Monte Carlo Simulations

James M. Flegal

Agenda

- Ordinary Monte Carlo
- Examples
- Monte Carlo integration
- Bootstrap
- ► Toy collector exercise

The "Monte Carlo method" refers to the theory and practice of learning about probability distributions by simulation rather than calculus. In ordinary Monte Carlo (OMC) we use IID simulations from the distribution of interest. Suppose X_1, X_2, \ldots are IID simulations from some distribution, and suppose we want to know an expectation

$$\theta = E[Y_1] = E[g(X_1)].$$

The law of large numbers (LLN) then says

$$\bar{y}_n = \frac{1}{n} \sum_{i=1}^n Y_i = \frac{1}{n} \sum_{i=1}^n g(X_i)$$

converges in probability to θ .

The central limit theorem (CLT) says

$$\frac{\sqrt{n}(\bar{y}_n-\theta)}{\sigma}\stackrel{d}{\to} \mathsf{N}(0,1).$$

That is, for sufficiently large n,

$$\bar{y}_n \sim N(\theta, \sigma^2/n)$$
.

Further, we can estimate the standard error σ/\sqrt{n} with s_n/\sqrt{n} where s_n is the sample standard deviation.

We can also use the CLT form a confidence interval with

$$Pr(\bar{y}_n - 1.96s_n/\sqrt{n} < EY_1 < \bar{y}_n + 1.96s_n/\sqrt{n}) \approx 0.95.$$

Or we could simulate until a half-width (or width) of this confidence interval is sufficiently small, say less than $\epsilon>0$. That is, simulate until

$$1.96s_n/\sqrt{n}<\epsilon.$$

The theory of OMC is just the theory of frequentist statistical inference. The only differences are that

- ▶ the "data" $X_1, ..., X_n$ are computer simulations rather than measurements on objects in the real world
- ▶ the "sample size" *n* is the number of computer simulations rather than the size of some real world data
- the unknown parameter θ is in principle completely known, given by some integral, which we are unable to do.

Everything works just the same when the data X_1, X_2, \ldots , which are computer simulations are vectors. But the functions of interest $g(X_1), g(X_2), \ldots$ are scalars.

OMC works great, but it can be very difficult to simulate IID simulations of random variables or random vectors whose distribution is not brand name distributions

Approximating the Binomial

Suppose we flip a coin 10 times and we want to know the probability of getting more than 3 heads. This is trivial for the Binomial distribution, which we'll ignore.

```
runs <- 10000
one.trial <- function(){
    sum(sample(c(0,1),10,replace=T)) > 3
}
mc.binom <- sum(replicate(runs,one.trial()))/runs
mc.binom</pre>
```

```
## [1] 0.8339
pbinom(3,10,0.5,lower.tail=FALSE)
```

[1] 0.828125

Exercise: Program this example and estimate the Monte Carlo standard error

Approximating π

- Area of a circle is πr^2
- If we draw a square containing that circle its area will be $4r^2$
- ▶ So the ratio of the area of the circle to the area of the square is

$$\frac{\pi r^2}{4r^2} = \frac{\pi}{4}$$

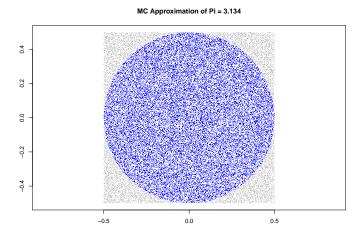
- ▶ Given this fact, we can empirically determine the ratio of the area of the circle to the area of the square we can simply multiply this number by 4 and we'll get our approximation of π .
- ► How?

Approximating π

- ▶ Randomly sample x and y values on the unit square centered at 0
- ▶ If $x^2 + y^2 \le .5^2$ then the point is in the circle
- ▶ The ratio of points in the circle multiplied by 4 is our estimate of π

```
runs <- 50000
xs <- runif(runs,min=-0.5,max=0.5)
ys <- runif(runs,min=-0.5,max=0.5)
in.circle <- xs^2 + ys^2 <= 0.5^2
mc.pi <- (sum(in.circle)/runs)*4
```

Approximating π



Example: Integration

Let $X \sim \Gamma(3/2, 1)$, i.e.

$$f(x) = \frac{2}{\sqrt{\pi}}\sqrt{x}e^{-x}I(x>0).$$

Suppose we want to find

$$\theta = \mathsf{E}\left[\frac{1}{(X+1)\log(X+3)}\right]$$
$$= \int_0^\infty \frac{1}{(x+1)\log(x+3)} \frac{2}{\sqrt{\pi}} \sqrt{x} e^{-x} dx.$$

- ▶ The expectation (or integral) θ is intractable, we don't know how to compute it analytically
- ► Further, suppose we want to estimate this quantity such that a 95% CI length is less than 0.002

Example: Integration

[1] 0.3456032 0.3702658

```
n <- 1000
x <- rgamma(n, 3/2, scale=1)
mean(x)

## [1] 1.524377
y <- 1/((x+1)*log(x+3))
est <- mean(y)
est

## [1] 0.3579345
mcse <- sd(y) / sqrt(length(y))
interval <- est + c(-1,1)*1.36*mcse
interval</pre>
```

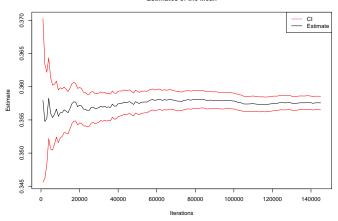
Example: Sequential stopping rule

```
eps <- 0.002
len <- diff(interval)
plotting.var <- c(est, interval)
while(len > eps){
    new.x \leftarrow rgamma(n, 3/2, scale=1)
    new.y <- 1/((\text{new.x+1})*\log(\text{new.x+3}))
    y <- cbind(y, new.y)
    est <- mean(y)
    mcse <- sd(v) / sgrt(length(v))
    interval <- est + c(-1.1)*1.96*mcse
    len <- diff(interval)
    plotting.var <- rbind(plotting.var, c(est, interval))</pre>
list(interval, length(y))
## [[1]]
## [1] 0.3565673 0.3585629
##
## [[2]]
## [1] 145000
temp <- seq(1000, length(y), 1000)
```

Example: Sequential stopping rule

```
plot(temp, plotting.var[,1], type="l", ylim=c(min(plotting.var), max(plotting.var)),
    main="Estimates of the Mean", xlab="Iterations", ylab="Estimate")
points(temp, plotting.var[,2], type="l", col="red")
points(temp, plotting.var[,3], type="l", col="red")
legend("topright", legend=c("CI", "Estimate"), lty=c(1,1), col=c(2,1))
```

Estimates of the Mean



High-dimensional examples

- ► FiveThirtyEight's Election Forecast
- ► FiveThirtyEight also predicts NBA, MLB, NCAAF, NFL, ... for the regular season and playoffs
- Vanguard's Retirement Nest Egg Calculator
- ► Fisher's Exact Test in R

Permutations with sample()

- sample() is powerful it works on any object that has a defined length().
- Permutations

```
## [1] 5 4 3 2 1

sample(1:6)

## [1] 6 2 4 5 3 1

replicate(3,sample(c("Curly","Larry","Moe","Shemp")))

## [,1] [,2] [,3]

## [1,] "Shemp" "Shemp" "Moe"

## [2,] "Curly" "Larry" "Curly"

## [4,] "Larry" "Curly" "Shemp"
```

Resampling with sample()

 Resampling from any existing distribution gives bootstrap estimators

```
bootstrap.resample <- function (object) sample (object, length(object), replace=TRUE)

## [,1] [,2] [,3] [,4] [,5]

## [1,] 9 7 9 7 9

## [2,] 6 6 6 8 9

## [3,] 9 10 8 8 10

## [4,] 9 8 10 8 8 8

## [5,] 10 9 7 10 7
```

▶ Recall: the *jackknife* removed one point from the sample and recalculated the statistic of interest. Here we resample the same length with replacement.

Bootstrap test

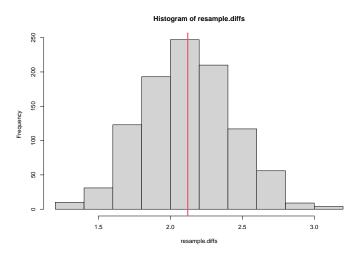
- ► The 2-sample t-test checks for differences in means according to a known null distribution.
- ▶ Let's resample and generate the sampling distribution under the bootstrap assumption.

```
library(MASS)

## Warning: package 'MASS' was built under R version 4.0.2

diff.in.means <- function(df) {
    mean(df[df$Sex=="M","Hwt"]) - mean(df[df$Sex=="F","Hwt"])
}
resample.diffs <- replicate(1000, diff.in.means(cats[bootstrap.resample(1:nrow(cats)),]))</pre>
```

Bootstrap test



Summary

- Ordinary Monte Carlo
- Repeated random sampling to obtain numerical results
- Using randomness to solve problems
- Most useful when it is difficult or impossible to use other approaches
- Can you solve The Riddler?

Exercise: Toy Collector

Children are frequently enticed to buy cereal in an effort to collect all the action figures. Assume there are 15 figures and each box contains exactly one with each figure being equally likely.

- ▶ Find the expected number of boxes needed to collect all 15.
- ► Find the standard deviation of the number of boxes needed to collect all 15 action figures.
- ▶ Now suppose we no longer have equal probabilities, instead let

Figure	Α	В	С	D	Ε	F	G	Н	I	J	K	L	М	N	0
Probability	.2	.1	.1	.1	.1	.1	.05	.05	.05	.05	.02	.02	.02	.02	.02

- ▶ Estimate the expected number of boxes needed to collect all 15.
- What is the uncertainty of your estimate?
- ▶ What is the probability you bought more than 50 boxes? 100 boxes? 200 boxes?