## Probability Modelling with Applications

## Mid Semester Practice Examination

This exam contains 4 problems, each worth 25 points. Answer all questions.

**Problem 1.** An urn contains 4 red balls and 3 white balls, and we draw three balls from the urn without replacement.

- (a) (6 points) What is the probability that all 3 balls are red?
- (b) (9 points) What is the conditional probability that all 3 balls are red, given that at least 2 are red?
- (c) (10 points) What is the conditional probability that all 3 balls are red given that the third ball drawn is red?

Solution. (a): Let X be the total number of balls drawn of each color, so that X is a  $\mathbb{Z}_+^2$ -valued random variable. Since we draw the balls without replacement, X has the hypergeometric distribution  $\mathcal{H}_{3,(4,3)}$ . The relevant probability is equal

$$\mathbb{P}(X = (3,0)) = \mathcal{H}_{3,(4,3)}(\{(3,0)\}) = \frac{\binom{4}{3}\binom{3}{0}}{\binom{7}{3}} = \frac{4 \cdot 1}{\frac{7 \cdot 6 \times 5}{6}} = \frac{4}{35}$$

what is the r.v.

$$\mathbb{P}(X = (3,0)) = \frac{\mathcal{H}_{3,(4,3)}(\{(3,0)\})}{\binom{7}{3}} = \frac{\frac{4 \cdot 1}{7 \cdot 6 \times 5}}{\frac{7}{6}} = \frac{4}{35}.$$

P(A|B) =  $\frac{\mathbb{P}(A \cap B)}{\mathbb{P}(X = (2,1))} = \frac{\binom{4}{2}\binom{3}{1}}{\binom{7}{3}} = \frac{6 \cdot 3}{35} = \frac{18}{35},$ 

the relevant probability is

the relevant probability is

$$\mathbb{P}(X = (3,0) \mid X \in \{\underbrace{(2,1),(3,0)}\}) = \frac{\mathbb{P}(X = (3,0))}{\mathbb{P}(X = (3,0)) + \mathbb{P}(X = (2,1))} = \frac{4/35}{4/35 + 18/35} = \frac{4}{22} = \frac{2}{11}.$$

(c): Let Y be the ordered tuple of colors drawn, so that Y is an  $\{R,W\}^3$ -valued random variable. The relevant probability is 2 = (R, W3)

$$\mathbb{P}(Y = (R, R, R) \mid Y \in \{(R, R, R), (R, W, R), (W, R, R), (W, W, R)\})$$

$$=\frac{\mathbb{P}(Y=(R,R,R))}{\mathbb{P}(Y=(R,R,R))+\mathbb{P}(Y=(R,W,R))+\mathbb{P}(Y=(W,R,R))+\mathbb{P}(Y=(W,W,R))}.$$

$$\mathbb{P}(Y = (R, R, R)) + \mathbb{P}(Y = (R, W, R)) + \mathbb{P}(Y = (W, R, R)) + \mathbb{P}(Y = (W, W, R))$$
Write  $Y = (Y_1, Y_2, Y_3)$ . Then Lecture  $\mathbb{P}(Y_1, Y_2, Y_3)$ . Then Lecture  $\mathbb{P}(Y_1, Y_2, Y_3)$ . Then Lecture  $\mathbb{P}(Y_1, Y_2, Y_3)$ . Thus  $\mathbb{P}(Y_1, Y_2, Y_3)$ .  $\mathbb{P}(Y_1, Y_2, Y_3)$ .  $\mathbb{P}(Y_1, Y_2, Y_3)$ .  $\mathbb{P}(Y_1, Y_2, Y_3)$ . Similarly, one has

$$\mathbb{P}(Y = (R, W, R)) = \mathbb{P}(Y_1 = R)\mathbb{P}(Y_2 = W \mid Y_1 = R)\mathbb{P}(Y_3 = R \mid Y_1 = R, Y_2 = W) = \frac{4}{7} \cdot \frac{3}{6} \cdot \frac{3}{5} = \frac{36}{7 \cdot 6 \cdot 5}$$

$$\mathbb{P}(Y = (W, R, R)) = \mathbb{P}(Y_1 = W)\mathbb{P}(Y_2 = R \mid Y_1 = W)\mathbb{P}(Y_3 = R \mid Y_1 = W, Y_2 = R) = \frac{3}{7} \cdot \frac{4}{6} \cdot \frac{3}{5} = \frac{36}{7 \cdot 6 \cdot 5},$$

$$\mathbb{P}(Y = (W, W, R)) = \mathbb{P}(Y_1 = W)\mathbb{P}(Y_2 = W \mid Y_1 = W)\mathbb{P}(Y_3 = R \mid Y_1 = Y_2 = W) = \frac{3}{7} \cdot \frac{2}{6} \cdot \frac{4}{5} = \frac{24}{7 \cdot 6 \cdot 5}.$$

Hence the relevant probability is

$$\frac{24}{24+36+36+24} = \frac{2}{2+3+3+2} = \frac{1}{5}.$$

**Problem 2.** Let X, Y and Z be independent real-valued random variables on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ .

(a) (10 points) Suppose that the distribution of X has a Lebesgue density function  $\rho: \mathbb{R} \to [0, \infty)$  given by

 $\frac{\rho(x)}{\beta(a,b)} \begin{cases}
\rho(x) = \begin{cases}
0 & x \in (-\infty, \mu - s], \\
\frac{1}{2s}(1 + \cos(\frac{x - \mu}{s}\pi)) & x \in (\mu - s, \mu + s), \\
0 & x \in [\mu + s, \infty),
\end{cases}$   $\frac{1}{2s}(1 + \cos(\frac{x - \mu}{s}\pi)) \quad x \in [\mu + s, \infty),$   $\frac{1}{2s}(1 + \cos(\frac{x - \mu}{s}\pi)) \quad x \in [\mu + s, \infty),$ 

for certain parameters  $\mu \in \mathbb{R}$  and s > 0. Determine the cumulative distribution function  $F_X$  of X.

(b) (15 points) Suppose that X has the beta distribution  $\beta_{5,1}$ , Y the exponential distribution  $\mathcal{E}_{\alpha}$  for some  $\alpha > 0$ , and

$$t_{\alpha,r}(x) = e^{-\alpha x} \frac{\alpha^r \alpha^{r-1}}{(r-0)!}$$

$$\frac{\mathbb{P}(Y \le 1 \text{ and } X \le Y)}{\mathbb{P}(Z \le 1)}.$$

$$F_X(c) = \int_{-\infty}^{c} \rho(x) dx \qquad (c \in \mathbb{R}).$$

 $\frac{\sum_{x \in \mathbb{R}} \operatorname{gamma distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{Constant}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}}{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{\alpha,6}} = \frac{\sum_{x \in \mathbb{R}} \operatorname{distribution } \Gamma_{$ 

Hence we obtain F(c) = 0 for  $c \in (-\infty, \mu - s]$ , and

or 
$$c \in (-\infty, \mu - s]$$
, and
$$F_X(c) = \int_{-\infty}^c \rho(x) dx = \frac{1}{2s} \int_{\mu - s}^c (1 + \cos(\frac{x - \mu}{s}\pi)) dx$$

$$= \frac{1}{2\pi} \int_{-\pi}^{\pi(c - \mu)/s} (1 + \cos(x)) dx = (x + \sin(x)) \Big|_{-\pi}^{\pi(c - \mu)/s} = 1$$

$$= \frac{1}{2} \left(\frac{c - \mu}{s} + \frac{1}{\pi} \sin(\pi \frac{c - \mu}{s}) + 1\right) = \mu + 3$$

for  $c \in (\mu - s, \mu + s]$ . In particular,  $F_X(\mu + s) = 1$ , and then  $F_X(c) = 1$  for all  $c > \mu + s$ .

(b): The density  $\rho$  of the random variable (X,Y) is the product of the densities  $\rho_X$  of X and  $\rho_Y$  of Y, since X and

Y are independent. One has

$$\rho_X(x) = 5x^4 \qquad (1-x)^4 = (0,1)$$

 $\rho_{X}(x) = 5x^{4} \quad \frac{(1-x)^{1-1}}{(x \in (0,1))} \quad \mathcal{F}(r, n-r+1) = \frac{(r-y)!(n-r)!}{n!}$   $\rho_{Y}(y) = \alpha e^{-\alpha y} \quad (y > 0). \quad \mathcal{F}(x, y) = \frac{0!(5-1)!}{5!} = \frac{1}{5!}$   $\rho(x,y) = 5\alpha x^{4}e^{-\alpha y} \quad (x \in (0,1), y > 0). \quad n = 5$ 

Also, the density of Z is given by

$$\rho_{Z}(z) = \frac{\alpha^{6}}{\Gamma(6)} z^{5} e^{-\alpha z} = \underbrace{\frac{\alpha^{6}}{5!} z^{5} e^{-\alpha z}}_{} (z > 0).$$

The probability of  $\{Y \leq 1 \text{ and } X \leq Y\}$  is equal to the integral of  $\rho$  over the region

which is equal to  $D:=\{(x,y)\in\mathbb{R}^2\mid 0\leq x\leq y\leq 1\},$ 

 $\iint_{\Omega} \rho(x,y) dx dy = 5\alpha \int_{0}^{1} e^{-\alpha y} \int_{0}^{y} x^{4} dy dy = \alpha \int_{0}^{1} y e^{-\alpha y} dy$ 

 $=\frac{5!}{\alpha^5}\int_0^1 \frac{\alpha^6}{5!} y^5 e^{-\alpha y} dy = \frac{5!}{\alpha^5} \mathbb{P}(Z \le 1).$   $\mathbb{P}(Z \le 1) = \int_0^1 \frac{\partial}{\partial z} Z^5 e^{-\partial z} dz$ 

So

and

So

 $\frac{\mathbb{P}(Y \le 1 \text{ and } X \le Y)}{\mathbb{P}(Y \le 1)} = \frac{120}{0.5}$ 

**Problem 3.** Let X and Y be independent standard normal random variables.

NO1)

- (a) (10 points) Determine the probability density of X+2Y. (b) (15 points) Determine the probability that  $|X| \ge \sqrt{3}|Y|$ .

Solution. (a): Let  $\rho_Y$  be the probability density of Y. Then 2Y has density  $\rho_{2Y}: \mathbb{R} \to [0, \infty)$  given by

$$\rho_{2Y}(x) = \frac{1}{2\sqrt{2\pi}} e^{-\frac{(x/2)^2}{2}} = \frac{1}{\sqrt{2\pi 4}} e^{-x^2/8} \qquad (x \in \mathbb{R}),$$

which is the density of the normal distribution  $\mathcal{N}_{0,4}$ . So Y has the  $\mathcal{N}_{0,4}$  distribution, and basic results about sums of independent normally distributed random variables from the lectures show that X + 2Y has the  $\mathcal{N}_{0.5}$  distribution. The latter has density  $\rho_{X+2Y}: \mathbb{R} \to [0, \infty)$  given by

given by 
$$\rho_{X+2Y}(x) = \frac{1}{\sqrt{10\pi}} e^{-x^2/10} \qquad (x \in \mathbb{R}).$$
  $\mathcal{N}_{o,a} + \mathcal{N}_{o,b} = \mathcal{N}_{o,a+b}$ 

(b): Since X and Y are independent, the density function  $\rho_{(X,Y)}: \mathbb{R}^2 \to [0,\infty)$  of (X,Y) is given by

The probability of the event  $\{|X| \ge \sqrt{3|Y|}\}$  is given by the integral of  $\rho_{(X,Y)}$  over  $D := \{(x,y) \in \mathbb{R}^2 \mid |x| \ge \sqrt{3}|y|\}$ :

$$\mathbb{P}(|X| \ge \sqrt{3}Y) = \frac{1}{2\pi} \iint_D e^{-(x^2 + y^2)/2} dx dy.$$

To compute this integral we switch to polar coordinates  $(r,\theta) \in (0,\infty) \times (-\pi,\pi]$  such that  $x = r\cos(\theta)$  and  $y = r\sin(\theta)$ for  $(x,y) \neq (0,0)$ . Then  $x^2 + y^2 = r^2$ ,  $dxdy = rdrd\theta$ , and  $D \setminus \{0\}$  is given in the  $(r,\theta)$ -plane by

$$\{(r,\theta) \in (0,\infty) \times (-\pi,\pi] \mid |\cos(\theta)| \ge \sqrt{3}|\sin(\theta)|\} = \{(r,\theta) \in (0,\infty) \times (-\pi,\pi] \mid \theta \in (-\pi,-\frac{5}{6}\pi] \cup [-\frac{1}{6}\pi,\frac{1}{6}\pi] \cup [\frac{5}{6}\pi,\pi]\}.$$

It follows that the probability of  $\{|X| \ge \sqrt{3}Y\}$  is

$$\iint_{D} \rho_{(X,Y)}(x,y) \mathrm{d}x \mathrm{d}y = \frac{1}{2\pi} \Big( \int_{-\pi}^{-5\pi/6} \mathrm{d}\theta + \int_{-\pi/6}^{\pi/6} \mathrm{d}\theta + \int_{\pi-\pi/6}^{\pi} \mathrm{d}\theta \Big) \int_{0}^{\infty} r e^{-r^{2}/2} \mathrm{d}r = -\frac{1}{3} e^{-r^{2}/2} \Big|_{r=0}^{\infty} = \frac{1}{3}.$$

**Problem 4** (25 points). Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space. Call a subset N of  $\Omega$  a null set if there exists an  $\overline{N} \in \mathcal{F}$ with  $N \subseteq \overline{N}$  and  $\mathbb{P}(\overline{N}) = 0$ . Let  $\mathcal{N}$  be the collection of all null sets, and let  $\overline{\mathcal{F}}$  be the collection of subsets A of  $\Omega$  for which there exist  $\overline{B,C} \in \mathcal{F}$  with  $B \subseteq A \subseteq C$  and  $\mathbb{P}(C \setminus B) = 0$ .  $\overline{\mathcal{F}} := \{A \subseteq \Omega \mid \exists B,C \in \mathcal{F} \text{ s.t. } B \subseteq A \subseteq C \text{ Show that } \overline{\mathcal{F}} \text{ is a } \sigma\text{-algebra} \text{ and that } \overline{\mathcal{F}} = \sigma(\mathcal{F} \cup \mathcal{N}).$ 

Solution. First note that any  $A \in \mathcal{F}$  is contained in  $\overline{\mathcal{F}}$  Indeed, we trivially have  $A \subseteq A \subseteq A$  and  $\mathbb{P}(A \setminus A) = \mathbb{P}(\emptyset) = 0$ .

So  $\mathcal{F} \subseteq \overline{\mathcal{F}}$ .

In particular, one has  $\Omega \in \mathcal{F} \subseteq \overline{\mathcal{F}}$ .

Now let  $A \in \overline{\mathcal{F}}$  be given, and let  $B, C \in \mathcal{F}$  be such that  $B \subseteq A \subseteq C$  and  $\mathbb{P}(C \cap B^c) = 0$ . Then  $B^c, C^c \in \mathcal{F}$ ,  $C^c \subseteq A^c \subseteq B^c$  and  $\mathbb{P}(B^c \setminus C^c) = \mathbb{P}(B^c \cap C) = 0$ . So  $A^c \in \overline{\mathcal{F}}$ .

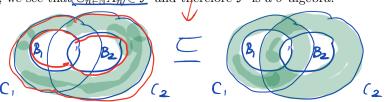
Finally, let  $(A_n)_{n \in \mathbb{N}}$  be a sequence of elements in  $\overline{\mathcal{F}}$  and  $(B_n)_{n \in \mathbb{N}}$ ,  $(C_n)_{n \in \mathbb{N}} \subseteq \mathcal{F}$  be such that  $B_n \subseteq A_n \subseteq C_n$  and  $\mathbb{P}(C_n \setminus B_n) = 0$  for all  $n \in \mathbb{N}$ . Then  $x \in C_m \setminus B_m$  for any  $m \in \mathbb{N}$  and  $x \in C_m \setminus (\cup_{n \in \mathbb{N}} B_n)$ . So

$$(\cup_{n\in\mathbb{N}}C_n)\setminus(\cup_{n\in\mathbb{N}}B_n)\subseteq(\cup_{n\in\mathbb{N}}C_n\setminus B_n).$$

Also,  $\bigcup_{n\in\mathbb{N}}B_n, \bigcup_{n\in\mathbb{N}}C_n\in\mathcal{F}$  and  $\bigcup_{n\in\mathbb{N}}B_n\subseteq\bigcup_{n\in\mathbb{N}}A_n\subseteq\bigcup_{n\in\mathbb{N}}C_n$ . Because each  $C_n\setminus B_n$  is measurable, so is their union, and we find

$$0 \leq \mathbb{P}((\cup_{n \in \mathbb{N}} C_n) \setminus (\cup_{n \in \mathbb{N}} B_n)) \leq \mathbb{P}((\cup_{n \in \mathbb{N}} C_n \setminus B_n) \leq \sum_{n=1}^{\infty} \mathbb{P}(C_n \setminus B_n) = 0.$$

Combining all this, we see that  $\bigcup_{n\in\mathbb{N}} A_n \in \overline{\mathcal{F}}$  and therefore  $\overline{\mathcal{F}}$  is a  $\sigma$ -algebra.



 $P(A \cap \phi^{c}) = P(A \cap \Omega)$   $(A \setminus \phi) = P(A)$ 

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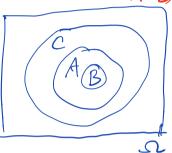
Now, if  $N \in \mathcal{N}$  is given, then there exists an  $N' \in \mathcal{F}$  containing N and satisfying  $\mathbb{P}(N') = 0$ . We then find  $\emptyset \subseteq N \subseteq N'$ and  $\mathbb{P}(N'\setminus\emptyset)=\mathbb{P}(N')=0$ . Since the empty set is contained in  $\mathbb{Z}$ , we find  $N\in\overline{\mathcal{F}}$  and thus  $\mathcal{N}\subseteq\overline{\mathcal{F}}$ . Since we have already shown that  $\overline{\mathcal{F}}$  is a  $\sigma$ -algebra and that  $\mathcal{F} \subseteq \overline{\mathcal{F}}$ , it follows that  $\sigma(\mathcal{F} \cup \mathcal{N}) \subseteq \overline{\mathcal{F}}$ .

Conversely, if  $A \in \overline{\mathcal{F}}$  is given, then there exist sets  $B, C \in \mathcal{F}$  such that  $B \subseteq A \subseteq C$  and  $\mathbb{P}(C \setminus B) = 0$ . Since  $A \setminus B$  is contained in  $C \setminus B$ , we have  $A \setminus B \in \mathcal{N}$  and thus  $A = B \cup (A \setminus B)$  is the union of an element in  $\mathcal{F}$  and an element in

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FUN

TO OF EVAMINATION  $\mathcal{N}$ . By the arbitrariness of  $A(\overline{\mathcal{F}} \subseteq \sigma(\mathcal{F} \cup \mathcal{N})$ . BEAEC (A-B) E(C-B) END OF EXAMINATION







BC





BC = AC = CE