

Joint Confidence Regions for Rankings based on Correlated Estimates

Matala, Shaine Rosewel

University of the Philippines

November 25, 2025

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- A direct assessment of the uncertainty in the estimated overall ranking would simultaneously involve all units and their relative standing to each other

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$$r_k = \sum_{j=1}^K I(\theta_j \leq \theta_k) = 1 + \sum_{j:j \neq k} I(\theta_j \leq \theta_k), \quad \text{for } k = 1, 2, \dots, K.$$

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- derive a joint confidence region for $\theta_1, \theta_2, \dots, \theta_K$
- obtain joint confidence intervals for the r_1, r_2, \dots, r_K using a result from Klein et al. (2020).

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Suppose that for each $k \in \{1, 2, \dots, K\}$ there exists values L_k and U_k ST

$$\theta_k \in (L_k, U_k), k = 1, 2, \dots, K. \quad (2)$$

If the condition in (2) holds, the main result from Klein et al. (2020) gives a range for the value of r_k for each $k \in \{1, 2, \dots, K\}$ as follows:

$$r_k \in \{|\Lambda_{Lk}| + 1, |\Lambda_{Lk}| + 2, |\Lambda_{Lk}| + 3, \dots, |\Lambda_{Lk}| + |\Lambda_{Ok}| + 1\} \quad (3)$$

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Other related studies

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Motivation

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Joint Confidence Regions for Rankings

Political setting

David & Legara (2015) demonstrated that candidates with a name-recall advantage, such as media celebrities, incumbents, and members of dynastic families, received majority of the votes in the 2010 senatorial elections.

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Measurement across geographies

- Klein et al. (2020) also noted that states with large unpopulated land areas and relatively few high-density population centers tend to report shorter travel times while longer travel times are typically observed in highly urbanized states with large populations and high population densities.

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Objective

This research aims to do the following:

- Develop a procedure to construct joint confidence intervals for the ranks and the ranked parameters when the estimates to be ranked may be correlated.

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Definitions and Assumptions

- Define $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_K)'$ and assume that $\hat{\theta} \sim N(\theta, \Sigma)$ where $\theta = (\theta_1, \theta_2, \dots, \theta_K)'$ is unknown and Σ is a known $K \times K$ positive definite matrix. The diagonal elements of Σ are $\sigma_1^2, \dots, \sigma_K^2$.

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Joint Confidence Regions for Rankings

Derive simultaneous confidence intervals for $\theta_1, \theta_2, \dots, \theta_K$ of the form
 $\mathfrak{R}_1 = [\hat{\theta}_1 \pm t \times \sigma_1] \times [\hat{\theta}_2 \pm t \times \sigma_2] \times \dots \times [\hat{\theta}_K \pm t \times \sigma_K]$. (6)

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- ② Once the confidence intervals in (6) have been obtained, we can then use the result of Klein et al. (2020) in (5) to get the lower and upper bounds on the ranks $r_k, k = 1, 2, \dots, K$. That is, we also get a joint confidence region for r_1, r_2, \dots, r_K .

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Proposed methodology to compute the joint confidence region for the unordered parameters:

Algorithm 1

Let the data be represented by $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_K)'$ and suppose that Σ is known

- 1: **for** $b = 1, 2, \dots, B$ **do**
- 2: Generate $\hat{\theta}_b^* \sim N_K(\hat{\theta}, \Sigma)$ and write $\hat{\theta}_b^* = (\hat{\theta}_{b1}^*, \hat{\theta}_{b2}^*, \dots, \hat{\theta}_{bK}^*)'$
- 3: Compute

$$t_b^* = \max_{1 \leq k \leq K} \left| \frac{\hat{\theta}_{bk}^* - \hat{\theta}_k}{\sigma_k} \right|$$

- 4: **end for**
- 5: Compute the $(1 - \alpha)$ -sample quantile of $t_1^*, t_2^*, \dots, t_B^*$, call this \hat{t} .
- 6: The joint confidence region for $\theta = (\theta_1, \theta_2, \dots, \theta_K)'$ is given by
$$\mathfrak{R}_1 = [\hat{\theta}_1 \pm \hat{t} \times \sigma_1] \times [\hat{\theta}_2 \pm \hat{t} \times \sigma_2] \times \dots \times [\hat{\theta}_K \pm \hat{t} \times \sigma_K].$$

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Joint Confidence Regions for Rankings

We want the joint confidence region in (6) to satisfy the following probability condition:

$$P(\hat{\theta}_k - t \cdot \sigma_k \leq \theta_k \leq \hat{\theta}_k + t \cdot \sigma_k, \forall k = 1, 2, \dots, K) = 1 - \alpha. \quad (7)$$

Equivalently, we require

$$P\left(\max_{k=1,2,\dots,K} \left|\frac{\hat{\theta}_k - \theta_k}{\sigma_k}\right| \leq t\right) = 1 - \alpha. \quad (8)$$

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Algorithm 1: Quantile Calculation

Proposed methodology to compute a joint confidence region for the ordered parameters

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- 2: Generate $\hat{\theta}_b^* = (\hat{\theta}_{b1}^*, \hat{\theta}_{b2}^*, \dots, \hat{\theta}_{bK}^*)'$ $\sim N_K(\hat{\theta}, \Sigma)$ and let $\hat{\theta}_{b(1)}^*, \hat{\theta}_{b(2)}^*, \dots, \hat{\theta}_{b(K)}^*$ be the corresponding ordered values
- 3: Compute $\hat{\sigma}_{b(k)}^*$ using:
 - asymptotic variance definition
 - second-level bootstrap
- 4: Compute

$$t_b^* = \max_{1 \leq k \leq K} \left| \frac{\hat{\theta}_{b(k)}^* - \hat{\theta}_k^*}{\hat{\sigma}_{b(k)}^*} \right|$$

- 5: **end for**
- 6: Compute the $(1 - \alpha)$ -sample quantile of t_1^*, \dots, t_B^* , call this \hat{t} .
- 7: The joint confidence region of $\theta_{(1)}, \theta_{(2)}, \dots, \theta_{(K)}$ is given by
$$\mathfrak{R}_2 = [\hat{\theta}_{(1)} \pm \hat{t} \times \hat{\sigma}_{(1)}] \times [\hat{\theta}_{(2)} \pm \hat{t} \times \hat{\sigma}_{(2)}] \times \cdots \times [\hat{\theta}_{(K)} \pm \hat{t} \times \hat{\sigma}_{(K)}]$$

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- 3: Compute $\hat{\sigma}_{b(k)}^*$ using:
 - asymptotic variance definition
 - second-level bootstrap
- 4: Compute $t_b^* = \max_{1 \leq k \leq K} \left| \frac{\hat{\theta}_{b(k)}^* - \hat{\theta}_k^*}{\hat{\sigma}_{b(k)}^*} \right|$
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Algorithm 2: Asymptotic Definition of Variance

$$\hat{\sigma}_{b(k)}^* = \sqrt{\left[\text{kth ordered value among } \{ \hat{\theta}_{b1}^{*2} + \sigma_1^2, \dots, \hat{\theta}_{bK}^{*2} + \sigma_K^2 \} \right] - \hat{\theta}_{(k)}^{*2}}$$

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Algorithm 3: Variance from Second-Level Bootstrap

```

1: for  $c = 1, 2, \dots, C$  do
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3:   Compute  $\hat{\sigma}_{b(k)}^* = \frac{\sum_{c=1}^C (\hat{\theta}_{bc(k)}^{**} - \bar{\hat{\theta}}_{b\cdot(k)}^{**})^2}{C - 1}$ ,  $\bar{\hat{\theta}}_{b\cdot(k)}^{**} = \frac{1}{C} \sum_{c=1}^C \hat{\theta}_{bc(k)}^{**}$ 
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Evaluation Algorithm

For given values of Σ and $\theta_1, \theta_2, \dots, \theta_K$ (with corresponding $\theta_{(1)}, \theta_{(2)}, \dots, \theta_{(K)}$ for rank-based methods)

- 1: **for** replications = 1, 2, ..., 5000 **do**
- 2: Generate $\hat{\theta} \sim N_K(\theta, \Sigma)$
- 3: Compute the confidence region \mathfrak{R}_1 for the unordered parameters using Algorithm 1 and the confidence region for the ordered parameters \mathfrak{R}_2 using Algorithms 2 and 3.
- 4: For the unordered parameters, check if $(\theta_1, \theta_2, \dots, \theta_K) \in \mathfrak{R}_1$ and compute T_1 , T_2 , and T_3 . For the ordered parameters, check if $(\theta_{(1)}, \theta_{(2)}, \dots, \theta_{(K)}) \in \mathfrak{R}_2$
- 5: **end for**
- 6: Compute the proportion of times that the condition in line 4 is satisfied and the average of T_1 , T_2 , and T_3 .

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Covariance Matrix Σ

- The covariance matrix Σ need not be a diagonal matrix.
- We assume that $V(\hat{\theta}) = \Sigma$ is known and express Σ as in (9), where \mathbf{R} is the population correlation matrix.

$$\Sigma = \Delta^{1/2} \mathbf{R} \Delta^{1/2}; \quad \Delta = \text{diag} \left\{ \sigma_1^2, \sigma_2^2, \dots, \sigma_K^2 \right\}. \quad (9)$$

- The diagonal elements of Σ , which are $\sigma_k^2 = V(\hat{\theta}_k)$ for $k = 1, 2, \dots, K$, are treated as known quantities in practice.

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Joint Confidence Regions for Rankings

• We assume certain correlation structures among the $\hat{\theta}$ s.

└ Correlation Structures: Equicorrelation

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- This assumes that the k variables are equally correlated, i.e., that $\rho_{jk} = \rho$ where $\rho \in [-1, 1]$ for $j \neq k \in \{1, \dots, K\}$.

$$\mathbf{R}_{\text{eq}} = (1 - \rho) \mathbf{I}_K + \rho \mathbf{1}_K \mathbf{1}'_K = \begin{bmatrix} 1 & \rho & \cdots & \rho \\ \rho & 1 & \cdots & \rho \\ \vdots & \vdots & \ddots & \vdots \\ \rho & \rho & \cdots & 1 \end{bmatrix}_{K \times K} \quad (10)$$

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- Useful in the context of pre-election surveys

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- In a block correlation matrix \mathbf{R}_{block} with G blocks, each diagonal block represents an equicorrelation structure within group g , denoted by

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where ρ_g is the within-block correlation and n_g is the number of variables in block g such that $\sum_{g=1}^G n_g = K$.

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- Useful in the mean travel time from the study of Klein et al. (2020).

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$$\rho_{\text{matern}} = \frac{2^{1-\nu}}{\Gamma(\nu)} (\kappa \parallel \mathbf{s}_i - \mathbf{s}_j \parallel)^\nu K_\nu(\kappa \parallel \mathbf{s}_i - \mathbf{s}_j \parallel)$$

where $\parallel \cdot \parallel$ denotes the Euclidean distance and K_ν is the second kind of the modified Bessel function. It has a scale parameter $\kappa > 0$ and a smoothness parameter $\nu > 0$. ρ_{matern} reduces to the exponential correlation when $\nu = 0.5$ and to Gaussian correlation function when $\nu = \infty$.

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Code blocks

Alert block

$$E = mc^2$$

Examples

Example blocks are automatically green in color

Blue block

- happens with level 2, 3 headings
- this is only true for ‘Madrid’ theme in R Markdown!!

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This works, incremental bullets

- Bullet 1

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└ This works, incremental bullets

• Bullet 1

This works, incremental bullets

- Bullet 1
- Bullet 2

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└ This works, incremental bullets

- Bullet 1
- Bullet 2

This nests, but does not increment

- Bullet 1
- Bullet 2
 - subbullet 1
 - subbullet 2

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└ This nests, but does not increment

- Bullet 1
- Bullet 2
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 - subbullet 2

This increments

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• Bullet 1

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└ This increments

- Bullet 1
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This increments too

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Algo Font Size Adjustments sample

```

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```

Today is going to be a great day

Today is...

A new beginning.

Your Reference

• slideshare

• themes

• incremental bullets

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Joint Confidence Regions for Rankings

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└ Text

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