Binary Data

Suppose Y is a binary response variable (taking values 0, 1).

 X_1,\ldots,X_p are the explanatory variables.

Goal:

model the relation between

$$\mathbb{E}(Y) = \pi = \Pr(Y = 1)$$

and X_1, \ldots, X_p , based on n independent samples.

Clinical Trial Example

Binary response variable Y: indicator of responder. Predictors $(X_1,...,X_p)$: Treatment group, Gender, etc.

Ungrouped data $(y_i, \mathbf{x}_{(i)})$:

Response	Treatment	Gender
y(1)	Α	male
y(m1)	A	male
y(m1+1)	В	male
y(m1+m2)	В	male
y(m1+m2+1	.) A	female
y(m1+m2+n	13) A	female

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Grouped representation (categorical predictors):

```
Treatment Gender
                   Group size # of responders
Α
         male
                      m 1
                                      7.1
         female
                      m2
                                      7.2
Α
В
         male
                      m3
                                      Z3
В
         female
                      m4
                                      7.4
```

- ▶ Data: $(m_i, Z_i, x_{(i)})$
- ► Even if grouped, individuals in the same group are assumed to be independent, and have the same probability of event occurrence.
- ► Estimates from the grouped and ungrouped (sparse) data are the same.
- ▶ If there are unique values (e.g., age) for each individual, data cannot be grouped.
- ▶ Note: there are differences in model diagnostics!

Binomial Distribution

For binary response $Y \in \{0,1\}$, the only distribution is the Bernoulli distribution.

▶ Bernoulli Distribution $Bin(1,\pi)$. $\mathbb{P}(Y=1)=\pi$, $\mathbb{P}(Y=0)=1-\pi$; or

$$\mathbb{P}(Y = y) = \pi^{y}(1 - \pi)^{1-y}, \ y = 0, 1.$$

If $Y_1, \ldots, Y_n \overset{i.i.d.}{\sim} Bin(1, \pi)$, then $Y = \sum_{i=1}^n Y_i$ has a Binomial distribution.

▶ Binomial Distribution: $Bin(n, \pi)$

$$\mathbb{P}(Y = y) = \binom{n}{y} \pi^{y} (1 - \pi)^{n-y}, \ y = 0, \dots, n.$$



Models for Binary data

Suppose $Y_i \sim Bin(1, \pi_i)$; X_i is a $p \times 1$ covariate vector for individual i; $\eta_i = X_i \beta$ is the linear predictor. The log-likelihood function is

$$\sum_{i=1}^n \left[Y_i \log \pi_i + (1-Y_i) \log (1-\pi_i) \right] \triangleq \sum_{i=1}^n \left[Y_i \theta_i - b(\theta_i) \right].$$

where

$$heta_i = \log rac{\pi_i}{1-\pi_i}, \ b(heta_i) = \log(1+\exp(heta_i))$$

The last piece of GLM is to link π_i with η_i , using some link function $g(\cdot)$

$$g(\pi_i) = \eta_i = X_i^T \beta$$



Models for Grouped Data

Suppose (Y_i, n_i, X_i) is the observed data, where $Y_i \sim Bin(n_i, \pi_i)$. The log likelihood function is

$$\sum_{i=1}^{m} \left\{ \log \binom{n_i}{Y_i} + \left[Y_i \log \pi_i + (n_i - Y_i) \log(1 - \pi_i) \right] \right\}$$

where

$$heta_i = \log rac{\pi_i}{1 - \pi_i}, \ b(heta_i) = n_i \log(1 + \exp(heta_i))$$

Again, we need to define a link function $g(\cdot)$ to link the mean $n_i\pi_i$ and η_i .

Link functions

$$0 \le \pi_i \le 1,$$
$$\eta_i = X_i^T \beta \in \mathbb{R},$$

which suggests that a link function must satisfy

$$g:[0,1]\mapsto\mathbb{R},$$

$$g(0) = -\infty$$
, and $g(1) = +\infty$.

Correspondingly, $g^{-1}: \mathbb{R} \mapsto [0,1]$ and monotone increasing.

1. logit/logistic (the canonical link):

$$g_1(\pi) = {b'}^{-1}(\pi) = \log rac{\pi}{1-\pi},$$
 $g_1^{-1}(\eta) = rac{e^{\eta}}{1+e^{\eta}}.$

2. probit/inverse Normal:

$$g_2(\pi) = \Phi^{-1}(\pi),$$

 $g_2^{-1}(\eta) = \Phi(\eta).$

3. complementary log-log:

$$g_3(\pi) = \log(-\log(1-\pi)),$$

 $g_3^{-1}(\eta) = 1 - e^{-e^{\eta}}.$

Probability of event vs. linear predictor

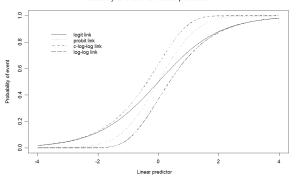


Figure: The choice of link functions depend on model fit and interpretation!

$$g_1(\pi) = -g_1(1-\pi), g_2(\pi) = -g_2(1-\pi)$$

• $g_3(\pi)$ is not symmetric.

Parameter Interpretation

Link function determines how the linear predictor exert its effect on the event rate/risk.

Consider a logistic model with two covariates X_1 and X_2 :

$$\log(\frac{\pi}{1-\pi}) = X\beta = \beta_0 + \beta_1 X_1 + \beta_2 X_2.$$

- $ightharpoonup \pi$ is usually called the **risk**
- $\blacktriangleright \pi/(1-\pi)$ is called the **odds**
- β_0 is the log odds for $X_1 = X_2 = 0$;
- β_1 is the unit change in log odds (or **log odds ratio**) per unit change of X_1 holding X_2 fixed (assuming X_1 , X_2 independent).
- if $\beta_1 > 0$, the risk increases with X_1 ; vice versa.



Examples

► Show/no-show:

Investigating the relation between show/no-show and appointment lag. $\label{eq:local_show} \ensuremath{\mathsf{lag}}.$

Peer reviewed publication:

Comparing urology fellows with and without time off in terms of their proportions of urological publications.