Outline

- I. States and observations
- II. Transitions and sensor models
- III. Filtering
- IV. Prediction
- V. Smoothing

^{*} Figures/images are from the <u>textbook site</u> or the instructor unless the source is cited specifically.

I. Time and Uncertainty

Probabilistic reasoning in static worlds discussed so far.

Every random variable has a single fixed value.

- Real situations are dynamic with evidence evolving with time and thus actions predicted (and chosen) based on the history of evidence:
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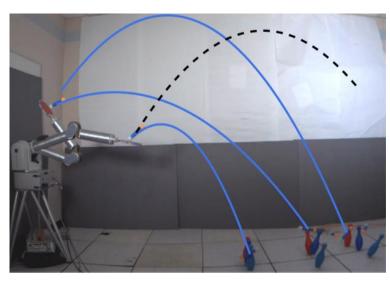
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Ximea MQ022CG-CM high speed camera

Frame rate: 170 fps (frames per second)

Resolution: 2048×1088 pixel



https://www.youtube.com/watch?v=dGBevZ54E3s

Batting an in-flight dumbbell-shaped object

Duration: 0.6 second with 90 frames

Motion of the object estimated by an extended Kalman filter (EKF).

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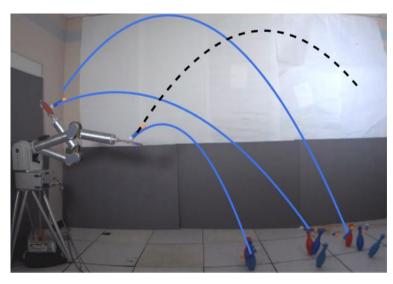
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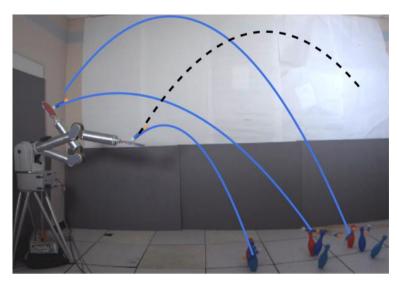
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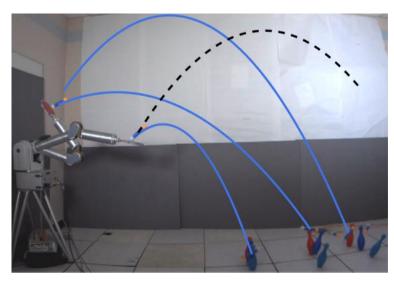
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 ${m E}_t = {m e}_t$ for some observed values ${m e}_t$

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The umbrella world is represented by

state variable sequence: R_0 , R_1 , R_2 , ... evidence variable sequence: U_0 , U_1 , U_2 , ...

 $U_{3:5}$ corresponds to U_3 , U_4 , U_5 .

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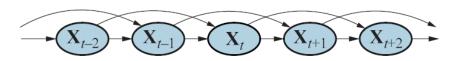
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1st order Markov process (transition model $P(X_t \mid X_{t-1})$)



 2^{nd} order Markov process (transition model $P(X_t \mid X_{t-1}, X_{t-2})$)

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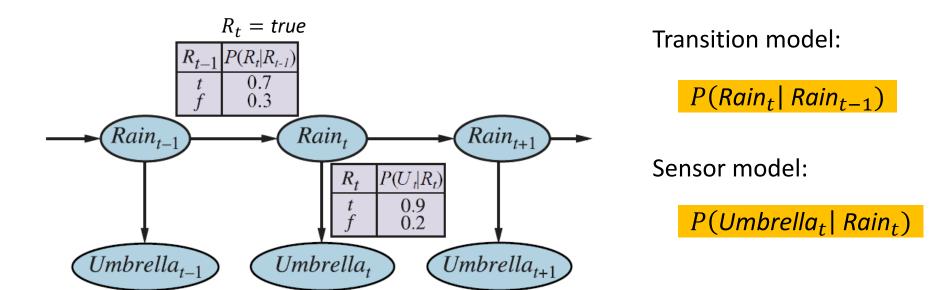
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$$P(E_t | X_{0:t}, E_{1:t-1}) = P(E_t | X_t)$$

sensor/observation model

Models for the Umbrella World



1st order Markov Process

The state (Rain) causes the sensor to take on a particular value (Umbrella).

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Given the prior distribution $P(X_0)$ at time 0, we have the complete joint distribution:

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- Discrete time models can handle an infinite set of variables due to
 - use of integer indices
 - use of implicit universal quantification to define sensor and transition models

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E.g., the probability of rain three days from now.

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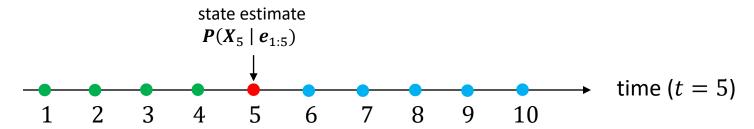
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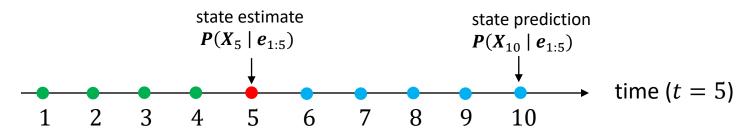
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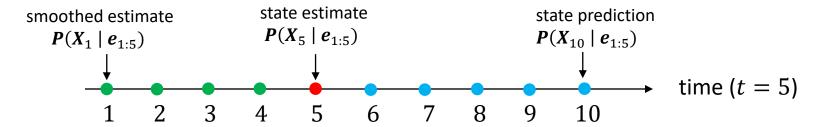
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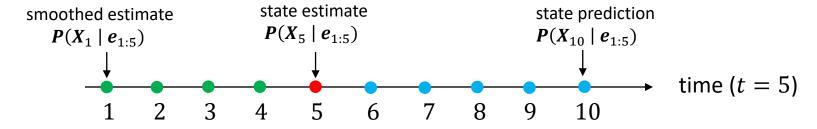
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E.g., the probability that it rained last Wednesday.



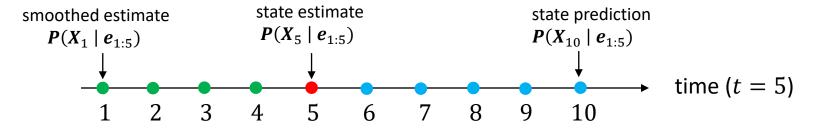
♦ Most likely explanation is the task of computing the sequence of states $x_{1:t}$ to maximize $P(X_{1:t} | e_{1:t})$.

Given a sequence of observations, find the sequence of states that is mostly likely to have generated those observations.

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E.g., If the umbrella appears on each of the first three days and is absent on the fourth, then the most likely explanation is that it rained on the first three days and did not rain on the fourth.

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$$1: t t 1: t$$
)

To be efficient (so usable in a real time scenario), a filtering algorithm

- needs to maintain current state estimate and update it on the fly, and
- ♠ should not go back over the entire history

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state estimate at $t+1$ state estimate at t

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Time and space for the update at *t* must be *constant* in order to keep track of the current state distribution indefinitely.

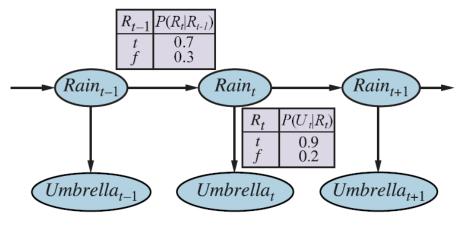
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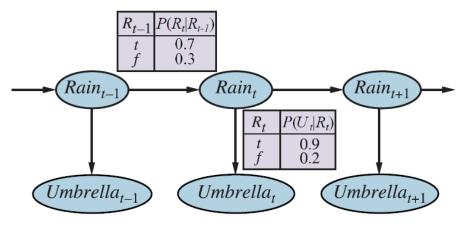
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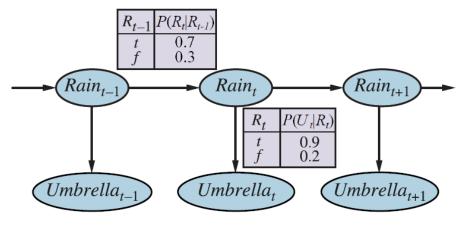
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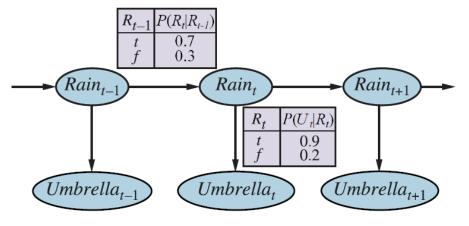
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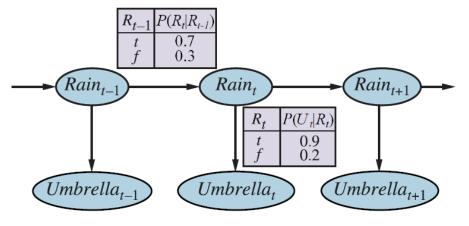
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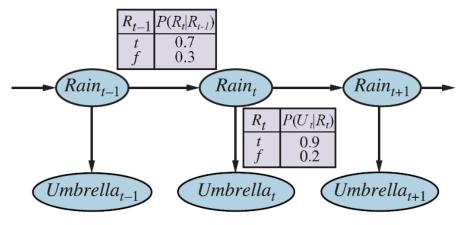
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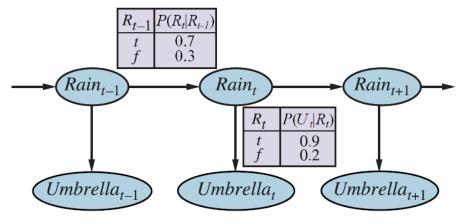
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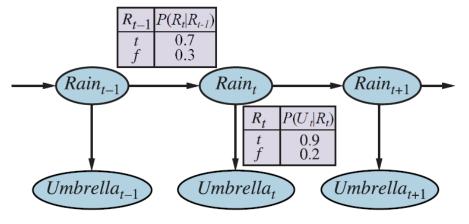
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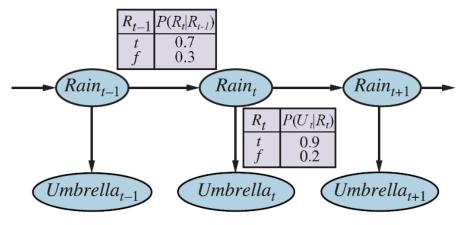
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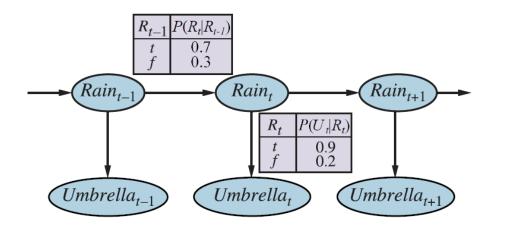
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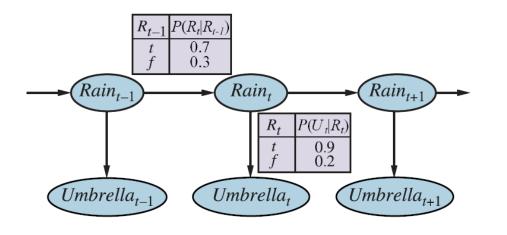
$$\approx \langle 0.818, 0.182 \rangle$$



$$P(R_1 \mid u_1) \approx \langle 0.818, 0.182 \rangle$$

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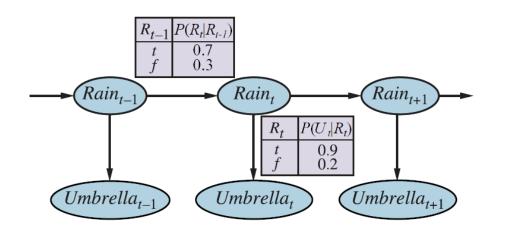
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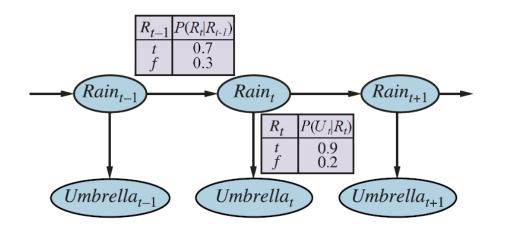
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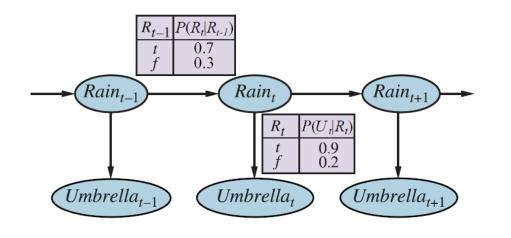
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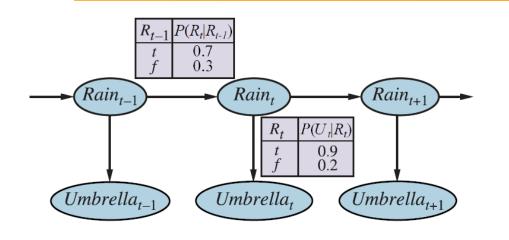
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$$\approx \alpha \langle 0.565, 0.075 \rangle$$

$$\approx \langle 0.883, 0.117 \rangle$$

Prediction is essentially filtering without the addition of new evidence.

Only prediction and no update at every time step.

For
$$k = 0,1,...$$

$$P(X_{t+k+1} \mid e_{1:t}) = \sum_{x_{t+k}} P(X_{t+k+1} \mid x_{t+k}) P(x_{t+k} \mid e_{1:t})$$
transition model recursion

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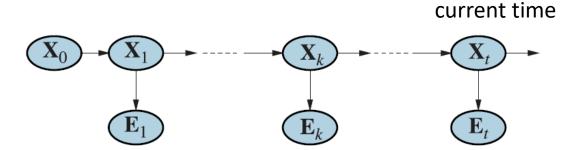
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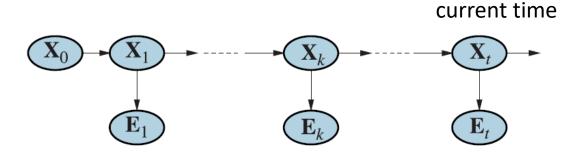
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Compute $P(X_k \mid e_{1:t})$ for some $0 \le k < t$.

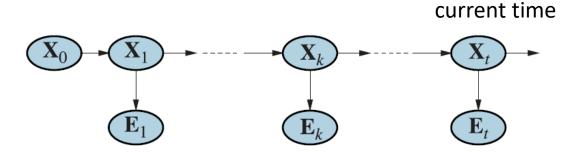


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 \clubsuit Split the evidence into two parts: up to k and from k+1 to t.

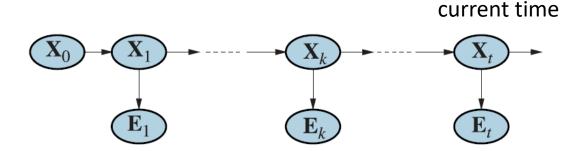
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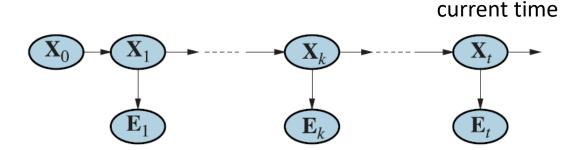


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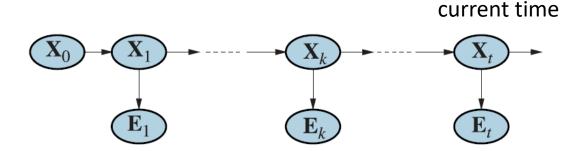
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$$\begin{split} \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:t}) &= \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}, \boldsymbol{e}_{k+1:t}) \\ &= \alpha \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}) \boldsymbol{P}(\boldsymbol{e}_{k+1:t} \mid \boldsymbol{X}_k, \boldsymbol{e}_{1:k}) \quad \text{(Bayes' rule, given } \boldsymbol{e} \mid 1:k \mid 1$$

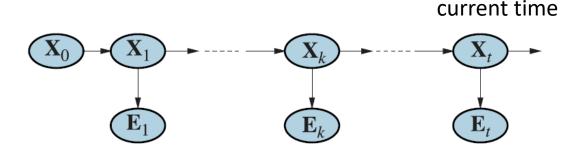
Compute $P(X_k \mid e_{1:t})$ for some $0 \le k < t$.



• Split the evidence into two parts: up to k and from k+1 to t.

$$\begin{split} \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:t}) &= \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}, \boldsymbol{e}_{k+1:t}) \\ &= \alpha \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}) \boldsymbol{P}(\boldsymbol{e}_{k+1:t} \mid \boldsymbol{X}_k, \boldsymbol{e}_{1:k}) \quad \text{(Bayes' rule, given } \boldsymbol{e} \mid 1:k \mid 1$$

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• Split the evidence into two parts: up to k and from k+1 to t.

$$\begin{aligned} \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:t}) &= \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}, \boldsymbol{e}_{k+1:t}) \\ &= \alpha \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}) \boldsymbol{P}(\boldsymbol{e}_{k+1:t} \mid \boldsymbol{X}_k, \boldsymbol{e}_{1:k}) \quad \text{(Bayes' rule, given } \boldsymbol{e} \mid 1:k \text{)} \\ &= \alpha \boldsymbol{P}(\boldsymbol{X}_k \mid \boldsymbol{e}_{1:k}) \boldsymbol{P}(\boldsymbol{e}_{k+1:t} \mid \boldsymbol{X}_k) \quad \text{(conditional independence)} \end{aligned}$$

. The backward message can also be computed recursively.

$$P(e_{k+1:t} \mid X_k) = \sum_{x_{k+1}} P(e_{k+1} \mid x_{k+1}) P(e_{k+2:t} \mid x_{k+1}) P(x_{k+1} \mid X_k)$$
sensor model recursion transition model

Equation for smoothing $(0 \le k < t)$:

$$P(X_k \mid \boldsymbol{e}_{1:t}) = \alpha P(X_k \mid \boldsymbol{e}_{1:k}) P(\boldsymbol{e}_{k+1:t} \mid X_k)$$

Equation for smoothing $(0 \le k < t)$:

$$P(X_k | e_{1:t}) = \alpha P(X_k | e_{1:k}) P(e_{k+1:t} | X_k)$$

Forward computation (filtering)

$$P(X_0) \rightarrow P(X_1 \mid e_{1:1}) \rightarrow \cdots \rightarrow P(X_k \mid e_{1:k})$$

where, for $0 \le i \le k - 1$,

$$P(X_{i+1} | e_{1:i+1}) = \alpha P(e_{i+1} | X_{i+1}) \sum_{x_i} P(X_{i+1} | x_i) P(x_i | e_{1:i})$$

Equation for smoothing $(0 \le k < t)$:

$$P(X_k \mid \boldsymbol{e}_{1:t}) = \alpha P(X_k \mid \boldsymbol{e}_{1:k}) P(\boldsymbol{e}_{k+1:t} \mid X_k)$$

Forward computation (filtering)

$$P(X_0) \rightarrow P(X_1 \mid e_{1:1}) \rightarrow \cdots \rightarrow P(X_k \mid e_{1:k})$$

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$$P(X_{i+1} | e_{1:i+1}) = \alpha P(e_{i+1} | X_{i+1}) \sum_{x_i} P(X_{i+1} | x_i) P(x_i | e_{1:i})$$

Backward computation

$$P(e_{t+1:t} \mid X_t) \rightarrow P(e_{t:t} \mid X_{t-1}) \rightarrow \cdots \rightarrow P(e_{k+1:t} \mid X_k)$$

$$= P(\mid X_t) = 1 \text{ (vector of 1s)}$$

Equation for smoothing $(0 \le k < t)$:

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$$P(X_{i+1} \mid e_{1:i+1}) = \alpha P(e_{i+1} \mid X_{i+1}) \sum_{x_i} P(X_{i+1} \mid x_i) P(x_i \mid e_{1:i})$$

Backward computation

$$P(e_{t+1:t} \mid X_t) \rightarrow P(e_{t:t} \mid X_{t-1}) \rightarrow \cdots \rightarrow P(e_{k+1:t} \mid X_k)$$

$$= P(\mid X_t) = 1 \text{ (vector of 1s)}$$

where, for $k \le j \le t$

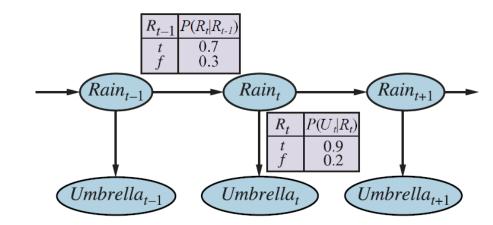
$$P(e_{j+1:t} | X_j) = \sum_{x_{j+1}} P(e_{j+1} | x_{j+1}) P(e_{j+2:t} | x_{j+1}) P(x_{j+1} | X_j)$$

Compute $P(R_1 | u_1, u_2)$ as follows:

probability of rain on day 1, given that umbrellas were observed on days 1 and 2.

$$u_1 = true$$

 $u_2 = true$

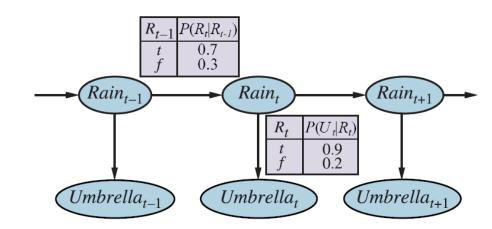


Compute $P(R_1 | u_1, u_2)$ as follows:

$$u_1 = true$$

$$u_2 = true$$

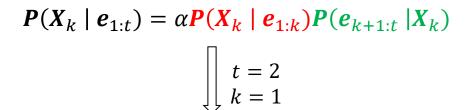




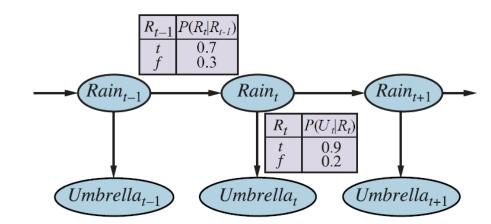
Compute $P(R_1 | u_1, u_2)$ as follows:

$$u_1 = true$$

$$u_2 = true$$



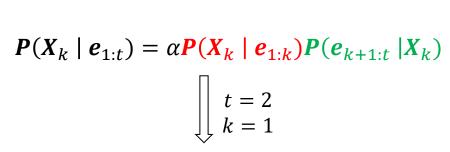
$$P(R_1 | u_1, u_2) = \alpha P(R_1 | u_1) P(u_2 | R_1)$$



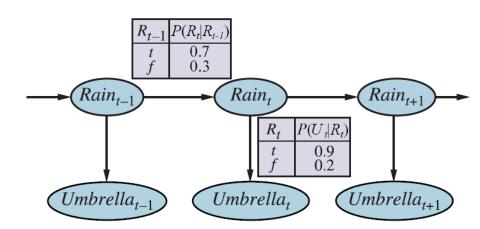
Compute $P(R_1 | u_1, u_2)$ as follows:

$$u_1 = true$$

$$u_2 = true$$



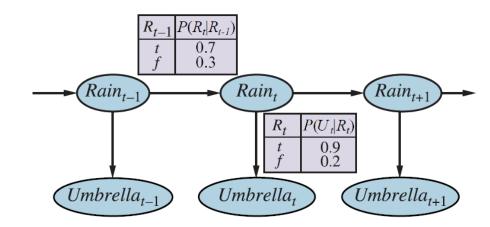
$$\begin{aligned} \textbf{\textit{P}}(\textit{R}_1 \mid \textit{u}_1, \textit{u}_2) &= \alpha \textbf{\textit{P}}(\textit{R}_1 \mid \textit{u}_1) \textbf{\textit{P}}(\textit{u}_2 \mid \textit{R}_1) \\ &\approx \langle 0.818, 0.182 \rangle \\ &\text{as computed earlier} \end{aligned}$$



Compute $P(R_1 \mid u_1, u_2)$ as follows:

$$u_1 = true$$

$$u_2 = true$$



$$P(X_k \mid e_{1:t}) = \alpha P(X_k \mid e_{1:k}) P(e_{k+1:t} \mid X_k)$$

$$\begin{bmatrix} t = 2 \\ k = 1 \end{bmatrix}$$

$$\begin{split} \textbf{\textit{P}}(\textit{R}_1 \mid \textit{u}_1, \textit{u}_2) &= \alpha \textbf{\textit{P}}(\textit{R}_1 \mid \textit{u}_1) \textbf{\textit{P}}(\textit{u}_2 \mid \textit{R}_1) \\ &\approx \langle 0.818, 0.182 \rangle \\ &\text{as computed earlier} \end{split}$$

$$P(u_2 \mid R_1) = \sum_{r_2} P(u_2 \mid r_2) P(\mid r_2) P(r_2 \mid R_1)$$

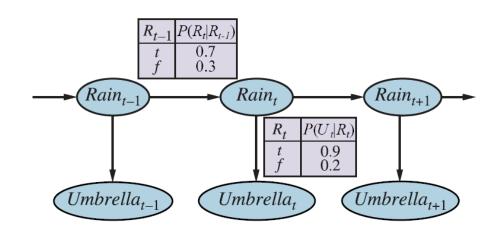
$$= (0.9 \cdot 1 \cdot \langle 0.7, 0.3 \rangle) + (0.2 \cdot 1 \cdot \langle 0.3, 0.7 \rangle)$$

$$= \langle 0.69, 0.41 \rangle$$

Compute $P(R_1 \mid u_1, u_2)$ as follows:

$$u_1 = true$$

$$u_2 = true$$



$$P(X_k \mid e_{1:t}) = \alpha P(X_k \mid e_{1:k}) P(e_{k+1:t} \mid X_k)$$

$$\begin{cases} t = 2 \\ k = 1 \end{cases}$$

$$\begin{split} \textbf{\textit{P}}(\textit{R}_1 \mid \textit{u}_1, \textit{u}_2) &= \alpha \textbf{\textit{P}}(\textit{R}_1 \mid \textit{u}_1) \textbf{\textit{P}}(\textit{u}_2 \mid \textit{R}_1) \\ &\approx \langle 0.818, 0.182 \rangle \\ \text{as computed earlier} \\ &\approx \alpha \langle 0.818, 0.182 \rangle \times \langle 0.69, 0.41 \rangle \\ &\approx \langle 0.883, 0.117 \rangle \end{split}$$

$$P(u_2 \mid R_1) = \sum_{r_2} P(u_2 \mid r_2) P(\mid r_2) P(r_2 \mid R_1)$$

$$= (0.9 \cdot 1 \cdot \langle 0.7, 0.3 \rangle) + (0.2 \cdot 1 \cdot \langle 0.3, 0.7 \rangle)$$

$$= \langle 0.69, 0.41 \rangle$$