

===== logic =====

conjunction: P and Q; **disjunction:** P or Q; **negation:** not P; **implication:** P \Rightarrow Q logically equivalent to (not P) or Q; **contrapositive:** not Q \Rightarrow not P; **converse:** Q \Rightarrow P;
Direct proof, proof by contraposition, proof by contradiction, proof by cases, induction, strong induction, well ordering principle,
Induction: Base case + Induction hypothesis (for some $k \geq 1$, assume ...) + Induction step (by the induction hypothesis).

===== number theory =====

gcd = (x, y) $\rightarrow y == 0 ? x : \text{gcd}(y, x \bmod y)$ (assume that $x \geq y \geq 0$ and $x > 0$)
def extended-gcd (x, y): return (x, 1, 0) if $y == 0$
 (d, a, b) = extended-gcd(y, x mod y); return (d, b, (a - int(x / y) * b))
RSA: Pick $N = pq$ (p, q primes), e such that $\text{gcd}(e, (p-1)(q-1)) = 1$. **public key:** (N, e);
private key: d. Property: $(x^e)^d = 1 \bmod N$
Fermat's Little Theorem: for any prime p and integer a in $\{1, 2, \dots, p-1\}$, we have $a^{(p-1)} = 1 \bmod p$. $\Rightarrow k^m = k^{(m \bmod (p-1))} \bmod p$

===== stable marriage =====

nonexistence of stable match of roommates

[A | B C D], [B | C A D], [C | A B D], [D | - - -]
 $\{(A, B), (C, D)\} \Rightarrow \{(A, D), (B, C)\} \Rightarrow \{(A, B), (C, D)\} \Rightarrow \dots$

Stable Marriage Algorithm: Every Morning: Each man goes to the first woman on his list not yet crossed off and proposes to her. **Every Afternoon:** Each woman says maybe to the man she likes best among the proposals (she now has him on a string) and never to all the rest.

Every Evening: Each of the rejected suitors crosses off the woman from his list.

Termination Lemma: guaranteed to terminate in n^2 days.

Improvement Lemma: If W has M on a string on the kth day, then on every subsequent day she has someone on a string whom she likes at least as much as M.

optimal woman: highest woman on a man's list whom he is paired with in any stable pairing.

Male optimal pairing: a stable pairing s.t. each man is paired with his optimal woman.

Theorem: The pairing output by the traditional propose & reject algorithm is male optimal.

Theorem: If a pairing is male optimal, then it is also female pessimal.

===== error correction =====

Lagrange interpolation: $\text{delta}_i(x) = \prod_{j \neq i} (x - x_j) / \prod_{j \neq i} (x_i - x_j)$

Secret sharing: at least k people \Rightarrow pick random polynomial with deg $k-1$ s.t. $P(0) = \text{secret}$. Give $P(1)$ to 1st, ..., $P(n)$ to nth.

Erasure errors: n packets, $\leq k$ packets lost \Rightarrow deg $= n-1$. send $n+k$ packets. Prime $q \geq n+k$

General errors: n packets, $\leq k$ packets lost \Rightarrow deg $= n-1$. send $n+2k$ packets. Prime $q \geq n+2k$

Error locator polynomial $E(x) = (x - e_1) \dots (x - e_k)$; $P(i)E(i) = R(i)E(i)$. Let $P(x)E(x) = Q(x)$. Solve $Q(x) = R(x)E(x)$. deg $n+k-1 + \text{deg } k$. But leading coefficient of E is 1.

$Q(x) = a_{n+k-1}x^{n+k-1} + \dots + a_0$; $E(x) = x^k + b_{k-1}x^{k-1} + \dots + b_0$

===== graph theory =====

Euler's Theorem: An undirected graph $G = (V, E)$ has an Eulerian tour if and only if the graph is connected (except possibly for isolated vertices) and every vertex has even degree. (for directed graph: for every vertex v, in-degree = out-degree)

de Bruijn sequence: a 2^n -bit **circular** sequence such that every string of length n occurs as a contiguous substring of the sequence exactly once.

de Bruijn graph: $V = \{0, 1\}^{(n-1)}$, $E =$ for each $a_1 a_2 \dots a_{n-1}$, $(a_1 a_2 \dots a_{n-1}, a_2 a_3 \dots a_{n-1} 0)$, $(a_1 a_2 \dots a_{n-1}, a_2 a_3 \dots a_{n-1} 1)$, $(0 a_1 a_2 \dots a_{n-2}, a_1 a_2 \dots a_{n-1})$, $(1 a_1 a_2 \dots a_{n-2}, a_1 a_2 \dots a_{n-1})$

n-dimensional hypercube graph: $|E| = n2^{(n-1)}$. Let $E_{\{S, V-S\}}$ denote the set of edges connecting vertices in S to vertices in $V-S$. Then $|E_{\{S, V-S\}}| \geq \min(|S|, |V-S|)$ (prove by induction on n, consider S_0 , S_1 both $\leq 2^{(n-1)}/2$; or $S_0 > 2^{(n-1)}/2$)

===== counting =====

place k balls into n bins with replacement, order does not matter: Represent each of the balls by a 0 and the separations between boxes by 1 $\Rightarrow k * 0 + (n-1) * 1 \Rightarrow n+k-1$ choose k

===== probability =====

Union bound: $\Pr[\text{union}] \leq \text{sum of } \Pr$

$E(X) = (\text{definition}) \sum_{a \in A} a * \Pr[X=a] = \sum_{w \in \Omega} X(w) * \Pr(w)$

variance: $\text{Var}(X) = E((X - E(X))^2)$ (definition) $= E(X^2) - (E(X))^2$

standard deviation: $\sigma(X) = \sqrt{\text{Var}(X)}$
covariance of X, Y: $E(XY) - E(X)E(Y)$
X,Y independent $\Rightarrow E(XY) = E(X)E(Y)$
X,Y independent $\Rightarrow \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$
Theorem: Let X be a random variable that takes on only non-negative integer values. Then $E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i]$.

===== **distribution** =====

binomial distribution: $\Pr[X=i] = \binom{n}{i} p^i (1-p)^{n-i}$. $X \sim \text{Bin}(n, p)$. $E[X] = np$.
 $E(X^2) = n^2 p^2 + np(1-p)$. $\text{Var} = np(1-p)$.
geometric distribution: $X \sim \text{Geom}(p)$. $\Pr[X=i] = (1-p)^{i-1} p$. $E[X] = 1/p$. $\Pr[X \geq i] = (1-p)^{i-1}$. $\text{Var} = (1-p)/p^2$.
Poisson distribution: $X \sim \text{Poiss}(\lambda)$. $\Pr[X=i] = \lambda^i / i! * e^{-\lambda}$. $E[X] = \lambda$.
 $E(X^2) = \lambda(\lambda+1)$, $\text{Var} = \lambda$.
fixed points in random permutation: $E(X) = 1$, $E(X^2) = 2$, $\text{Var} = 1$
random walk: $E(X) = 0$, $E(X^2) = n$, $\text{Var}(X) = n$

===== **example of probabilities** =====

error correction: encode n packets into n+k packets s.t. the recipient can reconstruct the original n packets from any n packets received. packets lost with prob p. X (packets received) $\sim \text{Bin}(n+k, 1-p)$
Hash table: n locations. A = event: insert m keys with no collisions, assuming uniformly distributed hash function.
 $\Pr[A] = (1 - 1/n) * (1 - 2/n) * \dots * (1 - (m-1)/n)$. Take \ln : $\ln(1-x) \sim -x$. $\ln(\Pr[A]) \sim -m^2/2n \Rightarrow \Pr[A] \sim e^{-m^2/2n}$
union bound: A_i : pair i has collision. $\Pr[\bar{A}] \leq \sum \Pr[A_i] = m(m-1)/2n$
throw m balls into n bins. Y = number of empty bins. $E[Y] = n(1-1/n)^m$
collect at least one copy of each of n different cards. X = boxes need to. X_i = #boxes buy while trying to get the i-th new card.
 $E(X_i) = n / (n - i + 1) \Rightarrow E(X) = n \sum_{i=1}^n 1/i \sim n (\ln n + \lambda)$.
 $\lambda = 0.5772\dots$ Euler constant

===== **approximation** =====

Markov's inequality: For a nonnegative r.v. X with $E(X) = \mu$, and any $\alpha > 0$, $\Pr[X \geq \alpha] \leq E(X) / \alpha$
proof technique: note $\sum_a a \Pr[X=a] \geq \sum_{a \geq \alpha} a \Pr[X=a]$
Chebyshev's inequality: For a random variable X with expectation $E(X) = \mu$, and for any $\alpha > 0$, $\Pr[|X-\mu| \geq \alpha] \leq \text{Var}(X) / \alpha^2$
proof technique: define r.v. $Y := (X-\mu)^2$. $\Pr[|X-\mu| \geq \alpha] = \Pr[Y \geq \alpha^2]$.
corollary: r.v. X with $E(X) = \mu$, $\sigma = \sqrt{\text{Var}(X)}$, $\Pr[|X-\mu| \geq \beta \sigma] \leq 1 / \beta^2$
i.i.d: independent, identically distributed
estimate a proportion p by taking a small sample
 $A_n = S_n / n = (X_1 + \dots + X_n) / n$
 $\text{Var}(A_n) = \text{Var}(X_i) / n = p(1-p) / n$
 $\Pr[|A_n - p| \geq \epsilon p] \leq \text{Var}(A_n) / (\epsilon p)^2 \leq \delta$
 $\Rightarrow n \geq (\sigma^2 / \mu^2) * 1 / (\epsilon^2 * \delta)$
 $\Rightarrow n \geq (1-p)/p * 1 / (\epsilon^2 * \delta)$
Law of Large Numbers: X_1, \dots, X_n i.i.d. r.v. with common $E(X_i) = \mu$. Define $A_n := 1/n \sum_i X_i$. Then for any $\alpha > 0$, we have $\Pr[|A_n - \mu| \geq \alpha] \rightarrow 0$ as $n \rightarrow \infty$

===== **conditional distribution** =====

conditional distribution: X given Y=b is collection of values $\{(a, \Pr[X=a|Y=b]): a \in A\}$
conditional expectation: $E(X|Y=b) = \sum_{a \in A} a * \Pr[X=a|Y=b]$
conditional independence: A and B are said to be conditionally independent given C if $\Pr[A, B|C] = \Pr[A|C] * \Pr[B|C]$
total expectation law: $E(X) = \sum_{b \in B} \Pr[Y=b] * E(X|Y=b)$
prior distribution: $\{(i, \Pr[X=i]): i = 1, \dots, n\}$
posterior distribution: $\{(i, \Pr[X=i|Y_1=H]): i = 1, \dots, n\}$
 $\Pr[Y_2=H|Y_1=H] = \sum_{i=1}^n \Pr[X=i|Y_1=H] * \Pr[Y_2=H|X=i, Y_1=H]$
MAP (maximum a posteriori) rule: guess the value a^* for which the conditional probability of $X=a^*$ given the observations is the largest
error probability analysis:
 $\Pr[E] = \Pr[\sum_i Z_i > n/2] = \sum_{k=\lceil n/2 \rceil}^n \binom{n}{k} p^k (1-p)^{n-k}$

$= \Pr[S > n/2] < \Pr[|S - np| > n(1/2 - p)] \leq \text{Var}(S) / (n^2(1/2 - p)^2) = p(1-p) / (1/2 - p)^2 * (1/n)$

===== continuous probability =====

probability density function: a function $f: \mathbb{R} \rightarrow \mathbb{R}$ s.t. $\Pr[a \leq X \leq b] = \int_a^b f(x) dx$ for all $a \leq b$

condition: $\int_{-\infty}^{\infty} f(x) dx = 1$

X discrete, Y = cX => distribution of Y: $\Pr[X=a] = \Pr[Y=a/c]$

X continuous, Y = cX => pdf of Y: $f_Y(x) = 1/c f_X(x/c)$

expectation: $E(X) = \int_{-\infty}^{\infty} xf(x) dx$

variance: $\text{Var}(X) = E((X - E(X))^2) = \int_{-\infty}^{\infty} x^2 f(x) dx - (\int_{-\infty}^{\infty} xf(x) dx)^2$

joint density: a function $f: \mathbb{R}^2 \rightarrow \mathbb{R}$ s.t. $\Pr[a \leq X \leq b, c \leq Y \leq d] = \int_c^d \int_a^b f(x,y) dx dy$ for all $a \leq b, c \leq d$

independence for continuous r.v.'s: events $a \leq X \leq b$ and $c \leq Y \leq d$ are independent for all a, b, c, d . or $f(x,y) = f_1(x)f_2(y)$

exponential distribution: $f(x) = \lambda e^{(-\lambda x)}$ if $x \geq 0$ else 0 ($\lambda > 0$) (with parameter λ)

$E(X) = 1/\lambda, E(X^2) = 2/\lambda^2, \text{Var}(X) = 1/\lambda^2$

$\Pr[X > t] = \int_t^{\infty} \lambda e^{(-\lambda x)} dx = e^{(-\lambda t)}$

memoryless property: exponential distribution & geometric distribution

$\Pr[X > s+t | X > t] = \Pr[X > s] \quad (s > 0, t > 0)$

X_1 exp with param λ_1, X_2 exp with param λ_2 => $Y = \min\{X_1, X_2\}$ exp with param $\lambda_1 + \lambda_2$

Normal distribution: $X \sim N(\mu, \sigma^2)$. $f(x) = 1/\sqrt{2\pi\sigma^2} * e^{-(x-\mu)^2 / (2\sigma^2)}$ ($\sigma > 0$)

when $\mu=0$ and $\sigma=1$, standard normal distribution

in general, $Y = (X - \mu) / \sigma$ has the standard normal distribution

Central Limit Theorem: Let X_1, \dots, X_n be i.i.d.r.v. with common expectation $\mu = E(X_i)$ and variance $\sigma^2 = \text{Var}(X_i)$ (both assumed to be $< \infty$). Define $A_n = (X_n - \mu) / (\sigma / \sqrt{n}) = ((\sum_i X_i) - n\mu) / (\sigma \sqrt{n})$. Then as $n \rightarrow \infty$, the distribution of A_n approaches the standard normal distribution in the sense that, for any real α ,

$\Pr[A_n \leq \alpha] \rightarrow 1/\sqrt{2\pi} \int_{-\infty}^{\alpha} e^{(-x^2/2)} dx$ as $n \rightarrow \infty$

===== cardinality and intractability =====

bijection from N to Z: $f(x) = x/2$ if x is even else $-(x+1)/2$

Cantor-Bernstein theorem: if $|A| \leq |B|$ and $|B| \leq |A|$ then $|A| = |B|$

injection from N to Q: $N \rightarrow \mathbb{Z} * \mathbb{Z} \rightarrow \mathbb{Q}$. The latter by mapping rational number to 2D points, and go by spiral

finite binary strings countable (prove by listing strings in increasing order of length, and then in lexicographic order)

$[0, 1]$ uncountable. prove by **diagonalization, add 2 to each digit**

cantor set: repeatedly remove $1/3 \sim 2/3$ of interval. **uncountable set of measure 0**

represent each number as ternary strings. $C = \{x \text{ in } [0,1]: x \text{ has ternary representation of only 0's and 2's}\}$

divide each ternary string by 2 \rightarrow onto $[0, 1]$

S countably infinite => $|P(S)| > |S|$

halting problem: Turing(P): if TestHalt(P, P) = "yes" then loop forever; else halt.

Turing(Turing) contradiction