

Generalized Linear Models for Longitudinal Data

Consider extensions of GLM to repeated measurements, or generalization of linear models for longitudinal data to non-Gaussian response variables (e.g., count-valued or binary responses).

- ▶ Marginal models: the basic premise is to make inferences about population averages.
- ▶ Mixed effects models: a subset of the regression coefficients vary from subject to subject.

Marginal Models

- ▶ The focus of marginal models is on inferences of population averages.
- ▶ Marginal models separately model the mean responses and within-subject association among the repeated responses.
- ▶ Marginal models do not require the joint distributional assumptions for the vector of responses, which may be difficult for the discrete data. The avoidance of distributional assumptions leads to a method of estimation of generalized estimating equations (GEE).

Features of Marginal Models

- ▶ Marginal expectation of Y_{ij} (i.e., μ_{ij}) depends on covariates through a known link function

$$g(\mu_{ij}) = X_{ij}\beta$$

- ▶ Marginal variance of Y_{ij} is a function of the marginal mean and a scale parameter

$$\text{var}(Y_{ij}) = \phi V(\mu_{ij})$$

- ▶ The “within-subject association” among the responses is a function of the means and of additional parameters, say α , that may need to be estimated.

$$\frac{\text{var}(Y_i) \cdot \alpha}{\text{cov}} \rightarrow \left(\right)$$

Example: Continuous Response

- ▶ $\mu_{ij} = X_{ij}\beta$ (i.e., linear regression)
 - ▶ $\text{var}(Y_{ij}) = \sigma_j^2$ (i.e., heterogenous variance for different visits, but no dependence on mean)
 - ▶ $\text{corr}(Y_{ij}, Y_{ik}) = \alpha^{|j-k|}$ ($0 \leq \alpha \leq 1$) (i.e., autoregressive correlation.
- Other correlation structures such as compound symmetry or unstructured are also possible.)

Example: Count Response

- ▶ $\log \mu_{ij} = X_{ij}\beta$ (i.e., log linear regression)
- ▶ $\text{var}(Y_{ij}) = \phi \mu_{ij}$ (i.e., Poisson variance with potential over dispersion)
- ▶ $\text{corr}(Y_{ij}, Y_{ik}) = \alpha$ (i.e., compound symmetry correlation structure)

Example: Binary Response

- ▶ $\text{logit}(\mu_{ij}) = \log \frac{\mathbb{P}(Y_{ij}=1)}{\mathbb{P}(Y_{ij}=0)} = X_{ij}\beta$ (i.e., logistic regression)
- ▶ $\text{var}(Y_{ij}) = \mu_{ij}(1 - \mu_{ij})$ (i.e., binary variance)
- ▶ $\text{corr}(Y_{ij}, Y_{ik}) = \alpha_{jk}$ (i.e., unstructured)
- ▶ Sometimes, people also use other measures (e.g., log odds ratio) to characterize associations for binary variables

Interpretation of Model Parameters

The regression parameters, β , have “population-averaged” interpretations (where “averaging” is over all individuals within subgroups of the population):

- ▶ describe effect of covariates on the average responses
- ▶ contrast the means in sub-populations that share common covariate values

For example, consider the logistic model

$$\text{logit}(\mu_{ij}) = \text{logit}(\mathbb{E}(Y_{ij}|X_{ij})) = \underline{X_{ij}\beta}$$

Handwritten annotations: An arrow points from the circled $X_{ij}\beta$ in the equation to the expression $X_{i1}\beta_1 + X_{i2}\beta_2$ above. Above $X_{i1}\beta_1$ is the word "age" with an arrow pointing to X_{i1} . Above $X_{i2}\beta_2$ is the word "sex" with an arrow pointing to X_{i2} . Arrows point from below to the underlined β_1 and β_2 terms.

Each element of β measures the change in the log odds of a “positive” response per unit change in the respective covariate, *for sub-populations defined by fixed and known covariate values.*

Parameter Estimation: GEE

GLM: $f(\theta) = y - b(\theta)$ $b'(\theta) = \mu$
 $g(\mu) = X\beta$

$\max_{\beta} f(\eta)$

Solve: $\frac{\partial f(\eta)}{\partial \beta} = 0$

$\frac{\partial f(\eta)}{\partial \beta} = \frac{\partial [y - b(\theta)]}{\partial \beta}$
 $= (y - b(\theta)) \cdot \frac{\partial \theta}{\partial \beta}$
 $= (y - b(\theta)) \frac{\partial \theta}{\partial \mu} \cdot \frac{\partial \mu}{\partial \beta}$
 $= (y - b(\theta)) \frac{1}{b'(\theta)} \cdot \frac{\partial \mu}{\partial \beta}$

$\downarrow \quad \downarrow$

$(y - \mu) \frac{1}{V} \cdot \frac{\partial \mu}{\partial \beta} = 0$

- ▶ It is difficult to derive a multivariate distribution for discrete response data.
- ▶ Thus, no “convenient” likelihood function to maximize.
- ▶ Instead, use Generalized Estimating Equations (GEE).
- ▶ No need to specify any distribution; just provide the mean function (link) and the association structure.

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$$\sum_{i=1}^m D_i^T V_i^{-1} (Y_i - \mu_i) = 0$$

where $V_i \approx \text{cov}(Y_i)$ and $D_i = \partial \mu_i / \partial \beta$

Example

Respiratory Illness

- ▶ Study how respiratory illness¹ (good or poor) is related to several factors (e.g., baseline status, age, sex, treatment, etc)⁰
- ▶ 111 subjects with 4 measurements per subject (1 baseline and 3 follow-ups)
- ▶ Each response variable is binary
- ▶ Fit marginal models with logit link and different correlation structures

GEE with logit link and exchangeable correlation structure

```
> resp_gee2 <- gee(nstat ~ centre + treatment + gender + baseline + age, data = resp, family = "binomial",  
+ id = subject, corstr = "exchangeable", scale.fix = TRUE, scale.value = 1)  
Beginning Cgee S-function, @(#) geeformula.d 4.13 98/01/27  
running glm to get initial regression estimate  
(Intercept)          centre2 treatmenttreatment      gendermale      baselinegood          age  
-0.90017133      0.67160098      1.29921589      0.11924365      1.88202860     -0.01816588  
> summary(resp_gee2)
```

GEE: GENERALIZED LINEAR MODELS FOR DEPENDENT DATA
gee S-function, version 4.13 modified 98/01/27 (1998)

Model:
Link: Logit
Variance to Mean Relation: Binomial
Correlation Structure: Exchangeable

Call:
gee(formula = nstat ~ centre + treatment + gender + baseline +
age, id = subject, data = resp, family = "binomial", corstr = "exchangeable",
scale.fix = TRUE, scale.value = 1)

Summary of Residuals:

Min	1Q	Median	3Q	Max
-0.93134415	-0.30623174	0.08973552	0.33018952	0.84307712

Coefficients:

	Estimate	Naïve S.E.	Naïve z	Robust S.E.	Robust z
(Intercept)	-0.90017133	0.4784634	-1.8813796	0.46032700	-1.9555041
centre2	0.67160098	0.3394723	1.9783676	0.35681913	1.8821889
treatmenttreatment	1.29921589	0.3356101	3.8712064	0.35077797	3.7038127
gendermale	0.11924365	0.4175568	0.2855747	0.44320235	0.2690501
baselinegood	1.88202860	0.3419147	5.5043802	0.35005152	5.3764332
age	-0.01816588	0.0125611	-1.4462014	0.01300426	-1.3969169

Estimated Scale Parameter: 1
Number of Iterations: 1

Working Correlation

	[,1]	[,2]	[,3]	[,4]
[1,]	1.0000000	0.3359883	0.3359883	0.3359883
[2,]	0.3359883	1.0000000	0.3359883	0.3359883
[3,]	0.3359883	0.3359883	1.0000000	0.3359883
[4,]	0.3359883	0.3359883	0.3359883	1.0000000

Example

Epileptic Seizure

- ▶ Clinical trial of 59 epileptics
- ▶ For each patient, the number of epileptic seizures was recorded during a baseline period of 8 weeks
- ▶ Patients were randomized to treatment with the antiepileptic drug progabide or placebo
- ▶ Number of seizures was then recorded in 4 consecutive 2-week intervals
- ▶ Question: Does seizure rate change over time? Is it related to baseline, age, or treatment group assignment?
- ▶ Question: Is the treatment effective?

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