CSE 250B: Homework 2 Solutions

- 1. Regression with one predictor variable
 - (a) We will predict the mean of the y-values: $\hat{y} = (1+3+4+6)/4 = 3.5$. The MSE of this prediction is exactly the variance of the y-values, namely:

$$MSE = \frac{(1-3.5)^2 + (3-3.5)^2 + (4-3.5)^2 + (6-3.5)^2}{4} = 3.25.$$

(b) If we simply predict x, the MSE is

$$\frac{1}{4} \sum_{i=1}^{4} (y^{(i)} - x^{(i)})^2 = \frac{1}{4} \left((1-1)^2 + (1-3)^2 + (4-4)^2 + (4-6)^2 \right) = 2.$$

(c) We saw in class that the MSE is minimized by choosing

$$a = \frac{\sum_{i} (y^{(i)} - \overline{y})(x^{(i)} - \overline{x})}{\sum_{i} (x^{(i)} - \overline{x})^{2}}$$
$$b = \overline{y} - a\overline{x}$$

where \overline{x} and \overline{y} are the mean values of x and y, respectively. This works out to a = 1, b = 1; and thus the prediction on x is simply x + 1. The MSE of this predictor is:

$$\frac{1}{4}\left(1^2+1^2+1^2+1^2\right)=1.$$

- 2. Lines through the origin
 - (a) The loss function is

$$L(a) = \sum_{i=1}^{n} (y^{(i)} - ax^{(i)})^2$$

(b) The derivative of this function is:

$$\frac{dL}{da} = -2\sum_{i=1}^{n} (y^{(i)} - ax^{(i)})x^{(i)}.$$

Setting this to zero yields

$$a = \frac{\sum_{i=1}^{n} x^{(i)} y^{(i)}}{\sum_{i=1}^{n} x^{(i)}^{2}}.$$

3. The loss induced by a linear predictor $w \cdot x + b$ is

$$L(w,b) = \sum_{i=1}^{n} |y^{(i)} - (w \cdot x^{(i)} + b)|.$$

4. Define

$$X = \begin{bmatrix} \leftarrow x^{(1)} \to \\ \leftarrow x^{(2)} \to \\ \vdots \\ \leftarrow x^{(n)} \to \end{bmatrix}$$

$$XX^{T} = \begin{bmatrix} x^{(1)} \cdot x^{(1)} & x^{(1)} \cdot x^{(2)} & \cdots & x^{(1)} \cdot x^{(n)} \\ x^{(2)} \cdot x^{(1)} & x^{(2)} \cdot x^{(2)} & \cdots & x^{(2)} \cdot x^{(n)} \\ x^{(n)} \cdot x^{(1)} & x^{(n)} \cdot x^{(2)} & \cdots & x^{(n)} \cdot x^{(n)} \end{bmatrix}$$

1

- 5. Discovering relevant features in regression.
 - (a) A sensible strategy is to do linear regression using the Lasso, and to choose a regularization constant λ that yields roughly 10 non-zero coefficients.
 - (b) The smallest value of λ we tried that gave nonzero coefficients for 10 features is 0.4. This yielded the following features (numbering starting at 1): 2, 3, 5, 7, 11, 13, 17, 19, 23, 29.
- 6. We want to find the $z \in \mathbb{R}^d$ that minimizes

$$L(z) = \sum_{i=1}^{n} ||x^{(i)} - z||^2 = \sum_{i=1}^{n} \sum_{j=1}^{d} (x_j^{(i)} - z_j)^2.$$

Taking partial derivatives, we have

$$\frac{\partial L}{\partial z_j} = \sum_{i=1}^n -2(x_j^{(i)} - z_j) = 2nz_j - 2\sum_{i=1}^n x_j^{(i)}.$$

Thus

$$\nabla L(z) = 2nz - 2\sum_{i=1}^{n} x^{(i)}.$$

Setting $\nabla L(z) = 0$ and solving for z, gives us

$$z^* = \frac{1}{n} \sum_{i=1}^{n} x^{(i)}.$$

7. Minimizing absolute loss. Pick any value v that is not identical to one of the data points $x^{(i)}$. Suppose that k of the data points are less than v while the remaining n-k are greater than v. Then, a small change $v \leftarrow v + \epsilon$, where ϵ may be positive or negative, will change the loss

$$L(v) = \sum_{i=1}^{n} |x^{(i)} - v|$$

by $+k\epsilon - (n-k)\epsilon = (2k-n)\epsilon$. This means that as long as $k \neq n/2$, it is always possible to change v in a way that reduces the loss. It follows that the minimum of L(v) is attained at values v for which k = n/2, that is, when v is the median of the data.

- 8. $L(w) = w_1^2 + 2w_2^2 + w_3^2 2w_3w_4 + w_4^2 + 2w_1 4w_2 + 4w_3^2 + 2w_1^2 + 2w_1^2 + 2w_2^2 + 2w_3^2 + 2w_3^2$
 - (a) The derivative is

$$\nabla L(w) = (2w_1 + 2, 4w_2 - 4, 2w_3 - 2w_4, -2w_3 + 2w_4)$$

(b) The derivative at w = (0,0,0,0) is (2,-4,0,0). Thus the update at this point is:

$$w_{new} = w - \eta \nabla L(w) = (0, 0, 0, 0) - \eta(2, -4, 0, 0) = (-2\eta, 4\eta, 0, 0).$$

- (c) To find the minimum value of L(w), we will equate $\nabla L(w)$ to zero:
 - $2w_1 + 2 = 0 \implies w_1 = -1$
 - $4w_2 4 = 0 \implies w_2 = 1$
 - $2w_3 2w_4 = 0 \implies w_3 = w_4$

The function is minimized at any point of the form (-1, 1, x, x).

- (d) No, there is not a unique solution.
- 9. We are interested in analyzing

$$L(w) = \sum_{i=1}^{n} (y^{(i)} - w \cdot x^{(i)})^{2} + \lambda ||w||^{2}.$$

(a) To compute $\nabla L(w)$, we compute partial derivatives.

$$\frac{\partial L}{\partial w_j} = \left(\sum_{i=1}^n -2x_j^{(i)}(y^{(i)} - w \cdot x^{(i)})\right) + 2\lambda w_j$$

Thus

$$\nabla L(w) = -2\sum_{i=1}^{n} (y^{(i)} - w \cdot x^{(i)})x^{(i)} + 2\lambda w.$$

(b) The update for gradient descent with step size η looks like

$$w_{t+1} = w_t - \eta \nabla L(w_t)$$

= $w_t (1 - 2\eta \lambda) + 2\eta \sum_{i=1}^n (y^{(i)} - w_t \cdot x^{(i)}) x^{(i)}$

(c) The update for stochastic gradient descent looks like the following.

$$w_{t+1} = w_t(1 - 2\eta\lambda) + 2\eta(y^{(i_t)} - w_t \cdot x^{(i_t)})x^{(i_t)}$$

where i_t is the index chosen at time t.

- 10. Closed form solution for ridge regression.
 - (a) The loss function can be rewritten thus:

$$L(w) = \sum_{i=1}^{n} (y^{(i)} - w \cdot x^{(i)})^{2} + \lambda ||w||^{2}$$
$$= ||y - Xw||^{2} + \lambda ||w||^{2}$$
$$= y^{T}y - 2w^{T}X^{T}y + w^{T}X^{T}Xw + \lambda w^{T}w$$

(b) Taking the derivative of the loss, we get:

$$\nabla L(w) = -2X^{T}y + 2X^{T}Xw + 2\lambda w = -2X^{T}y + 2(X^{T}X + \lambda I)w.$$

Setting this to zero yields $w = (X^T X + \lambda I)^{-1} (X^T y)$.

- 11. A case when Lasso finds a sparse solution.
 - (a) For any w, we can write

$$LS(w) = \|y - Xw\|^2 = \|y - Xw^* - X(w - w^*)\|^2$$

$$= \|y - Xw^*\|^2 + \|X(w - w^*)\|^2 - 2(w - w^*)X^T(y - Xw^*)$$

$$= LS(w^*) + (w - w^*)X^TX(w - w^*) - 2(w - w^*)(X^Ty - X^TXw^*).$$

The last term is zero since $X^TXw^* = X^TX(X^TX)^{-1}X^Ty = X^Ty$.

(b) The simplified Lasso problem is

$$\min ||w - w^*||^2 ||w||_1 \le 1$$

To solve this problem, imagine growing an ℓ_2 ball around the least-squares w^* until it touches the ℓ_1 unit ball. The point of first contact is the solution w.

If w^* is (say) (2,2), then this point will be (1/2,1/2), which is not sparse. If w^* is (1,3), then this point will be (0,1), which is sparse.

12. Form of the squashing function.

$$\Pr(y = 1|x) = \frac{\Pr(y = 1, x)}{\Pr(x)} = \frac{\exp(-\|x - \mu_1\|^2 / 2\sigma^2)}{\exp(-\|x - \mu_1\|^2 / 2\sigma^2) + \exp(-\|x - \mu_2\|^2 / 2\sigma^2)}$$
$$= \frac{1}{1 + \exp((\|x - \mu_1\|^2 - \|x - \mu_2\|^2) / 2\sigma^2)}$$
$$= \frac{1}{1 + \exp(2x \cdot (\mu_2 - \mu_1) + \|\mu_1\|^2 - \|\mu_2\|^2)}.$$

This is of the form s(z) where $s(\cdot)$ is the squashing function and z is linear in x.