



Probability & MC Simulation

End Sem

Unit - 4

Joint Probability Distribution Function \rightarrow

JPDF for the joint event $\{X \leq x, Y \leq y\}$ is defined as

$$F_{X,Y}(x,y) = P\{X \leq x, Y \leq y\} = P(A \cap B)$$

JPDF is

$$f_{X,Y}(x,y) = \sum_{n=1}^N \sum_{m=1}^M P(X_n, Y_m) u(x-x_n) u(y-y_m)$$

$Q.$	(X, Y)	$(0, 0)$	$(1, 2)$	$(2, 3)$	$(3, 2)$
	$P(x, y)$	0.2	0.3	0.4	0.1

Find dist. function $F_{X,Y}(x,y)$ &
Marginal density functions.

Solution \rightarrow

$$f_{X,Y}(x,y) = \sum_{n=1}^N \sum_{m=1}^M P(X_n, Y_m) u(x-x_n) u(y-y_m)$$

$$\begin{aligned} F_{X,Y}(x,y) &= 0.2u(x)u(y) + 0.3u(x-1)u(y-2) \\ &\quad + 0.4u(x-2)u(y-3) + 0.1u(x-3)u(y-2) \end{aligned}$$

Marginal Density functions

$$F_X(n) = F_{X,Y}(n, \infty)$$

$$F_X(n) = 0.2u(n) + 0.3u(n-1) + 0.4u(n-2) + 0.1u(n-3)$$

∴

$$F_X(y) = F_{X,Y}(\infty, y)$$

$$F_Y(y) = 0.2u(y) + 0.3u(y-2) + 0.4u(y-3) + 0.1u(y-2)$$

Q:

(X, Y)	-1	0	1
0	$\frac{3}{18}$	$\frac{2}{18}$	$\frac{3}{18}$
1	$\frac{1}{18}$	$\frac{1}{18}$	$\frac{1}{18}$
2	$\frac{2}{18}$	$\frac{1}{18}$	$\frac{2}{18}$

Für $k \rightarrow$

$$\frac{3}{18} + \frac{1}{18} + \frac{2}{18} + \frac{2}{18} + \frac{k}{18} + \frac{1}{18} + \frac{3}{18} + \frac{1}{18} + \frac{2}{18} = 1$$

$$\frac{8}{18} + \frac{k}{18} + \frac{7}{18} = 1 \Rightarrow \frac{15}{18} + \frac{k}{18} = 1$$

$$\frac{k}{18} = 1 - \frac{15}{18}$$

$$\frac{k}{18} = \frac{-3}{18}$$

$$k = -3$$

JDPF \rightarrow

$$F_{x,y}(n, y) = \sum_{n=1}^N \sum_{m=1}^M P(x_n, y_m) u(n-n_n) u(y-y_m)$$

$$\begin{aligned} F_{x,y}(n, y) = & \left[3u(n)u(y+1) + 2u(n)u(y) \right. \\ & + 3u(n)u(y-1) \\ & + u(n-1)u(y+1) + 3u(n-1)u(y) + \\ & \quad u(n-1)u(y-1) \\ & \left. + 2u(n-2)u(y+1) + u(n-2)u(y) + 2u(n-2)u(y-1) \right] \end{aligned}$$

Marginal \rightarrow

$$F_x(n) = F_{x,y}(n, \infty)$$

$$F_x(n) = \frac{1}{18} [3u(n) + 2u(n) + 3u(n-1) + u(n-1) \\ + 3u(n-1) + u(n-1) + 2u(n-2) \\ + u(n-2) + 2u(n-2)]$$

Same for $F_y(y)$

Joint Probability Density Function \rightarrow

$$f_{x,y}(n, y) = \frac{\partial^2 F_{x,y}(n, y)}{\partial n \partial y}$$

Joint Probability Density function is \rightarrow

$$f_{x,y}(n, y) = \sum_{n=1}^N \sum_{m=1}^M P(x_n, y_m) \delta(n - n_n) \delta(y - y_m)$$

$$F_{x,y}(n, y) = \int_{-\infty}^n \int_{-\infty}^y f_{x,y}(n, y) dndy$$

$$F_{X,Y}(n, y) = Ae^{-(2n+y)}, \quad n \geq 0, y \geq 0$$

Find value of A
Marginal density functions

$$\int_{n=0}^{\infty} \int_{y \geq 0} f_{X,Y}(n, y) dn dy = 1$$

$$\int_0^{\infty} \int_0^{\infty} Ae^{-(2n+y)} dn dy = 1$$

$$A \left(\int_0^{\infty} e^{-2n} dn \right) \left(\int_0^{\infty} e^{-y} dy \right) = 1$$

$$\int_0^{\infty} e^{-2n} dn = \frac{1}{2}, \quad \int_0^{\infty} e^{-y} dy = 1$$

$$\frac{A \times 1 \times 1}{2} = 2, \quad A = 2$$

$$\text{JPDF} = f_{X,Y}(n, y) = 2e^{-(2n+y)}, \quad n \geq 0, y \geq 0$$

Marginal Density $f_n \rightarrow$

$$F_x(x) = \int_{-\infty}^{\infty} f_{x,y}(n, y) dy$$

$$F_x(n) = \int_0^{\infty} 2e^{-(2n+y)} dy$$

$$\begin{aligned} f_x(n) &= 2e^{-2n} \int_0^{\infty} e^{-y} dy \\ &= 2e^{-2n} \left[-e^{-y} \right]_0^{\infty} \end{aligned}$$

$$= 2e^{-2n} [1 - 0] = 2e^{-2n}$$

$$\boxed{f_n(x) = 2e^{-2n}}, \quad n \geq 0$$

Same for $f_y(y) = \int_{-\infty}^{\infty} f_{x,y}(n, y) dn$

$$f_y(y) = \int_0^{\infty} 2e^{-(2n+y)} dn$$

$$f_y(y) = 2e^{-y} \int_0^{\infty} e^{-2n} dn$$

$$= 2e^{-y} \left[-\frac{1}{2} e^{-2n} \right]_0^\infty$$

$$= 2e^{-y} \left(\frac{1}{2} \right) = e^{-y}$$

$$\boxed{f_y(y) = e^{-y}} \quad y \geq 0$$

Marginal Density $f_x(n) \rightarrow$

$$\boxed{f_x(n) = 2e^{-2n}, \quad n \geq 0 \quad \& \quad f_y(y) = e^{-y}, \quad y \geq 0}$$

$$Q. f_{x,y}(n,y) = \begin{cases} be^{-(n+y)}, & 0 < n < a, 0 \\ & < y < \infty, \\ 0, & \text{elsewhere} \end{cases}$$

find constant b (in terms of a)
given it's a valid PDF.

Soln \rightarrow

$$\int \int f_{x,y}(n,y) dn dy = 1$$

$$\int_0^\infty \int_0^a be^{-(n+y)} dn dy = 1$$

$$b \left(\int_0^a e^{-n} dn \right) \left(\int_0^\infty e^{-y} dy \right) = 1$$

$$b \left(\left[-e^{-n} \right]_0^a \right) \left(\left[-e^{-y} \right]_0^\infty \right) = 1$$

$$b(1 - e^{-a})(1) = 1$$

$$b = \frac{1}{1 - e^{-a}}$$

Valid joint pdf \rightarrow

$$f_{x,y}(x, y) = \begin{cases} \frac{e^{-(x+y)}}{1 - e^{-a}}, & 0 < x < a, \\ & 0 < y < \infty, \\ 0 & \text{elsewhere} \end{cases}$$

Q: Roll a pair of unbiased dice. If X denotes the smaller outcome & Y denotes the larger outcome. What is the JPDF of X & Y ?

Soln →

outcomes = 36

$X = 2, Y = 3$

$$f(2,3) = P(X=2, Y=3) = \frac{2}{36} = \frac{1}{18}$$

6	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{1}{36}$
5	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{1}{36}$	0
4	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{1}{36}$	0	0
3	$\frac{2}{36}$	$\frac{2}{36}$	$\frac{1}{36}$	0	0	0
2	$\frac{2}{36}$	$\frac{2}{36}$	0	0	0	0
1	$\frac{2}{36}$	0	0	0	0	0
	1	2	3	4	5	6

$$f(x, y) = \begin{cases} 1/36 & \text{if } 1 \leq x = y \leq 6, \\ 2/36, & \text{if } 1 \leq x < y \leq 6, \\ 0, & \text{otherwise.} \end{cases}$$

Q. $f(x, y) = \begin{cases} kxy^2, & \text{if } 0 < x < y < 1, \\ 0, & \text{otherwise.} \end{cases}$

value of constant k ?

sols -

$$1 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy$$

$$1 = \int_0^1 \int_0^y kxy^2 dx dy$$

$$l = \int_0^1 k y^2 \left(\int_y^\infty n \, dn \right) dy$$

$$l = \int_0^1 k y^2 \left([n^2]_0^y \right) dy$$

$$l = \int_0^1 k y^2 \times y^2 \, dy$$

$$l = \frac{k}{2} \int_0^1 y^4 \, dy$$

$$l = \frac{k}{10} [y^5]_0^1$$

$$l = \frac{k}{10} \Rightarrow \boxed{k = 10}$$

Conditional Distribution &

Density fn

The distribution of random variable X when the distribution function of a random variable Y is known at some value of y is defined as the Conditional Distribution function of X .

$$F_x(x | Y = y) = \frac{\int_{-\infty}^x f_{x,y}(x, y) dx}{f_y(y)}$$

Conditional Density fn of $X \rightarrow$

$$= \frac{\int_{-\infty}^x \frac{d}{dx} f_{x,y}(x, y) dx}{f_y(y)}$$

Q: Find the conditional density functions for the joint density function.

$$f_{x,y}(x, y) = 4xye^{-(x^2+y^2)} u(x)u(y)$$

Solu →

Conditional density functions →

$$f_x(x|y) = \frac{f_{x,y}(x, y)}{f_y(y)}$$

and

$$f_y(y|x) = \frac{f_{x,y}(x, y)}{f_x(x)}$$

$$f_x(x) = \int_0^\infty f_{x,y}(x, y) dy$$

$$= \int_0^\infty 4nye^{-n^2} e^{-y^2} u(n) dy$$

$$= 4ne^{-n^2} u(n) \int_0^\infty ye^{-y^2} dy$$

$$= 4ne^{-n^2} u(n) \left[\frac{1}{2} e^{-y^2} \right]_0^\infty$$

$$= 4ne^{-n^2} u(n) \left[\frac{1}{2} \right]$$

$f_x(n) = 2ne^{-n^2} u(n)$	\rightarrow Conditional density functions.
$f_x(y) = 2ye^{-y^2} u(y)$	

→ Statistical Independence of random variables

The two random variables are said to be statistically independent if and only if the joint probability is equal to the product of individual probab.

→ Independence from joint PDF →

$$f_{x,y}(n, y) = \frac{1}{18} e^{(-n/6 - y/3)} (u(n) u(y))$$

Solu→

Computing the marginal of x .

$$f_y(y) = \int_{-\infty}^{\infty} f_{x,y}(n, y) dn$$

$$= \int_0^{\infty} \frac{1}{18} e^{-n/6} e^{-y/3} dn$$

$$= \frac{1}{18} e^{-y/3} \int_0^\infty e^{-n/6} dn$$

$$= \frac{1}{18} e^{-y/3} \left[-6 e^{-n/6} \right]_0^\infty$$

$$= \frac{1}{18} e^{-y/3} (-6)$$

$$= +\frac{6}{18} e^{-y/3} = \boxed{+\frac{1}{3} e^{-y/3}} \quad y \geq 0$$

Marginal of $X \rightarrow$

$$f_X(n) = \int_0^\infty \frac{1}{18} e^{-n/6} e^{-y/3}$$

$$= \frac{1}{18} e^{-n/6} \int_0^\infty e^{-y/3} dy$$

$$= \frac{1}{18} e^{-n/6} \left[-3 e^{-y/3} \right]_0^\infty$$

$$= \frac{1}{18} e^{-x/6} (13)$$

$= + \frac{1}{6} e^{-x/6}$	$x \geq 0.$
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Multiplying the marginals \rightarrow

$$f_X(x) f_X(y) = \left(\frac{1}{6} e^{-x/6} \right) \left(\frac{1}{3} e^{-y/3} \right)$$

$$= \boxed{\frac{1}{18} e^{-x/6 - y/3} = f_{X,Y}(x, y)}$$

$\therefore X \& Y$ are independent.

→ Covariance

Let x and y be any 2 random variables,
Then.

$$\text{Cov}(x, y) = E(xy) - E(x)E(y)$$

Proof →

$$\text{Cov}(x, y) = E((x - \mu_x)(y - \mu_y))$$

$$= E(xy - \mu_y x - \mu_x y + \mu_x \mu_y)$$

$$= E(xy) - \mu_x E(y) - \mu_y E(x) + \mu_x \mu_y$$

$$= E(xy) - \mu_x \mu_y - \mu_y \mu_x + \mu_x \mu_y$$

$$= E(xy) - \mu_x \mu_y$$

$$= \boxed{E(xy) - E(x)E(y)}$$

Corollary \rightarrow

$$\text{Cov}(X, X) = \sigma_X^2$$

$$\text{Cov}(X, X) = E(XX) - E(X)E(X)$$

$$= E(X^2) - \mu_n^2$$

$$= \text{Var}(X)$$

$$= \sigma_X^2$$

Q:

$$f(n, y) = \begin{cases} \frac{n+2y}{18}, & \text{for } n=1,2; y=1,2. \\ 0, & \text{elsewhere} \end{cases}$$

Covariance σ_{XY} between X & Y ?

solve \rightarrow

Marginal of X \rightarrow

$$f_1(n) = \sum_{y=1}^2 \frac{n+2y}{18}$$

$$= \frac{n+2}{18} + \frac{n+2(2)}{18}$$

$$= \frac{2n+6}{18}$$

Expected value of n \rightarrow

$$E(X) = \sum_{n=1}^2 n f_1(n) = (f_1(1) + 2 f_1(2))$$

$$= \frac{2(1)+6}{18} + \frac{2(2)+6}{18}$$

$$= \frac{8}{18} + 2 \left(\frac{10}{18} \right) = \frac{28}{18}$$

Marginal of X \rightarrow

$$f_2(y) = \sum_{n=1}^2 \frac{n+2y}{18} = \frac{1}{18} (3 + 4y)$$

Expected value of y \rightarrow

$$\begin{aligned} E(y) &= \sum_{y=1}^2 y f_1(y) = 1f_1(1) + 2f_1(2) \\ &= \frac{3+4}{18} + 2 \left(\frac{3+8}{18} \right) \\ &\approx \frac{7}{18} + \frac{22}{18} = \frac{29}{18} \end{aligned}$$

Product moment of X & Y \rightarrow

$$E(XY) = \sum_{n=1}^2 \sum_{y=1}^2 n y f(n, y)$$

$$\begin{aligned} &= f(1, 1) + 2f(1, 2) + 2f(2, 1) + \\ &\quad 4f(2, 2) \end{aligned}$$

$$= \frac{3}{18} + 2 \times \frac{5}{18} + 2 \frac{4}{18} + 4 \frac{6}{18}$$

$$= \frac{45}{18}$$

Covariance b/w X & $Y \rightarrow$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$= \frac{45}{18} - \left(\frac{28}{18} \right) \left(\frac{29}{18} \right)$$

$$= \frac{45}{18} - \frac{28 \times 29}{18 \times 18}$$

$$= \frac{45}{18} - \frac{812}{324}$$

$$= \frac{45(18) - 812}{324}$$

$$= \frac{810 - 812}{324} = \frac{-2}{324} = \frac{-1}{162}$$

$= -0.00617$

$$f(x, y) = \begin{cases} n+y, & 0 < n, y < 1, \\ 0, & \text{elsewhere} \end{cases}$$

Covariance b/w X & Y?

Soln \rightarrow

Marginal function of X \rightarrow

$$f_1(n) = \int_0^1 (n+y) dy$$

$$= \left[ny + \frac{y^2}{2} \right]_0^1$$

$$= n + \frac{1}{2}$$

$$E(X) = \int_0^1 xf_1(n) dn$$

$$= \int_0^1 n \left(n + \frac{1}{2} \right) dn$$

$$= \int_0^1 \left(n^2 + \frac{n^2}{2} \right) dn$$

$$= \left[\frac{n^3}{3} + \frac{n^2}{4} \right]_0^1$$

$$= \frac{1}{3} + \frac{1}{4}$$

$$= \frac{4+3}{12} = \boxed{\frac{7}{12}}$$

Marginal fn of y \rightarrow

$$f_y(y) = \int_0^1 (n+y) dn$$

$$= \int_0^1 \frac{n^2}{2} + ny dn$$

$$= \left[\frac{n^3}{6} + ny \right]_0^1 = \left[\frac{1}{6} + y \right]$$

$$E(y) = \int_0^1 y f_r(y) dy$$

$$= \int_0^1 y \left(\frac{1}{2} + y\right) dy$$

$$= \int_0^1 \left(\frac{y}{2} + y^2\right) dy$$

$$= \left[\frac{y^2}{4} + \frac{y^3}{3} \right]_0^1$$

$$= \left[\frac{1}{4} + \frac{1}{3} \right]$$

$$= \boxed{\frac{7}{12}}$$

$$\boxed{E(x) = E(y)}$$

Product moment of X & $X \rightarrow$

$$\begin{aligned} E(XY) &= \int_0^1 \int_0^1 xy(n+ny) dndy \\ &= \int_0^1 \int_0^1 (n^2y + ny^2) dndy \\ &= \int_0^1 \left[\frac{n^3y}{3} + \frac{n^2y^2}{2} \right]_0^1 dy \\ &\sim \int_0^1 \left[\frac{y}{3} + \frac{y^2}{3} \right] dy \\ &\sim \left[\frac{y^2}{6} + \frac{y^3}{6} \right]_0^1 \end{aligned}$$

$$= \frac{1}{6} + \frac{1}{6} = \frac{2}{6} = \frac{1}{3}$$

Covariance b/w $X \& Y \rightarrow$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$= \frac{1}{3} - \frac{7}{12} \times \frac{7}{12}$$

$$= \frac{1}{3} - \frac{49}{144}$$

$$= \frac{48}{144} - \frac{49}{144}$$

$$= \boxed{\frac{-1}{144}}$$

Ans.

→ Central Limit Theorem

It states that probability density function of a sum of N independent random variables approaches a Gaussian distribution as $N \rightarrow \infty$

$$f(n, y) = \begin{cases} k(n+y), & 0 \leq n \leq 2 \\ 0 & \text{otherwise} \end{cases}$$

$$\int_0^1 \int_0^2 k(n+y) \, dy \, dn = 1$$

$$k \int_0^1 \int_0^2 \frac{n^2}{2} + ny \, dy \, dn = 1$$

$$k \int_0^1 [2 + 2y] \, dy = 1$$

$$k \int_0^1 2y + \frac{2y^2}{2} = 1$$

$$k \left[2y + y^2 \right]_0^1 = 1$$

$$k [3] = 1$$

$$\boxed{k^2 \frac{1}{3}}$$

$$f_n(n) = \int_0^1 \frac{1}{3} (n+y) dy$$

$$= \frac{1}{3} \int_0^1 ny + \frac{y^2}{2}$$

$$= \frac{1}{3} \left[ny + \frac{1}{2} y^2 \right]$$

$$= \boxed{\frac{1}{3}n + \frac{1}{6}}$$

Marginal of y

$$f_y(y) = \int_0^2 \frac{1}{3} (n+y) \, dn$$

$$= \frac{1}{3} \int_0^2 \frac{n^2}{2} + ny$$

$$= \frac{1}{3} \left[\frac{n^2}{2} + ny \right]_0^2$$

$$= \frac{1}{3} \left[\frac{4}{2} + 2y \right]$$

$$= \boxed{\frac{2}{3} + \frac{2y}{3}}$$

$S = \{(1, 2), (1, 3), (1, 4), (1, 5), (1, 6)$
 $(2, 3), (2, 4), (2, 5), (2, 6)$
 $(3, 4), (3, 5), (3, 6)$
 $(4, 5), (4, 6)$
 $(5, 6)\}$

Marginal of n

1 (36

$k_{n,y}$, $n = 1, 2, 3$

$y = 1, 2, 3$

$$\sum \sum p(n, y)$$

$$k \left[\sum_{n=1}^3 (n + 2n + 3n) \right] = 1$$

$$k \left[6 \sum_{n=1}^3 n \right] = 1$$

$$6k (1+2+3) = 1$$

$$\boxed{k = \frac{1}{36}}$$

$$f(x, y) = \begin{cases} kxy^2 & , 0 < x < y < 1 \\ 0 & , \text{otherwise} \end{cases}$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy$$

$$\int_{y=0}^{y=1} \int_{x=0}^{x=y} kxy^2 dx dy = 1$$

$$K \int_{y=0}^{y=1} \left(\frac{x^2}{2} \right)_{x=0}^{x=y} y^2 dy = 1$$

$$\frac{k}{2} \int_{y=0}^{y=1} y^4 dy = 1$$

$$\frac{k}{2} \left(\frac{y^5}{5} \right)_{y=0}^{y=1} = 1$$

$$\frac{k}{2} \left(\frac{1}{5} \right) = 1 \Rightarrow \boxed{k = 10}$$

$$f(x, y) = \begin{cases} 6/5 (x^2 + 2xy) & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$P(x) := \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy$$

$$= \int_0^{y=1} \int_0^{x=y} \frac{6}{5} (x^2 + 2xy) dx dy$$

$$= \frac{6}{5} \int_0^1 \left(\frac{x^3}{3} + 2x \cdot \frac{x^2}{2} y \right) dy$$

$$= \frac{6}{5} \int_0^1 \left(\frac{x^3 + 3x^2 y}{3} \right) dy$$

$$= \frac{6}{5} \int_0^1 \frac{4}{3} y^3 dy$$

$$= \frac{6}{5} \left[\frac{4y^4}{12} \right]_0^{y=1}$$

$$= \frac{6}{5} \left[\frac{4}{12} \right]$$

$$= \frac{6 \times 4}{5 \times 12} = \frac{4}{10} = \boxed{\frac{2}{5}} \text{ Any}$$

Marginal

$$f_n(n) = \int_{-\infty}^{\infty} f(n, y) dy$$

$$f_y(y) = \int_{-\infty}^{\infty} f(n, y) dn$$

$$f_n(n) = \int_{-\infty}^{\infty} f(n, y) dy$$

$$= \int_{-\sqrt{n}}^{\sqrt{n}} \frac{3}{4} dy$$

$$= \frac{3}{4} \left[y \Big|_{-\sqrt{n}}^{\sqrt{n}} \right] = \frac{3}{4} \left[\sqrt{n} - (-\sqrt{n}) \right]$$

$$= \frac{3}{4} 2\sqrt{n}$$

$$= \frac{3\sqrt{n}}{2}$$

$$f_x(n) = \int_{y=n}^{y=\infty} 2e^{-n-y} dy$$

$$= 2 \left[e^{-n-y} \right]_{y=n}^{y=\infty}$$

$$= 2e^{-2n}$$

$$f_y(x) = \int_{n=0}^{n=y} 2e^{-n-y} dn$$

$$= 2 \left[e^{-y-y} - e^{0-y} \right]$$

$$= 2 \left[e^{-2y} - e^{-y} \right]$$

$$= 2e^{-2y} - 2e^{-y}$$

If CDF is given then,
PDF is,

$$f(n, y) = \frac{\partial^2 F}{\partial n \partial y}$$

$$f(n, y) = \frac{\partial^2 F}{\partial n \partial y}$$

$$= \frac{1}{5} \frac{\partial}{\partial n} \frac{\partial}{\partial y} (2n^3 y + 3n^2 y^2)$$

$$= \frac{1}{5} \frac{\partial}{\partial n} (2n^3 + 6n^2 y)$$

$$= \frac{1}{5} (6n^2 + 12ny)$$

$$= \frac{6}{5} (n^2 + 2ny)$$

$$f_X(n) = \int_0^\infty 6 e^{-2n} e^{-3y} dy$$

$$= 6 \int_0^\infty e^{-2n} e^{-3y} dy$$

$$= 6 e^{2n} \left(\frac{e^{-3y}}{-3} \right)_0^\infty$$

$$= 6 e^{2n} \times \left(\frac{e^{-\infty} - e^0}{-3} \right)$$

$$\stackrel{?}{=} \frac{2e^{2n}}{-3} = -2e^{2n}$$

$$f_Y(y) = 6 \int_0^\infty e^{-2n} e^{-3y} dn$$

$$= 6 e^{-3y} \left[\frac{\bar{e}^{2n}}{-2} \right]_0^\infty$$

$$= -3e^{-3y} [e^{2\infty} - e^0]$$

$$= -3e^{-3y}$$

$$PDF: f(n,y) = kny \quad 0 < n < 1$$

mf, cdf, dep: ind. value of k

$$F_X(n) = \int_0^1 kny \, dy$$

$$= k \int_0^1 ny \, dy$$

$$= kn \left[\frac{y^2}{2} \right]_0^1$$

$$= kn \times \frac{1}{2} = \frac{kn}{2}$$

$$F_y(y) = \int_0^1 k n y \, dn$$

$$= k y \left[\frac{n^2}{2} \right]_0^1$$

$$= \frac{k y}{2}$$

$$(CDF) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(n) \, dn \, dy = 1$$

$$= \int_0^1 \int_0^{n=y} k n y \, dn \, dy$$

$$= k \int_0^1 \left[\frac{n^2}{2} \right]_0^{n=1} y = k \int_0^1 \frac{1}{2} y$$

$$= \frac{k}{2} \left[\frac{y^2}{2} \right]_0^1 = \frac{k}{2} \times \frac{1}{2} = k = 4$$

Cond. Dist. Density

$$E_x(x|y=y) = \frac{\int_{-\infty}^y f_{x,y}(x,y) dx}{f_y(y)}$$

Density \rightarrow

$$= \frac{\int_{-\infty}^y f_{x,y}(x,y) dx}{f_y(y)}$$

$$f_{x,y}(x,y) = 4xye^{-(x^2+y^2)} u(x) u(y)$$

$$f_x(x|y) = \frac{f_{x,y}(x,y)}{f_y(y)}$$

$$f_x(x|y) = \frac{f_{x,y}(x,y)}{f_x(x)}$$

$$f_x(x) = \int_0^\infty 4xye^{-x^2} e^{-y^2} u(x) dy$$

$$= 4xe^{-x^2} u(x) \int_0^\infty ye^{-y^2} dy$$

$$\approx (1) \left[\frac{e^{-y^2}}{2} \right]_0^\infty$$

$$= 4xe^{-x^2} u(x) \times \frac{1}{2}$$

$$f_y(y) = 2ye^{-y^2} u(y)$$

$$\text{Cov}(x, y) = E(xy) - E(x)E(y)$$

Proof ->

$$\begin{aligned}\text{Cov}(x, y) &= E((x - \mu_x)(y - \mu_y)) \\ &= E(xy - \mu_y x - \mu_x y + \mu_x \mu_y) \\ &= E(xy) - \mu_x \mu_y - \mu_y \mu_x + \mu_x \mu_y \\ &= E(xy) - \mu_x \mu_y\end{aligned}$$

$$= \boxed{E(xy) - E(x)E(y)}$$

