

Lecture 20:  
Continuous Random Variables - Part IV

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# Continuous random variables

- ▶ The normal or Gaussian distribution has a PDF

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

for  $-\infty < x < \infty$ , depending upon two parameters, the mean and the variance

$$E(X) = \mu \text{ and } \text{Var}(X) = \sigma^2$$

of the distribution.

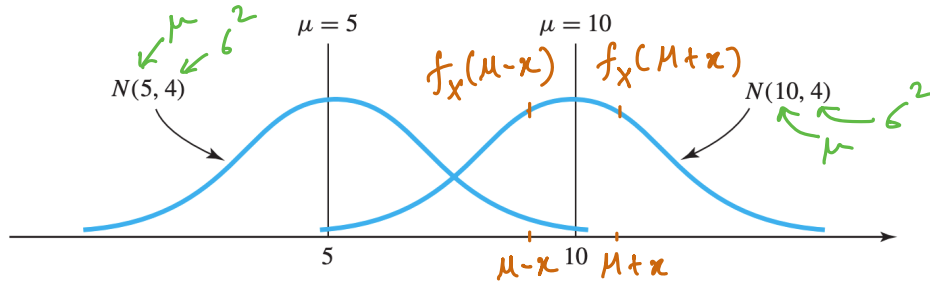
- ▶ Notation:  $X \sim N(\mu, \sigma^2)$  means that  $X$  has a normal distribution with mean  $\mu$  and variance  $\sigma^2$ .

# Continuous random variables

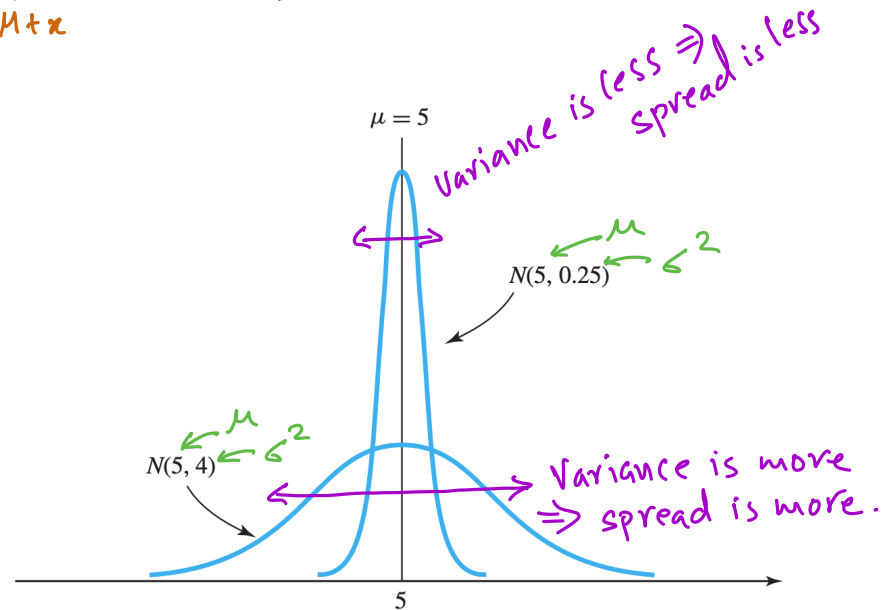
verify: homework

$$f_X(\mu+x) = f_X(\mu-x), \forall x$$

- The PDF is a bell-shaped curve that is symmetric about  $\mu$ :



- Effect of change in variance to the PDF:

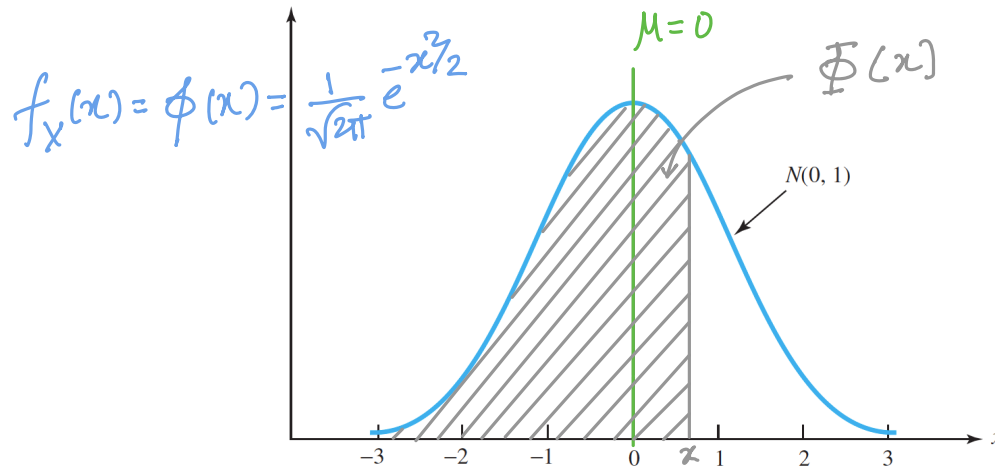


# Continuous random variables

- If  $X \sim N(0, 1)$  then  $X$  is said to have the **standard normal distribution**.

- Its PDF is denoted  $\phi(x)$ : 
$$\phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

- Its CDF is denoted  $\Phi(x)$ : 
$$\Phi(x) = \int_{-\infty}^x \phi(t) dt$$
 *closed form?*



By symmetry,  
 $\phi(x) = \phi(-x)$ .

# Continuous random variables

- ▶ The **exponential distribution** is often used to model waiting times (e.g., of a customer in a queue).
- ▶ Its PDF is

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0; \\ 0, & x < 0, \end{cases}$$

where  $\lambda > 0$  is the parameter for the distribution.

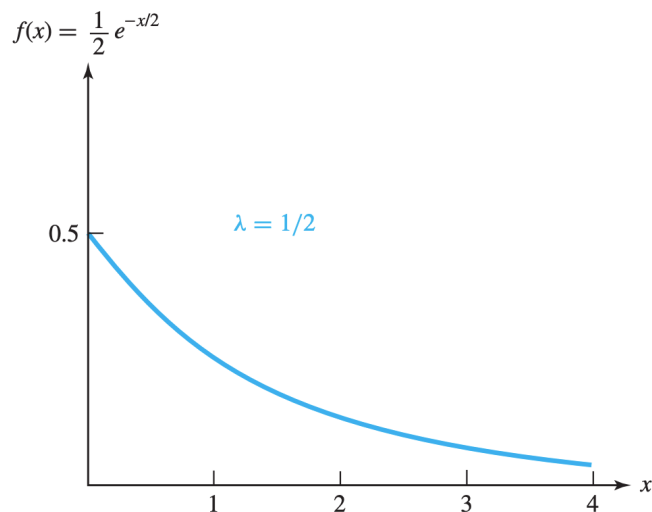
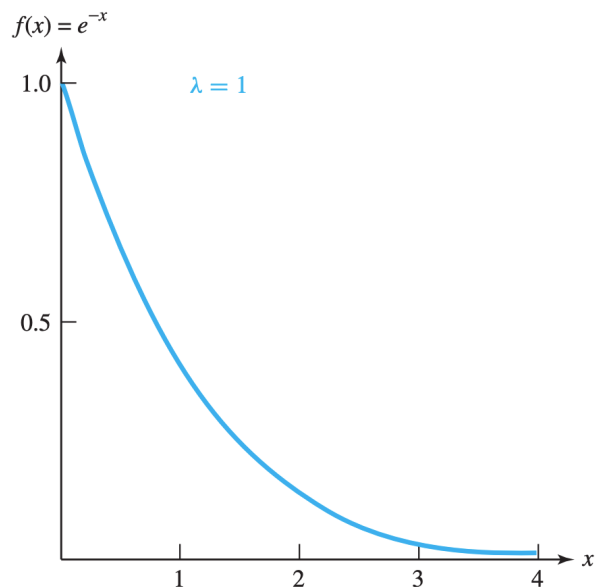
- ▶ The CDF is

$$\begin{aligned} F_X(x) &= \int_0^x \lambda e^{-\lambda t} dt, & 0 \leq x < \infty \\ &= \left. \frac{\lambda e^{-\lambda t}}{-\lambda} \right|_{t=0}^{t=x} = 1 - e^{-\lambda x} \end{aligned}$$

$$\Rightarrow F_X(x) = \begin{cases} 0 & x < 0, \\ 1 - e^{-\lambda x} & x \geq 0. \end{cases}$$

# Continuous random variables

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0. \end{cases}$$



# Continuous random variables

- Mean of an r.v. with exponential distribution:

$$\begin{aligned} E(X) &= \int_0^{\infty} \underbrace{t \cdot \lambda}_{u} \underbrace{e^{-\lambda t}}_{dv} dt \\ &= t \cdot (-e^{-\lambda t}) \Big|_0^{\infty} - \int_0^{\infty} -e^{-\lambda t} dt \\ &= 0 - \frac{1}{+\lambda} e^{-\lambda t} \Big|_0^{\infty} \\ &= 0 - \left[ 0 - \frac{1}{\lambda} \right] \\ &= \frac{1}{\lambda} \end{aligned}$$

Integration by parts:

$$\int u dv = uv - \int v du$$

$$u = t \Rightarrow du = dt$$

$$v = -e^{-\lambda t} \Rightarrow dv = \lambda e^{-\lambda t} dt$$

Definite integral:

$$\int_a^b u dv = uv \Big|_a^b - \int_a^b v du$$

$$\int_a^{\infty} u dv = \lim_{b \rightarrow \infty} \left[ uv \Big|_a^b - \int_a^b v du \right]$$

# Continuous random variables

- Variance of an r.v. with exponential distribution:

$$\begin{aligned} E(X^2) &= \int_0^{\infty} \underbrace{t^2}_{u} \cdot \underbrace{\lambda e^{-\lambda t}}_{dv} dt \\ &= t^2 \cdot (-e^{-\lambda t}) \Big|_0^{\infty} - \int_0^{\infty} -2t e^{-\lambda t} dt \\ &= 0 + \frac{2}{\lambda} \underbrace{\int_0^{\infty} \lambda t e^{-\lambda t} dt}_{E(X)} = \frac{2}{\lambda} \cdot \frac{1}{\lambda} = \frac{2}{\lambda^2} \end{aligned}$$

Integration by parts:

$$\int u dv = uv - \int v du$$

$$u = t^2 \Rightarrow du = 2t dt$$

$$v = -e^{-\lambda t} \Rightarrow dv = \lambda e^{-\lambda t} dt$$

$$\begin{aligned} \Rightarrow \text{Var}(X) &= E(X^2) - E(X)^2 \\ &= \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2} \end{aligned}$$



# Continuous random variables

- Median of an r.v. with exponential distribution:

We need to solve  $F_X(x) = 0.5$

$$\Rightarrow 1 - e^{-\lambda x} = 0.5$$

$$\Rightarrow e^{-\lambda x} = 0.5$$

$$\Rightarrow -\lambda x = \ln 0.5$$

$$\Rightarrow -\lambda x = -0.693$$

$$\Rightarrow x = \frac{0.693}{\lambda}$$

$$\Rightarrow x = 0.693 \cdot E(X)$$

# Note

- ▶ Source of figures: reference books