

The original constrained problem

The original constrained problem is stated as

$$\begin{aligned} \min_{\tau} \quad & \iint -\text{sgn}(v) u f_Y(u, v; \tau) du dv \\ \text{subject to } & \kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau) dw dv \geq 0, \text{ and } \tau^\top \tau = 1. \end{aligned}$$

Suppose that the strict feasible set $\text{strict}(\mathcal{F})$ is non-empty, and let τ^0 denote a constrained minimizer of this original problem. For simplicity, we let $g(\tau) = -\iint \text{sgn}(v) u f_Y(u, v; \tau) du dv$, $c_1(\tau) = \kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau) dw dv$, and $c_2(\tau) = \tau^\top \tau$. The value of $g(\tau)$ at $\tau = \tau^0$, $g(\tau^0)$, is denoted by g^0 . Similarly, c_i^0 denotes the value of $c_i(\tau)$ at $\tau = \tau^0$, $c_i(\tau^0)$, for $i = 1, 2$. Also, \mathcal{A}^0 denotes the set of active constraint at τ^0 , $\mathcal{A}(\tau^0)$. In our current case, it is either $\mathcal{A}^0 = \{c_1^0, c_2^0\}$, or $\mathcal{A}^0 = \{c_2^0\}$.

Penalty-barrier function

One of the method to solve equality-inequality mixed constrained optimization problem is use penalty-barrier function, which is a composite measure of the objective function and the penalty of violating the constraints. The penalty-barrier function is then formalized as

$$\begin{aligned} \Phi_{BP}(\tau, \mu) = & \iint -\text{sgn}(v) u f_Y(u, v; \tau) du dv \\ & - \mu \ln \left[\kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau) dw dv \right] + \frac{1}{2\mu} \|\tau^\top \tau - 1\|_2^2, \end{aligned}$$

where μ is a sequence of positive decreasing small constants converging to zero. An unconstrained minimizer of $\Phi_{BP}(\tau, \mu)$ is denoted by $\tau^*(\mu)$ for emphasizing that it is a vector function of μ , or τ_μ^* for short. It can be proven that the constraint is strictly satisfied, i.e., $c(\tau_\mu^*) = \kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau_\mu^*) dw dv > 0$.

There is theorem which gives the conditions under which, for sufficiently small μ , the sequence $\{\tau_\mu^*\}$ defines a differentiable penalty-barrier trajectory converging to τ_μ^0 .

To find τ_μ^* , we exploit its stationarity. The gradient of $\Phi_{PB}(\tau, \mu)$ is

$$\begin{aligned} \nabla_\tau \Phi_{PB}(\tau, \mu) = & \iint -\text{sgn}(v) u \nabla_\tau f_Y(u, v; \tau) du dv \\ & + \mu \frac{\iint \text{sgn}(v) w \nabla_\tau f_Z(w, v; \tau) dw dv}{\kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau) dw dv} + \frac{2}{\mu} (\tau^\top \tau - 1) \tau, \end{aligned}$$

noting that ∇_τ represents the first order derivative with respect to τ . If we are willing to assume that $\Phi_{PB}(\tau, \mu)$ is twice-continuously differentiable, it must hold that $\nabla \Phi_{PB}(\tau_\mu^*, \mu) = 0$ to satisfy the stationarity, i.e.,

$$\iint \text{sgn}(v) u \nabla_\tau f_Y(u, v; \tau_\mu^*) du dv = \mu \frac{\iint \text{sgn}(v) w \nabla_\tau f_Z(w, v; \tau_\mu^*) dw dv}{\kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau_\mu^*) dw dv} + \frac{2}{\mu} (\tau_\mu^{*\top} \tau_\mu^* - 1) \tau_\mu^*,$$

with $\tau_\mu^{*\top} \tau_\mu^* - 1 = 0$. The barrier multiplier, the coefficient in this linear relationship above, denoted by λ_μ , is defined as

$$\lambda_\mu \triangleq \frac{\mu}{\kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau_\mu) dw dv}.$$

This relationship can be re-written as

$$\lambda_\mu \left[\kappa - \iint \text{sgn}(v) w f_Z(w, v; \tau_\mu) dw dv \right] = \mu.$$

This relationship between the barrier multiplier, the constraint value, and the barrier parameter, called perturbed complementarity, is analogous as $\mu \rightarrow 0$ to the complementarity condition $c(\tau^*)\lambda^* = 0$ that holds at a KKT point.

Estimation of the log-barrier penalty function

To estimate the log-barrier penalty function, we use kernel density estimators, denoted by $\hat{f}_Y(u, v; \tau)$ and $\hat{f}_Z(w, v; \tau)$, to estimate the corresponding density functions. Hence, the estimated log-barrier function is

$$\hat{B}(\tau, \mu) = \iint -\text{sgn}(v) u \hat{f}_Y(u, v; \tau) du dv - \mu \ln \left[\kappa - \iint \text{sgn}(v) w \hat{f}_Z(w, v; \tau) dw dv \right],$$

and the gradient of the estimator is

$$\begin{aligned} \nabla \hat{B}(\tau, \mu) &= \iint -\text{sgn}(v) u \nabla \hat{f}_Y(u, v; \tau) du dv + \mu \frac{\iint \text{sgn}(v) w \nabla \hat{f}_Z(w, v; \tau) dw dv}{\kappa - \iint \text{sgn}(v) w \hat{f}_Z(w, v; \tau) dw dv} \\ &= \iint -\text{sgn}(v) u \nabla \hat{f}_Y(u, v; \tau) du dv + \hat{\lambda}_\mu \iint \text{sgn}(v) w \nabla \hat{f}_Z(w, v; \tau) dw dv, \end{aligned}$$

where $\hat{\lambda}_\mu(\tau) = \mu / \left[\kappa - \iint \text{sgn}(v) w \hat{f}_Z(w, v; \tau) dw dv \right]$.

Consistency of $\hat{\tau}^k$ and $\hat{\lambda}_\mu$.

We need to prove that $\hat{\tau}^k$ is a consistent estimator of τ^{*k} .

$\hat{\tau}^k - \tau^{*k} = O_p(n^{1/2})$, and $\hat{\lambda}^k - \lambda^{*k} = O_p(n^{1/2})$.

Theorem proved that λ_μ is bounded.

Asymptotic distribution of $\hat{\tau}^k$

Estimating equations:

$$\nabla \hat{B}(\tau, \mu) = \iint -\text{sgn}(v)u \nabla \hat{f}_Y(u, v; \tau) du dv + \hat{\lambda}_\mu(\tau) \iint \text{sgn}(v)w \nabla \hat{f}_Z(w, v; \tau) dw dv = 0$$

where $\hat{\lambda}_\mu(\tau) = \mu / [\kappa - \iint \text{sgn}(v)w \hat{f}_Z(w, v; \tau) dw dv]$.

$$\begin{aligned} \nabla \hat{B}(\tau, \mu) &= \iint -\text{sgn}(v)u \nabla \hat{f}_Y(u, v; \tau) du dv + \hat{\lambda}_\mu \iint \text{sgn}(v)w \nabla \hat{f}_Z(w, v; \tau) dw dv \\ &= -\frac{2}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{Y1} k \left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h} \right) \mathbf{X}_i + \hat{\lambda}_\mu(\tau) \frac{2}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{Z1} k \left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h} \right) \mathbf{X}_i \\ &= N(\mu_1, \Sigma_1) + C_p N(\mu_2, \Sigma_2) \end{aligned}$$

$$\begin{aligned} \nabla^2 \hat{B}(\tau, \mu) &= \iint -\text{sgn}(v)u \nabla^2 \hat{f}_Y(u, v; \tau) du dv + \\ &\quad \nabla \hat{\lambda}_\mu(\tau) \iint \text{sgn}(v)w \nabla \hat{f}_Z(w, v; \tau) dw dv + \hat{\lambda}_\mu(\tau) \iint \text{sgn}(v)w \nabla^2 \hat{f}_Z(w, v; \tau) dw dv \\ &= -\frac{2}{nh} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \left(\hat{\lambda}_\mu(\tau) \boldsymbol{\beta}_{Z1} - \boldsymbol{\beta}_{Y1} \right) k' \left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h} \right) \mathbf{X}_i \mathbf{X}_i^\top + \\ &\quad \frac{2}{n} \sum_{i=1}^n \nabla \hat{\lambda}_\mu(\tau) \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{Z1} k \left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h} \right) \mathbf{X}_i \end{aligned}$$

$$\begin{aligned} \nabla \hat{\lambda}_\mu(\tau) &= \frac{\mu}{\left(\kappa - \iint \text{sgn}(v)w \hat{f}_Z(w, v; \tau) dw dv \right)^2} \iint \text{sgn}(v)w \nabla \hat{f}_Z(w, v; \tau) dw dv \\ &= \mu \left[\kappa - \frac{1}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{Z1} \left\{ 1 - 2K \left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h} \right) \right\} \right]^{-1} \left[\frac{2}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{Z1} k \left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h} \right) \mathbf{X}_i \right] \end{aligned}$$

[Notation: k and κ looks to similar]

[Need to estimate $\hat{\beta}$ too]

Need to prove that the difference between $\hat{B}_n(\tau, \hat{\beta}, \mu)$ and $\hat{B}_n(\tau, \beta^*, \mu)$ is negligible? i.e., $\hat{B}_n(\beta^*) - \hat{B}_n(\hat{\beta}) = O_p(n^{-1/2})$

Taylor expansion of $\nabla \hat{B}(\tau^{*k}, \mu)$ at $\tau = \hat{\tau}^k$ shows that

$$\nabla \hat{B}(\tau^{*k}, \mu) = \nabla \hat{B}(\hat{\tau}^k, \mu) - \nabla^2 \hat{B}(\tilde{\tau}^k, \mu)(\hat{\tau}^k - \tau^{*k}),$$

where $\tilde{\tau}^k$ is between τ^{*k} and $\hat{\tau}^k$. As $\hat{\tau}^k$ is the minimizer of $B(\tau, \mu)$, it satisfies the first order condition that $\nabla B(\hat{\tau}^k, \mu) = 0$. Therefore, we have

$$\sqrt{n} \nabla \hat{B}(\tau^{*k}, \mu) = -\sqrt{n} \nabla^2 \hat{B}(\tilde{\tau}^k, \mu)(\hat{\tau}^k - \tau^{*k}).$$

Derivation of the integrations

The integration we need

$$\begin{aligned} & \iint \text{sgn}(v) u f(u, v; \tau, \beta_{\cdot 1}) du dv \\ &= 2 \iint u \mathbb{I}(v \geq 0) f(v, u; \tau, \beta_{\cdot 1}) dv du - \int u f(u; \beta_{\cdot 1}) du \end{aligned}$$

The estimator is

$$\begin{aligned} & \iint \text{sgn}(v) u \hat{f}_n(u, v; \tau, \beta_{\cdot 1}) du dv \\ &= 2 \iint u \mathbb{I}(v \geq 0) \hat{f}_n(v, u; \tau, \beta_{\cdot 1}) dv du - \int u \hat{f}_n(u; \tau, \beta_{\cdot 1}) du \\ &= \frac{2}{nh^2} \iint u \mathbb{I}(v \geq 0) \sum_{i=1}^n k\left(\frac{v - V_i}{h}\right) k\left(\frac{u - U_i}{h}\right) du dv - \\ & \quad \frac{1}{nh} \int u \sum_{i=1}^n k\left(\frac{u - U_i}{h}\right) du \\ &= \frac{2}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{\cdot 1} \left\{ 1 - K\left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h}\right) \right\} - \frac{1}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{\cdot 1} \\ &= \frac{1}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{\cdot 1} \left\{ 1 - 2K\left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h}\right) \right\} \end{aligned}$$

where $\hat{f}_n(u_1, u_2; \tau, \hat{\beta}_{\cdot 1})$ are the kernel density estimator for $(X^\top \tau, X^\top \beta_{\cdot 1})$ with the forms of

$$\hat{f}_n(u, v; \boldsymbol{\tau}, \hat{\boldsymbol{\beta}}_{\cdot 1}) = \frac{1}{nh^2} \sum_{i=1}^n k\left(\frac{u - U_i}{h}\right) k\left(\frac{v - V_i}{h}\right).$$

Moreover, $K(s)$ is the corresponding CDF of the kernel function $k(s)$, which is chosen to be a symmetric probability density. More precisely, $k(s)$ satisfies the following assumptions:

1. $\int_{-\infty}^{\infty} k(s) ds = 1$.
2. $k(s) > 0$ for all s .
3. $k(-s) = k(s)$ for all s .

4. The first order derivative of the kernel, $k'(s)$, exists and is bounded.

The last equality above holds by following the derivation.

We first derive $\frac{2}{h^2} \iint u_2 \mathbb{I}(u_1 \geq 0) k\left(\frac{u_1 - U_{i,1}}{h}\right) k\left(\frac{u_2 - U_{i,2}}{h}\right) du_1 du_2$. Let $s = \frac{u_1 - U_{i,1}}{h}$ and $t = \frac{u_2 - U_{i,2}}{h}$. Then, $u_1 = U_{i,1} + sh$ and $u_2 = U_{i,2} + th$. Also, $du_1 = h ds$ and $du_2 = h dt$.

$$\begin{aligned}
& \frac{2}{h^2} \iint u_2 \mathbb{I}(u_1 \geq 0) k\left(\frac{u_1 - U_{i,1}}{h}\right) k\left(\frac{u_2 - U_{i,2}}{h}\right) du_1 du_2 \\
&= 2 \iint (U_{i,2} + th) \mathbb{I}(U_{i,1} + sh \geq 0) k(s) k(t) ds dt \\
&= 2 \int U_{i,2} \mathbb{I}\left(s \geq -\frac{U_{i,1}}{h}\right) k(s) ds \\
&= 2U_{i,2} \left\{ 1 - K\left(-\frac{U_{i,1}}{h}\right) \right\} \\
&= 2\mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{\cdot 1} \left\{ 1 - K\left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h}\right) \right\},
\end{aligned}$$

where $K(s) = \int k(s) ds + c$. The second equality holds, as $\int k(t) dt = 1$ and $\int t k(t) dt = 0$. The third equality holds as $\int \mathbb{I}\left(s \geq -\frac{U_{i,1}}{h}\right) k(s) ds = 1 - \int_{-\infty}^{-U_{i,1}/h} k(s) ds = 1 - K\left(-\frac{U_{i,1}}{h}\right)$, where $U_{i,1} = \mathbf{X}_i^\top \boldsymbol{\tau}$.

Then, we derive $\frac{1}{h} \int u_2 k\left(\frac{u_2 - U_{i,2}}{h}\right) du_2$ by changing variable similarly. Let $t = \frac{u_2 - U_{i,2}}{h}$, and we get $u_2 = U_{i,2} + th$, and $du_2 = h dt$.

$$\begin{aligned}
& \frac{1}{h} \int u_2 k\left(\frac{u_2 - U_{i,2}}{h}\right) du_2 \\
&= \int (U_{i,2} + th) k(t) dt \\
&= U_{i,2} = \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{\cdot 1}.
\end{aligned}$$

Again, the second equality holds as $\int k(t) dt = 1$, and $\int t k(t) dt = 0$. The integration over the first-order derivative

$$\begin{aligned}
& \iint \text{sgn}(v) u \nabla \widehat{f}_n(u, v; \tau, \beta_{\cdot 1}) du dv \\
&= \frac{\partial}{\partial \tau} \iint \text{sgn}(v) u \widehat{f}_n(u, v; \tau, \beta_{\cdot 1}) du dv \\
&= \frac{2}{n} \sum_{i=1}^n \mathbf{X}_{i,1}^\top \boldsymbol{\beta}_{\cdot 1} k\left(-\frac{\mathbf{X}_i^\top \boldsymbol{\tau}}{h}\right) \mathbf{X}_i
\end{aligned}$$