

# AI 1103 Assignment-2

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Download all latex-tikz codes from

[https://github.com/Shantanu467/AI1103/blob/main/Assignment\\_2/Assignment\\_2.tex](https://github.com/Shantanu467/AI1103/blob/main/Assignment_2/Assignment_2.tex)

## Proof 0.1

$$E(X) = \int_{-\infty}^{+\infty} x f_X(x) dx \quad (3)$$

$$\mu = \int_{-\infty}^0 0 \times x dx + \int_0^{+\infty} \lambda e^{-\lambda x} \times x dx \quad (4)$$

$$= 0 + \left( -\frac{(\lambda x + 1) e^{-\lambda x}}{\lambda} \right) \Big|_0^{+\infty} \quad (5)$$

$$= 0 - \left( -\frac{(0 + 1) e^0}{\lambda} \right) \quad (6)$$

$$\mu = \frac{1}{\lambda} \quad (7)$$

Now, using (7) and  $\mu = 1$  (Given), we get;

$$\mu = \frac{1}{\lambda} = 1 \quad (8)$$

$$\implies \lambda = 1 \quad (9)$$

## Problem

### GATE EC: Question-74

Let  $X_1$  be an exponential random variable with mean 1 and  $X_2$  a gamma random variable with mean 2 and variance 2. If  $X_1$  and  $X_2$  are independently distributed, then  $P(X_1 < X_2)$  is equal to....

So,

$$X_1 \sim f(x; 1) = e^{-x} \quad (10)$$

**Definition 2** The probability density function (pdf) of an **gamma distribution** is

$$f(x; \lambda) = \begin{cases} \frac{a^\lambda}{\Gamma \lambda} e^{-ax} x^{\lambda-1}, & x > 0 \\ 0, & x \leq 0 \end{cases} \quad (11)$$

## Solution

**Definition 1** The probability density function (pdf) of an **exponential distribution** is Here,  $\lambda > 0$ ,  $a > 0$  and

$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0 \end{cases} \quad (1)$$

Here  $\lambda > 0$  is the parameter of the distribution, often called the rate parameter.

**Lemma 0.2** The **mean** and **variance** of an exponentially distributed random variable  $X$  with  $\lambda$  and  $a$  is given by

$$\Gamma n = a^n \int_0^\infty e^{-ax} x^{n-1} dx \quad (12)$$

$$\Gamma n = (n-1)! \quad (13)$$

**Lemma 0.1** The **mean** of an exponentially distributed random variable  $X$  with rate parameter  $\lambda$  is given by

$$E(X) = \frac{1}{\lambda} \quad (2)$$

$$\mu = \frac{\lambda}{a} \quad (14)$$

$$\sigma^2 = \frac{\lambda}{a^2} \quad (15)$$

**Proof 0.2**

Upon solving the definite integral, We get :

$$E(X) = \int_{-\infty}^{+\infty} x f_X(x) dx \quad (16)$$

$$P(X_1 < X_2) = \frac{3}{4} \quad (38)$$

$$\mu = 0 + \int_0^{\infty} x \times \frac{a^\lambda}{\Gamma\lambda} e^{ax} x^{\lambda-1} dx \quad (17)$$

$$= 0 + \frac{a^\lambda}{\Gamma\lambda} \int_0^{\infty} e^{ax} x^\lambda dx \quad (18)$$

$$\text{Now, Using } n = \lambda + 1 \text{ in (12)} \quad (19)$$

$$\text{We get: } \Gamma(\lambda + 1) = a^{\lambda+1} \int_0^{\infty} e^{ax} x^\lambda dx \quad (20)$$

$$\text{Now, } \mu = \frac{a^\lambda}{\Gamma\lambda} \times \frac{\Gamma(\lambda + 1)}{a^{\lambda+1}} \quad (21)$$

$$= \frac{1}{a} \times \frac{\Gamma(\lambda + 1)}{\Gamma\lambda} \quad (22)$$

$$\text{Using (13)} \quad (23)$$

$$= \frac{1}{a} \times \frac{\lambda!}{(\lambda - 1)!} \quad (24)$$

$$\mu = \frac{\lambda}{a} \quad (25)$$

Similarly we can find  $E(X^2)$  .

$$E(X) = \frac{\lambda(\lambda + 1)}{a^2} \quad (26)$$

$$\sigma^2 = (E(X))^2 - E(X^2) \quad (27)$$

$$= \frac{\lambda}{a^2} \quad (28)$$

Now by using (25) (28) And given info;

$$\text{Mean} = \frac{\lambda}{a} = 2 \text{ (Given)} \implies \lambda = 2a \quad (29)$$

$$\text{Variance} = \frac{\lambda}{a^2} = \frac{2a}{a^2} = 2 \text{ (Given)} \quad (30)$$

$$\implies a = 1, \lambda = 2 \quad (31)$$

$$X_2 \sim G(1, 2) = \frac{1}{\Gamma 2} e^{-x} x \quad (32)$$

$$= e^{-x} x \quad (33)$$

Now by the help of (10) and (33), we can solve for  $P(X_1 < X_2)$  .

$$P(X_1 < X_2) = P(X_1 < X_2 \mid X_1 = X_2) \quad (34)$$

$$= \int_0^{\infty} f_{X_2}(x_2) \times \int_0^{x_2} f_{X_1}(x_1) dx_1 dx_2 \quad (35)$$

$$= \int_0^{\infty} x_2 e^{-x_2} \times \int_0^{x_2} e^{-x_1} dx_1 dx_2 \quad (36)$$

$$= \int_0^{\infty} x_2 e^{-x_2} \times (1 - e^{-x_2}) dx_2 \quad (37)$$