Lecture 16: Expectation and Variance - Part III

Satyajit Thakor IIT Mandi

$$Var(X) \ge 0. \ Var(X) = E[(X - M_X)^2]$$

$$= \sum_{x} (x - M_X)^2 P_x(x)$$

$$\ge 0 \ (x - M_X)^2 \ge 0, V$$

 $\geq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right), p_{X}(x) \geq 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right), p_{X}(x) \geq 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right), p_{X}(x) \geq 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $\leq 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$ $= 0 \left(\frac{(x-\mu_{X})^{2}}{(x-\mu_{X})^{2}} \right) = 0 \text{ for all } x$

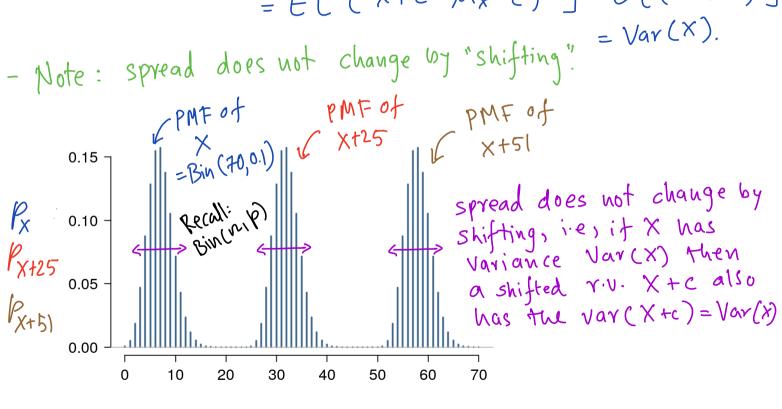
Note: Var(c) = 0 we and that when the r.v. X is a constant C(i.e., P(X=C)=1) then the spread of PMF is zero:

Shifting X by c E(X+c)=Mx+c. $\operatorname{Var}(X+c) = \operatorname{Var}(X)$ for any constant c.

$$Var(X+c) = E[(X+c-E(X+c))^{2}]$$

$$= E[(X+c-M_{X}-c)^{2}] = E[(X-M_{X})^{2}]$$

$$= Var(X).$$



Var(
$$cX$$
) = $c^2 \text{Var}(X)$ for any constant c .

$$Var(cX) = $E[(cX - E(cX))^2]$

$$= E[c^2(X - M_X)^2]$$

$$= c^2 E[(X - M_X)^2] \therefore \text{ linearity of } E(\cdot).$$

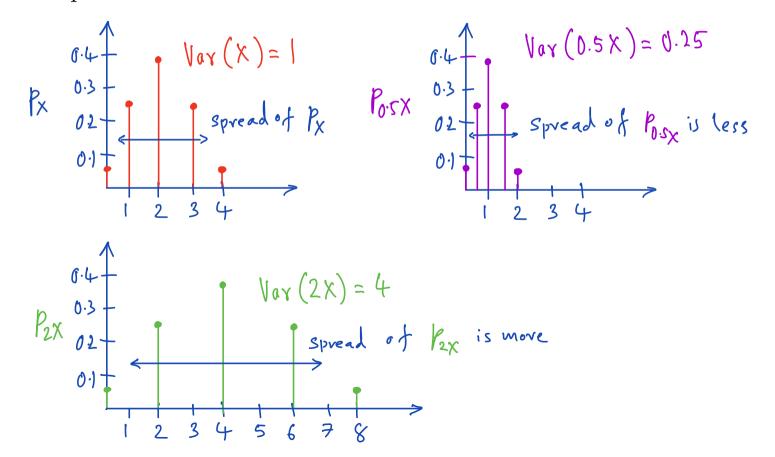
$$= c^2 \text{ Var}(X).$$$$

- That is, it we scale X by c then the variance Var(cX) of the resulting r.V. cX is Var(X) scalled by c?.

 So, depending on the value of c, the spread for cX either reduces or increases (it c=1 then the spread remains the come)
- reduces or increases (if c=1 then the spread remains the same.)

 The next example demonstrates this.

Example: Let $X \sim \text{Bin}(4,.5)$. Plot PMFs of X, 2X, .5X and compare their variance.



Functions of independent random variables

 \triangleright Theorem: Suppose that X and Y are independent r.v.s. Then

$$W = g(X) \text{ and } Z = h(Y) \text{ are also independent.}$$

$$PYOVF:$$

$$P(W = W, Z = Z) = P(W = W, Z = Z) = P\left(\begin{bmatrix} y & \{X = x\} \} \\ x : g(x) = W \end{bmatrix} \cap \begin{bmatrix} y & \{Y = y\} \} \\ y : h(y) = Z \end{bmatrix}\right)$$

Note: You may skip the

Proof (will not be asked in exam). But remember

$$x: g(x) = \omega$$
 $y: h(y) = Z$
 $x: g(x) = \omega$
 $y: h(y) = Z$

the theorem statement. Il since Pxx (x,y) = Px(2). Px(y) The proof is just to $= \sum p_{x}(x) \sum p_{y}(y) = p_{y}(\omega) \cdot p_{z}(z).$ Satisfy your curiosity. 2: 9(x)=w y:h(y)=Z

ightharpoonup If X and Y are independent, then

$$Var(X+Y) = Var(X) + Var(Y).$$

$$Var(X+Y) = E \left[(X+Y-E(X+Y))^{2} \right]$$

$$= E \left[(X+Y-M_{X}-M_{Y})^{2} \right] \therefore \text{ linearity of } E().$$

$$= E \left[(X-M_{X})^{2} + (Y-M_{Y})^{2} + 2(X-M_{X})(Y-M_{Y}) \right]$$

$$= Var(X) + Var(Y) + 2 E \left[(X-M_{X})(Y-M_{Y}) \right]$$

$$= Var(X) + Var(Y) + 2 E(X-M_{X})(Y-M_{Y})$$

$$= Var(X) + Var(Y) + 2 E(X-M_{X}).E(Y-M_{Y})$$

$$= Var(X) + Var(Y) + 2 (M_{X}-M_{X}).(M_{Y}-M_{Y})$$

$$= Var(X) + Var(Y)$$

Example

Let X be a random variable with the PMF
$$p_X(-3) = 1/6, p_X(6) = 1/2, p_X(9) = 1/3.$$
 (a) Find $\mu_{g(X)} = E(g(X))$, where $g(X) = (2X + 1)^2$. (b) Find SD $(g(X))$.

(a)
$$M_g(x) = E[(2x+1)^2]$$

$$= \frac{25}{6} + \frac{169}{2} + \frac{361}{3} = 209.$$
(b) $\delta_g(x) = E[(2x+1)^2 - 209)^2]$

$$= \sum [(2x+1)^2 - 209] \cdot P_x(x)$$

$$= (25 - 209)^2 \cdot \frac{1}{6} + (169 - 209)^2 \cdot \frac{1}{2} + (361 - 209)^2 \cdot \frac{1}{3}$$

= $14144 \implies SD(9(x)) = \sqrt{6_{g(x)}^2} = 118.9$.