

Linear Models for Classification

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Classification

- Given an observation vector \mathbf{x} , assign it to one of the K classes
- The target can take one of the K discrete labels \mathcal{C}_k , $k = 1, 2, \dots, K$
 - Classify $\mathbf{x} = (45\text{Kg}, 4.8\text{ft})$ as adult or kid
 - Classify a given image as face or nonface
 - Classify a speech waveform as one of the 40 phonemes
- Discriminant function to create separating hyperplane between classes
Eg. Least squares, Fisher discriminant, perceptron, SVM
- Probabilistic approaches to estimate posterior probabilities $P[\mathcal{C}_k/\mathbf{x}]$
 - Discriminative models: Directly estimate $P[\mathcal{C}_k/\mathbf{x}] = f(\mathbf{x}, \mathbf{w})$
Eg. Logistic regression, DNN classifiers
 - Generative models: Arrive at the posterior from joint density $p(\mathbf{x}, \mathcal{C}_k)$
Eg. PDF Estimation: GMM (RV), HMM (RP)

Discriminant Functions (Binary Classification)

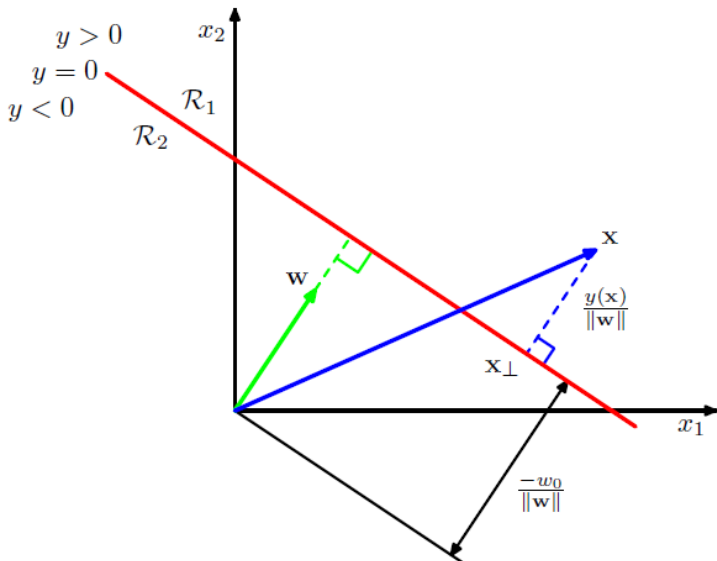
- For an input vector $\mathbf{x} \in \mathbb{R}^D$, Linear Discriminant Function is given by

$$y(\mathbf{x}) = \mathbf{w}^T \mathbf{x} + w_0$$

- Decision Rule: Assign \mathbf{x} to \mathcal{C}_1 if $y(\mathbf{x}) \geq 0$, otherwise assign \mathbf{x} to \mathcal{C}_2
- Decision boundary, $y(\mathbf{x}) = 0$, corresponds to a $D - 1$ dim. hyperplane
- Weight vector \mathbf{w} is orthogonal to every vector on the decision surface
- Weight vector \mathbf{w} determines the orientation of the decision surface
- Bias parameter w_0 determines the location of the decision surface
- Normal distance from origin to the decision surface is given by

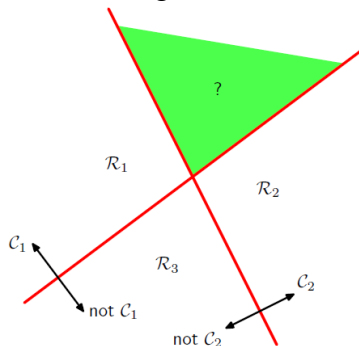
$$\frac{\mathbf{w}^T \mathbf{x}}{\|\mathbf{w}\|} = -\frac{w_0}{\|\mathbf{w}\|}$$

Geometry of Decision Boundary in 2D



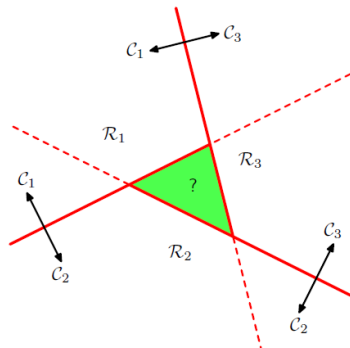
Extending to Multiple Classes

One Against Rest



$K-1$ Decision Surfaces

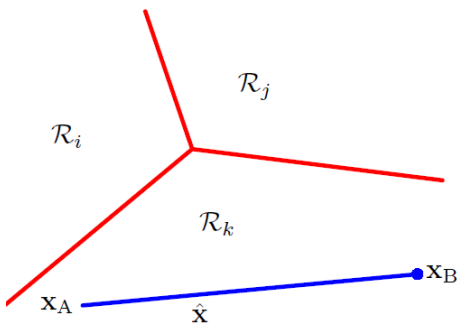
One Against One



$K(K-1)/2$ Decision Surfaces

- Both the approaches lead to ambiguous regions!
- Decision regions should be singly connected and convex.

Multiclass Linear Discriminant



- K-Linear discriminant functions

$$y_k(\mathbf{x}) = \mathbf{w}_k^T \mathbf{x} + w_{k0} \quad k = 1, 2, \dots, K$$

- Decision rule: Assign \mathbf{x} to \mathcal{C}_k if

$$y_k(\mathbf{x}) > y_j(\mathbf{x}), \quad \forall j \neq k$$

- Decision boundary \mathcal{C}_k & \mathcal{C}_j :

$$y_k(\mathbf{x}) = y_j(\mathbf{x})$$

$$(\mathbf{w}_k - \mathbf{w}_j)^T \mathbf{x} + (w_{k0} - w_{j0}) = 0$$

- Decision regions are convex
- For two classes, we can either employ a single discriminant or two discriminants

Least Squares for Classification

- Let us apply Least Squares Regression approach to classification
- Training data: $\{(\mathbf{x}_1, \mathbf{t}_1), (\mathbf{x}_2, \mathbf{t}_2), \dots, (\mathbf{x}_N, \mathbf{t}_N)\}$
 - $\mathbf{x}_n \in \mathbb{R}^D$ denotes input observation vectors
 - \mathbf{t}_n , with 1-of-K binary coding, denotes the class label (K - Classes)
- Each class \mathcal{C}_k is described by its own linear discriminant

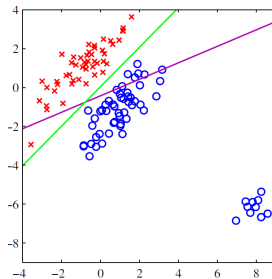
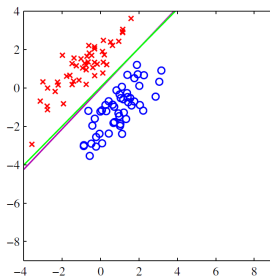
$$y_k(\mathbf{x}) = \mathbf{w}_k^\top \mathbf{x} + w_{k0} \quad k = 1, 2, \dots, K$$

- Regressed target for \mathbf{x}_n : $\hat{\mathbf{t}}_n = \mathbf{y}(\mathbf{x}_n) = \mathbf{W}^\top \mathbf{x}_n$ (Augmented)
- Estimate \mathbf{W} to minimize sum of squares error

$$J(\mathbf{W}) = \frac{1}{2} \text{Tr} \left\{ (\mathbf{XW} - \mathbf{T})^\top (\mathbf{XW} - \mathbf{T}) \right\}$$

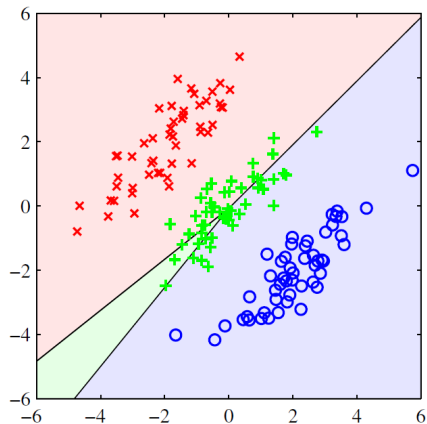
- Setting $\nabla_{\mathbf{W}} J(\mathbf{W}) = \mathbf{0}$, we get $\mathbf{W}_* = \left(\mathbf{X}^\top \mathbf{X} \right)^{-1} \mathbf{X}^\top \mathbf{T}$

Issues with LS

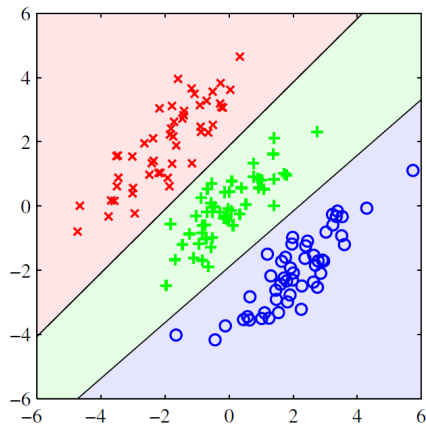


- Least squares solutions lack robustness to outliers
- Additional data-points resulted in significant change in boundary
- Sum of squares error penalizes predictions that are "too correct"
- Attempts to achieve "many-to-one" mapping through linearity!
- LS approach failed even for linearly separable classes

Issues with LS



Least Squares Classifier



Logistic Regression

Homework

- **Property of LS:** If every target in the training set satisfies some linear constraint

$$\mathbf{a}^T \mathbf{t}_n + b = 0, \quad \forall n$$

for some arbitrary constants \mathbf{a} and b , then the model prediction for any value of \mathbf{x} satisfies the same constraint.

$$\mathbf{a}^T \mathbf{y}(\mathbf{x}) + b = 0$$

- If we use 1-of-K coding for targets, then the predictions sum to 1.

$$\sum_{k=1}^K y_k(\mathbf{x}) = 1$$

- However, $y_k(\mathbf{x})$ cannot be interpreted as posterior probability. They can be negative!

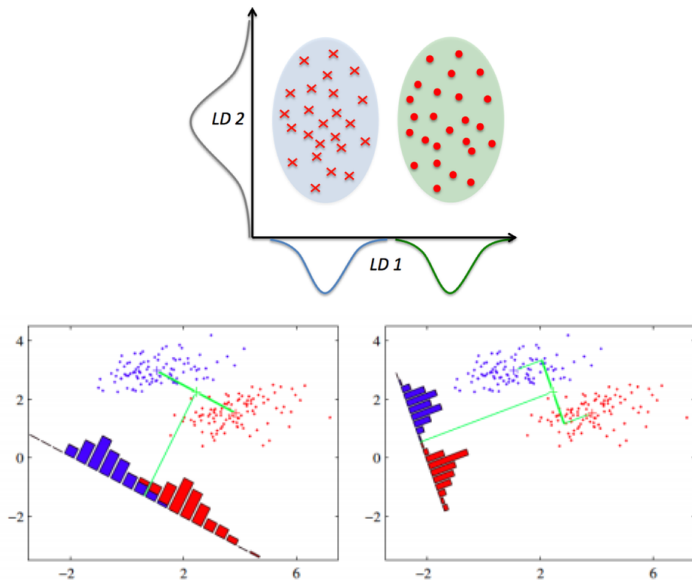
Linear Discriminant Analysis

- Dimensionality reduction interpretation to LS classifier (2 classes)
 - Project the $\mathbf{x} \in \mathbb{R}^D$ on to real line (1-D): $y(\mathbf{x}) = \mathbf{w}^T \mathbf{x}$
 - Assign \mathbf{x} to \mathcal{C}_1 if $y(\mathbf{x}) > -w_0$ and to \mathcal{C}_2 otherwise
 - Projecting to 1-D, in general, leads to loss of information
 - Adjust \mathbf{w} to select a projection that maximizes the class separation.
- Consider N_1 points from \mathcal{C}_1 and N_2 points from \mathcal{C}_2 in D -dim space
 - Mean vectors in original space: $\mathbf{m}_1 = \frac{1}{N_1} \sum_{n \in \mathcal{C}_1} \mathbf{x}_n$ $\mathbf{m}_2 = \frac{1}{N_2} \sum_{n \in \mathcal{C}_2} \mathbf{x}_n$
 - Let the projected means be $\mu_1 = \mathbf{w}^T \mathbf{m}_1$ and $\mu_2 = \mathbf{w}^T \mathbf{m}_2$
 - Choose \mathbf{w} to maximize the separation between projected means

$$\mu_2 - \mu_1 = \mathbf{w}^T (\mathbf{m}_2 - \mathbf{m}_1) \quad \text{st.} \quad \mathbf{w}^T \mathbf{w} = 1$$

- The direction for \mathbf{w} can be shown to be $\mathbf{w} \propto (\mathbf{m}_2 - \mathbf{m}_1)$
- This approach is not optimal for nondiagonal covariances

LDA - Illustration



Fisher Discriminant Analysis

- For nondiagonal covariances, spread of data should also be considered
- Project the data in a direction that
 - maximizes separation between the means of the projected classes
 - minimizes the variance within each projected class
- The objective function is given by

$$J(\mathbf{w}) = \frac{(\mu_2 - \mu_1)^2}{\sigma_1^2 + \sigma_2^2} = \frac{\mathbf{w}^T \mathbf{S}_B \mathbf{w}}{\mathbf{w}^T \mathbf{S}_W \mathbf{w}}$$

- \mathbf{S}_B is *between-class* covariance: $\mathbf{S}_B = (\mathbf{m}_2 - \mathbf{m}_1)(\mathbf{m}_2 - \mathbf{m}_1)^T$
- \mathbf{S}_W is *within-class* covariance: $\mathbf{S}_W = \sum_{k=1}^2 \sum_{n \in \mathcal{C}_k} (\mathbf{x}_n - \mathbf{m}_k)(\mathbf{x}_n - \mathbf{m}_k)^T$
- Optimal direction of \mathbf{w} can be obtained by maximizing $J(\mathbf{w})$

Solution to Fisher Criterion

- Equating $\nabla J(\mathbf{w}) = \mathbf{0}$, we obtain

$$(\mathbf{w}^T \mathbf{S}_B \mathbf{w}) \mathbf{S}_W \mathbf{w} = (\mathbf{w}^T \mathbf{S}_W \mathbf{w}) \mathbf{S}_B \mathbf{w}$$

- $\mathbf{w}^T \mathbf{S}_B \mathbf{w}$ and $\mathbf{w}^T \mathbf{S}_W \mathbf{w}$ contribute to only scalar multiples
- The direction of \mathbf{w} is given by $\mathbf{S}_W \mathbf{w} \propto \mathbf{S}_B \mathbf{w}$
- $\mathbf{S}_B \mathbf{w}$ is always in the direction of $\mathbf{m}_2 - \mathbf{m}_1$
- The direction of the fisher discriminant: $\mathbf{w} \propto \mathbf{S}_W^{-1}(\mathbf{m}_2 - \mathbf{m}_1)$
- Multiplication by \mathbf{S}_W^{-1} can be interpreted as data whitening.
- FDA also is sensitive to outliers even if they are "too correct"
- FD is strictly not a "discriminant" in true sense.
 - It just projects binary class data to 1-D
 - We need to arrive at a threshold w_0 to perform classification

Homework: Relation to Least Squares

- In LS approach, linear discriminant is determined to make model predictions as close as possible to target values
- In FDA, the discriminant is derived to achieve maximum class separation in the projected space
- If we take targets for \mathcal{C}_1 and \mathcal{C}_2 as $\frac{N}{N_1}$ and $-\frac{N}{N_2}$, respectively, where $N = N_1 + N_2$, show that LS approach yields the same solution as FD.

Extending FDA to Multiple Classes

- Assumption: Data dimensionality $D > K$ Number of classes
- For multi-class case, it is not enough to project data to 1-D
- Project the data to a lower dimensional space $D' < D$
 - Let $\mathbf{W} : \mathbb{R}^{D \times D'}$ be the linear map to achieve dimensionality reduction
 - The lower dimensional representation of $\mathbf{x}_n \in \mathbb{R}^D$ is $\mathbf{y}_n = \mathbf{W}^T \mathbf{x}_n \in \mathbb{R}^{D'}$
 - Maximize between-class spread and minimize within-class spread in the projected space
 - Let \mathbf{S}_W and \mathbf{S}_B denote within-class and between-class covariance in original D-dim space
 - Let $\boldsymbol{\Sigma}_W$ and $\boldsymbol{\Sigma}_B$ be their counterparts in projected space
 - Estimate \mathbf{W} to maximize $J(\mathbf{W}) = \text{Tr} \{ \boldsymbol{\Sigma}_W^{-1} \boldsymbol{\Sigma}_B \}$
- The columns of \mathbf{W} are given by eigenvectors corresponding to the D' largest eigenvalues of $\mathbf{S}_W^{-1} \mathbf{S}_B$

Statistics in Original D -dim Space

- Class specific statistics in the original D -dim space

$$\mathbf{m}_k = \frac{1}{N_k} \sum_{n \in \mathcal{C}_k} \mathbf{x}_n \quad \mathbf{S}_k = \sum_{n \in \mathcal{C}_k} (\mathbf{x}_n - \mathbf{m}_k)(\mathbf{x}_n - \mathbf{m}_k)^\top$$

- Define *within-class* covariance as $\mathbf{S}_W = \sum_{k=1}^K \mathbf{S}_k$
- Overall data statistics (without considering class labels)

$$\mathbf{m}_T = \frac{1}{N} \sum_{n=1}^N \mathbf{x}_n = \frac{1}{N} \sum_{k=1}^K N_k \mathbf{m}_k \quad \mathbf{S}_T = \sum_{n=1}^N (\mathbf{x}_n - \mathbf{m}_T)(\mathbf{x}_n - \mathbf{m}_T)^\top$$

- Assumption: Define *between-class* covariance using $\mathbf{S}_T = \mathbf{S}_W + \mathbf{S}_B$

$$\mathbf{S}_B = \sum_{k=1}^K N_k (\mathbf{m}_k - \mathbf{m}_T)(\mathbf{m}_k - \mathbf{m}_T)^\top$$

Statistics in Projected D' -dim Space

- The projected data-points in $\mathbb{R}^{D'}$ are given by $\mathbf{y}_n = \mathbf{W}^T \mathbf{x}_n$
- The class-specific statistics in the projected space are

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n \in \mathcal{C}_k} \mathbf{y}_n \quad \boldsymbol{\Sigma}_k = \sum_{n \in \mathcal{C}_k} (\mathbf{y}_n - \boldsymbol{\mu}_k)(\mathbf{y}_n - \boldsymbol{\mu}_k)^T$$

- Relation between the statistics of original and projected space

$$\boldsymbol{\mu}_k = \mathbf{W}^T \mathbf{m}_k \quad \boldsymbol{\Sigma}_k = \mathbf{W}^T \mathbf{S}_k \mathbf{W} \quad \boldsymbol{\Sigma}_W = \mathbf{W}^T \mathbf{S}_W \mathbf{W} \quad \boldsymbol{\Sigma}_B = \mathbf{W}^T \mathbf{S}_B \mathbf{W}$$

- Estimate \mathbf{W} to maximize $J(\mathbf{W}) = \text{Tr} \left\{ (\mathbf{W}^T \mathbf{S}_W \mathbf{W})^{-1} (\mathbf{W}^T \mathbf{S}_B \mathbf{W}) \right\}$
 - Solution leads to the famous eigenvalue problem: $\mathbf{S}_W \mathbf{W} \propto \mathbf{S}_B \mathbf{W}$
 - Eigenvectors corresponding to D' largest eigenvalues forms \mathbf{W}
- D' is bounded by rank of \mathbf{S}_B , and can be at most $K - 1$

The Perceptron

- Issues with least squares and linear discriminants
 - Linear relation cannot achieve *many-to-one* mapping
 - No inbuilt mechanism to ignore *too-correct* outliers
- Perceptron employs a *nonlinearity* to estimate target: $y(\mathbf{x}) = f(\mathbf{w}^T \mathbf{x})$
- $f(\cdot)$ is a hard-limiting nonlinear function given by

$$f(a) = \begin{cases} +1, & a \geq 0 \\ -1, & a < 0 \end{cases}$$

- Desired targets are encoded as +1 for \mathcal{C}_1 and -1 for \mathcal{C}_2 .
- Estimate \mathbf{w} s.t. $\mathbf{w}^T \mathbf{x}_n \geq 0$ for $\mathbf{x}_n \in \mathcal{C}_1$, and $\mathbf{w}^T \mathbf{x}_n < 0$ for $\mathbf{x}_n \in \mathcal{C}_2$
- Both the targets and estimates are discrete
- Number of misclassified examples is piece-wise constant

The Perceptron Learning

- Perceptron criterion: defined over set of misclassified examples \mathcal{M}

$$J_p(\mathbf{w}) = - \sum_{n \in \mathcal{M}} (\mathbf{w}^T \mathbf{x}_n) t_n$$

- SGD can be employed to update the weight vector

$$\mathbf{w}^{new} = \mathbf{w}^{old} - \eta \nabla J_p(\mathbf{w}) = \mathbf{w}^{old} + \eta \sum_{n \in \mathcal{M}} \mathbf{x}_n t_n$$

- If x_n is correctly classified, do not update the weight vector
 - If $\mathbf{x}_n \in \mathcal{C}_1$ is misclassified: $\mathbf{w}^{new} = \mathbf{w}^{old} + \mathbf{x}_n$
 - If $\mathbf{x}_n \in \mathcal{C}_2$ is misclassified: $\mathbf{w}^{new} = \mathbf{w}^{old} - \mathbf{x}_n$
- Perceptron algorithm is guaranteed to converge in finite steps, for linearly separable classes. [\[Demo\]](#)
- Perceptron algorithm is not vulnerable to *too-correct* outliers

Perceptron Convergence Theorem

- Starting from a random initial guess \mathbf{w}_0 , perform K updates:

$$\mathbf{w}_K = \mathbf{w}_{K-1} + \mathbf{x}_K t_K = \mathbf{w}_0 + \sum_{k=1}^K \mathbf{x}_k t_k$$

where \mathbf{x}_k is the randomly selected misclassified point at k^{th} iteration

- Letting $\mathbf{w}_0 = \mathbf{0}$, the norm of \mathbf{w}_K is bounded by

$$\|\mathbf{w}_K\| = \left\| \sum_{k=1}^K \mathbf{x}_k t_k \right\| \leq \sum_{k=1}^K \|\mathbf{x}_k\|$$

- Let α be the maximum norm in the training data: $\alpha = \max_n \|\mathbf{x}_n\|$
- The norm of the weight vector at K^{th} iteration is upper bounded by

$$\|\mathbf{w}_K\| \leq K\alpha$$

Perceptron Convergence Theorem

- Let \mathbf{w}_* be one of the solutions that separates classes exactly

$$\mathbf{w}_*^T \mathbf{w}_K = \sum_{k=1}^K \mathbf{w}_*^T \mathbf{x}_k t_k$$

- Since \mathbf{w}_* is a solution, $\mathbf{w}_* \mathbf{x}_n t_n \geq 0$ for all n in the dataset
- Let β be the minimum projection onto \mathbf{w}_* : $\beta = \min_n \mathbf{w}_* \mathbf{x}_n t_n$
- Norm of the weight vector \mathbf{w}_K is lower bounded by

$$\mathbf{w}_*^T \mathbf{w}_K \geq K\beta \quad \|\mathbf{w}_K\| \geq \frac{1}{\|\mathbf{w}_*\|} K^2 \beta^2$$

- The range of the norm of the weight vector is given by

$$K^2 \beta' \leq \|\mathbf{w}_K\| \leq K\alpha$$

- Lower bound is quadratic in K , upper bound is linear in K .

Limitations of Perceptron

- Perceptron algorithm converges only for linearly separable data
 - Number of iterations K could be very large
 - Error does not decrease monotonically $J_P(\mathbf{w}_{k+1}) \not\leq J_P(\mathbf{w}_k)$
 - The final solution need not be optimal- one of the decision boundaries.
- Does not converge if the classes are not linearly separable - XOR
- Does not offer probabilistic interpretation or confidence intervals.
- Cannot be readily extended to multiclass problems

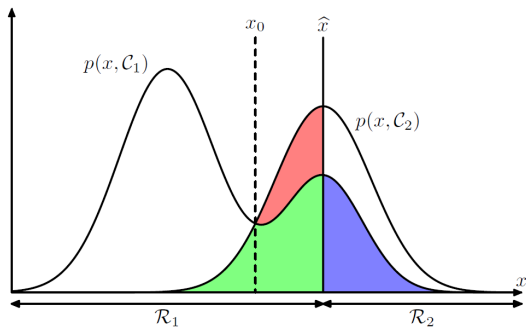
Decision Theory

- Probability theory offers a mathematical tool to deal with uncertainty.
- Supervised learning: Predict target \mathbf{t} from observed input \mathbf{x}
- In such a case, $p(\mathbf{x}, \mathbf{t})$ provides a complete summary of uncertainty
- In classification task, \mathbf{t} can take discrete labels $\mathcal{C}_k \quad k = 1, 2, \dots, K$
 - Estimating $p(\mathbf{x}, \mathcal{C}_k)$ from training set is referred to as *Inference* step
 - Assigning a test-point \mathbf{x}_t to one of the classes is *Decision* step
- Decision rule divides the input space into K decision regions \mathcal{R}_k
- Choose *decision boundaries* to minimize the probability of error
- The probability of error for binary classification is given by

$$p(\text{mistake}) = \int_{\mathcal{R}_1} p(\mathbf{x}, \mathcal{C}_2) d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathbf{x}, \mathcal{C}_1) d\mathbf{x}$$

- Divide input space into \mathcal{R}_1 and \mathcal{R}_2 s.t. $p(\text{mistake})$ is minimized

Choice of Decision Regions (Boundaries)



- Let $x = \hat{x}$ be decision boundary. $C_1 : x \leq \hat{x}$ and $C_2 : x > \hat{x}$
- Errors arise from area under the blue, red and green regions
 - Blue: Points from C_1 but misclassified as C_2
 - Green + Red: Points from C_2 , but misclassified as C_1
 - Adjust the boundary \hat{x} to minimize the area under blue+green+red
 - Area under blue+green remains constant irrespective of choice of \hat{x}

Minimizing Classification Error

$$\begin{aligned} p(\text{mistake}) &= \int_{\mathcal{R}_1} p(\mathbf{x}, C_2) d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathbf{x}, C_1) d\mathbf{x} \\ &= 1 + \int_{\mathcal{R}_1} p(\mathbf{x}, C_2) d\mathbf{x} - \int_{\mathcal{R}_1} p(\mathbf{x}, C_1) d\mathbf{x} \\ &= 1 - \int_{\mathcal{R}_1} (p(\mathbf{x}, C_1) - p(\mathbf{x}, C_2)) d\mathbf{x} \end{aligned}$$

- Choose region \mathcal{R}_1 such that $p(\mathbf{x}, C_1) > p(\mathbf{x}, C_2)$
- Decision rule: Assign \mathbf{x} to C_1 if $p[C_1/\mathbf{x}] > p[C_2/\mathbf{x}]$
- For general case of K classes, it is easier to maximize $p(\text{correct})$

$$p(\text{correct}) = \sum_{k=1}^K \int_{\mathcal{R}_k} p(\mathbf{x}, C_k) d\mathbf{x}$$

- Decision Rule: Assign \mathbf{x} to the class with highest posterior $p[C_k/\mathbf{x}]$

Expected Loss

- In reality, the errors may have varying degrees of consequences

- Penalty for Healthy vs Cancer classification could be
$$\begin{matrix} & H & C \\ \begin{matrix} H \\ C \end{matrix} & \begin{pmatrix} 0 & 1 \\ 100 & 0 \end{pmatrix} \end{matrix}$$

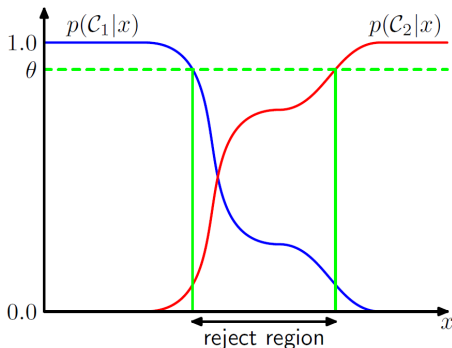
- Expected loss can be obtained by weighing error with penalty

$$\mathbb{E}[L] = \sum_k \sum_j \int_{\mathcal{R}_j} L_{kj} p(\mathbf{x}, \mathcal{C}_k) d\mathbf{x}$$

- Choose region \mathcal{R}_j to minimize $\sum_k L_{kj} p(\mathbf{x}, \mathcal{C}_k)$
- Decision Rule: Assign \mathbf{x} to \mathcal{C}_j that minimizes $\sum_k L_{kj} p[\mathcal{C}_k/\mathbf{x}]$
- This is a trivial assignment once posterior probabilities are estimated

Reject Option

- Errors arises from regions where $\max_k p[\mathcal{C}_k/\mathbf{x}] \ll 1$
 - That means, all the posteriors are in similar range
 - In those regions the classifier is relatively uncertain
- In such cases, it is better to avoid decision making
 - Reject the test samples \mathbf{x} for which $\max_k p[\mathcal{C}_k/\mathbf{x}] < \theta$



Homework: Expected Loss with Reject Option

- Consider a classification problem in which the loss incurred when an input vector from class \mathcal{C}_k is classified as belonging to class \mathcal{C}_j is given by the loss matrix L_{kj} , and for which the loss incurred in selecting the reject option is λ . Find the decision criterion that will give the minimum expected loss. Verify that this reduces to the reject criterion discussed earlier when the loss matrix is given by $L_{kj} = 1 - I_{kj}$. What is the relationship between λ and the rejection threshold θ ?

Inference & Decision Stages

- Classification problem can be broken into inference and decision stages
- Generative models
 - Inference: Estimate posterior $p[C_k/\mathbf{x}]$ from the joint density $p(\mathbf{x}, C_k)$

$$p[C_k/\mathbf{x}] = \frac{p[C_k]p(\mathbf{x}/C_k)}{p(\mathbf{x})}$$

- Such a model can generate synthetic examples of a class from $p(\mathbf{x}, C_k)$
- Discriminative models
 - Inference: Estimate posterior $p[C_k/\mathbf{x}]$ as a parametric function $y(\mathbf{x}, \mathbf{w})$
 - Decision: $p[C_k/\mathbf{x}]$ can be used for decision with loss and reject option
- Discriminant functions
 - Find a function $y(\mathbf{x}, \mathbf{w})$ that maps input \mathbf{x} to a class label
 - Inference and decision stages cannot be separated

Pros & Cons

Feature	Generate	Discriminative	Discriminant
Computation	High	Moderate	Low
Data Req.	Very high	Moderate	Low
Outlier detection	Yes $p(\mathbf{x})$	No	No
Accuracy	Reasonable	Higher	Low
Minimizing Risk	Easy	Easy	Not SF
Reject option	Easy	Easy	Not SF
Modifying Priors	Easy	Not SF	No
Model Fusion	Easy	Easy	Not SF

$$\begin{aligned}
 P[C_k/\mathbf{x}_A, \mathbf{x}_B] &\propto p[C_k]p(\mathbf{x}_A, \mathbf{x}_B/C_k) \\
 &\propto p[C_k]p(\mathbf{x}_A/C_k)p(\mathbf{x}_B/C_k) && \text{Cond. Ind.} \\
 &\propto \frac{p[C_k/\mathbf{x}_A]p[C_k/\mathbf{x}_B]}{p[C_k]}
 \end{aligned}$$

Maximum Likelihood Density Estimation

- Consider data $X = \{x_1, x_2, \dots, x_N\}$ drawn from unknown distribution
- Underlying distribution $p(x_n)$ be approximated by a parametric form
 - Gaussian assumption: $p(x_n) \sim \mathcal{N}(x_n/\mu, \sigma^2)$
 - Laplacian assumption: $p(x_n) \sim \mathcal{L}(x_n/\mu, b)$
- Assuming i.i.d, the likelihood function can be written as

$$\begin{aligned} p(X/\mu, \sigma^2) &= \prod_{n=1}^N p(x_n/\mu, \sigma^2) \\ &= \prod_{n=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_n - \mu)^2}{2\sigma^2}\right) \end{aligned}$$

- Estimate μ and σ to maximize the likelihood function, or eqv.

$$J(\mu, \sigma^2) = \log p(X/\mu, \sigma^2)$$

Parameter Estimation

- μ and σ can be determined by equating the derivatives to zero

$$\hat{\mu} = \frac{1}{N} \sum_{n=1}^N x_n \quad \hat{\sigma}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \hat{\mu})^2$$

- Checking for *bias* in estimation: Taking expectation over estimates

$$\mathbb{E}[\hat{\mu}] = \mathbb{E} \left[\frac{1}{N} \sum_{n=1}^N x_n \right] = \frac{1}{N} \sum_{n=1}^N \mathbb{E}[x_n] = \mu \quad \text{unbiased}$$

$$\mathbb{E}[\hat{\sigma}^2] = \frac{1}{N} \sum_{n=1}^N \mathbb{E}[(x_n - \hat{\mu})^2] = \frac{N-1}{N} \sigma^2 \quad \text{biased}$$

- Variance is underestimated in ML approach. $\hat{\sigma}^2 < \sigma^2$
- The variance estimation can be corrected as $\tilde{\sigma}^2 = \frac{1}{N-1} \sum_{n=1}^N (x_n - \hat{\mu})^2$

Generative Models

- Training Data: $\{(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_n, t_n), \dots, (\mathbf{x}_N, t_N)\}$
 - $\mathbf{x}_n \in \mathbb{R}^D$ are input variables, $t_n \in \{0, 1\}$ are targets
 - Let N_1 points are from \mathcal{C}_1 and N_2 points are from \mathcal{C}_2 : $N = N_1 + N_2$
- Inference stage requires estimation of posterior $P[\mathcal{C}_k/\mathbf{x}_n]$ for $k = 1, 2$
 - Priors determined from training data be: $p[\mathcal{C}_1] = \frac{N_1}{N}$ and $p[\mathcal{C}_2] = \frac{N_2}{N}$
 - Let the CCDs $p(\mathbf{x}/\mathcal{C}_k)$ are Gaussian with shared covariance

$$p(\mathbf{x}/\mathcal{C}_1) \sim \mathcal{N}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}) \quad \text{and} \quad p(\mathbf{x}/\mathcal{C}_2) \sim \mathcal{N}(\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

- ML can be used to estimate parameters $\theta = (\boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma})$

$$\boldsymbol{\mu}_1 = \frac{1}{N_1} \sum_{n \in \mathcal{C}_1} \mathbf{x}_n \quad \boldsymbol{\mu}_2 = \frac{1}{N_2} \sum_{n \in \mathcal{C}_2} \mathbf{x}_n \quad \boldsymbol{\Sigma} = \frac{N_1 \boldsymbol{\Sigma}_1 + N_2 \boldsymbol{\Sigma}_2}{N_1 + N_2}$$

- Posterior $P[\mathcal{C}_k/\mathbf{x}]$ can be evaluated from the priors and CCDs

Generative Models (Decision Boundary)

- The decision boundary is locus of points satisfying $P[C_1/\mathbf{x}] = P[C_2/\mathbf{x}]$

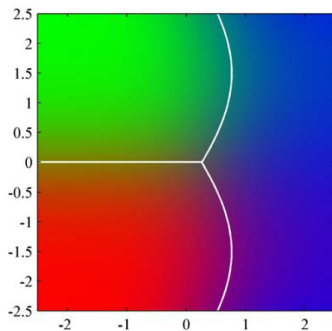
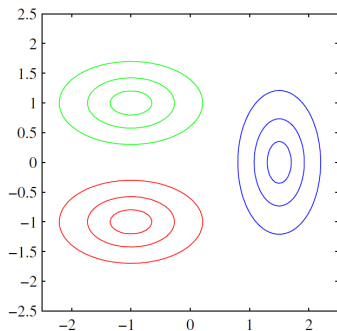
$$P[C_1] \mathcal{N}(\mathbf{x}/\boldsymbol{\mu}_1, \boldsymbol{\Sigma}) = P[C_2] \mathcal{N}(\mathbf{x}/\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

- The quadratic terms cancel because of shared covariance
- The decision boundary is a linear in \mathbf{x} : $\mathbf{w}^T \mathbf{x} + w_0$

$$\mathbf{w} = \boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2) \quad w_0 = -\frac{1}{2}\boldsymbol{\mu}_1^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_2^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_2 + \log \frac{P[C_1]}{P[C_2]}$$

- Priors affect only the bias parameters, not the orientation
- The posterior density of C_1 is given by $P[C_1/\mathbf{x}] = \sigma(\mathbf{w}^T \mathbf{x} + w_0)$
- The decision boundary would be quadratic, if covariance is not shared

Illustration of Decision Boundaries



- Red and Green classes share the same covariance matrix
 - Decision boundary is linear
- Blue has a different covariance - decision boundaries are quadratic
- Nonlinear decision boundaries can be modeled with pdfs having higher order moments!

Probabilistic Discriminative Models

- Discriminative models impose a parametric function on posterior

$$P[C_k/\mathbf{x}] = y(\mathbf{x}, \mathbf{W})$$

- The function $y(\cdot)$ should satisfy the axioms of probability
- The posterior probability of the k^{th} class is given by (Bayes)

$$P[C_k/\mathbf{x}] = \frac{P[C_k]p(\mathbf{x}/C_k)}{\sum_{j=1}^K P[C_j]p(\mathbf{x}/C_j)}$$

- Let $a_k = \log P[C_k]p(\mathbf{x}/C_k)$ be parameterized as $a_k = \mathbf{w}_k^T \mathbf{x}$
- Posterior probability can be expressed as a softmax over activations a_k

$$P[C_k/\mathbf{x}] = \frac{\exp(a_k)}{\sum_{j=1}^K \exp(a_j)} \quad \text{Softmax Fn.}$$

Binary Classifier

- For binary classifier, we need not evaluate two weight vectors.
- The posterior probability can be expressed as

$$p[\mathcal{C}_1/\mathbf{x}] = \frac{P[\mathcal{C}_1]p(\mathbf{x}/\mathcal{C}_1)}{P[\mathcal{C}_1]p(\mathbf{x}/\mathcal{C}_1) + P[\mathcal{C}_2]p(\mathbf{x}/\mathcal{C}_2)}$$

- Let $a = -\log \frac{P[\mathcal{C}_1]p(\mathbf{x}/\mathcal{C}_1)}{P[\mathcal{C}_2]p(\mathbf{x}/\mathcal{C}_2)}$ be parameterized as $a = \mathbf{w}^T \mathbf{x}$
- Posterior probability of \mathcal{C}_1 can be expressed as Sigmoid over a

$$P[\mathcal{C}_1/\mathbf{x}] = \frac{1}{1 + \exp(-a)} = \sigma(a)$$

- Posterior probability of \mathcal{C}_2 is given by $P[\mathcal{C}_2/\mathbf{x}] = 1 - \sigma(a)$
- Derivative of sigmoid function:

$$\frac{d\sigma}{da} = \sigma(a)(1 - \sigma(a))$$

Logistic Regression

- Training Data: $\{(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_n, t_n), \dots, (\mathbf{x}_N, t_N)\}$
 - $\mathbf{x}_n \in \mathbb{R}^D$ are input variables, $t_n \in \{0, 1\}$ are targets: $P[\mathcal{C}_1/\mathbf{x}_n]$
 - Let the target for \mathcal{C}_1 be 1 and \mathcal{C}_2 be 0
- Let the posterior probability be estimated as

$$\hat{P}[\mathcal{C}_1/\mathbf{x}_n] = y(\mathbf{x}_n, \mathbf{w}) = \sigma(\mathbf{w}^T \mathbf{x}_n) \quad \hat{P}[\mathcal{C}_2/\mathbf{x}_n] = 1 - \sigma(\mathbf{w}^T \mathbf{x}_n)$$

- Assuming that data points are i.i.d., the likelihood of data is given by

$$P(\mathbf{t}/\mathbf{X}, \mathbf{w}) = \prod_{n=1}^N y_n^{t_n} (1 - y_n)^{1-t_n}$$

- \mathbf{w} can be estimated by minimizing negative log of the likelihood, also referred to as *cross-entropy* loss

$$J(\mathbf{w}) = -\log P(\mathbf{t}/\mathbf{X}, \mathbf{w}) = -\sum_{n=1}^N \{t_n \log y_n + (1 - t_n) \log(1 - y_n)\}$$

Parameter Estimation

- Model parameters \mathbf{w} can be updated using gradient descent

$$\mathbf{w}^{new} = \mathbf{w}^{old} - \eta \nabla J(\mathbf{w})$$

- The 1st and 2nd order derivatives of the loss function are given by

$$\nabla J(\mathbf{w}) = \sum_{n=1}^N \mathbf{x}_n (y_n - t_n) = \mathbf{X}^T (\mathbf{y} - \mathbf{t})$$

$$\mathbf{H} = \nabla^2 J(\mathbf{w}) = \sum_{n=1}^N \mathbf{x}_n y_n (1 - y_n) \mathbf{x}_n^T = \mathbf{X}^T \mathbf{R} \mathbf{X}$$

where \mathbf{R} is a diagonal matrix with entries $R_{nn} = y_n(1 - y_n)$

- Hessian matrix varies depending on parameters \mathbf{w} through \mathbf{R}
- Since $0 < y_n < 1$, Hessian matrix \mathbf{H} is positive definite: $\mathbf{u}^T \mathbf{H} \mathbf{u} > 0$
- Error function $J(\mathbf{w})$ is convex in $\mathbf{w} \implies$ admits unique minimum

Iterative Reweighted Least Squares

$$\begin{aligned}\mathbf{w}^{new} &= \mathbf{w}^{old} - \left(\mathbf{X}^T \mathbf{R} \mathbf{X}\right)^{-1} \mathbf{X}^T (\mathbf{y} - \mathbf{t}) \\ &= \left(\mathbf{X}^T \mathbf{R} \mathbf{X}\right)^{-1} \left(\mathbf{X}^T \mathbf{R} \mathbf{X} \mathbf{w}^{old} - \mathbf{X}^T (\mathbf{y} - \mathbf{t})\right) \\ &= \left(\mathbf{X}^T \mathbf{R} \mathbf{X}\right)^{-1} \mathbf{X}^T \mathbf{R} \left(\mathbf{X} \mathbf{w}^{old} - \mathbf{R}^{-1} (\mathbf{y} - \mathbf{t})\right) \\ &= \left(\mathbf{X}^T \mathbf{R} \mathbf{X}\right)^{-1} \mathbf{X}^T \mathbf{R} \mathbf{z}\end{aligned}$$

where $\mathbf{z} = (\mathbf{X} \mathbf{w}^{old} - \mathbf{R}^{-1} (\mathbf{y} - \mathbf{t}))$.

- The solution takes the form of normal equations of weighted LS.
- However, the weighing matrix \mathbf{R} is not constant, but depends on \mathbf{w}
- Hence, the normal equations need to be applied iteratively.

Multiclass Logistic Regression (Homework)

- Consider multiclass data examples denoted by $X = \{(\mathbf{x}_{1:N}, \mathbf{t}_{1:N})\}$
 - $\mathbf{x}_n \in \mathbb{R}^D$ represents input observations
 - \mathbf{t}_n is a K -dim one-hot vector denoting the class posteriors
- For this case, the posterior probabilities can be estimated as

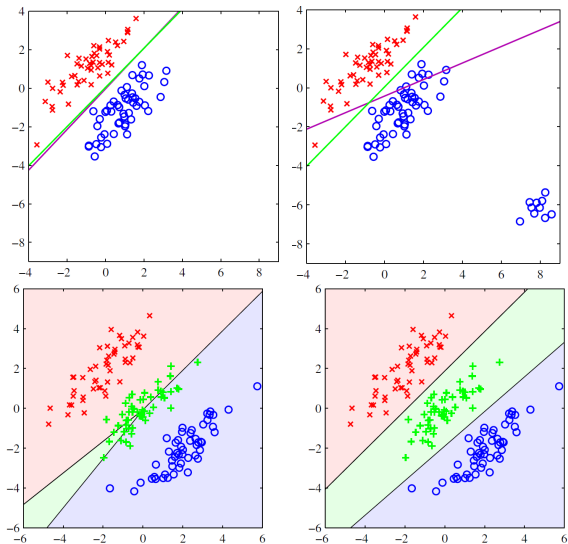
$$y(\mathbf{x}_n, \mathbf{w}_k) = \hat{P}[C_k/\mathbf{x}_n] = \frac{\exp(a_{nk})}{\sum_{j=1}^K \exp(a_{nj})} \quad a_{nk} = \mathbf{w}_k^T \mathbf{x}_n$$

- The likelihood function can be written as

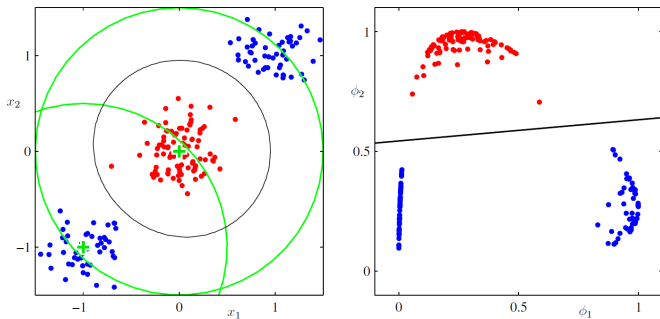
$$P[\mathbf{T}/\mathbf{X}, \mathbf{w}_1, \mathbf{w}_2, \dots, \mathbf{w}_K] = \prod_{n=1}^N \prod_{k=1}^K y_{nk}^{t_{nk}}$$

$$J(\mathbf{W}) = -P[\mathbf{T}/\mathbf{X}, \mathbf{W}] = -\sum_{n=1}^N \sum_{k=1}^K t_{nk} \log y_{nk} \quad (\text{CE})$$

Illustration of Logistic Regression



Nonlinear Decision Boundary (Transformed Space $\phi(.)$)



- Two Gaussian kernels are used to transform the data
- In general, we need to design the kernel ϕ from the data.
- DNN can be used to learn the optimal transformation from the data
- Last layer of a DNN classifier performs logistic regression

Summary of Linear Classifiers

- Assumption: Classes are separable by linear hyperplanes
- Linear discriminant functions model the separating hyperplanes
 - Least squares, Fisher discriminant, Perceptron, SVM (later)
 - May not work even if classes are separable because of outliers
- Generative models estimate posteriors from priors and CCDs
 - ML or MAP estimators are employed to model CCD
 - Don't consider other class examples while estimating CCD
- Discriminative models directly estimate posterior probabilities
 - Binary classes - logistic activation - binary cross entropy
 - Multiple classes - softmax activation - cross entropy
 - Rely on discriminative features - vulnerable to adversarial examples
- If decision boundary is not linear, apply these techniques on $\phi(\mathbf{x})$
 - Neural networks offer *a way* of learning $\phi(\mathbf{x})$ from data

Thank You!