## Machine Learning

Homework 3 solution

- Shyam Arumugaswamy

1. 
$$\frac{d}{d\theta}\theta^{t}(1-\theta)^{h} = \theta^{t}.h.(1-\theta)^{h-1}.(-1) + (1-\theta)^{h}.t.\theta^{t-1} -- chain rule$$

$$= -\theta^{t}.h.(1-\theta)^{h-1} + t.\theta^{t-1}.(1-\theta)^{h}$$

$$= -h.\theta.\theta^{t-1}.(1-\theta)^{h-1} + t.(1-\theta)^{h-1}.\theta^{t-1}(1-\theta)$$

$$= \theta^{t-1}.(1-\theta)^{h-1}(t(1-\theta) - h\theta)$$

$$\frac{d^{2}}{d^{2}\theta}\theta^{t}(1-\theta)^{h} = \theta^{t-1}.(1-\theta)^{h-1}(-t-h) + ((t(1-\theta)-h\theta)[-\theta^{t-1}.(1-\theta)^{h-2}.(h-1) + \theta^{t-2}.(1-\theta)^{h-1}.(t-1)])$$

$$= -\theta^{t-1}.(1-\theta)^{h-1}(t+h) + \theta^{t-2}.(1-\theta)^{h-2}.(t.(1-\theta)-\theta h).((t-1).(1-\theta)-(h-1)\theta)$$

let 
$$f(\theta) = \log \theta^{t} (1-\theta)^{h} = t \log \theta + h \cdot \log(1-\theta)$$

$$\frac{d}{d\theta}f(\theta) = t/\theta - h/(1-\theta)$$

$$\frac{d^2}{d^2\theta}f(\theta) = -t/\theta^2 - h/(1-\theta)^2$$

2. Let  $p(\theta) = \log f(\theta)$  and  $\theta^t$  be any local maximum of  $p(\theta)$ , so we have  $p(\theta^t) >= p(\theta)$  for  $\theta$  in small neighbourhood of  $\theta^t$ . for a monotonic transform like exponential function, we can say

$$f(\theta^t) = \exp(p(\theta^t)) >= \exp(p(\theta)) = f(\theta)$$
 ... so  $\theta^t$  is a max of f from the previous problem, we can see that computational becomes simpler using log and maximum or minimum is preserved.

3.  $\theta_{MLE} = \arg \max \theta p(D | \theta)$  and  $\theta_{MAP} = \arg \max \theta p(D | \theta) p(\theta)$ 

If we consider uniform distribution as prior, then  $p(\boldsymbol{\theta})$  will be constant for all  $\boldsymbol{\theta}$ . So arg max  $\boldsymbol{\theta}$   $p(D|\boldsymbol{\theta})$   $p(\boldsymbol{\theta})$  = arg max  $\boldsymbol{\theta}$   $p(D|\boldsymbol{\theta})$  therefore,  $\boldsymbol{\theta}_{MAP} = \boldsymbol{\theta}_{MLE}$ 

4.  $p(x=m \mid N, \theta) = (^{N}_{m}) \theta^{m} (1 - \theta)^{N-m}$  if a, b are hyper parameters of prior beta distribution, then the posterior is Beta(m+a,l+b) distributed.

so, the posterior mean is

$$E(\boldsymbol{\theta} \mid D) = \frac{m+a}{m+a+l+b}$$
$$= \frac{m}{m+a+l+b} + \frac{a}{m+a+l+b}$$

$$=\frac{m}{m+a+1+b} \cdot \frac{m+1}{m+1} + \frac{a}{m+a+1+b} \cdot \frac{a+b}{a+b}$$

$$=\frac{m+1}{m+a+1+b}\cdot\frac{m}{m+1}+\frac{a+b}{m+a+1+b}\cdot\frac{a}{a+b}$$

if 
$$\frac{a+b}{m+a+l+b} = \lambda$$

then 
$$1 - \lambda = 1 - \frac{a+b}{m+a+l+b}$$

$$= \frac{m+a+l+b-a-b}{m+a+l+b}$$

$$= \frac{m+l}{m+a+l+b}$$

therefore, 
$$E(\boldsymbol{\theta} \mid D) = (1 - \lambda) \cdot \frac{m}{m+1} + \lambda \cdot \frac{a}{a+b}$$

the max likelihood estimate is  $\frac{m}{m+1}$  and posterior mean is  $\frac{a}{a+b}$ 

5. Let X be Poison distributed

Poi 
$$(x|\lambda) = \frac{e^{-\lambda}}{x!} \cdot \lambda^{x}$$

$$\ln(\lambda) = \ln \prod_{x=1}^{n} \frac{e^{-\lambda}}{x!} \cdot \lambda^{x} = \sum_{i=1}^{n} \ln e^{-\lambda} + \sum_{i=1}^{n} \ln \frac{\lambda^{xi}}{x!}$$
$$= -n\lambda + \sum_{i=1}^{n} (xi \ln \lambda - \ln x!)$$

$$d (\ln x) / dx = -n + nx / x = 0$$

$$E[\lambda] = \sum_{i=1}^{n} \ln \frac{E[x!]}{n} = \lambda$$
  
prior for  $\lambda$ :

alpha & beta are denoted as a & b in this problem

$$p(\lambda \mid \alpha, \beta) = \frac{\beta^{\alpha}}{\tau(\alpha)} \lambda^{\alpha-1} \exp(-\beta \lambda)$$

so P(\(\lambda\) | D)  $\,\,\varpropto$  P(D | \(\lambda\)) P(\(\lambda\)) and is Gamma ( $\sum_{i=1}^{n} xi + \alpha$  , n +  $\beta$ )

for MAP soln, arg max  $_{\lambda}$  p( $\lambda$ |D) = arg max  $_{\lambda}$  ln p( $\lambda$ |D)

$$\label{eq:definition} \mbox{ln p($\lambda$ | D) = } (\sum_{i=1}^{n} xi \ + \ \alpha - 1) \mbox{ln $\lambda$} \ - (n + \beta) \lambda \ + c$$

taking first derivative

$$\frac{d}{d\lambda}\ln p(\lambda|D) = \left(\sum_{i=1}^{n} xi + \alpha - 1\right)/\lambda - (n+\beta) = 0$$

$$\lambda = \frac{(\sum_{i=1}^{n} xi + \alpha - 1)}{(n+\beta)}$$