#### Introduction

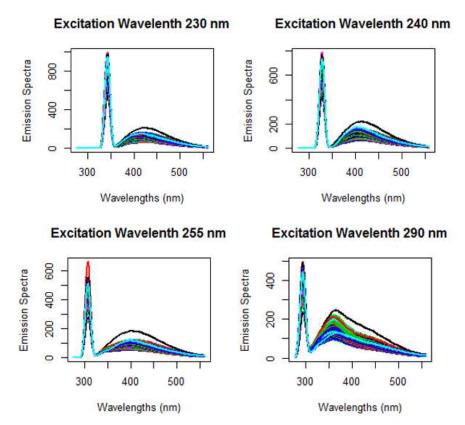
There are two parts in this project. Both are analysis of real data sets. The first part is an analysis of the sugar process data along with two response variables, which was generated by Bro in 1999. The sugar was dissolved in un-buffered water (2.25g/15mL) and the solution was measured spectrofluorometrically in a 10 by 10 mm cuvette on a PE LS50B spectrofluorometer. Raw non-smoothed data was output from the spectrofluorometer. For every sample the emission spectra from 275-560 nm were measured in 0.5 nm intervals (571 wavelengths) at seven excitation wavelengths (230, 240, 255, 290, 305, 325, 340 nm). The smoothed FPCA is used to build a regression model to study the association between the fluorescence data and the response variables.

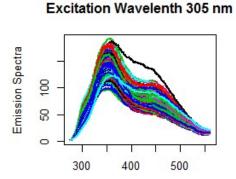
For the second part we analyze the pharmaceutical tablets data generated by Dyrby et al. in 2002. The spectral curves of Near Infrared Transmittance for 310 tablets were measured. These 310 tablets belong to four classes. Linear Discriminant Analysis is used to build a classification rule which assigns a tablet to one of four classes based on its spectral curve.

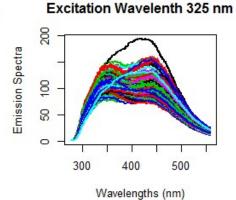
#### First Part: Sugar Process Data

The first data set is a 197 by 3997 matrix, where n = 197 is the sample size and 3997 is the number of observation points for the seven curves. That is, the first 571 number in each row is the observed values of the sample curves for the excitation wavelength 230, the next 571 numbers for the excitation wavelength 240, and so on. The aforementioned response variables are "color" and "ash measurements".

We start by plotting the data for each excitation wavelength:

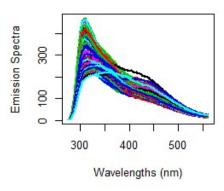






#### Excitation Wavelenth 340 nm

Wavelengths (nm)

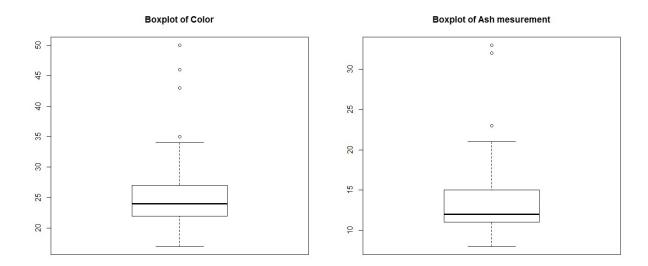


All the graphs have some similarities. Overall, they all start increasing to a peak and eventually decrease to 0 at the end. In particular, first four graphs are really similar to each other. They all have a high peak at the beginning followed by a small round bump. The last three graphs do not have the small round bump like the first four.

Furthermore, we make a summary and boxplots of the two response variables:

Table 1.1: Summary of response variables

| Statistics     | Color | Ash  |
|----------------|-------|------|
| Minimum        | 17    | 8    |
| First Quartile | 22    | 11   |
| Median         | 24    | 12   |
| Mean           | 24.5  | 13.2 |
| Third Quartile | 27    | 15   |
| Maximum        | 50    | 33   |



From the boxplots, there are a few outliers for both response variables. It seems that the response variable, color is more evenly distributed compared to ash measurement.

We use the smooth FPCA to calculate the PC scores for each seven functional data as predictor variables to build a regression model for the both response variables. We decided to first use 1000 basis functions to smooth the data. Then, we choose smoothing parameters and number of PC components by cross-validation. The following tables summarize the cross-validation errors.

Table 1.2: Cross-Validation Errors for Response Variable "Color"

| λ \ M      | 1    | 2    | 3    | 4    | 5    | 6    | 7    | 8    | 9    | 10   |
|------------|------|------|------|------|------|------|------|------|------|------|
| $10^{-10}$ | 1427 | 1479 | 1715 | 1513 | 1824 | 2408 | 2143 | 2632 | 2101 | 2470 |
| $10^{-8}$  | 1427 | 1479 | 1715 | 1513 | 1824 | 2408 | 2143 | 2632 | 2101 | 2470 |
| $10^{-6}$  | 1427 | 1479 | 1715 | 1513 | 1824 | 2408 | 2142 | 2632 | 2101 | 2470 |
| $10^{-4}$  | 1427 | 1479 | 1715 | 1513 | 1824 | 2407 | 2141 | 2632 | 2101 | 2468 |
| $10^{-2}$  | 1427 | 1479 | 1717 | 1516 | 1902 | 2375 | 1981 | 2324 | 2175 | 2353 |
| 1          | 1427 | 1479 | 1721 | 1522 | 1841 | 2234 | 1630 | 1897 | 2003 | 2344 |
| 10         | 1427 | 1492 | 1739 | 1521 | 1835 | 2058 | 1567 | 1570 | 1962 | 2005 |

Table 1.3: Cross Validation Errors for Response Variable "Ash Measurement"

| λ \ M      | 1    | 2   | 3   | 4   | 5   | 6   | 7   | 8   | 9   | 10  |
|------------|------|-----|-----|-----|-----|-----|-----|-----|-----|-----|
| $10^{-10}$ | 1277 | 705 | 416 | 455 | 443 | 474 | 528 | 612 | 730 | 781 |
| 10-8       | 1277 | 705 | 416 | 455 | 443 | 474 | 528 | 612 | 730 | 781 |
| $10^{-6}$  | 1277 | 705 | 416 | 455 | 443 | 474 | 528 | 612 | 730 | 781 |
| $10^{-4}$  | 1277 | 705 | 416 | 455 | 443 | 474 | 528 | 612 | 728 | 781 |
| $10^{-2}$  | 1278 | 705 | 416 | 455 | 472 | 492 | 501 | 568 | 636 | 730 |
| 1          | 1278 | 705 | 416 | 456 | 477 | 474 | 511 | 564 | 617 | 578 |
| 10         | 1278 | 711 | 427 | 456 | 477 | 466 | 570 | 638 | 560 | 756 |

We found that the optimal tuning parameter and number of PC components for response variable "color" are  $10^{-2}$  and 1, respectively. Also, optimal tuning parameter and number of PC components for response variable "ash measurement" is  $10^{-10}$  and 3, respectively. Using these optimal values, we calculate the PC scores for each of seven functional data and use these scores as predictor variables to build a regression model for both response variables.

Diagram 1.2: Summary of Regression with Response Variable "Color"

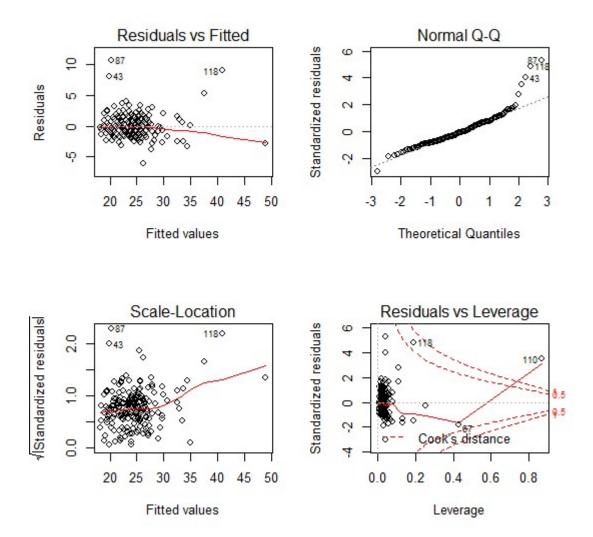
```
Call:
lm(formula = Y1 ~ pX11)
Residuals:
           1Q Median
                         3Q
-6.1473 -1.3400 -0.2046 1.0937 10.7282
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 2.446e+01 1.476e-01 165.719 < 2e-16 ***
pX111
          4.064e-03 1.415e-03 2.871 0.00456 **
pX112
           9.337e-05 2.237e-03 0.042 0.96675
pX113
           2.762e-03 2.038e-03 1.355 0.17694
pX114
pX115
           9.238e-03 3.069e-03 3.010 0.00297 **
           2.257e-03 6.081e-03 0.371 0.71096
pX116
           1.428e-02 2.818e-03 5.067 9.60e-07 ***
pX117 -9.311e-03 1.042e-03 -8.934 3.67e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 2.071 on 189 degrees of freedom
Multiple R-squared: 0.7905, Adjusted R-squared: 0.7827
F-statistic: 101.9 on 7 and 189 DF, p-value: < 2.2e-16
```

Diagram 1.2: Summary of Regression with Response Variable "Ash Measurement"

```
Call:
lm(formula = Y2 ~ pX12)
Residuals:
   Min 1Q Median 3Q
-3.2865 -0.8660 -0.0538 0.7957 4.3491
Coefficients:
          Estimate Std. Error t value Pr(>|t|)
 (Intercept) 13.177665 0.092991 141.709 < 2e-16 ***
pX121 -0.003828 0.001732 -2.210 0.028380 *
pX122
         0.037012 0.011859 3.121 0.002108 **
         0.008237 0.011388 0.723 0.470437
pX123
          0.007525 0.003738 2.013 0.045632 *
pX124
pX125
        -0.052099 0.014149 -3.682 0.000308 ***
         -0.058023 0.021247 -2.731 0.006963 **
pX126
pX127
          0.030996 0.006981 4.440 1.59e-05 ***
          0.050974 0.014460 3.525 0.000540 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 1.305 on 175 degrees of freedom
Multiple R-squared: 0.8896, Adjusted R-squared: 0.8764
F-statistic: 67.16 on 21 and 175 DF, p-value: < 2.2e-16
```

We plotted both regressions. The plots are below.

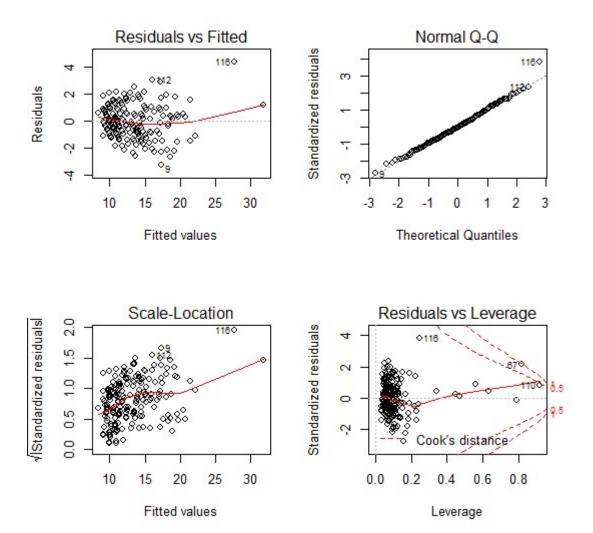
Diagram 1.3: Regression with Response Variable "Color"



Based on the Residue vs Fitted plot above, there does not seem to be a non-linear relationship between predictor variables and the response variable. Also, Q-Q plot suggests that the residuals are not normally distributed because of the heavy tails. Furthermore, the Scale-Location plot does not support the assumption of equal variance. Overall, the regression model is a pretty good fit.

Diagram 1.1 shows that those significant coefficients come from first, fourth, sixth, and seventh curves. In other words, the emission spectra are affected mostly by excitation wavelengths of 230 nm, 290 nm, 325 nm, and 340 nm.

Diagram 1.4: Regression with Response Variable "Ash Measurement"



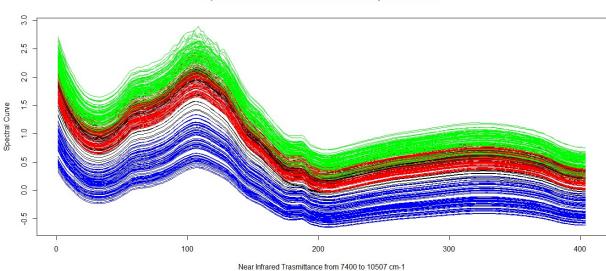
Based on the Residue vs Fitted plot above, there seem to be a non-linear relationship between predictor variables and the response variable, which might be a quadratic relationship. However, the line is extremely close to a straight line. Also, Q-Q plot suggests that the residuals are normally distributed. Furthermore, the Scale-Location plot does not support the assumption of equal variance. Overall, the regression model is a pretty good fit.

Diagram 1.2 shows that those significant coefficients come from first, second, third, fourth, and seventh curves. In other words, the emission spectra are affected mostly by excitation wavelengths of 230 nm, 240 nm, 255 nm, 290 nm, and 340 nm.

#### **Second Part: Pharmaceutical Tablets Data**

The pharmaceutical tablets data are 310 by 404 matrix, where 310 is the same size and 404 is the number of equally spaced observation points from 7400 to 10507  $cm^{-1}$ . These tablets belong to four classes. First, we plot all 310 curves and color them based on their classes, where class 1 is colored black, class 2 is colored blue, class 3 is colored red, and class 4 is colored green.

Diagram 2.1: Spectral Curves by Colors



Spectral Curves for 310 tablets colored by their classes

Each curve clearly has the same pattern. Class 2, 3, and 4 seem to be clearly separated from each other. Class 2 curves are in the bottom, class 3 curves are in the middle and class 4 curves are at the top. On the other hand, class 1 curves are mixed in between class 2 and class 4.

We use the smooth FPCA to calculate the PC scores as predictor values for the linear discriminant analysis. We decided to first use 1000 basis functions to smooth the data. Then, we choose smoothing parameters and number of PC components by cross-validation. The following tables summarize the cross-validation classification errors:

| $\lambda \setminus M$ | 1     | 2     | 3     | 4     | 5     | 6     | 7     | 8     | 9     | 10    |
|-----------------------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| $10^{-10}$            | 1.452 | 0.758 | 0.565 | 0.290 | 0.242 | 0.258 | 0.242 | 0.226 | 0.177 | 0.194 |
| $10^{-8}$             | 1.452 | 0.758 | 0.565 | 0.290 | 0.242 | 0.258 | 0.242 | 0.226 | 0.177 | 0.194 |
| $10^{-6}$             | 1.452 | 0.758 | 0.565 | 0.290 | 0.242 | 0.258 | 0.242 | 0.226 | 0.177 | 0.194 |
| $10^{-4}$             | 1.452 | 0.758 | 0.565 | 0.290 | 0.242 | 0.258 | 0.242 | 0.226 | 0.177 | 0.194 |
| $10^{-2}$             | 1.452 | 0.758 | 0.565 | 0.290 | 0.242 | 0.258 | 0.242 | 0.226 | 0.177 | 0.194 |
| 1                     | 1.452 | 0.758 | 0.565 | 0.290 | 0.242 | 0.258 | 0.194 | 0.194 | 0.194 | 0.194 |
| 10                    | 1.452 | 0.758 | 0.565 | 0.323 | 0.210 | 0.177 | 0.210 | 0.210 | 0.194 | 0.210 |

Table 2.1: Cross-Validation Classification Errors

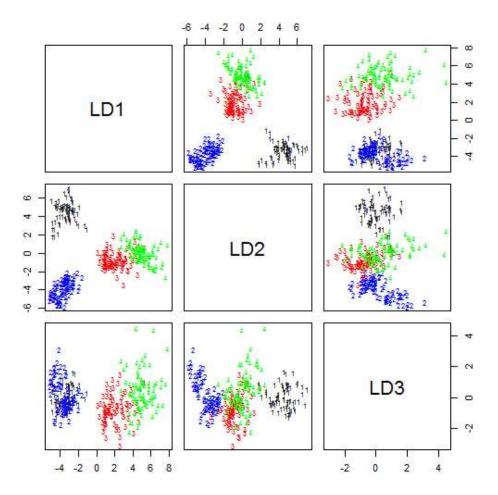
We chose  $10^{-6}$  and 9 to be the optimal tuning parameter and number of PC components, respectively. Using these optimal values, we calculate the PC scores and use these scores as predictor variables to build a classification model using Linear Discriminant Analysis.

Diagram 2.2: Summary of Linear Discriminant Analysis

```
Call:
lda(Y ~ ., data = data)
Prior probabilities of groups:
0.2258065 0.2580645 0.2580645 0.2580645
Group means:
                    V3
                                V4
                                            V5
                                                         V6
         V2
   2.254320 -0.37708079 0.25272882 0.093099079 0.004408347 0.020293226
2 -13.570510 0.16527278 -0.02060730 0.002843224 0.001577569 -0.002001701
  2.367212 0.01259795 -0.08405141 -0.063933927 0.014914541 -0.004229207
4 9.230768 0.15207495 -0.11647901 -0.020370991 -0.020349413 -0.011525665
           V8
                        V9
                                     V10
1 0.003893078 -3.162811e-03 0.0019893621
2 0.008041494 1.060995e-02 -0.0035921739
3 -0.005096534 4.042236e-05 0.0013110285
4 -0.006351403 -7.882915e-03 0.0005404536
Coefficients of linear discriminants:
           LD1
                     LD2
V2
     0.3033702 0.2283782 -0.006932079
V3
     2.0868775 -3.5416534 0.712065605
    -6.4954666 5.3424779 0.278184058
    -9.4467909 7.6884430 4.062276642
V6 -16.9518499 2.4160361 -25.371691617
V7 -29.4373314 32.5216062 -3.240334075
V8 -36.0315592 16.1063218 11.856818738
V9
     9.3979447 0.4608267 -9.237565840
V10 -9.6204113 7.8303482 -7.059613525
Proportion of trace:
  LD1
        LD2
                LD3
0.6068 0.3840 0.0092
```

To calculate the classification error, we use the leave-one-out method. We get that the classification error is 0.04516129. Also, we plot the class of the linear discriminant analysis using the color representation from above.

Diagram 2.3: Linear Discriminant Analysis Graphs



From Diagram 2.3, there are some overlaps between classes but they are pretty distinctively separated most of the times. I think the reason for this phenomenal is that the original curves have some overlaps as mentioned before. With a relatively small classification error, I would say this classification rule is a decent one.

#### **Appendix**

#### Part 1:

```
X=read.table("project.one.1.X.txt")
s=seq(275,560,0.5)
X1=as.matrix(X[,1:571])
X2=as.matrix(X[,572:1142])
X3=as.matrix(X[,1143:1713])
X4=as.matrix(X[,1714:2284])
X5=as.matrix(X[,2285:2855])
X6=as.matrix(X[,2856:3426])
X7=as.matrix(X[,3427:3997])
matplot(s,t(X1),type="1",lwd=2,lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 230 nm")
matplot(s,t(X2),type="1",lwd=2,lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 240 nm")
matplot(s, t(X3), type="l", lwd=2, lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 255 nm")
matplot(s, t(X4), type="l", lwd=2, lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 290 nm")
matplot(s,t(X5),type="1",lwd=2,lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 305 nm")
matplot(s, t(X6), type="l", lwd=2, lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 325 nm")
matplot(s,t(X7),type="1",lwd=2,lty=1,
xlab="Wavelengths (nm)", ylab="Emission Spectra",
main="Excitation Wavelenth 340 nm")
Y=read.table("project.one.1.Y.txt")
Y1=as.vector(as.matrix(Y[,1]))
Y2=as.matrix(Y[,2])
n=1:197
summary(Y)
```

```
par(mfrow=c(1,2))
boxplot(Y1, main="Boxplot of Color")
boxplot(Y2, main="Boxplot of Ash mesurement")
t=0:570
basis.obj=create.bspline.basis(c(0,571), nbasis=1000)
fdParobj=fdPar(basis.obj, 2, 1e-10)
X=scale(X,center=TRUE, scale=FALSE)
X1=as.matrix(X[,1:571])
X2=as.matrix(X[,572:1142])
X3=as.matrix(X[,1143:1713])
X4=as.matrix(X[,1714:2284])
X5=as.matrix(X[,2285:2855])
X6=as.matrix(X[,2856:3426])
X7=as.matrix(X[,3427:3997])
fit.list.1=smooth.basis(t,t(X1),fdParobj)
X1.fd=fit.list.1$fd
fit.list.2=smooth.basis(t,t(X2),fdParobj)
X2.fd=fit.list.2$fd
fit.list.3=smooth.basis(t,t(X3),fdParobj)
X3.fd=fit.list.3$fd
fit.list.4=smooth.basis(t,t(X4),fdParobj)
X4.fd=fit.list.4$fd
fit.list.5=smooth.basis(t,t(X5),fdParobj)
X5.fd=fit.list.5$fd
fit.list.6=smooth.basis(t,t(X6),fdParobj)
X6.fd=fit.list.6$fd
fit.list.7=smooth.basis(t,t(X7),fdParobj)
X7.fd=fit.list.7$fd
lambda=10^{(c(-10,-8,-6,-4,-2,0,1))}
M=1:10
cv.fold=split(sample(197), rep(1:5, length=197))
CV.error.1=array(0, c(length(lambda),length(M)))
CV.error.2=array(0, c(length(lambda),length(M)))
{for(j in 1:length(lambda))
{for(k in 1:length(M))
{ print(c("j,k=",j,k))
pca.Par=fdPar(basis.obj, 2, lambda[j])
pca.list.1=pca.fd(X1.fd, M[k], pca.Par)
pX1=pca.list.1$scores
```

```
pca.list.2=pca.fd(X2.fd, M[k], pca.Par)
pX2=pca.list.2$scores
pca.list.3=pca.fd(X3.fd, M[k], pca.Par)
pX3=pca.list.3$scores
pca.list.4=pca.fd(X4.fd, M[k], pca.Par)
pX4=pca.list.4$scores
pca.list.5=pca.fd(X5.fd, M[k], pca.Par)
pX5=pca.list.5$scores
pca.list.6=pca.fd(X6.fd, M[k], pca.Par)
pX6=pca.list.6$scores
pca.list.7=pca.fd(X7.fd, M[k], pca.Par)
pX7=pca.list.7$scores
pX=cbind(pX1,pX2,pX3,pX4,pX5,pX6,pX7)
for(1 in 1:5)
{test=cv.fold[[1]]
train=setdiff(1:197, test)
fit.1=lm(Y1[train]~pX[train,])
Y1.test.pred=cbind(1,pX[test,])%*%coef(fit.1)
CV.error.1[j,k]=CV.error.1[j,k]+sum((Y1.test.pred-Y1[test])^2)
print(c("CV.error.1[j,k]=",CV.error.1[j,k]))
fit.2=lm(Y2[train]~pX[train,])
Y2.test.pred=cbind(1,pX[test,])%*%coef(fit.2)
CV.error.2[j,k]=CV.error.2[j,k]+sum((Y2.test.pred-Y2[test])^2)
print(c("CV.error.2[j,k]=",CV.error.2[j,k]))
}
}
}
cv1=which(CV.error.1==min(CV.error.1), arr.ind=TRUE)
cv2=which(CV.error.2==min(CV.error.2), arr.ind=TRUE)
pca.Par=fdPar(basis.obj, 2, lambda[5])
pca.list.1=pca.fd(X1.fd, M[1], pca.Par)
pX1=pca.list.1$scores
```

```
pca.list.2=pca.fd(X2.fd, M[1], pca.Par)
pX2=pca.list.2$scores
pca.list.3=pca.fd(X3.fd, M[1], pca.Par)
pX3=pca.list.3$scores
pca.list.4=pca.fd(X4.fd, M[1], pca.Par)
pX4=pca.list.4$scores
pca.list.5=pca.fd(X5.fd, M[1], pca.Par)
pX5=pca.list.5$scores
pca.list.6=pca.fd(X6.fd, M[1], pca.Par)
pX6=pca.list.6$scores
pca.list.7=pca.fd(X7.fd, M[1], pca.Par)
pX7=pca.list.7$scores
pX11=cbind(pX1,pX2,pX3,pX4,pX5,pX6,pX7)
fit.1=lm(Y1~pX11)
summary(fit.1)
par(mfrow=c(2,2))
plot(fit.1)
pca.Par=fdPar(basis.obj, 2, lambda[1])
pca.list.1=pca.fd(X1.fd, M[3], pca.Par)
pX1=pca.list.1$scores
pca.list.2=pca.fd(X2.fd, M[3], pca.Par)
pX2=pca.list.2$scores
pca.list.3=pca.fd(X3.fd, M[3], pca.Par)
pX3=pca.list.3$scores
pca.list.4=pca.fd(X4.fd, M[3], pca.Par)
pX4=pca.list.4$scores
pca.list.5=pca.fd(X5.fd, M[3], pca.Par)
pX5=pca.list.5$scores
pca.list.6=pca.fd(X6.fd, M[3], pca.Par)
pX6=pca.list.6$scores
pca.list.7=pca.fd(X7.fd, M[3], pca.Par)
```

```
pX7=pca.list.7$scores
pX12=cbind(pX1,pX2,pX3,pX4,pX5,pX6,pX7)
fit.2=lm(Y2\sim pX12)
summary(fit.2)
par(mfrow=c(2,2))
plot(fit.2)
Part 2:
X=as.matrix(read.table("project.one.2.X.txt"))
Y=as.vector(as.matrix(read.table("project.one.2.Y.txt")))
color=rep("black",length(Y))
color[Y==2]="blue"
color[Y==3]="red"
color[Y==4]="green"
par(mfrow=c(1,1))
matplot(t(X),type="1",lty=1,col=color,main="Spectral Curves for 310
tablets colored by their classes
", ylab="Spectral Curve", xlab="Near Infrared Trasmittance from 7400 to
10507 cm-1")
t=0:403
X=scale(X,center=TRUE,scale=FALSE)
basis.obj=create.bspline.basis(c(0,404), nbasis=1000)
fdParobj=fdPar(basis.obj, 2, 1e-10)
fit.list=smooth.basis(t,t(X),fdParobj)
X.fd=fit.list$fd
lambda=10^{(c(-10,-8,-6,-4,-2,0,1))}
cv.fold=split(sample(310),rep(1:5,length=310))
CV.error=array(0,c(length(lambda),length(M)))
for(j in 1:length(lambda))
{for(k in 1:length(M))
{ print(c("j,k=",j,k))
pca.Par=fdPar(basis.obj,2,lambda[j])
pca.list=pca.fd(X.fd,M[k],pca.Par)
pX=pca.list$scores
data=as.data.frame(cbind(Y,pX))
for(1 in 1:5)
{test=cv.fold[[1]]
train=setdiff(1:310,test)
```

```
fit=lda(Y~.,data=data)
CV.error[j,k]=CV.error[j,k]+
sum(predict(fit, data[-train, ])$class!=data[-train,1])/length(test)
print(c("CV.error[j,k]=",CV.error[j,k]))
}
}
cv=which(CV.error==min(CV.error), arr.ind=TRUE)

pca.Par=fdPar(basis.obj,2,lambda[3])
pca.list=pca.fd(X.fd,M[9],pca.Par)
pX=pca.list$scores
data=as.data.frame(cbind(Y,pX))
fit=lda(Y~.,data=data)
plot(fit)

fit=lda(Y~.,data=data,CV=TRUE)
sum(fit$class!=data$Y)/length(Y)
```