

Homework 9

Name: 詹远瞩, Number: 17300180094

2020 年 12 月 22 日

1 1.1

$$\therefore Y = \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_m \end{pmatrix} = \begin{pmatrix} \alpha_1^T \\ \alpha_2^T \\ \vdots \\ \alpha_m^T \end{pmatrix} X = \begin{pmatrix} \alpha_{11} & \alpha_{12} & \cdots & \alpha_{1m} \\ \alpha_{21} & \alpha_{22} & \cdots & \alpha_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ \alpha_{m1} & \alpha_{m2} & \cdots & \alpha_{mm} \end{pmatrix} X \triangleq AX$$

$$\therefore X = A^T Y = \begin{pmatrix} \alpha_{11} & \alpha_{21} & \cdots & \alpha_{m1} \\ \alpha_{12} & \alpha_{22} & \cdots & \alpha_{m2} \\ \vdots & \vdots & \ddots & \vdots \\ \alpha_{1m} & \alpha_{2m} & \cdots & \alpha_{mm} \end{pmatrix} Y$$

$$\therefore X_i = \sum_{k=1}^m \alpha_{ki} Y_k$$

$$\text{Var}(X_i) = \text{Var}\left(\sum_{k=1}^m \alpha_{ki} Y_k\right)$$

$$= \sum_{k=1}^m \alpha_{ki}^2 \text{Var}(Y_k)$$

$$= \sum_{k=1}^m \alpha_{ki}^2 \lambda_k = \sigma_{ii}$$

$$\therefore \sum_{k=1}^m \rho^2(Y_k, X_i) = \sum_{k=1}^m \frac{\lambda_k \alpha_{ki}^2}{\sigma_{ii}}$$

$$= \frac{\sum_{k=1}^m \lambda_k \alpha_{ki}^2}{\sigma_{ii}}$$

$$= 1$$

2 1.2

首先对样本阵 X 进行标准化处理:

$$\begin{aligned}\bar{x}_1 &= \frac{\sum_{i=1}^n x_{1i}}{n} = 4 & s_{11} &= \frac{\sum_{i=1}^n (x_{1i} - \bar{x}_1)^2}{n-1} = \frac{16}{5} \\ \bar{x}_2 &= \frac{\sum_{i=1}^n x_{2i}}{n} = 5 & s_{22} &= \frac{\sum_{i=1}^n (x_{2i} - \bar{x}_2)^2}{n-1} = 4 \\ x_{ij}^* &= \frac{x_{ij} - \bar{x}_i}{\sqrt{s_{ii}}} \\ \therefore X^* &= (x_{ij}^*) = \begin{pmatrix} -\frac{\sqrt{5}}{2} & -\frac{\sqrt{5}}{4} & -\frac{\sqrt{5}}{4} & 0 & \frac{\sqrt{5}}{4} & \frac{3\sqrt{5}}{4} \\ -\frac{3}{2} & -\frac{1}{2} & 0 & 0 & \frac{1}{2} & \frac{3}{2} \end{pmatrix}\end{aligned}$$

为方便起见, 下面的 X 代表已进行标准化的样本阵

$$S = \frac{XX^T}{n-1} = \begin{pmatrix} 1 & \frac{17\sqrt{5}}{40} \\ \frac{17\sqrt{5}}{40} & 1 \end{pmatrix}$$

对 S 进行特征值分解, 得到:

$$\begin{aligned}\lambda_1 &= \frac{40 + 17\sqrt{5}}{40} = 1.95 & \alpha_1 &= \left(\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}}\right)^T \\ \lambda_2 &= \frac{40 - 17\sqrt{5}}{40} = 0.05 & \alpha_2 &= \left(\frac{1}{\sqrt{2}}, -\frac{1}{\sqrt{2}}\right)^T \\ \therefore y_1 &= \frac{1}{\sqrt{2}}x_1 + \frac{1}{\sqrt{2}}x_2 & y_2 &= \frac{1}{\sqrt{2}}x_1 - \frac{1}{\sqrt{2}}x_2 \\ \frac{\lambda_1}{\lambda_1 + \lambda_2} &= 0.975\end{aligned}$$

样本的绝大部分信息可以由 y_1 来表示

3 1.3

只要证明对 S 的任一元素 S_{ij} 都有 $\mathbb{E}S_{ij} = \Sigma_{ij}$ 即可

$$\begin{aligned}
S_{ij} &= \frac{1}{n-1} \sum_{k=1}^n (x_{ik} - \bar{x}_i)(x_{jk} - \bar{x}_j) \\
&= \frac{1}{n-1} \sum_{k=1}^n (x_{ik} - \mu_i + \mu_i - \bar{x}_i)(x_{jk} - \mu_j + \mu_j - \bar{x}_j) \\
&= \frac{1}{n-1} \sum_{k=1}^n ((x_{ik} - \mu_i)(x_{jk} - \mu_j) + (x_{ik} - \mu_i)(\mu_j - \bar{x}_j) + (\mu_i - \bar{x}_i)(x_{jk} - \mu_j) + (\mu_i - \bar{x}_i)(\mu_j - \bar{x}_j)) \\
&= \frac{1}{n-1} \left(\sum_{k=1}^n (x_{ik} - \mu_i)(x_{jk} - \mu_j) + (\mu_j - \bar{x}_j) \sum_{k=1}^n (x_{ik} - \mu_i) + (\mu_i - \bar{x}_i) \sum_{k=1}^n (x_{jk} - \mu_j) + n(\mu_i - \bar{x}_i)(\mu_j - \bar{x}_j) \right)
\end{aligned}$$

$$\begin{aligned}
\mathbb{E} \sum_{k=1}^n (x_{ik} - \mu_i)(x_{jk} - \mu_j) &= \sum_{k=1}^n \mathbb{E}(x_{ik} - \mu_i)(x_{jk} - \mu_j) \\
&= \sum_{k=1}^n Cov(X_i, X_j) = n\Sigma_{ij} \\
\mathbb{E}(\mu_j - \bar{x}_j) \sum_{k=1}^n (x_{ik} - \mu_i) &= -\frac{1}{n} \sum_{k=1}^n \sum_{l=1}^n \mathbb{E}(x_{jl} - \mu_j)(x_{ik} - \mu_i) \\
&= -\frac{1}{n} \sum_{k=1}^n \mathbb{E}(x_{jk} - \mu_j)(x_{ik} - \mu_i) \\
&= -\Sigma_{ij} \\
\mathbb{E}(\mu_i - \bar{x}_i) \sum_{k=1}^n (x_{jk} - \mu_j) &= -\frac{1}{n} \sum_{k=1}^n \sum_{l=1}^n \mathbb{E}(x_{il} - \mu_i)(x_{jk} - \mu_j) \\
&= -\frac{1}{n} \sum_{k=1}^n \mathbb{E}(x_{ik} - \mu_i)(x_{jk} - \mu_j) \\
&= -\Sigma_{ij} \\
\mathbb{E}n(\mu_i - \bar{x}_i)(\mu_j - \bar{x}_j) &= \frac{1}{n} \sum_{k=1}^n \sum_{l=1}^n \mathbb{E}(x_{il} - \mu_i)(x_{jk} - \mu_j) \\
&= \frac{1}{n} \sum_{k=1}^n \mathbb{E}(x_{ik} - \mu_i)(x_{jk} - \mu_j) \\
&= \Sigma_{ij}
\end{aligned}$$

$$\begin{aligned}\therefore \mathbb{E}S_{ij} &= \frac{1}{n-1}(n\Sigma_{ij} - \Sigma_{ij} - \Sigma_{ij} + \Sigma_{ij}) \\ &= \Sigma_{ij}\end{aligned}$$

因此 S 是 Σ 的无偏估计

4 1.4

设 $A \in \mathbb{R}^{m \times n}, m \geq n$, A 的奇异值分解为 $A = U\Sigma V^T = \sum_{i=1}^n \sigma_i u_i v_i^T$, 令 $A_r = \sum_{i=1}^r \sigma_i u_i v_i^T$, 我们首先证明: $\arg \min_{\text{rank}(L) \leq r} \|A - L\|_2 = A_r$, $\arg \min_{\text{rank}(L) \leq r} \|A - L\|_F = A_r$:

由于 $r(L) \leq r$, 所以 $\exists X, Y \in \mathbb{R}^{n \times r}$ 满足 $L = XY^T$, 由于 Y 的秩最多为 r , 所以 $\exists w = \sum_{i=1}^{r+1} a_i v_i$ 满足 $Y^T w = 0, w^T w = 1$

$$\begin{aligned}\|A - L\|_2^2 &= \max_{\|x\|_2=1} \|(A - L)x\|_2^2 \\ &\geq \|(A - L)w\|_2^2 \\ &= \|Aw - XY^T w\|_2^2 \\ &= \|Aw\|_2^2 \\ \because w^T w &= 1 \quad \therefore \sum_{i=1}^{r+1} a_i^2 = 1 \\ \therefore Aw &= \left(\sum_{i=1}^n \sigma_i u_i v_i^T\right) \left(\sum_{j=1}^{r+1} a_j v_j\right) \\ &= \sum_{i=1}^{r+1} a_i \sigma_i u_i \\ \therefore \|Aw\|_2^2 &= \sum_{i=1}^{r+1} a_i^2 \sigma_i^2 \\ &\geq \sigma_{r+1}^2 \\ \therefore \|A - A_r\|^2 &= \sigma_{r+1}^2 \\ \therefore \arg \min_{\text{rank}(L) \leq r} \|A - L\|_2 &= A_r\end{aligned}$$

设有矩阵 C, D, C_k, D_k , 其中 $C = \sum_{i=1}^n \sigma_i u_i v_i^T, C_k = \sum_{i=1}^k \sigma_i u_i v_i^T$, C_k 是 C 的截断奇异值分解, 类似定义 D 和 D_k , 则 $\sigma_i(C) = \sigma_1(C - C_{i-1})$, 其中 $\sigma_i(C)$ 代表 C 的第 i 大的奇异值

$$\begin{aligned}\sigma_i(C) + \sigma_j(D) &= \sigma_1(C - C_{i-1}) + \sigma_1(D - D_{j-1}) \\ &= \|C - C_{i-1}\|_2 + \|D - D_{j-1}\|_2 \\ &\geq \sigma_1((C + D) - (C_{i-1} + D_{j-1}))\end{aligned}$$

记 $A = C + D$

$$\begin{aligned}\sigma_1((C + D) - (C_{i-1} + D_{j-1})) &= \sigma_1(A - (C_{i-1} + D_{j-1})) \\ \because \text{rank}(C_{i-1} + D_{j-1}) &\leq \text{rank}(C_{i-1}) + \text{rank}(D_{j-1}) \leq i + j - 2 \\ \therefore \sigma_1(A - (C_{i-1} + D_{j-1})) &\geq \sigma_1(A - A_{i+j-2})\end{aligned}$$

令 $C = A - L, D = L, i \geq 1, j = r + 1$

$$\begin{aligned}\sigma_i(A - L) + \sigma_{r+1}(L) &\geq \sigma_1(A - A_{i+r-1}) \\ \because \text{rank}(L) &\leq r \quad \therefore \sigma_{r+1}(L) = 0 \\ \therefore \sigma_i(A - L) &\geq \sigma_1(A - A_{i+r-1}) \\ \therefore \sigma_i(A - L) &\geq \sigma_{r+i}(A) \\ \therefore \|A - L\|_F^2 &= \sum_{i=1}^n \sigma_i^2(A - L) \\ &\geq \sum_{i=1}^{n-r} \sigma_{r+i}^2(A - L) \\ &= \sum_{i=r+1}^n \sigma_i^2(A) \\ &= \|A - A_r\|_F^2 \\ \therefore \arg \min_{\text{rank}(L) \leq r} \|A - L\|_F &= A_r\end{aligned}$$

设 X 的奇异值分解为 $X = U\Sigma V^T$, X 的奇异值由大到小为: $\sigma_1 \geq \sigma_2 \geq \dots \geq \sigma_m \geq 0$, 则使该最优化问题得到最优解的 $L = \sum_{i=1}^k \sigma_i u_i v_i^T$, 记 $U_k = (u_1, u_2, \dots, u_k)$ 是 U 的前 k 列构成的截断矩阵, $V_k = (v_1, v_2, \dots, v_k)$ 是 V 的前 k 列构成的截断矩阵, $\Sigma_k = \text{diag}\{\sigma_1, \sigma_2, \dots, \sigma_k\}$, 则 $L = U_k \Sigma_k V_k^T$, 而根据教材中的奇异值分解求主成分的算法, 最终的主成分为 $Y = V_k^T X^T = \Sigma_k U_k^T$, L 与 Y 只相差一个正交变换, 因此两者包含的信息是相同的, 所以主成分等价于求此最优化问题。