** for detecting conditional independencies
- **Start from the complete undirected graph**
- **Recursive conditional independence tests** for deleting edges
- The **output is a CPDAG** where the edges should be transformed into arcs

# Testing conditional independencies

## Some considerations

- $X_i$  and  $X_j$  are independent given  $\mathbf{Z}$  iff  $2N MI(X_i, X_j | \mathbf{Z}) \rightarrow \chi^2_{(r_i-1)(r_j-1) | \mathbf{Z}}$
- The **reliability** of the test:
  - Increases with  $N$ , the number of cases (it is an asymptotic test)
  - Reduces dramatically with the order of the test (number of variables in  $\mathbf{Z}$ )

# Testing conditional independencies

## Completed Partially DAG (CPDAG)

- Using only conditional independence tests: not possible to obtain a **unique DAG**
- Usually a **completed partially DAG (CPDAG)** is obtained
- Each **CPDAG** represents an **equivalent class of DAGs**
- Two DAGs,  $S_1$  and  $S_2$  are **equivalent** (or Markov equivalent) if for all  $W, Y, Z$

$$I_{S_1}(W, Y|Z) \iff I_{S_2}(W, Y|Z)$$

- Two DAGs,  $S_1$  and  $S_2$  are **equivalent** iff they have the **same edges** (no direction) and the **same head to head patterns** (arcs  $X \rightarrow Z$  and  $Y \rightarrow Z$  and  $X$  and  $Y$  are not adjacent)

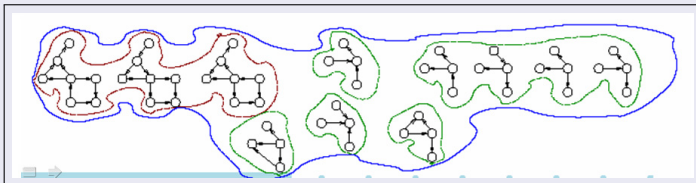
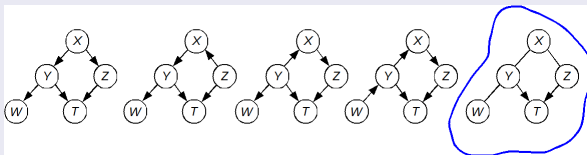
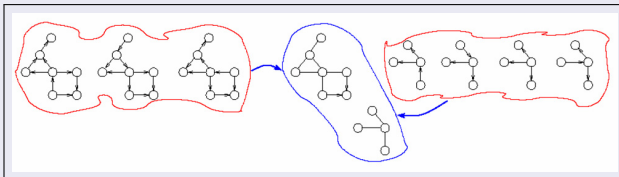


Figure: Equivalent DAGs

# Testing conditional independencies

## Completed Partially DAG (CPDAG)



- **Arcs in the CPDAG** appear in all DAGs of its equivalent class
- **Edges in the CPDAG** can be orientated in different ways in each DAG of its class (without new head to head patterns or cycles)

# Testing conditional independencies

## PC algorithm (Spirtes et al. 1993)

Form complete, undirected graph  $S$

$t = -1$

**repeat**

$t = t + 1$

**repeat**

select ordered pair of adjacent nodes  $A, B$  in  $S$

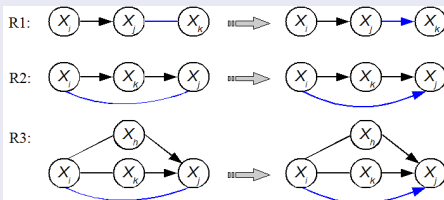
select neighborhood  $C$  of  $A$  of size  $t$  (if possible)

delete edge  $A - B$  in  $S$  if  $A$  and  $B$  cond. ind. given  $C$

**until** all ordered pairs have been tested

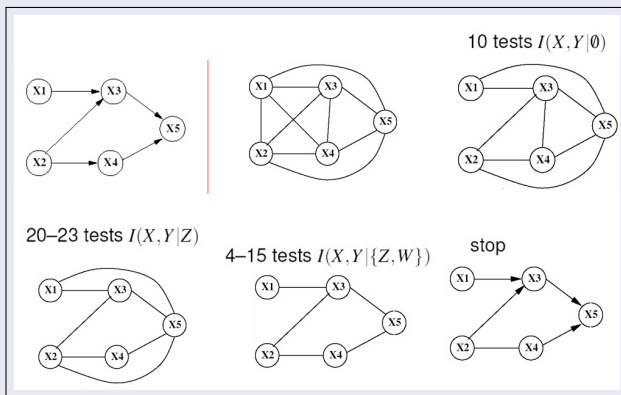
**until** all neighborhood are of size smaller than  $t$

Transform edges in arcs by applying some simple rules



# Testing conditional independencies

## PC algorithm (Spirtes et al. 1993). Example with $t = 2$



**Figure:** Example of the PC algorithm with  $t = 2$

# Score+search approaches

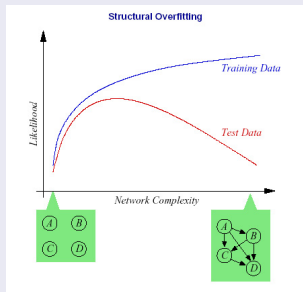
## Introduction

- They try to find the structure that best “fit” the data
- They are characterized by:
  - A **score** (metric or evaluation function) to measure the fitness of each candidate structure
    - Penalized log-likelihood
    - Bayesian metrics
  - A **space of structures** where the search is carried out
    - Directed acyclic graphs
    - Equivalence classes
    - Order between the variables
  - A **search method** to explore in an intelligent manner the space of possible solutions
    - Local search
    - Heuristics

# Score+search approaches

## Score metrics. Penalized log-likelihood

- **Log-likelihood of the data:**  $\log P(D : S, \theta) = \sum_{i=1}^n \sum_{j=1}^{q_i} \sum_{k=1}^{r_i} \log(\theta_{ijk})^{N_{ijk}}$
- $\log P(D : S, \hat{\theta}) = \sum_{i=1}^n \sum_{j=1}^{q_i} \sum_{k=1}^{r_i} N_{ijk} \log \frac{N_{ijk}}{N_{ij}}$   
with  $\hat{\theta}_{ijk} = \frac{N_{ijk}}{N_{ij}}$  (maximum likelihood estimate)



**Figure:** Likelihood increases monotonically wrt model complexity



# Score+search approaches

## Score metrics. Penalized log-likelihood

- Avoid overfitting **penalizing the complexity** of the Bayesian network in the log-likelihood :

$$\sum_{i=1}^n \sum_{j=1}^{q_i} \sum_{k=1}^{r_i} N_{ijk} \log \frac{N_{ijk}}{N_{ij}} - \dim(S) \text{pen}(N)$$

- $\dim(S) = \sum_{i=1}^n q_i(r_i - 1)$  **model dimension**
- $\text{pen}(N)$  **no negative penalization function**
  - $\text{pen}(N) = 1$ : Akaike's information criterion (AIC) (Akaike, 1974)
  - $\text{pen}(N) = \frac{1}{2} \log N$ : Bayesian information criterion (BIC) (Schwarz, 1978). It is equivalent to the minimum description length (MDL) (Lam and Bacchus, 1994) criterion

# Score+search approaches

## Score metrics. Bayesian model selection

- Try to obtain the **structure with maximum a posteriori probability given the data**: that is  $\arg \max_S P(S|D)$
- Using Bayes formula:

$$P(S|D) = \frac{P(D|S)P(S)}{P(D)}$$

$$P(S|D) \propto P(D|S)P(S)$$

- $P(D|S)$  is the **marginal likelihood** of the data
- $P(S)$  denotes the **prior distribution** over structures
- If  $P(S)$  is **uniform** ( $\max P(S|D) \equiv \max P(D|S)$ ) we try to obtain the **structure with maximum marginal likelihood**

# Score+search approaches

## Score metrics. Bayesian model selection. K2 metric

- Accounts for **uncertainty also in the parameters**:

$$P(D|S) = \int P(D|S, \theta) p(\theta|S) d\theta$$

- $P(D|S)$  **posterior probability** of the data given the structure
- $P(D|S, \theta)$  **likelihood** of the data given the Bayesian network (structure + parameters)
- $p(\theta|S)$  **prior** distribution over the parameters

# Score+search approaches

## Score metrics. Bayesian model selection. K2 metric

- Assuming that  $p(\theta_{ij}|S)$  is uniform, it is possible to obtain a closed formula for  $P(D|S)$  (Cooper and Herskovits, 1992)

$$P(D|S) = \prod_{i=1}^n \prod_{j=1}^{q_i} \frac{(r_i - 1)!}{(N_{ij} + r_i - 1)!} \prod_{k=1}^{r_i} N_{ijk}!$$

- $n$ : number of variables
- $r_i$ : number of states  $X_i$  can have
- $q_i$ : number of possible state combinations of  $\mathbf{Pa}_i$
- $N_{ijk}$ : number of cases in  $D$  where  $X_i$  takes its  $k$ -th value and the parent set of  $X_i$  are on their  $j$ -th combination of values
- $N_{ij} = \sum_{k=1}^{r_i} N_{ijk}$

# Score+search approaches

## Score metrics. Bayesian model selection. K2 algorithm

- An **ordering** between the nodes is assumed
- An **upper bound** is set **on the number of parents** for any node
- **For every node**,  $X_i$ , K2 searches for **the set of parent nodes that maximizes**:

$$g(X_i, \mathbf{Pa}_i) = \prod_{j=1}^{q_i} \frac{(r_i - 1)!}{(N_{ij} + r_i - 1)!} \prod_{k=1}^{r_i} N_{ijk}!$$

- K2 assumes **initially that a node does not have parents**
- **At each step** K2 incrementally **adds the parent** whose addition **provides the best value for**  $g(X_i, \mathbf{Pa}_i)$
- K2 **stops** when adding a single parent to any node cannot increase  $g(X_i, \mathbf{Pa}_i)$
- K2 is a **greedy** algorithm

# Score+search approaches

## Score metrics. Bayesian model selection. BDe metric

- Assuming that  $p(\theta_{ij}|S)$  follows a Dirichlet distribution, it is possible to obtain a closed formula for  $P(D|S)$

$$P(D|S) = \prod_{i=1}^n \prod_{j=1}^{q_i} \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} \prod_{k=1}^{r_i} \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})}$$

- This is called the Bayesian Dirichlet (BD) score
- $\alpha_{ijk}$  denotes the parameters of the Dirichlet distribution
  - $\alpha_{ijk} = 1$ : K2 metric (Cooper and Herskovits, 1992)
  - $\alpha_{ijk} = \alpha P(x_i^k, \mathbf{Pa}_i = \mathbf{pa}_i^j | S)$ : likelihood-equivalent Bayesian Dirichlet (BDe) score (Heckerman et al., 1995)
  - $\alpha_{ijk} = \alpha / q_i r_i$ : BDeu score (Buntine, 1991)
- Decomposable score = can be expressed as a sum of values that depend on only one node and its parents. All (estimated log-likelihood, AIC, BIC/MDL, BD, K2, BDe and BDeu)
- Score equivalence property = two Markov equivalent graphs score the same. All but K2 and BD are score equivalent

# Score+search approaches

## Different spaces for search

- **Space of directed acyclic graphs**

$$d(n) = \sum_{i=1}^n (-1)^{i+1} \binom{n}{i} 2^{i(n-i)} d(n-i); \quad d(0) = 1; \quad d(1) = 1$$

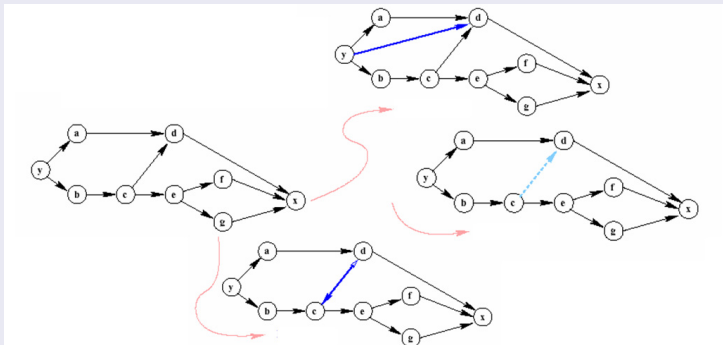
E.g.,  $d(10) \simeq 4.2 \times 10^{18}$

- **Space of equivalence classes** (each class reflects the same set of conditional independencies)
  - Scores: score equivalent (Chickering, 1996)
- **Ordering between the variables** (Larrañaga et al., 1996, Friedman and Koller, 2002): cardinality of the search space  $n!$

# Score+search approaches

## Search algorithms. Local search. B algorithm (Buntine, 1991)

- Local operators: **insert, delete and invert an arc**
- Efficient search** due to the decomposability of the scores





# Score+search approaches

## Search algorithms. Metaheuristics and exact methods

- **Greedy** search (Buntine, 1991; Cooper and Herskovits, 1992), simulated **annealing** (Heckerman et al., 1995), **genetic** algorithms (Larrañaga et al., 1996), **MCMC** methods (Giudici and Green, 1999; Friedman and Koller, 2003; Grzegorzcyk and Husmeier, 2008) and **estimation of distribution algorithms** (Larrañaga et al., 2000; Blanco et al., 2003)
- **Exact** methods (several dozens of variables only): dynamic programming (Koivisto and Sood, 2004; Silander and Myllymäki, 2006; Malone et al., 2011), branch and bound (de Campos and Ji, 2011), and mathematical programming (Martínez-Rodríguez et al., 2008; Jaakkola et al., 2010)

# Outline

- 1 Introduction
- 2 Learning Parameters
- 3 Learning Structures
- 4 Summary**

# Learning Bayesian networks

## Structure + parameters

- Learning parameters
  - Maximum likelihood estimation
  - Bayesian estimation (Dirichlet distribution)
- Learning structures
  - Detecting conditional independencies (PC algorithm)
  - Score + search (penalized log-likelihood (AIC, BIC, MDL), Bayesian metrics (K2, BD, BDe, BDeu); local, metaheuristics)
  - Hybrid methods